

- Jeffery 先验是一无信息先验分布，计算方法是根据下面的 Fisher information 矩阵对应的行列式。

$$P(\theta) \propto \sqrt{\det I(\theta)}.$$

- 具体好处就是，不受到原分布的参数选择差异的影响，所以说是 non-informative 的，这一性质对于 Bayesian 方法中很重要，因为先验的选择往往会影响到后验参数的估计，
- 减少参数量。

- Fish information \leftarrow 似然
- $$I(\theta) = -E \left[\frac{d^2 \log f}{d\theta^2} \right]$$
- $$\frac{d^2 \log f}{d\theta^2} = \frac{d}{d\theta} \left(\frac{\frac{df}{d\theta}}{f} \right)$$

- Jeffery 先验.

$f(b|\theta)$ 为似然函数.

Fish information.
$$I(\theta) = E \left[\left(\frac{d \log f(b, \theta)}{d\theta} \right)^2 \right]$$
$$= \int \left(\frac{d \log f(b, \theta)}{d\theta} \right)^2 f(b, \theta) db.$$

$p(\theta) \propto \sqrt{I(\theta)}$. 为 Jeffery 先验.

- $p(\varphi)$ 是 -- 映射

$$p(\varphi) = p(\theta) \left| \frac{d\theta}{d\varphi} \right|$$

$$\propto \sqrt{I(\theta)} \cdot \left(\frac{d\theta}{d\varphi} \right)^2$$

$$= \sqrt{E \left(\frac{d \log f}{d\theta} \right)^2 \left(\frac{d\theta}{d\varphi} \right)^2}$$

$$= \sqrt{E \left(\left(\frac{d \log f}{d\theta} \cdot \frac{d\theta}{d\varphi} \right)^2 \right)}$$

$$= \sqrt{E \left(\frac{d \log f}{d\varphi} \right)^2}$$

$$= \sqrt{I(\varphi)}$$

先验是不变量, 并且没有新的参数.

- $X \sim N(\mu, \sigma^2)$

① μ 不定, σ 固定.

$$\log f = -\frac{1}{2} \left(\frac{x - \mu}{\sigma} \right)^2$$

$$\sqrt{I(\mu)} = \sqrt{E \left(\frac{(x - \mu)^2}{\sigma^4} \right)} = \sqrt{\frac{1}{\sigma^2}}$$

不是先验

$$\int_{-\infty}^{+\infty} db = \infty$$

$$I(\mu) \propto 1$$

没有信息的先验.

② μ 固定, 求 δ .

$$\text{令 } \tau = \frac{1}{\delta^2}$$

$$f(\tau) = \frac{1}{\sqrt{2\pi}\delta} \exp\left(-\frac{(x-\mu)^2}{2\delta^2}\right) = \frac{\tau^{\frac{1}{2}}}{\sqrt{2\pi}} \exp\left(-\frac{\tau(x-\mu)^2}{2}\right)$$

$$\ln f = \frac{1}{2} \ln \tau - \frac{\tau}{2} (x-\mu)^2$$

$$\frac{d \ln f}{d \tau} = \frac{1}{2\tau} - \frac{1}{2} (x-\mu)^2$$

$$E\left[\frac{d \ln f}{d \tau}\right] = E\left[\frac{1}{4\tau^2} - \frac{(x-\mu)^2}{2\tau} + \frac{(x-\mu)^4}{4}\right]$$

$$= \frac{1}{4\tau^2} - \frac{1}{2\tau^2} + \frac{3}{4\tau^2}$$

$$= \frac{1}{2\tau^2}$$

$$= \sqrt{I(\tau)} \quad \text{这里假设了 } \tau \text{ 为 Jefferys 先验}$$

• 泊松分布.

$$f(x|\lambda) = e^{-\lambda} \frac{\lambda^n}{n!}$$

$$\pi(\lambda) = \lambda \text{ 的 Jefferys 先验}$$

$$\pi(\lambda) = E\left[\left(\frac{n}{\lambda} - 1\right)^2\right]$$

$$= E\left(1 + \frac{n^2}{\lambda^2} - \frac{2n}{\lambda}\right)$$

$$= \sqrt{\frac{1}{\lambda}}$$

• 优点: 不需要任何额外参数.

