10

11

12

13

#### Inferring dark matter substructure with astrometric lensing beyond the power spectrum

Anonymous author(s)

ABSTRACT

Astrometry—the precise measurement of positions and motions of celestial objects—has emerged as a promising avenue for characterizing the dark matter population in our Galaxy. By leveraging recent advances in simulation-based inference and neural network architectures, we introduce a novel method to search for global dark matter-induced gravitational lensing signatures in astrometric datasets. Our method based on neural likelihood-ratio estimation shows significantly enhanced sensitivity to a cold dark matter population and more favorable scaling with measurement noise compared to existing approaches based on two-point correlation statistics, establishing machine learning as a powerful tool for characterizing dark matter using astrometric data.

Keywords: astrostatistics techniques (1886) — cosmology (343) — dark matter (353) — gravitational lensing (670) — convolutional neural networks (1938) — astrometry (80)

#### 1. INTRODUCTION AND BACKGROUND

Although there exists plenty of evidence for dark matter (DM) on galactic scales and above (see Green (2021) for a recent overview), the distribution of DM clumps—17 subhalos—on sub-galactic scales is less well-understood and remains an active area of cosmological study. This distribution additionally correlates with and may provide clues about the underlying particle physics nature of dark matter (see e.g., Schutz (2020); Bode et al. (2001); Dalcanton & Hogan (2001)), highlighting its rel22 evance across multiple domains.

While more massive dark matter subhalos can be de-25 tected and studied through their association with lu-26 minous tracers such as bound stellar populations, sub- $_{27}$  halos with smaller masses  $\lesssim\,10^9\,\mathrm{M}_\odot$  are not generally 28 associated with luminous matter (Fitts et al. 2017; Read <sup>29</sup> et al. 2017), rendering their characterization challenging. 30 Gravitational effects provide one of the few avenues to 31 probe the distribution of these otherwise-invisible sub-<sup>32</sup> halos (Buckley & Peter 2018). Gravitational lensing i. e., 33 the bending of light from a background source due to a 34 foreground mass, is one such effect and has been pro-35 posed in various incarnations as a probe of dark subha-36 los. Strong gravitational lensing, for example, has been 37 used to infer the presence of dark matter substructure 38 in galaxies outside of our own (Hezaveh et al. 2016; 39 Vegetti et al. 2010; Gilman et al. 2020; Vegetti et al. 40 2012). Astrometric lensing, on the other hand, has re-41 cently emerged as a promising way to characterize the 42 dark matter subhalo population within the Milky Way.

Astrometry refers to the precise measurement of the 44 positions and motions of luminous celestial objects like 45 stars and galaxies. Gravitational lensing of these back-46 ground objects by a moving foreground mass, such as a 47 dark matter subhalo, can imprint a characteristic, cor-48 related signal on their measured kinematics (angular ve-49 locities and/or accelerations). Van Tilburg et al. (2018) 50 introduced several methods for extracting this signature, 51 including computing convolutions of the expected lens-52 ing signal on astrometric datasets and detecting local 53 kinematic outliers. Mondino et al. (2020) applied the for-54 mer method to data from the Gaia satellite, obtaining 55 constraints on the abundance of dark compact objects 56 in the Milky Way and showcasing the applicability of 57 astrometric dark matter searches in a practical setting. 58 Finally, Mishra-Sharma et al. (2020) proposed using the 59 angular power spectrum of the astrometric field as an 60 observable to infer the population properties of subha-61 los in our Galaxy, leveraging the collective, correlated 62 signal of a large subhalo sample.

Astrometric datasets are inherently high-dimensional, consisting of positions and kinematics of potentially millions of objects. Especially when the expected signal consists of the collective imprint of a large number of lenses, characterizing their population properties involves marginalizing over all possible configurations of subhalos, rendering the likelihood intractable and usually necessitating the use of simplified data representations like the power spectrum. While effective, such simplification can result in loss of information compared to that contained in the original dataset when the ex-

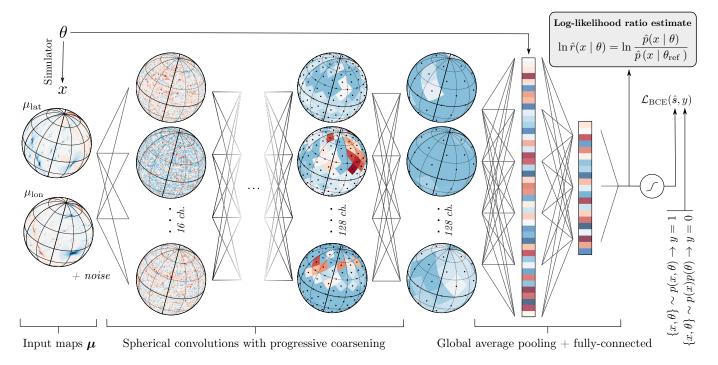


Figure 1. A schematic illustration of the method and neural network architecture used in this work.

105

<sup>74</sup> pected signal is non-Gaussian in nature. The existence <sup>75</sup> of systematic effects that are degenerate with a puta-<sup>76</sup> tive signal in the low-dimensional summary domain can <sup>77</sup> further inhibit sensitivity.

