## **SESSION 3**

# RANDOM DISTRIBUTIONS

# DISCRETE VS. CONTINUOUS DISTRIBUTIONS

- Random Variable -- a variable which contains the outcomes of a chance experiment
- **Discrete Random Variable** -- the set of all possible values is at most a finite or a countably infinite number of possible values
  - Number of new subscribers to a magazine
  - Number of bad checks received by a restaurant
  - Number of absent employees on a given day
- Continuous Random Variable -- takes on values at every point over a given interval
  - Elapsed time between arrivals of bank customers
  - Percent of the labor force that is unemployed

# CUMULATIVE DISTRIBUTION FUNCTION

<u>Def.</u> The cumulative distribution function (or cdf) F(.) of the random variable X for any real number b in the set of possible values is defined as

$$F(b) = P(X \le b)$$

## Properties of cumulative distribution function

- 1. F(b) is a nondecreasing function of b.
- 2.  $F(\infty) = 1, F(-\infty) = 0$
- 3.  $P(a < X \le b) = F(b) F(a)$  for all a < b

## Requirements for a Discrete Probability Function

• Probabilities are between 0 and 1, inclusively

$$0 \le P(X) \le 1$$
 for all  $X$ 

Total of all probabilities equals

$$\sum_{\text{over all } X} P(X) = 1$$

## **Examples**

❖ Let X denote the random variable that is defined as the sum of two fair dice; then

$$P(X = 2) = 1/36$$
;  $P(X = 3) = 2/36$ ;  $P(X = 4) = 3/36$ ;  $P(X = 5) = 4/36$ ;  $P(X = 6) = 5/36$ ;  $P(X = 7) = 6/36$ ;  $P(X = 8) = 5/36$ ;  $P(X = 9) = 4/36$ ;  $P(X = 10) = 3/36$ ;  $P(X = 11) = 2/36$ ;  $P(X = 12) = 1/36$ 

 $\bullet$  Tossing a coin having the probability p of landing head until the first head appears. Let N denote the number of flips required; then

$$P(N = 1) = p$$
  
 $P(N = 2) = (1 - p)p$   
...  
 $P(N = n) = (1 - p)^{n-1}p$   
Note that  $\sum_{n=1}^{\infty} P(N = n) = 1$ 

#### **Probability Mass Function**

<u>Def.</u> The probability mass function of a discrete random variable X, denoted by P(.), is defined as:

$$P(a) = P(X = a)$$

• The cumulative distribution function F(.) can be expressed as:

$$F(a) = \sum_{\text{All } x_i \le a} P(x_i)$$

Example  $\Leftrightarrow$  Let X have a probability mass function given by P(1) = 1/2; P(2) = 1/3; P(3) = 1/6 then

$$F(a) = \begin{cases} 0 & a < 1 \\ \frac{1}{2} & 1 \le a < 2 \\ \frac{5}{6} & 2 \le a < 3 \\ 1 & a \ge 3 \end{cases}$$

#### Mean (Expectation) of a Discrete Distribution

$$\mu = E(X) = \sum_{\text{All } x_i} x_i P(x_i)$$

### <u>Example</u>

X	P(X)	XP(X)
-1	0.1	-0.1
0	0.2	0.0
1	0.4	0.4
2	0.2	0.2
3	0.1	0.3

$$Mean = 1.0$$

#### Variance and Standard Deviation of a Discrete Distribution

$$\sigma^{2} = \sum_{\text{All } x_{i}} (x_{i} - \mu)^{2} P(x_{i}) \qquad \boxed{\sigma = \sqrt{\sigma^{2}}}$$

$$\sigma = \sqrt{\sigma^2}$$

#### <u>Example</u>

X	P(X)	$X-\mu$	$(X-\mu)^2$	$(X-\mu)^2 P(X)$
-1	0.1	-2	4	0.4
0	0.2	-1	1	0.2
1	0.4	0	0	0.0
2	0.2	1	1	0.2
3	0.1	2	4	0.2

Variance = 1.2

## **Binomial Distribution**

- Experiment involves *n* identical trials
- Each trial has exactly two possible outcomes: success and failure
- Each trial is independent of the previous trials
  - p is the probability of a success on any one trial
  - q = (1-p) is the probability of a failure on any one trial
  - p and q are constant throughout the experiment
  - X is the number of successes in the n trials

