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Inferring dark matter substructure with astrometric lensing beyond the power spectrum

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ABSTRACT

Astrometry—the precise measurement of positions and motions of celestial objects—has emerged as a promising avenue for characterizing the dark matter population in our Galaxy. By leveraging recent advances in simulation-based inference and neural network architectures, we introduce a novel method to search for global dark matter-induced gravitational lensing signatures in astrometric datasets. Our method based on neural likelihood-ratio estimation shows significantly enhanced sensitivity to a cold dark matter population and more favorable scaling with measurement noise compared to existing approaches based on two-point correlation statistics, establishing machine learning as a powerful tool for characterizing dark matter using astrometric data.

Keywords: astrostatistics techniques (1886) — cosmology (343) — dark matter (353) — gravitational lensing (670) — convolutional neural networks (1938) — astrometry (80)

1. INTRODUCTION AND BACKGROUND

Although there exists plenty of evidence for dark matter (DM) on galactic scales and above (see Green (2021) for a recent overview), the distribution of DM clumps—17 subhalos—on sub-galactic scales is less well-understood and remains an active area of cosmological study. This distribution additionally correlates with and may provide clues about the underlying particle physics nature of dark matter (see e.g., Schutz (2020); Bode et al. (2001); Dalcanton & Hogan (2001)), highlighting its rel22 evance across multiple domains.

While more massive dark matter subhalos can be de-25 tected and studied through their association with lu-26 minous tracers such as bound stellar populations, sub- $_{27}$ halos with smaller masses $\lesssim\,10^9\,\mathrm{M}_\odot$ are not generally 28 associated with luminous matter (Fitts et al. 2017; Read ²⁹ et al. 2017), rendering their characterization challenging. 30 Gravitational effects provide one of the few avenues to 31 probe the distribution of these otherwise-invisible sub-³² halos (Buckley & Peter 2018). Gravitational lensing i. e., 33 the bending of light from a background source due to a 34 foreground mass, is one such effect and has been pro-35 posed in various incarnations as a probe of dark subha-36 los. Strong gravitational lensing, for example, has been 37 used to infer the presence of dark matter substructure 38 in galaxies outside of our own (Hezaveh et al. 2016; 39 Vegetti et al. 2010; Gilman et al. 2020; Vegetti et al. 40 2012). Astrometric lensing, on the other hand, has re-41 cently emerged as a promising way to characterize the 42 dark matter subhalo population within the Milky Way.

Astrometry refers to the precise measurement of the 44 positions and motions of luminous celestial objects like 45 stars and galaxies. Gravitational lensing of these back-46 ground objects by a moving foreground mass, such as a 47 dark matter subhalo, can imprint a characteristic, cor-48 related signal on their measured kinematics (angular ve-49 locities and/or accelerations). Van Tilburg et al. (2018) 50 introduced several methods for extracting this signature, 51 including computing convolutions of the expected lens-52 ing signal on astrometric datasets and detecting local 53 kinematic outliers. Mondino et al. (2020) applied the for-54 mer method to data from the Gaia satellite, obtaining 55 constraints on the abundance of dark compact objects 56 in the Milky Way and showcasing the applicability of 57 astrometric dark matter searches in a practical setting. 58 Finally, Mishra-Sharma et al. (2020) proposed using the 59 angular power spectrum of the astrometric field as an 60 observable to infer the population properties of subha-61 los in our Galaxy, leveraging the collective, correlated 62 signal of a large subhalo sample.

Astrometric datasets are inherently high-dimensional, consisting of positions and kinematics of potentially millions of objects. Especially when the expected signal consists of the collective imprint of a large number of lenses, characterizing their population properties involves marginalizing over all possible configurations of subhalos, rendering the likelihood intractable and usually necessitating the use of simplified data representations like the power spectrum. While effective, such simplification can result in loss of information compared to that contained in the original dataset when the ex-