The dawn of the era of precision astrometry, with 79 the Gaia satellite (Gaia Collaboration 2016) having re-80 cently delivered the most precise astrometric dataset 81 to-date (Gaia Collaboration 2018a; Lindegren et al. 82 2018; Gaia Collaboration 2021) and surveys includ-83 ing the Square Kilometer Array (SKA) (Fomalont & 84 Reid 2004; Jarvis et al. 2015) and Roman Space Tele-85 scope (WFIRST Astrometry Working Group 2019) set 86 to achieve further leaps in sensitivity over the next 87 decade, calls for methods that can extract more infor-88 mation from these datasets than is possible using ex-89 isting techniques. In this direction, Vattis et al. (2020) 90 proposed using a binary classifier in order to detect ei-91 ther the presence or absence of a substructure signal in 92 astrometric maps. In this paper, we introduce an infer-93 ence approach that uses spherical convolutional neural 94 networks—exploiting the symmetry structure of the sig-95 nal and data domain—in conjunction with parameter-96 ized classifiers (Cranmer et al. 2015; Baldi et al. 2016) 97 in order to estimate likelihood ratios associated with 98 the abundance of a cold dark matter population directly 99 from a binned map of the astrometric velocity field. We 100 show that our method outperforms established propos-101 als based on the two-point correlation statistics of the

102 astrometric field, both in absolute sensitivity as well as 103 its scaling with measurement noise.

#### 2. MODEL AND INFERENCE

## $2.1. \ \ The \ forward \ model$

We consider a population of Navarro-Frenk-White (NFW) (Navarro et al. 1996) subhalos following a power-law mass function,  ${\rm d}n/{\rm d}m \propto m^{\alpha}$ , with slope  $\alpha=-1.9$  as expected if the population is sourced from nearly scale-invariant primordial fluctuations in the canonical  $\Lambda$  Cold Dark Matter ( $\Lambda$ CDM) scenario. The concentration-mass relation from Sánchez-Conde & Prada (2014) is used to model the concentrations associated with density profiles of individual subhalos.

Subhalos between  $10^7-10^{10} \,\mathrm{M}_{\odot}$  are simulated, assuming the influence of lighter subhalos to be too small to be discernable (Mishra-Sharma et al. 2020). The subhalo fraction  $f_{\mathrm{sub}}$ , quantifying the expected fraction of the mass of the Milky Way contributed by subhalos in the range  $10^{-6}-10^{10} \,\mathrm{M}_{\odot}$ , is taken to be the parameter of interest. The spatial distribution of subhalos in the Galactocentric frame is modeled using results from the Aquarius simulation following Hütten et al. (2016); Springel et al. (2008). Since this spatial distribution actounts for the depletion of subhalos towards the Galactic Center due to gravitational tidal effects, the angular number density of subhalos looking out from the Sun frame can be considered to be effectively isotropic.

220

221

The asymptotic velocities of subhalos in the Galacto-130 centric frame are taken to follow a truncated Maxwell-131 Boltzmann distribution (Chandrasekhar 1939; Lisanti <sub>132</sub> 2017)  $f_{\text{Gal}}(\mathbf{v}) \propto e^{-\mathbf{v}^2/v_0^2} \cdot H(v_{\text{esc}} - |\mathbf{v}|)$ , where  $v_{\text{esc}} =$  $_{133}$  550 km s<sup>-1</sup> is the Galactic escape velocity (Piffl et al.  $v_0 = 220 \,\mathrm{km}\,\mathrm{s}^{-1}$  (Kerr & Lynden-Bell 1986), and H is the Heaviside step function. Once instanti-136 ated, the positions and velocities of subhalos are transformed into the Galactic frame, assuming  $R_{\odot} = 8.2 \,\mathrm{kpc}$ 138 to be the distance of the Sun from the Galactic Cen-139 ter (Gravity Collaboration 2019; Bovy 2020) and  $\mathbf{v}_{\odot} =$ (11, 232, 7) km s<sup>-1</sup> its Galactocentric velocity (Schönrich et al. 2010). Note that the asymmetry in the direction of motion of the Sun in the Milky Way introduces a pre-143 ferred direction for the Sun-frame velocities of subhalos, breaking strict rotation invariance in the forward model. 145 Although not explicitly pursued here, this asymmetry 146 can be used as an additional distinguishing handle for 147 the lensing signal, as was done in Mishra-Sharma et al. (2020).148

Our datasets consist of the 2-dimensional angular velocity map of background sources on the celestial sphere. Given a spherically-symmetric subhalo lens moving with transverse velocity  $\mathbf{v}_l$ , the expected lens-induced velocity for a quasar at impact parameter  $\mathbf{b}$  is given by (Van Tilburg et al. 2018)

$$\mu(\mathbf{b}) = 4G_{N} \left\{ \frac{M(b)}{b^{2}} \left[ 2\hat{\mathbf{b}} \left( \hat{\mathbf{b}} \cdot \mathbf{v}_{l} \right) - \mathbf{v}_{l} \right] - \frac{M'(b)}{b} \hat{\mathbf{b}} \left( \hat{\mathbf{b}} \cdot \mathbf{v}_{l} \right) \right\}$$
(1)

where M(b) and M'(b) are the projected mass of the subhalo at a given impact parameter distance  $b = |\mathbf{b}|$  and its gradient. An example of the induced velocity signal on part of the celestial sphere, projected along the Galactic latitudinal and longitudinal directions and exhibiting dipole-like structures, is shown in the leftmost column of Fig. 1.

We take our source population to consist of remote, point-like galaxies known as quasars which, due to their large distances from the Earth, are not expected to have significant intrinsic angular velocities. We assume the sources to be isotropically-distributed, although this assumption can be easily relaxed for a realistic source sample. The velocity maps are assumed to be spatially binned, and we use a HEALPix binning (Gorski et al. 2005) with resolution parameter nside=64, corresponding to  $N_{pix}=49{,}152$  pixels over the full sky with pixel area  $\sim 0.8 \ deg^2$ . The values within each pixel then quantify the average latitudinal and longitudinal velocity components of quasars within that pixel, with the impact parameter  $\bf b$  representing the vector from the center of a subhalo to the center of the pixel.

In order to enable a comparison with traditional approaches—which are generally not expected to be sen-

180 sitive to a cold dark matter subhalo population with 181 next-generation astrometric surveys (Van Tilburg et al. 2018; Mishra-Sharma et al. 2020)—we benchmark using 183 an optimistic observational configuration corresponding to measuring the proper motions of  $N_q=10^8$  quasars with noise  $\sigma_\mu=0.1~\mu{\rm as\,yr}^{-1}$ .

