

Introduction to fitting and optimization using Python libraries

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Non-linear least squares

 \triangleright Set of m observations

$$(x_1, y_1), (x_2, y_2), (x_3, y_3), \dots, (x_m, y_m)$$

Model function

$$\hat{y} = f(x, \beta_1, \beta_2, \beta_3, \dots, \beta_n) = f(x, \beta)$$

ightharpoonup Vector of n < m fitting parameters

$$\boldsymbol{\beta} = (\beta_1, \beta_2, \beta_3, \dots, \beta_n)$$

Residuals

$$r_i = y_i - f(x_i, \boldsymbol{\beta})$$

Minimization sum of squared residuals

$$S = \sum_{i=1}^{m} r_i^2$$

Gradient of objective function

$$\frac{\partial S}{\partial \beta_j} = 2 \sum_{i=1}^m \frac{\partial r_i}{\partial \beta_j} r_i$$

$$= -2 \sum_{i=1}^m J_{ij} r_i = 0$$

$$j = 0, 1, 2, \dots, n$$

Non-linear least squares

Change in residual

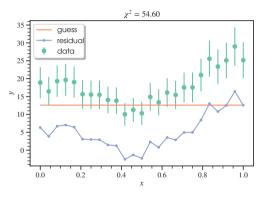
$$\Delta y_i = y_i - f(x_i, \boldsymbol{\beta}^k)$$

Residual approximation

$$r_i \approx \Delta y_i - \sum_{s=1}^n J_{is} \Delta \beta_s$$

Normal equations

$$\sum_{i=1}^{m} \sum_{s=1}^{n} J_{ij} J_{is} \Delta \beta = \sum_{i=1}^{m} J_{ij} \Delta y_{i} \quad j = 0, 1, 2, \dots, n$$
$$(\boldsymbol{J}^{\mathsf{T}} \boldsymbol{J}) \Delta \beta = \boldsymbol{J}^{\mathsf{T}} \Delta \boldsymbol{y}$$



Fitting a one-dimensional function with uncertainties: initial guess

^{*}If Jacobian function cannot be analytically derived, finite difference methods can be used as approximation. Accuracy (e.g. step size or higher order differences) typically needs to be balanced with the cost of evaluating difference equations.

Non-linear least squares

Jacobian matrix *

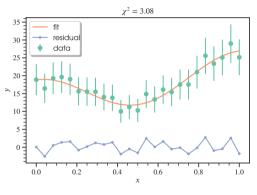
$$\boldsymbol{J}_{ij} = -\frac{\partial r_i}{\partial \beta_j} = \frac{\partial f(\boldsymbol{x}_i, \boldsymbol{\beta})}{\partial \beta_j} \quad \boldsymbol{J}_i = \frac{\partial f(\boldsymbol{x}_i, \boldsymbol{\beta})}{\partial \boldsymbol{\beta}}$$

lterative refinement shifting parameter $\Delta \beta_j$ from step k to step k+1

$$\beta_j pprox \beta_j^{k+1} = \beta_j^k + \Delta \beta_j$$

First-order Taylor expansion

$$f(x_i, \boldsymbol{\beta}) \approx f(x_i, \boldsymbol{\beta}^k) + \sum_{j=1}^n \frac{\partial f(x_i, \boldsymbol{\beta}^k)}{\partial \beta_j} \Delta \beta_j$$
$$= f(x_i, \boldsymbol{\beta}^k) + \sum_{j=1}^n J_{ij} \Delta \beta_j$$



Fitting a one-dimensional function with uncertainties: final fit

Non-linear weighted least squares

 \blacktriangleright Set of m observations with uncertainties

$$(x_1, y_1, \sigma_1), (x_2, y_2, \sigma_2), \dots, (x_m, y_m, \sigma_m)$$

Mean squared weighted deviation

$$\chi^{2} = \sum_{i=1}^{m} \frac{[y_{i} - f(x_{i}, \beta)]^{2}}{\sigma_{i}^{2}}$$

Diagonal weight matrix †

$$W_{ii} = \frac{1}{\sigma_i^2}$$

Weighted normal equations

$$(\mathbf{J}^{\mathsf{T}}\mathbf{W}\mathbf{J})\Delta\mathbf{\beta} = \mathbf{J}^{\mathsf{T}}\mathbf{W}\Delta\mathbf{v}$$

$$C = (J^{\mathsf{T}}WJ)$$

lacksquare Parameter uncertainties $eta_j \pm \epsilon_j$

$$\epsilon_j = \sqrt{C_{jj}}$$

Reduced chi-squared statistic v = m - n

$$\chi_{\nu}^{2} = \frac{\chi^{2}}{\nu} = \frac{1}{m-n} \sum_{i=1}^{m} \frac{[y_{i} - f(x_{i}, \boldsymbol{\beta})]^{2}}{\sigma_{i}^{2}}$$

Standard errors $^{\dagger}\chi^2(\beta_j \pm \hat{\epsilon}_j) = \chi^2(\beta_j) + 1$

$$\hat{\epsilon}_j = \sqrt{\chi_{\nu}^2 C_{jj}} = \frac{\chi \epsilon_j}{\sqrt{m-n}}$$

Covariance matrix

[†]Assumes uncertainties are independent and uncorrelated.