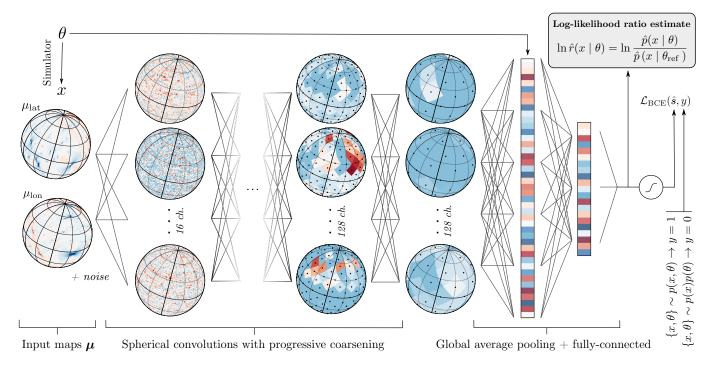


Figure 1. A schematic illustration of the method and neural network architecture used in this work.

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⁷⁴ pected signal is non-Gaussian in nature. The existence ⁷⁵ of systematic effects that are degenerate with a puta-⁷⁶ tive signal in the low-dimensional summary domain can ⁷⁷ further inhibit sensitivity.

The dawn of the era of precision astrometry, with 79 the Gaia satellite (Gaia Collaboration 2016) having re-80 cently delivered the most precise astrometric dataset 81 to-date (Gaia Collaboration 2018a; Lindegren et al. 82 2018; Gaia Collaboration 2021) and surveys includ-83 ing the Square Kilometer Array (SKA) (Fomalont & 84 Reid 2004; Jarvis et al. 2015) and Roman Space Tele-85 scope (WFIRST Astrometry Working Group 2019) set 86 to achieve further leaps in sensitivity over the next 87 decade, calls for methods that can extract more infor-88 mation from these datasets than is possible using ex-89 isting techniques. In this direction, Vattis et al. (2020) 90 proposed using a binary classifier in order to detect ei-91 ther the presence or absence of a substructure signal 92 in astrometric maps. In this paper, we introduce an in-93 ference approach that uses spherical convolutional neu-94 ral networks—exploiting the symmetry structure of the 95 signal and data domain—in conjunction with param-96 eterized classifiers (Cranmer et al. 2015; Baldi et al. 97 2016) in order to estimate likelihood ratios associated 98 with the abundance of a cold dark matter population 99 directly from a binned map of the astrometric velocity 100 field. We show that our method outperforms established proposals based on the two-point correlation statistics of 102 the astrometric field, both in absolute sensitivity as well

103 as its scaling with measurement noise. Besides the spe104 cific domain application, this paper showcases how neu105 ral network architectures suited to processing real-world
106 data structures—in our cases, pixelated vector fields of
107 velocities on the celestial sphere—can be combined with
108 advancements in simulation-based inference in order to
109 directly perform inference on complex, high-dimensional
110 datasets without having to resorting to simplified sum111 mary statistics.

2. MODEL AND INFERENCE

2.1. The forward model

We consider a population of Navarro-Frenk-White (NFW) (Navarro et al. 1996) subhalos following a power-law mass function, ${\rm d}n/{\rm d}m \propto m^{\alpha}$, with slope ${\rm d}n = -1.9$ as expected if the population is sourced from nearly scale-invariant primordial fluctuations in the canonical Λ Cold Dark Matter (Λ CDM) scenario. The concentration-mass relation from Sánchez-Conde & Prada (2014) is used to model the concentrations associated with density profiles of individual subhalos.

Subhalos between $10^7-10^{10} \,\mathrm{M_\odot}$ are simulated, assuming the influence of lighter subhalos to be too small to be discernable (Mishra-Sharma et al. 2020). The subhalo fraction f_{sub} , quantifying the expected fraction of the mass of the Milky Way contributed by subhalos in the range $10^{-6}-10^{10} \,\mathrm{M_\odot}$, is taken to be the parameter of interest. The spatial distribution of subhalos in the Galactocentric frame is modeled using results from

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the Aquarius simulation following Hütten et al. (2016); Springel et al. (2008). Since this spatial distribution accounts for the depletion of subhalos towards the Galactic Center due to gravitational tidal effects, the angular number density of subhalos looking out from the Sun frame can be considered to be effectively isotropic.