## 2.2. The power spectrum approach

Mishra-Sharma et al. (2020) introduced an approach 188 for extracting the astrometric signal due to a dark mat-189 ter subhalo population by decomposing the observed 190 map into its angular (vector) power spectrum. The 191 power spectrum is a summary statistic ubiquitous in as-192 trophysics and cosmology and quantifies the amount of 193 correlation contained at different spatial scales. In the 194 case of data on a sphere, the basis of spherical harmon-195 ics is often used, and the power spectrum then encodes 196 the correlation structure on different multipoles  $\ell$ . The 197 power spectrum effectively captures the linear compo-198 nent of the signal and, when the underlying signal is 199 a Gaussian random field, captures all of the relevant 200 information contained in the map(s) (Tegmark 1997). 201 The expected signal in the power spectrum domain can 202 be evaluated semi-analytically using the formalism de-203 scribed in Mishra-Sharma et al. (2020) and, assuming 204 a Gaussian likelihood, the expected sensitivity can be 205 computed using a Fisher forecasting approach. We use 206 this prescription as a comparison point to the method 207 introduced here.

While effective, reduction of the full astrometric map to its power spectrum results in loss of information; this can be seen from the fact that the signal in the leftmost column of Fig. 1 is far from Gaussian. Furthermore, the existence of correlations on large angular scales due to e.g., biases in calibration of celestial reference frames (Gaia Collaboration 2018b) or systematic variations in measurements taken over different regions of the sky introduces degeneracies with a putative signal and precludes their usage in the present context. For this reason multipoles  $\ell < 10$  were discarded in Mishra-Sharma et al. (2020), degrading the projected sensitivity.

# 2.3. Likelihood-ratio estimation using parameterized classifiers

Recent advances in machine learning have enabled methods that can be used to efficiently perform inference on models defined through complex simulations; see Cranmer et al. (2020) for a recent review. Here, we make use of neural likelihood-ratio estimation (Cranmer et al. 2015; Baldi et al. 2016; Brehmer et al. 2018a, 2020, 2018b; Hermans et al. 2019), previously applied to the problem of inferring dark matter substructure using ob-

317

318

252 objective.

254

2019) and cold stellar streams (Hermans et al. 2020). Given a classifier that can distinguish between sam-233 ples  $\{x\} \sim p(x \mid \theta)$  drawn from parameter points  $\theta$  and those from a fixed reference hypothesis  $\{x\} \sim p(x \mid \theta_{ref})$ , 235 the decision function output by the optimal classifier 236  $s(x,\theta) = p(x \mid \theta)/(p(x \mid \theta) + p(x \mid \theta_{ref}))$  is one-to-one with the likelihood ratio,  $r(x \mid \theta) \equiv p(x \mid \theta)/p(x \mid \theta_{ref}) =$ 238  $s(x,\theta)/(1-s(x,\theta))$ , a fact appreciated as the likelihood-239 ratio trick (Cranmer et al. 2015; Mohamed & Lakshmi-240 narayanan 2017). The classifier  $s(x,\theta)$  in this case is neural network that can work directly on the highdimensional data x, and is parameterized by  $\theta$  by having included as an input feature. In order to improve nu-<sup>244</sup> merical stability and reduce dependence on the fixed ref-245 erence hypothesis  $\theta_{\rm ref}$ , we follow Hermans et al. (2019) 246 and train a classifier to distinguish between data-sample pairs from the joint distribution  $\{x,\theta\} \sim p(x,\theta)$  and those from a product of marginal distributions  $\{x,\theta\}$ 

230 servations of strong gravitational lenses (Brehmer et al.

# 2.4. Extracting information from high-dimensional astrometric maps

 $p(x)p(\theta)$  (defining the reference hypothesis and in prac-

250 tice obtained by shuffling samples within a batch) using

251 the binary cross-entropy (BCE) loss as the optimization

Since our dataset consists of a velocity field sampled 255 256 on a sphere, we use a spherical convolutional neural network in order to directly learn useful representations 258 from these maps that are efficiently suited for the down-259 stream classification task. Specifically, we make use of <sup>260</sup> DeepSphere (Defferrard et al. 2020; Perraudin et al. 2019), a graph-based convolutional neural network tai-262 lored to data sampled on a sphere. For this purpose, the HEALPix grid can be cast as a weighted undirected graph with  $N_{\rm pix}$  vertices and edges connecting each pixel vertex to its set of 8 neighboring pixels. The weighted adjacency matrix over neighboring pixels (i, j) is given <sub>267</sub> by  $A_{ij} = \exp\left(-\Delta r_{ij}^2/\rho^2\right)$  where  $\Delta r_{ij}$  specifies the 3-268 dimensional Euclidean distance between the pixel centers and the widths  $\rho$  are obtained from Defferrard et al. 270 (2020). DeepSphere then efficiently performs convolu-271 tions in the spectral domain using a basis of Chebychev 272 polynomials as convolutional kernels (Defferrard et al. 273 2016); here, we set K = 4 as the maximum polynomial 274 order.

All inputs are normalized to zero mean and unit standard deviation across the training sample. Starting with 277 2 scalar input channels representing the two orthogo-278 nal (Galactic latitude and longitude) components of the <sup>279</sup> velocity vector map, <sup>1</sup> we perform a graph convolution 280 operation, increasing the channel dimension to 16 fol-281 lowed by a batch normalization, ReLU nonlinearity, and 282 downsampling the representation by a factor of 4 with 283 max pooling into the next coarser HEALPix resolution. 284 Pooling leverages scale separation, preserving important 285 characteristics of the signal across different resolutions. 286 Four more such layers are employed, increasing the chan-<sup>287</sup> nel dimension by a factor of 2 at each step until a maxi-288 mum of 128, with maps after the last convolutional layer 289 having resolution nside=2 corresponding to 48 pixels. 290 At this stage, we average over the spatial dimension 291 (known as global average pooling (Lin et al. 2014)) in or-292 der to encourage approximate rotation invariance, out-293 putting 128 features onto which the parameter of inter- $_{294}$  est  $f_{
m sub}$  is appended. These features are passed through <sup>295</sup> a fully-connected network with (1024, 128) hidden units 296 and ReLU activations outputting the classifier decision  $\hat{s}$  by applying a sigmoidal projection.