[†]Useful when the original weights are unreliable.

Non-linear weighted least squares

Various methods

- Gauss-Newton method
- Levenberg-Marquardt algorithm

Matrix methods

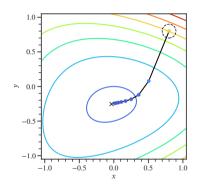
- QR decomposition
- Singular Value decomposition

Gradient methods

- Newton's method
- Davidon-Fletcher-Powell method
- Steepest descent
- Conjugate gradient search

Direct search methods

- ▶ Alternating variable search
- ▶ Nelder-Mead (simplex) search



Gauss-Newton method for minimizing a two-dimensional function

Levenberg-Marquardt algorithm

Interpolation between Gauss–Newton algorithm and steepest descent

Weighted normal equations

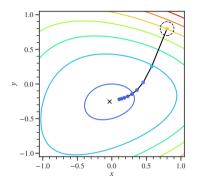
$$(\boldsymbol{J}^{\intercal}\boldsymbol{W}\boldsymbol{J} + \lambda\boldsymbol{I})\Delta\boldsymbol{\beta} = \boldsymbol{J}^{\intercal}\boldsymbol{W}\Delta\boldsymbol{y}$$

Steepest descent §

$$\lambda \boldsymbol{I} \gg \boldsymbol{J}^{\mathsf{T}} \boldsymbol{W} \boldsymbol{J} \quad \Delta \boldsymbol{\beta} = \frac{1}{\lambda} \boldsymbol{J}^{\mathsf{T}} \boldsymbol{W} \Delta \boldsymbol{y}$$

Gauss-Newton method ¶

$$J^{\mathsf{T}}WJ \gg \lambda I \quad (J^{\mathsf{T}}WJ)\Delta\beta = J^{\mathsf{T}}W\Delta y$$



Steepest descent method for minimizing a two-dimensional function

Sconverges slowly, but solution will converge with appropriate step size.

Converges rapidly in close proximity to minima, but solution may diverge.

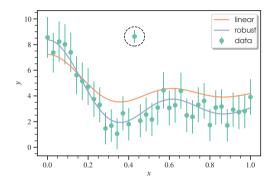
Robust non-linear optimization

Robust optimization using loss functions

$$\sum_{i=1}^m \rho(W_{ii} r_i^2))$$

Loss functions

- Maximum likelihood (least squares) $\rho(z) = z$
- Soft approximation of L_1 norm $\rho(z) = 2(\sqrt{1+z} 1)$
- $\text{Huber } \rho(z) = \begin{cases} z & z < 1\\ \sqrt{z} 1 & z > 1 \end{cases}$
- Arctan $\rho(z) = \arctan z$



Robust optimization using soft approximation of $L_{\rm I}$ norm improves the fit when data contains outliers

Constrained optimization

Minimize a function $\chi^2(\beta)$ subject to

$$g_k(\boldsymbol{\beta}) = c_k \quad k = 1, 2, \dots$$

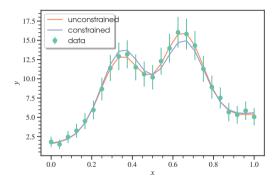
 $h_l(\boldsymbol{\beta}) \ge d_l \quad l = 1, 2, \dots$

Equality constraints

- Substitution
- Lagrange multipliers

Inequality constraints

- Penalty methods
- Linear and nonlinear programming
- Quadratic programming
- Branch and bound
- ▶ First-choice bounding functions
- Bucket elimination



Fitting of two overlapping Gaussian peaks with equal widths, equal amplitudes, and peak overlap limited to two standard deviations between centers

Python libraries

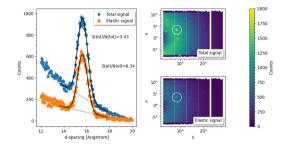
Many fitting Python libraries are available

SciPy |

- scipy.optimize.curve_fit
- scipy.optimize.least_squares
- scipy.optimize.leastsq
- scipy.optimize.minimize
- LMfit **
 - lmfit.minimize

Mantid †† (GNU scientific library)

Fit



Fitting single crystal peaks from a protein collected at CORELLI using Mantid

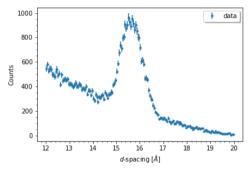
https://docs.scipy.org/doc/scipy/reference/optimize.html