The asymptotic velocities of subhalos in the Galacto-138 centric frame are taken to follow a truncated Maxwell-139 Boltzmann distribution (Chandrasekhar 1939; Lisanti $_{140}$ 2017) $f_{\rm Gal}(\mathbf{v}) \propto e^{-\mathbf{v}^2/v_0^2} \cdot H(v_{\rm esc} - |\mathbf{v}|)$, where $v_{\rm esc} = 141$ 550 km s⁻¹ is the Galactic escape velocity (Piffl et al. $v_0 = 220 \,\mathrm{km \, s^{-1}}$ (Kerr & Lynden-Bell 1986), and H is the Heaviside step function. Once instanti-144 ated, the positions and velocities of subhalos are transformed into the Galactic frame, assuming $R_{\odot} = 8.2 \,\mathrm{kpc}$ 146 to be the distance of the Sun from the Galactic Center (Gravity Collaboration 2019; Bovy 2020) and $\mathbf{v}_{\odot} =$ 148 (11, 232, 7) km s⁻¹ its Galactocentric velocity (Schönrich 149 et al. 2010). Note that the asymmetry in the direction 150 of motion of the Sun in the Milky Way introduces a pre-151 ferred direction for the Sun-frame velocities of subhalos, breaking strict rotation invariance in the forward model. 153 Although not explicitly pursued here, this asymmetry can be used as an additional distinguishing handle for 155 the lensing signal, as was done in Mishra-Sharma et al. 156 (2020).

Our datasets consist of the 2-dimensional angular velocity map of background sources on the celestial sphere. Given a spherically-symmetric subhalo lens moving with transverse velocity \mathbf{v}_l , the expected lens-induced velocity for a quasar at impact parameter \mathbf{b} is given by (Van Tilburg et al. 2018)

$$\boldsymbol{\mu}(\mathbf{b}) = 4G_{\mathrm{N}} \left\{ \frac{M(b)}{b^{2}} \left[2\hat{\mathbf{b}} \left(\hat{\mathbf{b}} \cdot \mathbf{v}_{l} \right) - \mathbf{v}_{l} \right] - \frac{M'(b)}{b} \hat{\mathbf{b}} \left(\hat{\mathbf{b}} \cdot \mathbf{v}_{l} \right) \right\}$$
(1)

where M(b) and M'(b) are the projected mass of the subhalo at a given impact parameter distance $b = |\mathbf{b}|$ and its gradient. An example of the induced velocity signal on part of the celestial sphere, projected along the Galactic latitudinal and longitudinal directions and exhibiting dipole-like structures, is shown in the leftmost column of Fig. 1.

We take our source population to consist of remote, point-like galaxies known as quasars which, due to their large distances from the Earth, are not expected to have significant intrinsic angular velocities. We assume the sources to be isotropically-distributed, although this assumption can be easily relaxed for a realistic source sample. The velocity maps are assumed to be spatially binned, and we use a HEALPix binning (Gorski et al. 2005) with resolution parameter nside=64, corresponding to $N_{pix}=49,152$ pixels over the full sky with pixel area $\sim 0.8 \ \text{deg}^2$. The values within each pixel then quantized.

182 tify the average latitudinal and longitudinal velocity 183 components of quasars within that pixel, with the im-184 pact parameter **b** representing the vector from the center 185 of a subhalo to the center of the pixel.

In order to enable a comparison with traditional approaches—which are generally not expected to be sensitive to a cold dark matter subhalo population with next-generation astrometric surveys (Van Tilburg et al. 2018; Mishra-Sharma et al. 2020)—we benchmark using an optimistic observational configuration corresponding to measuring the proper motions of $N_q = 10^8$ quasars with noise $\sigma_{\mu} = 0.1~\mu{\rm as\,yr}^{-1}$.