#### 2.5. Model training and evaluation

 $10^5$  maps from the forward model were produced, with 15% of these held out for validation. The estimator was trained using a batch size of 64 for up to 50 epochs with early stopping if the validation loss had not improved after 10 epochs. The ADAM optimizer (Kingma & Ba 2017) was used with initial learning rate  $10^{-3}$  decayed through cosine annealing. A coarse grid search was used to inform the architecture and hyperparameter choices in this work.

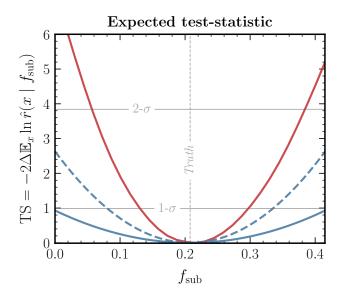
For a given test map, the log-likelihood ratio profile can be obtained by evaluating the trained estimator for different values of  $f_{\rm sub}$  while keeping the input map fixed. The network output prior to the final sigmoidal projection directly gives the required log-likelihood ratio estimate:  $\ln \hat{r} = S^{-1}(\hat{s})$ , where S is the sigmoid function (Hermans et al. 2019, 2020). Figure 1 presents an illustrative summary of the neural network architecture and method used in this work.

#### 3. EXPERIMENTS ON SIMULATED DATA

### 3.1. Baseline results and diagnostics

We evaluate our trained likelihood-ratio estimator on maps drawn from a benchmark configuration moti-

We note that by representing the input angular velocity vector field in terms of two input scalar channels, we break the desired rotation equivariance of spherical convolutions due to differences in how scalar and vector representations transform under rotations. Although this will have a downstream effect on rotation invariance, a detailed study of how this influences the performance of our model is beyond the scope of this paper.



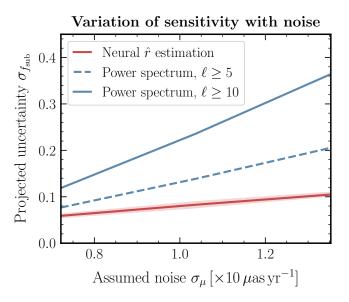


Figure 2. (Left) The expected log-likelihood ratio test-statistic (TS) profile for a cold dark matter population as a function of substructure fraction  $f_{\rm sub}$  obtained using the neural likelihood-ratio estimation method introduced in this work (red line) compared with the corresponding profiles for existing approaches using power spectrum summaries with different multipole thresholds  $\ell \gtrsim 5$  (dashed blue line) and  $\ell \gtrsim 10$  (solid blue line). The vertical dotted line indicates the true benchmark value of the parameter  $f_{\rm sub}$  in the test dataset. Our method shows enhanced sensitivity to a cold dark matter population compared to traditional approaches. (Right) Scaling of the expected sensitivities, quantified by the respective 1- $\sigma$  uncertainties, with perobject instrumental noise. For the machine learning-based approach, the band quantifies the middle-95% containment of the inferred 1- $\sigma$  uncertainty. Our method shows a more favorable scaling with assumed measurement noise.

321 vated by Hütten et al. (2016); Springel et al. (2008),  $_{322}$  containing 150 subhalos in expectation between  $10^{8}$  $_{323}$   $10^{10}\,\mathrm{M}_{\odot}$  and corresponding to  $f_{\mathrm{sub}}\simeq0.2$ . The left panel 324 of Fig. 2 shows the expected log-likelihood ratio teststatistic (TS) as a function of substructure fraction  $f_{\text{sub}}$ 326 for this nominal configuration. This is obtained by evaluating the trained estimator on 100 test maps over a uniform grid in  $f_{\text{sub}}$  and taking the point-wise mean. Cor-329 responding curves using the power spectrum approach 330 are shown in blue, using minimum multipoles of  $\ell > 5$ dashed) and  $\ell \geq 10$  (solid). Thresholds corresponding to 1- and 2- $\sigma$  significance assuming a  $\chi^2$ -distributed TS 333 are shown as the horizontal grey lines. We see that sensitivity gains of over a factor of  $\sim 2$  can be expected 335 for this particular benchmark when using the machine 336 learning approach compared to the traditional power 337 spectrum approach. No significant bias on the central value of the inferred DM abundance relative to the overall uncertainty scale is observed.

The right panel of Fig. 2 shows the scaling of expected  $^{341}$  1- $\sigma$  uncertainty on substructure fraction  $f_{\rm sub}$  with assumed noise per quasar, keeping the number of quasars fixed (red, with the line showing the median and shaded  $^{344}$  band corresponding to the middle-95% containment of the uncertainty inferred over 50 test datasets) compared to the power spectrum approach (blue lines). A far more  $^{346}$  favorable scaling of the machine learning approach is

<sup>348</sup> seen compared to the power spectrum approach, sug-<sup>349</sup> gesting that it is especially advantageous in low signal-<sup>350</sup> to-noise regimes that are generally most relevant for <sup>351</sup> dark matter searches.