## <u>X is said to be a binomial random variable</u>

## **Binomial Distribution**

- Applications
  - Sampling with replacement
  - Sampling without replacement -- n < 5% N
- Probability function

$$P(X=i) = C_i^n p^i q^{n-i}$$

where

$$C_i^n = \frac{n!}{i!(n-i)!}$$

### **Binomial Distribution**

• Mean value

$$\mu = np$$

Variance and standard deviation

$$\sigma^2 = npq$$

$$\sigma = \sqrt{npq}$$

## **Binomial Distribution**

### <u>Examples</u>

❖ Four fair coin are flipped. What is the probability that two heads and two tails are obtained.

Letting X equal the number of heads (successes) that appear, then X is a binomial random variable with parameter (4, 1/2). Therefore,

$$P(X=2) = C_2^4 \left(\frac{1}{2}\right)^2 \left(1 - \frac{1}{2}\right)^{4-2} = \frac{3}{8}$$

### **Binomial Distribution**

### **Examples**

❖ If all items produced by a machine will be defective with probability 0.1, independently of each other. What is the probability that in a sample of three items, at most one will be defective?

If X is the number of defective items in the sample then X is a binomial random variable with parameters (3, 0.1). Hence, the desired probability is given by

$$P(X = 0) + P(X = 1) = C_0^3(0.1)^0(0.9)^3 + C_1^3(0.1)^1(0.9)^2 = 0.972$$

## **Poisson Distribution**

- Describes discrete occurrences over a continuum or interval
- A discrete distribution
- Describes rare events
- Each occurrence is independent any other occurrences
- The number of occurrences in each interval can vary from zero to infinity
- The expected number of occurrences must hold constant throughout the experiment.

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## **Poisson Distribution**

### **Probability Function**

$$P(X=i) = e^{-\lambda} \frac{\lambda^{i}}{i!}$$
  $\lambda$ : long-run average the base of natural logarithms  $(e = 2.718282...)$ 

## **Applications**

- Arrivals at queueing systems
  - airports -- people, airplanes, automobiles, baggage
  - banks -- people, automobiles, loan applications
  - computer file servers -- read and write operations
- Defects in manufactured goods
  - number of defects per 1,000 feet of extruded copper wire
  - number of blemishes per square foot of painted surface
  - number of errors per typed page

## **Poisson Distribution**

• Mean value

$$\mu = \lambda$$

Variance and standard deviation

$$\sigma^2 = \lambda$$

$$\sigma = \sqrt{\lambda}$$

### **Poisson Distribution**

### **Examples**

Suppose that the number of typographical errors on a single page of a book has a Poisson distribution with parameter  $\lambda = 1$ . Calculate the probability that there is at least one error on a certain page.

Let X denote the number of errors on a certain page of the book

$$P(X \ge 1) = 1 - P(X = 0) = 1 - e^{-1} \frac{1^0}{0!} = 1 - e^{-1} = 1 - \frac{1}{e} = 0.633$$

\* If the number of accidents occurring on a highway each day is a Poisson random variable with parameter  $\lambda = 3$ , what is the probability that no accidents occur today?

$$P(X = 0) = e^{-3} \frac{3^0}{0!} = e^{-3} = \frac{1}{e^3} = 0.05$$

## **Poisson Distribution**

### Poisson Approximation of the Binomial Distribution

- Binomial probabilities are difficult to calculate when *n* is large.
- Under certain conditions binomial probabilities may be approximated by Poisson probabilities
- Poisson approximation

If 
$$n > 20$$
 and  $np \le 7 \implies \text{Use } \lambda = np$ 

## **Probability Density Function**

<u>Def.</u> The probability density function of a discrete random variable X, denoted by f(.), is a function such that:

$$P(X \in B) = \int_{B} f(x) dx$$

From the above definition, it is noted that:

$$P\{X \in (-\infty, \infty)\} = \int_{-\infty}^{\infty} f(x)dx = 1$$

### **Probability Density Function**

#### *Notes*:

$$F(a) = P(X \le a) = \int_{-\infty}^{a} f(x) dx$$

$$\Rightarrow F(-\infty) = 0, F(\infty) = 1$$

$$P(a \le X \le b) = \int_a^b f(x)dx = F(b) - F(a)$$

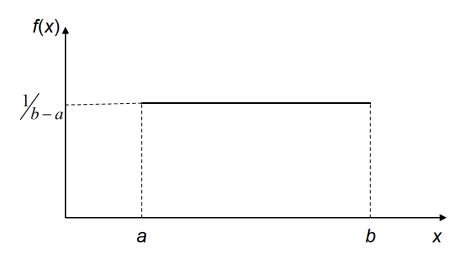
$$P(X=a) = \int_{a}^{a} f(x)dx = 0$$

$$P\left(a - \frac{\varepsilon}{2} \le X \le a + \frac{\varepsilon}{2}\right) = \int_{a - \varepsilon/2}^{a + \varepsilon/2} f(x) dx = \varepsilon f(a)$$

## **Uniform Distribution**

<u>Def.</u> A random variable X is said to be uniformly distributed over the interval (a, b) if its probability density function is given by

$$f(x) = \begin{cases} \frac{1}{b-a} & a \le x \le b \\ 0 & \text{otherwise} \end{cases}$$
  $(b > a)$ 



## **Uniform Distribution**

#### **Cumulative Distribution Function**

$$F(x) = \begin{cases} 0 & \text{if} & x \le a \\ \frac{x-a}{b-a} & \text{if} & a < x < b \\ 1 & \text{if} & x \ge b \end{cases}$$

Mean:

$$\mu = \frac{a+b}{2}$$

• Standard Deviation:  $\sigma = \frac{b-a}{\sqrt{12}}$ 

## Normal Distribution

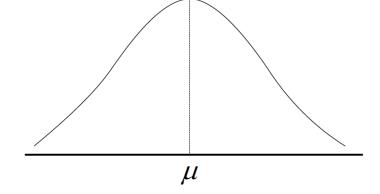
<u>Def.</u> A random variable X, with parameter  $\mu$  and  $\sigma^2$ , whose probability density function is defined by:

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{(x-\mu)^2}{2\sigma^2}} \qquad (-\infty < x < \infty)$$

where:  $\mu$ : Mean of X.

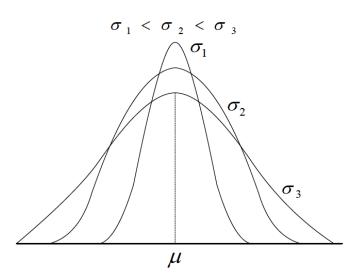
 $\sigma$ : Standard Deviation of X.

 $\pi = 3.14159$  e = 2.718282



## Normal Distribution

#### **Normal Curves for Different Standard Deviations**



- The cumulative distribution function of a normal distribution does not have an explicit expression.
- The cumulative distribution function of a normal distribution is given in tabular form for the case  $\mu = 0, \sigma = 1$  (which is called the *standardized normal distribution*)

## Normal Distribution

#### **Standardized Normal Distribution**

- A normal distribution with a mean of zero, and a standard deviation of one
- Z Formula standardizes any normal distribution
- Z Score
- computed by the Z Formula
- the number of standard deviations which a value is away from the mean

$$Z = \frac{X - \mu}{\sigma}$$

#### **Use of Normal Table**

## Normal Distribution

### Normal Approximation of the Binomial Distribution

- The normal distribution can be used to approximate binomial probabilities
- Procedure
  - 1. Convert binomial parameters to normal parameters
  - 2. Does the interval  $\mu \pm 3\sigma$  lie between 0 and n? If so, continue; otherwise, do not use the normal approximation.
  - 3. Correct for continuity
  - 4. Solve the normal distribution problem

## Normal Distribution

#### Normal Approximation of the Binomial Distribution

#### Conversion equations

$$\mu = np$$

$$\sigma = \sqrt{npq}$$

#### **Correcting for Continuity**

Value being determined	Correction
X >	+.50
$X \ge$	50
X <	50
$X \leq$	+.05
$\leq X \leq$	50 and $+.50$
< <i>X</i> <	+.50 and50