2.2. The power spectrum approach

Mishra-Sharma et al. (2020) introduced an approach 195 196 for extracting the astrometric signal due to a dark mat-197 ter subhalo population by decomposing the observed 198 map into its angular (vector) power spectrum. The 199 power spectrum is a summary statistic ubiquitous in as-200 trophysics and cosmology and quantifies the amount of 201 correlation contained at different spatial scales. In the 202 case of data on a sphere, the basis of spherical harmon-203 ics is often used, and the power spectrum then encodes 204 the correlation structure on different multipoles ℓ . The 205 power spectrum effectively captures the linear compo-206 nent of the signal and, when the underlying signal is 207 a Gaussian random field, captures all of the relevant 208 information contained in the map(s) (Tegmark 1997). 209 The expected signal in the power spectrum domain can 210 be evaluated semi-analytically using the formalism de-211 scribed in Mishra-Sharma et al. (2020) and, assuming 212 a Gaussian likelihood, the expected sensitivity can be 213 computed using a Fisher forecasting approach. We use 214 this prescription as a comparison point to the method 215 introduced here.

While effective, reduction of the full astrometric map to its power spectrum results in loss of information; this can be seen from the fact that the signal in the leftmost column of Fig. 1 is far from Gaussian. Furthermore, the existence of correlations on large angular scales due to e.g., biases in calibration of celestial reference frames (Gaia Collaboration 2018b) or systematic variations in measurements taken over different regions of the sky introduces degeneracies with a putative signal and precludes their usage in the present context. For this reason multipoles $\ell < 10$ were discarded in Mishra-Sharma et al. (2020), degrading the projected sensitivity.

2.3. Likelihood-ratio estimation using parameterized classifiers

Recent advances in machine learning have enabled methods that can be used to efficiently perform inference on models defined through complex simulations;

236 2018b; Hermans et al. 2019), previously applied to the 237 problem of inferring dark matter substructure using ob-238 servations of strong gravitational lenses (Brehmer et al. 2019) and cold stellar streams (Hermans et al. 2020). Given a classifier that can distinguish between samples $\{x\} \sim p(x \mid \theta)$ drawn from parameter points θ and those from a fixed reference hypothesis $\{x\} \sim p(x \mid \theta_{\text{ref}})$, 243 the decision function output by the optimal classifier $s(x,\theta) = p(x \mid \theta)/(p(x \mid \theta) + p(x \mid \theta_{ref}))$ is one-to-one with the likelihood ratio, $r(x \mid \theta) \equiv p(x \mid \theta)/p(x \mid \theta_{ref}) =$ $s(x,\theta)/(1-s(x,\theta))$, a fact appreciated as the likelihood-²⁴⁷ ratio trick (Cranmer et al. 2015; Mohamed & Lakshmi-248 narayanan 2017). The classifier $s(x,\theta)$ in this case is neural network that can work directly on the highdimensional data x, and is parameterized by θ by having included as an input feature. In order to improve nu-²⁵² merical stability and reduce dependence on the fixed reference hypothesis $\theta_{\rm ref}$, we follow Hermans et al. (2019) 254 and train a classifier to distinguish between data-sample 255 pairs from the joint distribution $\{x,\theta\} \sim p(x,\theta)$ and those from a product of marginal distributions $\{x, \theta\}$ $p(x)p(\theta)$ (defining the reference hypothesis and in prac-258 tice obtained by shuffling samples within a batch) using 259 the binary cross-entropy (BCE) loss as the optimization 260 objective.

233 see Cranmer et al. (2020) for a recent review. Here, we

235 et al. 2015; Baldi et al. 2016; Brehmer et al. 2018a, 2020,

make use of neural likelihood-ratio estimation (Cranmer

2.4. Extracting information from high-dimensional astrometric maps

Since our dataset consists of a velocity field sampled 263 on a sphere, we use a spherical convolutional neural network in order to directly learn useful representations 266 from these maps that are efficiently suited for the downstream classification task. Specifically, we make use of ²⁶⁸ DeepSphere (Defferrard et al. 2020; Perraudin et al. 269 2019), a graph-based convolutional neural network tai-270 lored to data sampled on a sphere. For this purpose, the HEALPix grid can be cast as a weighted undirected $_{272}$ graph with $N_{
m pix}$ vertices and edges connecting each pixel 273 vertex to its set of 8 neighboring pixels. The weighted 274 adjacency matrix over neighboring pixels (i, j) is given ₂₇₅ by $A_{ij}=\exp\left(-\Delta r_{ij}^2/\rho^2\right)$ where Δr_{ij} specifies the 3-276 dimensional Euclidean distance between the pixel centers and the widths ρ are obtained from Defferrard et al. 278 (2020). DeepSphere then efficiently performs convolu-279 tions in the spectral domain using a basis of Chebychev 280 polynomials as convolutional kernels (Defferrard et al. 281 2016); here, we set K=4 as the maximum polynomial 282 order.