Finally, we assess the quality of the approximate 353 likelihood-ratio estimator through a test of statistical 354 coverage. Within a hypothesis testing framework, this 355 is necessary in order to ensure that the learned esti-356 mator is conservative over the parameter range of in-357 terest and does not produced overly confident or biased results (Hermans et al. 2021). We obtain the es-359 timated TS profile for 1000 simulated samples with 360 true substructure fraction values drawn from the range  $f_{\text{sub}} \in [0.1, 0.3]$ . In doing so, we exclude parameter 362 points towards the edges of our parameter space since 363 the corresponding confidence intervals in these cases 364 would extend outside of the tested parameter range, as 365 can also be inferred from the baseline analysis shown 366 in Fig. 2. For nominal confidence levels in the range  $_{367}$   $1-\alpha \in [0.05, 0.95]$  we compute the empirical coverage over the set of samples, defined as the fraction of samples 369 whose true parameter value falls within the TS confi-370 dence interval, computed for a given confidence level un-<sub>371</sub> der the assumption that the TS is  $\chi^2$ -distributed (Wilks <sup>372</sup> 1938). The procedure is repeated for 10 different sets of 373 1000 samples in order to estimate the statistical uncer-374 tainty associated with the empirical coverage.

The results of the coverage test are shown in Fig. 3, illustrating the median (solid red) and middle-68% containment (red band) of the empirical coverage. We see that the empirical coverage has the desired property of being slightly conservative while still being close to the perfectly-calibrated regime indicated by the dashed-grey line. We emphasize that this diagnostic tests the quality of the likelihood-ratio estimator over the entire evaluation parameter range of interest  $f_{\rm sub} \in [0.1, 0.3]$  rather than the baseline value  $f_{\rm sub} \simeq 0.2$  in isolation.

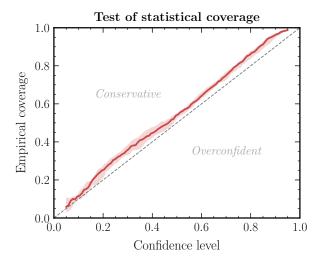


Figure 3. Calibration test

# 3.2. Experiments with unmodeled noise correlated on large scales

Since the existence of measurement noise correlated 388 on large spatial scales is a potential source of system-389 atic uncertainty when working with astrometric maps, re test the susceptibility of our method to such ef-391 fects by creating simulated data containing large-scale 392 noise not previously seen by the trained estimator. Instead of assuming a scale-invariant noise power spectrum  $C_\ell^{\rm noise} = 4\pi\sigma_\mu^2/N_q$  (Mishra-Sharma et al. 2020), 395 in this case we model noise with an order of magni- $_{396}$  tude excess in power on scales  $\ell \, \lesssim \, 10,$  parameterized as  $C_\ell^{
m noise}=4\pi\sigma_\mu^2/N_q\cdot(10-9S(\ell-10))$  where S de-398 notes the sigmoid function. The left panel of Fig. 4 illus-399 trates this noise model (thicker green line) as well as the 400 power spectrum of one simulated realization from this 401 model (thinner green line, obtained using the HEALPix 402 module anafast) contrasted with the standard scale-403 invariant noise case (red lines). The right panel of Fig. 4 shows the expected log-likelihood ratio test-statistic pro-405 file for the two cases. Although a bias in the maximum-406 likelihood estimate of  $f_{\rm sub}$  is seen when the test data  $_{407}$  has unmodeled noise (green line), the true test parame- $_{408}$  ter value (dashed vertical line) is seen to lie well within  $_{409}$  the inferred 1- $\sigma$  confidence interval. This suggests that  $_{410}$  the method is only marginally susceptible to substantive  $_{411}$  amounts of correlated noise on large spatial scales.

### 3.3. Experiments with a data-driven noise model

We finally assess the performance of our model using are alistic, data-drive noise model obtained using the astrometric catalog of quasars in *Gaia*'s second data release (DR2). The catalog contains the measured 2-dimensional positions, proper motions, as well as proper motion uncertainties of 555,934 quasars. We bin the proper motion uncertainties using the *HEALPix* pix-delization used in the analysis, down-weighting the mean variance in each bin by the number of quasars.

#### 4. CONCLUSIONS AND OUTLOOK

We have introduced a method to analyze astrometric 424 datasets over large regions of the sky using techniques 425 based on machine learning with the aim of inferring the 426 lensing signature of a dark matter substructure. We have 427 shown our method to be significantly more sensitive to 428 a cold dark matter subhalo population compared to es-429 tablished methods based on global summary statistics, 430 with more favorable scaling as a function of measure-431 ment noise. Since the collection and reduction of astro-432 metric data is an expensive endeavor, the use of methods 433 that can take advantage of more of the available infor-434 mation can be equated to long periods of data-taking, 435 underscoring their importance. Additionally, unlike the 436 power spectrum approach, the current method does not 437 require the construction of a numerically-expensive esti-438 mator to account for non-uniform exposure, selection ef-439 fects, and instrumental noise in realistic datasets. These, 440 as well as any other modeled observational effects, can be 441 incorporated directly at the level of the forward model. We have focused in this work on assessing sensitiv-443 ity to a cold dark matter-like subhalo population with 444 quasar velocity astrometry, which is within the scope of 445 upcoming radio surveys like the SKA (Fomalont & Reid 446 2004; Jarvis et al. 2015). Our method can also be applied 447 in a straightforward manner to look for the acceleration 448 lensing signal imprinted on Milky Way stars, in partic-449 ular sourced by a population of more compact subhalos 450 than those expected in the cold dark matter scenario. 451 These features are expected to imprint a larger degree 452 of non-Gaussianity compared to the signal explored here 453 (as can be seen, e.g., from Fig. 1 of Mishra-Sharma 454 et al. (2020)), and machine learning methods may pro-455 vide larger relative sensitivity gains when deployed in 456 that context. Such analyses are within purview of the