## Normal Distribution

### Normal Approximation of the Binomial Distribution

## <u>Example</u>

The binomial probability  $P(X \ge 25 | n = 60 \text{ and } p = .30)$ 

• Convert parameters: 
$$\mu = np = 18$$
,  $\sigma = \sqrt{npq} = 3.55$ 

• Check: 
$$\mu \pm 3\sigma = (7.35, 28.65) \in (0,60)$$

• Approximation: 
$$P(X \ge 24.5 | \mu = 18 \text{ and } \sigma = 3.55)$$

## **Exponential Distribution**

- Continuous
- Family of distributions
- Skewed to the right
- X varies from 0 to infinity
- Steadily decreases as *X* gets larger

#### **Probability function**

$$f(x) = \begin{cases} \lambda e^{-\lambda x} & if \quad x \ge 0 \\ 0 & if \quad x < 0 \end{cases}$$

## **Exponential Distribution**

#### **Cumulative Distribution Function**

$$F(x) = P(X \le x) = \int_{0}^{x} f(x) dx = 1 - e^{-\lambda x} \quad (x > 0)$$

## **Example**

• 
$$P(X \ge 2 | \lambda = 1.2) = 1 - P(X < 2 | \lambda = 1.2) = e^{-\lambda x} = e^{-1.2*2} = .0907$$

## Gamma Distribution

<u>Def.</u> A random variable *X* whose probability density function is defined by:

$$f(x) = \begin{cases} \frac{\lambda e^{-\lambda x} (\lambda x)^{\alpha - 1}}{\Gamma(\alpha)} & \text{if } x \ge 0\\ 0 & \text{if } x < 0 \end{cases}$$
$$(\lambda, \alpha > 0)$$

is said to be a Gamma distribution with parameters  $\alpha, \lambda$ .

The Gamma function: 
$$\Gamma(\alpha) = \int_{0}^{\infty} e^{-x} x^{\alpha-1} dx$$

When  $\alpha$  is an integer, say  $\alpha = n$ , the Gamma distribution becomes the Erlang distribution and it is easy to show, by induction, that

$$\Gamma(n) = (n-1)!$$

Chi-square ( $\chi^2$ ) Distribution

<u>Def.</u> If  $z_1, z_2, ..., z_k$  are normally and independently N(0,1) and

$$x = z_1^2 + z_2^2 + \dots + z_k^2$$

then x follows chi-square distribution with k degrees of freedom.

$$f(x) = \frac{1}{2^{\frac{k}{2}} \Gamma\left(\frac{k}{2}\right)} x^{\frac{k}{2}-1} e^{-\frac{x}{2}} \qquad (x > 0)$$

Mean:  $\mu = k$ 

Variance:  $\sigma^2 = 2k$ 

Chi-square ( $\chi^2$ ) Distribution

<u>Example</u>: If  $y_1, y_2, ..., y_n$  is a random sample from an  $N(\mu, \sigma^2)$  then

$$\frac{SS}{\sigma^2} = \frac{\sum_{i=1}^{n} (y_i - \overline{y})^2}{\sigma^2} \sim \chi_{n-1}^2$$

## Student (t) Distribution

<u>Def.</u> If z and  $\chi_k^2$  are independent standard normal and chi-square random variables then the random variable

$$t_k = \frac{z}{\sqrt{\chi_k^2/k}}$$

follows the t distribution with k degrees of freedom

$$f(x) = \frac{\Gamma\left(\frac{k+1}{2}\right)}{\sqrt{k\pi}\Gamma\left(\frac{k}{2}\right)\left[\frac{t^2}{k+1}\right]^{(k+1)/2}} \qquad (-\infty < t < +\infty) \quad \text{Mean:} \quad \mu = 0$$
Variance:  $\sigma^2 = \frac{k}{k-2} \quad \text{for } k > 2$ 

• If  $k \to \infty$ : the *t* distribution becomes the standard normal distribution

## Fisher (F) Distribution

<u>Def.</u> If  $\chi_u^2$  and  $\chi_v^2$  are independent chi-square random variables then

$$F_{u,v} = \frac{\chi_u^2/u}{\chi_v^2/v}$$

follows the F distribution with u numerator degrees of freedom and v denominator degrees of freedom

$$f(x) = \frac{\Gamma\left(\frac{u+v}{2}\right)\left(\frac{u}{v}\right)^{\frac{u}{2}} x^{\frac{u}{2}-1}}{\Gamma\left(\frac{u}{2}\right)\Gamma\left(\frac{v}{2}\right)\left[\left(\frac{u}{v}\right)x+1\right]^{\frac{u+v}{2}}} \qquad (0 < x < \infty)$$

## Fisher (F) Distribution

<u>Example</u>: If  $y_{11}, y_{12}, ..., y_{1n_1}$  and  $y_{21}, y_{22}, ..., y_{2n_2}$  are random samples from two independent normal populations with common variance  $\sigma^2$  then

$$\frac{S_1^2}{S_2^2} \sim F_{n_1 - 1, n_2 - 1}$$

where  $S_1^2$ ,  $S_2^2$  are the two sample variances.