All inputs are normalized to zero mean and unit standard deviation across the training sample. Starting with 285 2 scalar input channels representing the two orthogo-286 nal (Galactic latitude and longitude) components of the ²⁸⁷ velocity vector map, ¹ we perform a graph convolution 288 operation, increasing the channel dimension to 16 fol-289 lowed by a batch normalization, ReLU nonlinearity, and 290 downsampling the representation by a factor of 4 with 291 max pooling into the next coarser HEALPix resolution. 292 Pooling leverages scale separation, preserving important 293 characteristics of the signal across different resolutions. ²⁹⁴ Four more such layers are employed, increasing the chan-295 nel dimension by a factor of 2 at each step until a maxi-²⁹⁶ mum of 128, with maps after the last convolutional layer 297 having resolution nside=2 corresponding to 48 pixels. 298 At this stage, we average over the spatial dimension 299 (known as global average pooling (Lin et al. 2014)) in or-300 der to encourage approximate rotation invariance, out-301 putting 128 features onto which the parameter of inter- $_{302}$ est $f_{\rm sub}$ is appended. These features are passed through 303 a fully-connected network with (1024, 128) hidden units 304 and ReLU activations outputting the classifier decision \hat{s} by applying a sigmoidal projection.

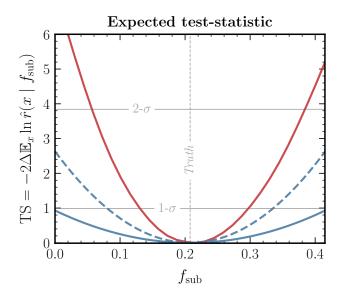
2.5. Model training and evaluation

10⁵ maps from the forward model were produced, with 15% of these held out for validation. The estimator was trained using a batch size of 64 for up to 50 epochs with early stopping if the validation loss had not improved after 10 epochs. The ADAM optimizer (Kingma & Ba 2017) was used with initial learning rate 10⁻³ decayed through cosine annealing. A coarse grid search was used to inform the architecture and hyperparameter choices in this work.

For a given test map, the log-likelihood ratio profile can be obtained by evaluating the trained estimator for different values of $f_{\rm sub}$ while keeping the input map fixed. The network output prior to the final sigmoidal projection directly gives the required log-likelihood ratio estimate: $\ln \hat{r} = S^{-1}(\hat{s})$, where S is the sigmoid function (Hermans et al. 2019, 2020). Figure 1 presents an illustrative summary of the neural network architecture and method used in this work.

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We note that by representing the input angular velocity vector field in terms of two input scalar channels, we break the desired rotation equivariance of spherical convolutions due to differences in how scalar and vector representations transform under rotations. Although this will have a downstream effect on rotation invariance, a detailed study of how this influences the performance of our model is beyond the scope of this paper.



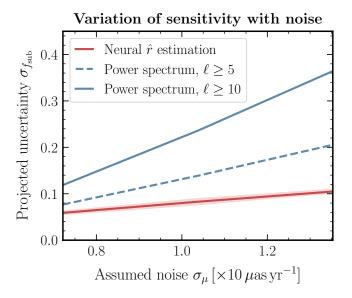


Figure 2. (Left) The expected log-likelihood ratio test-statistic (TS) profile for a cold dark matter population as a function of substructure fraction f_{sub} obtained using the neural likelihood-ratio estimation method introduced in this work (red line) compared with the corresponding profiles for existing approaches using power spectrum summaries with different multipole thresholds $\ell \gtrsim 5$ (dashed blue line) and $\ell \gtrsim 10$ (solid blue line). The vertical dotted line indicates the true benchmark value of the parameter $f_{\rm sub}$ in the test dataset. Our method shows enhanced sensitivity to a cold dark matter population compared to traditional approaches. (Right) Scaling of the expected sensitivities, quantified by the respective 1- σ uncertainties, with perobject instrumental noise. For the machine learning-based approach, the band quantifies the middle-95% containment of the inferred $1-\sigma$ uncertainty. Our method shows a more favorable scaling with assumed measurement noise.