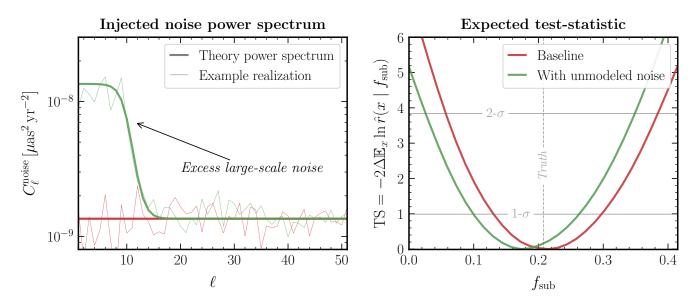


Figure 4. (Left) The power spectrum of the noise model (thicker green line) used to study the impact of correlated noise on large spatial scales, not modeled during training, on the performance of the likelihood-ratio estimator. The thinner green line shows the power spectrum of an example noise realization instantiated from this noise model. The red lines show corresponding power spectra for a scale-invariant noise model. (Right) The expected test-statistic profile for a model evaluated on maps containing excess large-scale noise (green line) compared to the model evaluated on maps with scale-invariant noise (red line). A bias in the maximum-likelihood estimate returned by the model is seen when substantial unaccounted-for noise is presented in the test maps.

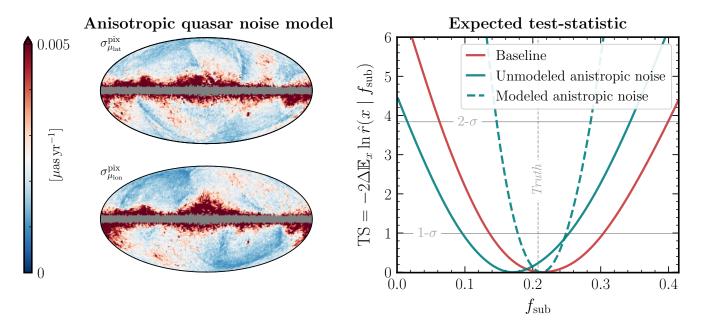


Figure 5. (Left) (Right)

555

doi: 10.1093/mnras/staa3165

457 upcoming Roman exoplanet microlensing survey (Pardo & Doré 2021) as well as future Gaia data releases. Several improvements and extensions to the method 460 presented in this paper are possible. The use of architectures that can equivariantly handle vector inputs (Esteves et al. 2020) can aid in learning more efficient rep-463 resentations of the astrometric map. Using convolutions 464 based on fixed rather than learned filters can addition-465 ally reduce model complexity and produce more inter-466 pretable representations (Cheng et al. 2020; Ha et al. 2021; Saydjari & Finkbeiner 2021; McEwen et al. 2021; Valogiannis & Dvorkin 2021). The use of methods for 469 likelihood-ratio estimation that can leverage additional 470 latent information in the forward model can significantly 471 enhance the sample efficiency of the analysis (Brehmer 472 et al. 2018a, 2020, 2018b; Stoye et al. 2018). We leave 473 the study of these extensions as well as application of 474 our method to other dark matter population scenarios 475 to future work. Astrometric lensing has been established as a promis-

477 ing way to characterize the Galactic dark matter popu-

478 lation, with theoretical progress in recent years going in 479 step with advances on the observational front. While this 480 work is a first attempt at bringing principled machine 481 learning techniques to this field, with the availability of 482 increasingly complex datasets we expect machine learn-483 ing to be an important general-purpose tool for future 484 astrometric dark matter searches.

### (Acknowledgments anonymized for review)

Software: Astropy (Robitaille et al. 2013; Price-520 Whelan et al. 2018), healpy (Gorski et al. 2005; 521 Zonca et al. 2019), IPython (Pérez & Granger 2007), <sup>522</sup> Jupyter (Kluyver et al. 2016), Matplotlib (Hunter 2007), 523 MLflow (Chen et al. 2020), NumPy (Harris et al. 524 2020), PyGSP (Defferrard et al. 2017), PyTorch (Paszke 525 et al. 2019), PyTorch Geometric (Fey & Lenssen 2019), 526 PyTorch Lightning (Falcon et al. 2020), sbi (Tejero-527 Cantero et al. 2020), SciPy (Virtanen et al. 2020), and 528 seaborn (Waskom et al. 2017).