## **FUNCTION OF RANDOM VARIABLES**

<u>Def.</u>: A function defined on values of a random variable

## **Expectation of a Function of a Random Variable**

In general,

$$E[g(X)] = \sum_{x} g(x) p(x)$$

$$E[g(X)] = \int_{x} g(x) f(x) dx$$

## **FUNCTION OF RANDOM VARIABLES**

## **Example**

1. Suppose X has the following probability mass function

$$p(0) = 0.2$$
  $p(1) = 0.5$   $p(2) = 0.3$ 

Then: 
$$E[X^2] = \sum_{x} x^2 p(x) = 0 * 0.2 + 1 * 0.5 + 4 * 0.3 = 1.7$$

2. Let X be uniformly distributed over (0,1)

Then 
$$E[X^3] = \int_0^1 x^3 f(x) dx = \int_0^1 x^3 dx = \frac{1}{4}$$

# PROPERTIES OF EXPECTATION & VARIANCE

1. 
$$E(c) = c$$

2. 
$$E(cy) = cE(y) = c\mu$$

3. 
$$V(c) = 0$$

4. 
$$V(cy) = c^2 V(y) = c^2 \sigma^2$$

5. 
$$E(y_1 + y_2) = E(y_1) + E(y_2)$$

6. 
$$Cov(y_1, y_2) = E[(y_1 - \mu_1)(y_2 - \mu_2)]$$

7. 
$$V(y_1 \pm y_2) = V(y_1) + V(y_2) \pm 2Cov(y_1, y_2)$$

8. If  $y_1, y_2$  are independent:

$$Cov(y_1, y_2) = 0$$
  
 $V(y_1 \pm y_2) = V(y_1) + V(y_2)$   
 $E(y_1, y_2) = E(y_1)E(y_2)$ 

# **MARKOV'S INEQUALTY**

If X is a random variable that takes only *nonnegative* values, then for any value a > 0

$$P\{X \ge a\} \le \frac{E[X]}{a}$$

# **CHEBYSHEV'S INEQUALTY**

If X is a random variable with mean  $\mu$  and variance  $\sigma^2$ , then, for any value K > 0

$$P(|X-\mu| \ge K) \le \frac{\sigma^2}{K^2}$$

Note that for k > 1:

$$\left| P(\mu - k\sigma < X < \mu + k\sigma) \ge 1 - \frac{1}{k^2} \right| \qquad (K = k\sigma)$$

## **EXAMPLE**

Suppose that the number of items produced in a factory during a week is a random variable with mean 500.

a. What can be said about the probability that this week's production will be at least 1000?

$$P\{X \ge 1000\} \le \frac{E[X]}{1000} = 0.5$$

b. If the variance of a week's production is 1000, then what can be said about the probability that this week's production will be between 400 and 600?

$$P(|X - 500| \ge 100) \le \frac{\sigma^2}{100^2} = 0.1 \Rightarrow P(400 \le X \le 600) \ge 0.9$$

# THE STRONG LAW OF LARGE NUMBER

Let  $X_1, X_2,...$  be independent random variables having a common distribution with mean  $\mu$ . Then, with probability 1

$$\lim_{n\to\infty} \frac{X_1 + X_2 + \dots + X_n}{n} = \mu$$

## **CENTRAL LIMIT THEOREM**

Let  $X_1, X_2,...$  be independent, identically distributed random variables with mean  $\mu$  and variance  $\sigma^2$ . Then the distribution of

$$\frac{X_1 + X_2 + \dots + X_n - n\mu}{\sigma\sqrt{n}}$$

tends to the standard normal as  $n \to \infty$ , i.e.,

$$P\left(\frac{X_1 + X_2 + \dots + X_n - n\mu}{\sigma\sqrt{n}} \le a\right) \to \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{a} e^{-x^2/2} dx$$