3. EXPERIMENTS ON SIMULATED DATA

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3.1. Baseline results and diagnostics

We evaluate our trained likelihood-ratio estimator on maps drawn from a benchmark configuration motivated by Hütten et al. (2016); Springel et al. (2008), $_{330}$ containing 150 subhalos in expectation between 10^{8} – $10^{10} \,\mathrm{M}_{\odot}$ and corresponding to $f_{\mathrm{sub}} \simeq 0.2$. The left panel 332 of Fig. 2 shows the expected log-likelihood ratio teststatistic (TS) as a function of substructure fraction f_{sub} 334 for this nominal configuration. This is obtained by eval-335 uating the trained estimator on 100 test maps over a uni $f_{
m sub}$ form grid in $f_{
m sub}$ and taking the point-wise mean. Cor-337 responding curves using the power spectrum approach 338 are shown in blue, using minimum multipoles of $\ell \geq 5$ (dashed) and $\ell \geq 10$ (solid). Thresholds corresponding to 1- and 2- σ significance assuming a χ^2 -distributed TS 341 are shown as the horizontal grey lines. We see that sensitivity gains of over a factor of ~ 2 can be expected for this particular benchmark when using the machine 344 learning approach compared to the traditional power 345 spectrum approach. No significant bias on the central value of the inferred DM abundance relative to the overall uncertainty scale is observed.

The right panel of Fig. 2 shows the scaling of expected 349 1- σ uncertainty on substructure fraction $f_{\rm sub}$ with as-350 sumed noise per quasar, keeping the number of quasars

351 fixed (red, with the line showing the median and shaded 352 band corresponding to the middle-95% containment of 353 the uncertainty inferred over 50 test datasets) compared 354 to the power spectrum approach (blue lines). A far more 355 favorable scaling of the machine learning approach is 356 seen compared to the power spectrum approach, sug-357 gesting that it is especially advantageous in low signal-358 to-noise regimes that are generally most relevant for 359 dark matter searches.

Finally, we assess the quality of the approximate 361 likelihood-ratio estimator through a test of statistical 362 coverage. Within a hypothesis testing framework, this 363 is necessary in order to ensure that the learned esti-364 mator is conservative over the parameter range of in-365 terest and does not produced overly confident or bi-366 ased results (Hermans et al. 2021). We obtain the es-367 timated TS profile for 1000 simulated samples with 368 true substructure fraction values drawn from the range $f_{\text{sub}} \in [0.1, 0.3]$. In doing so, we exclude parameter 370 points towards the edges of our parameter space since 371 the corresponding confidence intervals in these cases 372 would extend outside of the tested parameter range, as 373 can also be inferred from the baseline analysis shown 374 in Fig. 2. For nominal confidence levels in the range $_{375}$ $1-\alpha \in [0.05, 0.95]$ we compute the empirical coverage 376 over the set of samples, defined as the fraction of samples 377 whose true parameter value falls within the TS confi-

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 378 dence interval, computed for a given confidence level un- 379 der the assumption that the TS is χ^2 -distributed (Wilks 380 1938). The procedure is repeated for 10 different sets of 381 1000 samples in order to estimate the statistical uncer- 382 tainty associated with the empirical coverage.

The results of the coverage test are shown in Fig. 3, illustrating the median (solid red) and middle-68% constainment (red band) of the empirical coverage. We see that the empirical coverage has the desired property of being slightly conservative while still being close to the perfectly-calibrated regime indicated by the dashed-grey line. We emphasize that this diagnostic tests the quality of the likelihood-ratio estimator over the entire evaluation parameter range of interest $f_{\rm sub} \in [0.1, 0.3]$ rather than the baseline value $f_{\rm sub} \simeq 0.2$ in isolation.