#### REFERENCES

```
529 Baldi, P., Cranmer, K., Faucett, T., Sadowski, P., &
     Whiteson, D. 2016, Eur. Phys. J. C, 76, 235,
530
     arXiv: 1601.07913, doi: 10.1140/epjc/s10052-016-4099-4
531
532 Bode, P., Ostriker, J. P., & Turok, N. 2001, Astrophys. J.,
     556, 93, arXiv: astro-ph/0010389, doi: 10.1086/321541
533
534 Bovy, J. 2020, arXiv: 2012.02169
535 Brehmer, J., Cranmer, K., Louppe, G., & Pavez, J. 2018a,
     Phys. Rev. D, 98, 052004, arXiv: 1805.00020,
536
     doi: 10.1103/PhysRevD.98.052004
537
     -. 2018b, Phys. Rev. Lett., 121, 111801,
538
     arXiv: 1805.00013, doi: 10.1103/PhysRevLett.121.111801
539
540 Brehmer, J., Louppe, G., Pavez, J., & Cranmer, K. 2020,
     Proc. Nat. Acad. Sci., 117, 5242, arXiv: 1805.12244,
541
     doi: 10.1073/pnas.1915980117
542
  Brehmer, J., Mishra-Sharma, S., Hermans, J., Louppe, G.,
     & Cranmer, K. 2019, Astrophys. J., 886, 49,
     arXiv: 1909.02005, doi: 10.3847/1538-4357/ab4c41
545
546 Buckley, M. R., & Peter, A. H. G. 2018, Phys. Rept., 761,
     1, arXiv: 1712.06615, doi: 10.1016/j.physrep.2018.07.003
547
548 Chandrasekhar, S. 1939, An introduction to the study of
     stellar structure
549
550 Chen, A., et al. 2020, in Proceedings of the fourth
     international workshop on data management for
551
     end-to-end machine learning, 1-4
552
   Cheng, S., Ting, Y.-S., Ménard, B., & Bruna, J. 2020, Mon.
553
     Not. Roy. Astron. Soc., 499, 5902, arXiv: 2006.08561,
```

```
556 Cranmer, K., Brehmer, J., & Louppe, G. 2020, Proc. Nat.
     Acad. Sci., 117, 30055, arXiv: 1911.01429,
     doi: 10.1073/pnas.1912789117
  Cranmer, K., Pavez, J., & Louppe, G. 2015,
     arXiv: 1506.02169
560
561 Dalcanton, J. J., & Hogan, C. J. 2001, Astrophys. J., 561,
     35, arXiv: astro-ph/0004381, doi: 10.1086/323207
562
563 Defferrard, M., Bresson, X., & Vandergheynst, P. 2016,
     arXiv: 1606.09375
564
  Defferrard, M., Martin, L., Pena, R., & Perraudin, N. 2017,
565
     PyGSP: Graph Signal Processing in Python, v0.5.0,
566
     Zenodo, doi: 10.5281/zenodo.1003158.
567
     https://doi.org/10.5281/zenodo.1003158
  Defferrard, M., Milani, M., Gusset, F., & Perraudin, N.
     2020, arXiv: 2012.15000
570
571 Esteves, C., Makadia, A., & Daniilidis, K. 2020,
     Spin-Weighted Spherical CNNs, arXiv: 2006.10731
572
573 Falcon, W., et al. 2020,
     PyTorchLightning/pytorch-lightning: 0.7.6 release, 0.7.6,
574
     Zenodo, doi: 10.5281/zenodo.3828935.
575
     https://doi.org/10.5281/zenodo.3828935
576
577 Fey, M., & Lenssen, J. E. 2019, in ICLR Workshop on
     Representation Learning on Graphs and Manifolds
578
579 Fitts, A., et al. 2017, Mon. Not. Roy. Astron. Soc., 471,
     3547, arXiv: 1611.02281, doi: 10.1093/mnras/stx1757
```

```
581 Fomalont, E. B., & Reid, M. 2004, New Astron. Rev., 48,
     1473, arXiv: astro-ph/0409611,
582
     doi: 10.1016/j.newar.2004.09.037
583
584 Gaia Collaboration. 2016, Astron. Astrophys., 595, A1,
     arXiv: 1609.04153, doi: 10.1051/0004-6361/201629272
585
    -. 2018a, Astron. Astrophys., 616, A1, arXiv: 1804.09365,
586
     doi: 10.1051/0004-6361/201833051
    -. 2018b, Astron. Astrophys., 616, A14, arXiv: 1804.09377,
588
     doi: 10.1051/0004-6361/201832916
589
    -. 2021, Astron. Astrophys., 649, A1, arXiv: 2012.01533,
590
     doi: 10.1051/0004-6361/202039657
591
592 Gilman, D., Birrer, S., Nierenberg, A., et al. 2020, Mon.
     Not. Roy. Astron. Soc., 491, 6077, arXiv: 1908.06983,
593
     doi: 10.1093/mnras/stz3480
594
   Gorski, K. M., Hivon, E., Banday, A. J., et al. 2005,
595
     Astrophys. J., 622, 759, arXiv: astro-ph/0409513,
596
     doi: 10.1086/427976
597
598 Gravity Collaboration. 2019, Astron. Astrophys., 625, L10,
     arXiv: 1904.05721, doi: 10.1051/0004-6361/201935656
600 Green, A. M. 2021, in Les Houches summer school on Dark
601
602 Ha, W., Singh, C., Lanusse, F., Upadhyayula, S., & Yu, B.
     2021, arXiv: 2107.09145
603
604 Harris, C. R., et al. 2020, Nature, 585, 357,
     doi: 10.1038/s41586-020-2649-2
605
606 Hermans, J., Banik, N., Weniger, C., Bertone, G., &
     Louppe, G. 2020, arXiv: 2011.14923
607
608 Hermans, J., Begy, V., & Louppe, G. 2019,
     arXiv: 1903.04057
610 Hermans, J., Delaunoy, A., Rozet, F., Wehenkel, A., &
     Louppe, G. 2021, Averting A Crisis In Simulation-Based
611
     Inference, arXiv: 2110.06581
612
613 Hezaveh, Y. D., et al. 2016, Astrophys. J., 823, 37,
     arXiv: 1601.01388, doi: 10.3847/0004-637X/823/1/37
614
615 Hunter, J. D. 2007, Computing In Science & Engineering,
616
617 Hütten, M., Combet, C., Maier, G., & Maurin, D. 2016,
     JCAP, 09, 047, arXiv: 1606.04898,
618
     doi: 10.1088/1475-7516/2016/09/047
619
620 Jarvis, M. J., Bacon, D., Blake, C., et al. 2015,
     arXiv: 1501.03825
621
622 Kerr, F. J., & Lynden-Bell, D. 1986, Mon. Not. Rov.
     Astron. Soc., 221, 1023
623
624 Kingma, D. P., & Ba, J. 2017, Adam: A Method for
     Stochastic Optimization, arXiv: 1412.6980
625
626 Kluyver, T., et al. 2016, in ELPUB
627 Lin, M., Chen, Q., & Yan, S. 2014, Network In Network,
     arXiv: 1312.4400
628
629 Lindegren, L., et al. 2018, Astron. Astrophys., 616, A2,
```