$$\bullet \quad X = \theta + \varepsilon \quad \varepsilon \sim N(0, \tau^{-\frac{1}{2}})$$

① τ 固定,

$$P(\theta) = N(\theta | 0, \lambda^{\frac{1}{2}})$$

$$P(x|\theta)P(\theta) = \frac{1}{\sqrt{2\pi}\tau} \exp\left(-\frac{(x-\theta)^2}{2\tau}\right) \frac{1}{\sqrt{2\pi\lambda}} \exp\left(-\frac{\theta^2}{2\lambda}\right)$$

$$= \frac{1}{\sqrt{2\pi}\tau} \exp\left(-\frac{\lambda x^2}{\lambda + \tau}\right)$$

$$\hat{\theta} = \frac{\lambda x}{\lambda + \tau} \quad (\text{峰值})$$

② τ 不固定.

第一种估计法: $P(\theta, \tau) = P(\theta)P(\tau)$

$$P(\theta, \tau) \propto P(x|\theta, \tau)P(\theta, \tau)$$

$$= P(x|\theta, \tau)P(\theta)P(\tau)$$

$$P(\theta) \sim N(0, \lambda^{\frac{1}{2}})$$

$$P(\tau) = \text{Ga}(\alpha, \frac{1}{2}\beta)$$

$$\propto \frac{1}{\sqrt{2\pi}} \tau^{\frac{1}{2}} \exp\left(-\frac{\tau(x-\theta)^2}{2}\right) \frac{\lambda^{\frac{1}{2}}}{\sqrt{2\pi}} \exp\left(-\frac{\lambda\theta^2}{2}\right) \cdot \left(\frac{\beta}{2}\right)^{\frac{\alpha}{2}} \exp\left(-\frac{\beta\tau}{2}\right)$$

$$L = \tau^{\frac{\alpha+1}{2}-1} \exp\left(-\frac{\tau}{2}[(x-\theta)^2 + \beta]\right) \exp\left(-\frac{\lambda}{2}\theta^2\right)$$

$$\ln L = \left(\frac{\alpha+1}{2}-1\right) \ln \tau - \frac{\tau}{2}[(x-\theta)^2 + \beta] + \frac{1}{2} \ln \lambda - \frac{\lambda}{2} \theta^2$$

$$Q = 2 \ln L = (\alpha-1) \ln \tau + \tau[(x-\theta)^2 + \beta] - \ln \lambda + \lambda \theta^2$$

极大似然解法.

$$\begin{cases} \frac{\partial Q}{\partial \tau} = \frac{1}{\tau}(\alpha-1) + (x-\theta)^2 + \beta = 0 \end{cases}$$

$$\begin{cases} \frac{\partial Q}{\partial \theta} = 2\tau(x-\theta) + \lambda\theta = 0 \end{cases}$$

无法解

第二种做法 $P(\theta, \tau) = P(\theta|\tau)P(\tau)$

$P(\theta|\tau) \sim N(0, (\lambda\tau)^{\frac{1}{2}})$ τ 为 Gam 分布

代入原方程有

极大似然

$$\begin{cases} \frac{\partial Q}{\partial \theta} = (-x - \theta) + \lambda\theta = 0 \\ \frac{\partial Q}{\partial \tau} = -\frac{\alpha}{\tau} + \frac{1}{\tau}(\alpha - \theta)^2 + \beta + \lambda\theta^2 = 0 \end{cases}$$

可解

上面 (θ, τ) 共轭。但太多超参数。

第三种做法: $P(\theta, \tau) = P(\theta|\tau)P(\tau)$

$P(\tau) \propto 1$ 满足 Jeffery 分布

$P(\tau) \propto \pi(\tau) = \frac{1}{\tau}$

$L = \exp(-\frac{\tau}{2}x^2)$ 指数分布