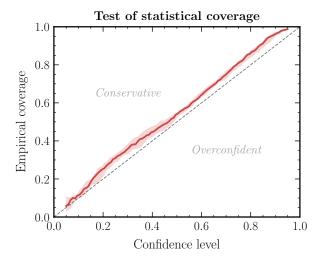


Figure 3. Calibration test

3.2. Experiments with unmodeled noise correlated on large scales

Since the existence of measurement noise correlated 395 on large spatial scales is a potential source of systematic uncertainty when working with astrometric maps, we test the susceptibility of our method to such ef-399 fects by creating simulated data containing large-scale 400 noise not previously seen by the trained estimator. Instead of assuming a scale-invariant noise power spec-402 trum $C_\ell^{\rm noise} = 4\pi\sigma_\mu^2/N_q$ (Mishra-Sharma et al. 2020), 403 in this case we model noise with an order of magni-404 tude excess in power on scales $\ell \lesssim$ 10, parameterized as $C_{\ell}^{\text{noise}} = 4\pi\sigma_{\mu}^2/N_q \cdot (10 - 9S(\ell - 10))$ where S de-406 notes the sigmoid function. The left panel of Fig. 4 illus-407 trates this noise model (thicker green line) as well as the 408 power spectrum of one simulated realization from this 409 model (thinner green line, obtained using the HEALPix module anafast) contrasted with the standard scalemodule anafast). The right panel of Fig. 4
shows the expected log-likelihood ratio test-statistic promodule for the two cases. Although a bias in the maximummodule for the two cases.

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3.3. Experiments with a data-driven noise model

We finally assess the real-world performance of our 421 422 method through a realistic, data-drive noise model ob-423 tained using the astrometric catalog of quasars in Gaia's 424 second data release (DR2). The catalog contains the 425 measured 2-dimensional positions, proper motions, as 426 well as proper motion uncertainties of 555,934 guasars. 427 We compute the pixel-wise proper motion uncertain-428 ties as the inverse-variance weighted values within each 429 HEALPix pixel; $\sigma_{\mu}^{\text{pix}} = \left(\sum_{q \in \text{pix}} \sigma_{\mu,q}^{-2}\right)^{-1/2}$, where $\sigma_{\mu,q}^2$ 430 are the provided variances of individual quasars in a 431 given pixel. This results in a highly anisotropic noise 432 model, shown in the left column of Fig. 5, additionally 433 having different uncertainties in the latitudinal and lon-434 gitudinal directions. As expected due to occlusion from 435 the Galactic disk, uncertainties are significantly higher 436 towards the Galactic plane, and the region closest to 437 the plane where no quasars are included in the catalog 438 (shown in grey) is masked, additionally testing the effect 439 of partial sky coverage.

4. CONCLUSIONS AND OUTLOOK

We have introduced a method to analyze astrometric 442 datasets over large regions of the sky using techniques 443 based on machine learning with the aim of inferring the 444 lensing signature of a dark matter substructure. We have 445 shown our method to be significantly more sensitive to 446 a cold dark matter subhalo population compared to es-447 tablished methods based on global summary statistics, 448 with more favorable scaling as a function of measure-449 ment noise. Since the collection and reduction of astro-450 metric data is an expensive endeavor, the use of methods 451 that can take advantage of more of the available infor-452 mation can be equated to long periods of data-taking, 453 underscoring their importance. Additionally, unlike the 454 power spectrum approach, the current method does not 455 require the construction of a numerically-expensive esti-456 mator to account for non-uniform exposure, selection ef-457 fects, and instrumental noise in realistic datasets. These, 458 as well as any other modeled observational effects, can be 459 incorporated directly at the level of the forward model.