arXiv: 1804.09366, doi: 10.1051/0004-6361/201832727

```
631 Lisanti, M. 2017, in Theoretical Advanced Study Institute
     in Elementary Particle Physics: New Frontiers in Fields
632
     and Strings, 399-446
633
634 McEwen, J. D., Wallis, C. G. R., & Mavor-Parker, A. N.
     2021, arXiv: 2102.02828
636 Mishra-Sharma, S., Van Tilburg, K., & Weiner, N. 2020,
     Phys. Rev. D, 102, 023026, arXiv: 2003.02264,
     doi: 10.1103/PhysRevD.102.023026
639 Mohamed, S., & Lakshminarayanan, B. 2017, Learning in
     Implicit Generative Models, arXiv: 1610.03483
640
641 Mondino, C., Taki, A.-M., Van Tilburg, K., & Weiner, N.
     2020, Phys. Rev. Lett., 125, 111101, arXiv: 2002.01938,
642
     doi: 10.1103/PhysRevLett.125.111101
643
644 Navarro, J. F., Frenk, C. S., & White, S. D. M. 1996,
     Astrophys. J., 462, 563, arXiv: astro-ph/9508025,
645
     doi: 10.1086/177173
646
647 Pardo, K., & Doré, O. 2021, arXiv: 2108.10886
648 Paszke, A., et al. 2019, in Advances in Neural Information
     Processing Systems 32, ed. H. Wallach, H. Larochelle,
649
     A. Beygelzimer, F. d'Alché-Buc, E. Fox, & R. Garnett
650
     (Curran Associates, Inc.), 8024-8035
651
652 Pérez, F., & Granger, B. E. 2007, Computing in Science
     and Engineering, 9, 21, doi: 10.1109/MCSE.2007.53
653
  Perraudin, N., Defferrard, M., Kacprzak, T., & Sgier, R.
654
     2019, Astronomy and Computing, 27, 130,
655
     arXiv: 1810.12186, doi: 10.1016/j.ascom.2019.03.004
656
   Piffl, T., et al. 2014, Astron. Astrophys., 562, A91,
     arXiv: 1309.4293, doi: 10.1051/0004-6361/201322531
  Price-Whelan, A., et al. 2018, Astron. J., 156, 123,
     arXiv: 1801.02634, doi: 10.3847/1538-3881/aabc4f
660
661 Read, J. I., Iorio, G., Agertz, O., & Fraternali, F. 2017,
     Mon. Not. Roy. Astron. Soc., 467, 2019,
662
     arXiv: 1607.03127, doi: 10.1093/mnras/stx147
663
664 Robitaille, T. P., et al. 2013, Astron. Astrophys., 558, A33,
     arXiv: 1307.6212, doi: 10.1051/0004-6361/201322068
665
666 Sánchez-Conde, M. A., & Prada, F. 2014, Mon. Not. Roy.
     Astron. Soc., 442, 2271, arXiv: 1312.1729,
667
     doi: 10.1093/mnras/stu1014
668
669 Saydjari, A. K., & Finkbeiner, D. P. 2021,
     arXiv: 2104.11244
671 Schönrich, R., Binney, J., & Dehnen, W. 2010, Mon. Not.
     Roy. Astron. Soc., 403, 1829, arXiv: 0912.3693,
672
     doi: 10.1111/j.1365-2966.2010.16253.x
673
674 Schutz, K. 2020, Phys. Rev. D, 101, 123026,
     arXiv: 2001.05503, doi: 10.1103/PhysRevD.101.123026
675
676 Springel, V., Wang, J., Vogelsberger, M., et al. 2008, Mon.
     Not. Roy. Astron. Soc., 391, 1685, arXiv: 0809.0898,
677
     doi: 10.1111/j.1365-2966.2008.14066.x
678
  Stoye, M., Brehmer, J., Louppe, G., Pavez, J., & Cranmer,
     K. 2018, arXiv: 1808.00973
```

```
681 Tegmark, M. 1997, Phys. Rev. D, 55, 5895,
```

- arXiv: astro-ph/9611174, doi: 10.1103/PhysRevD.55.5895
- Tejero-Cantero, A., et al. 2020, Journal of Open Source
- 684 Software, 5, 2505, doi: 10.21105/joss.02505
- 685 Valogiannis, G., & Dvorkin, C. 2021, arXiv: 2108.07821
- 686 Van Tilburg, K., Taki, A.-M., & Weiner, N. 2018, JCAP,
- 687 07, 041, arXiv: 1804.01991,
- doi: 10.1088/1475-7516/2018/07/041
- 689 Vattis, K., Toomey, M. W., & Koushiappas, S. M. 2020,
- arXiv: 2008.11577
- 691 Vegetti, S., Koopmans, L. V. E., Bolton, A., Treu, T., &
- 692 Gavazzi, R. 2010, Mon. Not. Roy. Astron. Soc., 408, 1969,
- 693 arXiv: 0910.0760, doi: 10.1111/j.1365-2966.2010.16865.x

- 694 Vegetti, S., Lagattuta, D. J., McKean, J. P., et al. 2012,
- Nature, 481, 341, arXiv: 1201.3643,
- doi: 10.1038/nature10669
- 697 Virtanen, P., et al. 2020, Nature Methods, 17, 261,
- doi: 10.1038/s41592-019-0686-2
- 699 Waskom, M., et al. 2017, mwaskom/seaborn: v0.8.1
- 700 (September 2017), v0.8.1, Zenodo,
- 701 doi: 10.5281/zenodo.883859.
- 702 https://doi.org/10.5281/zenodo.883859
- 703 WFIRST Astrometry Working Group. 2019, Journal of
- Astronomical Telescopes, Instruments, and Systems, 5,
- 705 044005, arXiv: 1712.05420,
- doi: 10.1117/1.JATIS.5.4.044005
- 707 Wilks, S. S. 1938, The Annals of Mathematical Statistics,
- 708 9, 60, doi: 10.1214/aoms/1177732360
- 709 Zonca, A., Singer, L., Lenz, D., et al. 2019, Journal of Open
- o Source Software, 4, 1298, doi: 10.21105/joss.01298