^{**}https://lmfit.github.io/lmfit-py/intro.html

tthttps://www.gnu.org/software/gsl/

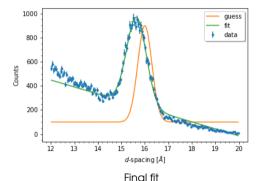
Loading and plotting data

```
import numpy as np
import matplotlib.pyplot as plt
import scipy.optimize
import lmfit
import h5py
# --- load data ---
f = h5py.File('../data/total.nxs', mode='r')
ws = f['mantid_workspace_1/workspace/']
d_spacing_bin_edges = ws['axis1'][()]
counts = ws['values'][()].flatten()
errors = ws['errors'][()].flatten()
f.close()
# --- plot data ---
d_spacing = 0.5*(d_spacing_bin_edges[1:]+d_spacing_bin_edges[:-1])
fig, ax = plt.subplots(1,1)
ax.errorbar(d_spacing, counts, yerr=errors, fmt='.', label='data')
ax.legend(shadow=True)
ax.minorticks on()
```



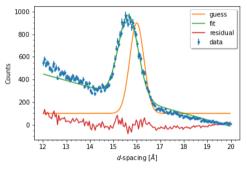
Plotted peak

Initializing and fitting model



Plotting residuals and calculating uncertanties

```
# --- plot residuals ---
def residual(d. counts, A. mu, sigma, B. c):
    return counts-model(d, A, mu, sigma, B, c)
ax.plot(d_spacing, residual(d_spacing, counts, *popt), label='residual')
ax.legend(shadow=True)
fig.savefig('scd-residual-total.png')
# --- calculate errors ---
perr = np.sqrt(np.diag(pcov))
print('Fitted uncertanties as 1-std using curve fit')
print('A
             = {:8.3f} ± {:5.3f}'.format(popt[0].perr[0]))
print('mu
             = {:8.3f} ± {:5.3f}'.format(popt[1].perr[1]))
print('sigma = {:8.3f} + {:5.3f}'.format(popt[2].perr[2]))
print('B
             = {:8.3f} ± {:5.3f}'.format(popt[3],perr[3]))
print('c
             = {:8.3f} ± {:5.3f}'.format(popt[4].perr[4]))
```



Final residuals

More examples

```
# --- define weighted least squares problem ---
def weighted deviations(x, d, counts, errors):
    A. mu. sigma. B. c = x
    return residual(d. counts, A. mu. sigma, B. c)/errors
sol = scipy.optimize.least_squares(weighted_deviations, x0=p0,
                                   args=(d spacing, counts, errors).
                                   method='lm')
vals = sol.x
J = sol.iac
cov = np.linalg.inv(.L.T.dot(.I))
err = np.sqrt(np.diag(cov))
print('Fitted uncertanties as 1-std using least_squares')
print('A
            = {:8.3f} ± {:5.3f}'.format(vals[0],err[0]))
print('mu = {:8.3f} ± {:5.3f}'.format(vals[1].err[1]))
print('sigma = {:8.3f} # {:5.3f}', format(vals[2], err[2]))
print('B
           = \{:8.3f\} \pm \{:5.3f\}', format(vals[3], err[3])\}
print('c
          = \{:8.3f\} \pm \{:5.3f\}', format(vals[4], err[4])\}
chi2dof = np.sum(sol.fun**2)/(sol.fun.size-sol.x.size)
cov *= chi2dof
stderr = np.sqrt(np.diag(cov))
print('Fitted uncertanties as 1-stderr using least squares')
print('A
             = {:8.3f} ± {:5.3f}'.format(vals[0],stderr[0]))
print('mu = {:8.3f} ± {:5.3f}'.format(vals[1].stderr[1]))
print('sigma = {:8.3f} ± {:5.3f}'.format(vals[2].stderr[2]))
print('B
             = {:8.3f} ± {:5.3f}'.format(vals[3].stderr[3]))
print('c
             = {:8.3f} ± {:5.3f}'.format(vals[4].stderr[4]))
```

```
# --- use constranied optimization ---
def weighted residual(params, d. counts, errors):
    A = params['A']
    mu = params['mu']
    sigma = params['sigma']
    B = params['B']
    c = params['c']
    return residual(d. counts, A. mu. sigma, B. c)/errors
params = lmfit.Parameters()
params.add('A', value=A, min=0, max=np.inf)
params.add('mu', value=mu, min=-np.inf, max=np.inf)
params.add('sigma', value=sigma, min=0, max=np.inf)
params.add('B', value=B, min=0, max=np.inf)
params.add('c', value=c, min=-np.inf, max=np.inf)
result = lmfit.minimize(weighted_residual, params,
                        args=(d_spacing, counts, errors))
print('Fitted uncertanties as 1-stderr using lmfit')
for key in params.keys():
    value, stderr = result.params[kev].value, result.params[kev].stderr
    print('{:7} = {:8.3f} ± {:5.3f}'.format(kev.value.stderr))
```