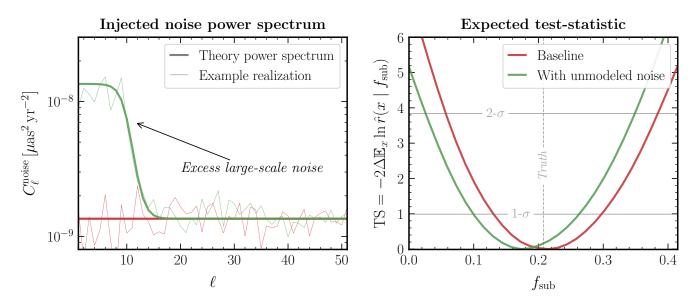


Figure 4. (Left) The power spectrum of the noise model (thicker green line) used to study the impact of correlated noise on large spatial scales, not modeled during training, on the performance of the likelihood-ratio estimator. The thinner green line shows the power spectrum of an example noise realization instantiated from this noise model. The red lines show corresponding power spectra for a scale-invariant noise model. (Right) The expected test-statistic profile for a model evaluated on maps containing excess large-scale noise (green line) compared to the model evaluated on maps with scale-invariant noise (red line). A bias in the maximum-likelihood estimate returned by the model is seen when substantial unaccounted-for noise is presented in the test maps.

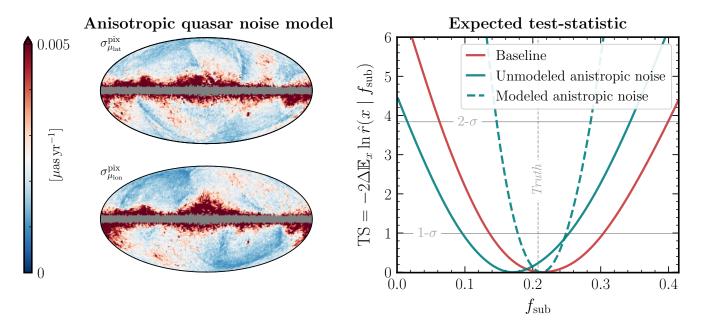


Figure 5. (Left) (Right)

We have focused in this work on assessing sensitiv-461 ity to a cold dark matter-like subhalo population with 462 quasar velocity astrometry, which is within the scope of 463 upcoming radio surveys like the SKA (Fomalont & Reid 464 2004; Jarvis et al. 2015). Our method can also be applied 465 in a straightforward manner to look for the acceleration 466 lensing signal imprinted on Milky Way stars, in partic-467 ular sourced by a population of more compact subhalos 468 than those expected in the cold dark matter scenario. 469 These features are expected to imprint a larger degree 470 of non-Gaussianity compared to the signal explored here 471 (as can be seen, e.g., from Fig. 1 of Mishra-Sharma 472 et al. (2020)), and machine learning methods may pro-473 vide larger relative sensitivity gains when deployed in 474 that context. Such analyses are within purview of the upcoming Roman exoplanet microlensing survey (Pardo 476 & Doré 2021) as well as future Gaia data releases.

Several improvements and extensions to the method presented in this paper are possible. The use of architectures that can equivariantly handle vector inputs (Esteves et al. 2020) can aid in learning more efficient representations of the astrometric map. Using convolutions based on fixed rather than learned filters can additionally reduce model complexity and produce more interpretable representations (Cheng et al. 2020; Ha et al. 2021; Saydjari & Finkbeiner 2021; McEwen et al. 2021; Valogiannis & Dvorkin 2021). The use of methods for likelihood-ratio estimation that can leverage additional

the study of these extensions as well as application of the supplication of the supplication of the study of these extensions as well as application of ur method to other dark matter population scenarios to future work.

Astrometric lensing has been established as a promising way to characterize the Galactic dark matter population, with theoretical progress in recent years going in step with advances on the observational front. While this work is a first attempt at bringing principled machine learning techniques to this field, with the availability of increasingly complex datasets we expect machine learning to be an important general-purpose tool for future strometric dark matter searches.

(Acknowledgments anonymized for review)

Software: Astropy (Robitaille et al. 2013; PriceWhelan et al. 2018), healpy (Gorski et al. 2005;
Donca et al. 2019), IPython (Pérez & Granger 2007),
Jupyter (Kluyver et al. 2016), Matplotlib (Hunter 2007),
MLflow (Chen et al. 2020), NumPy (Harris et al. 2020), PyGSP (Defferrard et al. 2017), PyTorch (Paszke et al. 2019), PyTorch Geometric (Fey & Lenssen 2019),
PyTorch Lightning (Falcon et al. 2020), sbi (TejeroCantero et al. 2020), SciPy (Virtanen et al. 2020), and
Seaborn (Waskom et al. 2017).

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