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Lecture 5

高斯噪声下的多样本检测

LECTURE4

- ✓ 假设检验方法：假设建立、概率映射、测量样本、统计判决
- ✓ 单样本参量检测：
 - Bayes平均风险最小准则：先验概率、代价因子已知
 - 极大极小准则：先验概率未知
 - NP准则：先验概率、代价因子均未知
- ✓ 检测性能指标：虚警概率、发现概率、错误概率等



检测场景

- 多样本
- 噪声高斯分布
- 白、非白
- 同分布、不同分布





1

多样本假设检验准则

2

IID高斯噪声下的相关接收

3

相关接收的检测性能

4

不等均值等协方差时的检测性能

5

等均值不等协方差时的检测性能



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多样本

- 实信号 $y_i = s_{ji} + n_i, j=0,1; i=1\dots M$
- 复信号 $y_i = \alpha e^{-i\beta} u_{ji} + z_i, j=0,1; i=1\dots M$
- $Y = [y_1, y_2, \dots, y_M]^T$
 - ✓ 时间采样
 - ✓ 空间采样
 - ✓ 频率采样



希尔伯特变换

$$\hat{s}(t) = \mathcal{H}\{s(t)\} = \frac{1}{\pi} \int \frac{s(\tau)}{t - \tau} d\tau = s(t) * \frac{1}{\pi t}$$

$$\Rightarrow \mathcal{F}\{s(t)\} \cdot \mathcal{F}\left\{\frac{1}{\pi t}\right\} = -jS(f) \cdot \text{sgn}(f)$$

$$\text{其中 } \text{sgn}(f) = \begin{cases} 1 & f > 0 \\ 0 & f = 0 \\ -1 & f < 0 \end{cases}$$

$$s(t) = \mathcal{H}^{-1}\{\hat{s}(t)\} = -\frac{1}{\pi} \int \frac{\hat{s}(\tau)}{t - \tau} d\tau = -\hat{s}(t) * \frac{1}{\pi t}$$



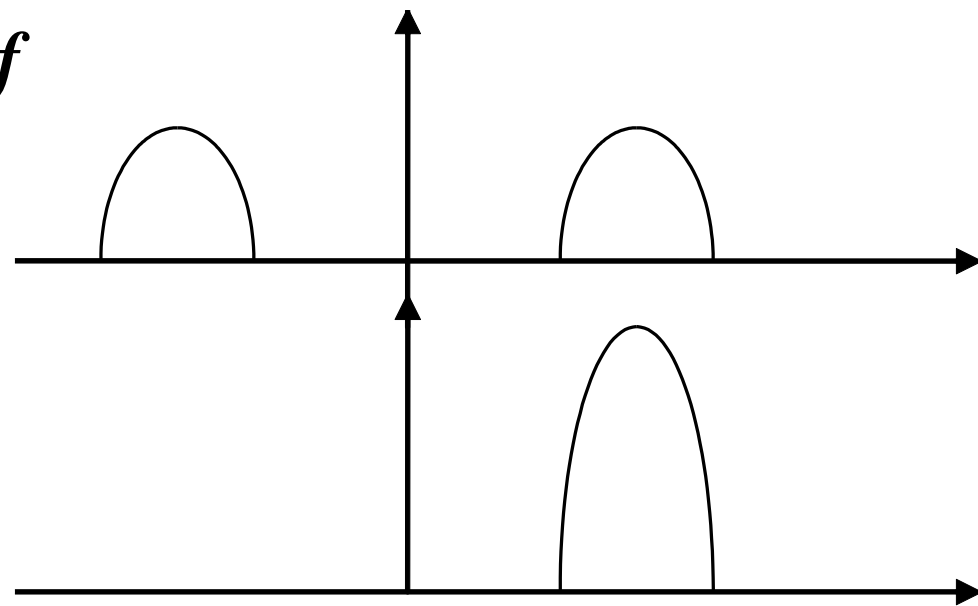
解析信号 (预包络信号)

$$\begin{aligned} s(t) &= \int_0^{\infty} S(f) e^{j2\pi ft} df + \int_{-\infty}^0 S(f) e^{j2\pi ft} df \\ &= \int_0^{\infty} S(f) e^{j2\pi ft} df + \int_0^{\infty} [S(f) e^{j2\pi ft}]^* df \\ &= \operatorname{Re} \int_{-\infty}^{\infty} 2S(f) U(f) e^{j2\pi ft} df \end{aligned}$$

$$S_P(f) = 2S(f)U(f) \Rightarrow s(t) = \operatorname{Re}[s_P(t)]$$

\Downarrow

$$s_P(t) = s(t) * \left[\delta(t) + j \frac{1}{\pi t} \right] = s(t) + j\hat{s}(t)$$



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检验准则

- Bayes平均风险最小准则

$$\frac{f(\vec{Y} \mid H_1)}{f(\vec{Y} \mid H_0)} \geq \frac{P(H_0)(C_{10} - C_{00})}{P(H_1)(C_{01} - C_{11})}, \text{ 判为 } H_1;$$

$$\frac{f(\vec{Y} \mid H_1)}{f(\vec{Y} \mid H_0)} < \frac{P(H_0)(C_{10} - C_{00})}{P(H_1)(C_{01} - C_{11})}, \text{ 判为 } H_0.$$

- 最小平均错误概率

$$\frac{f(\vec{Y} \mid H_1)}{f(\vec{Y} \mid H_0)} \geq \frac{P(H_0)}{P(H_1)} = \frac{P(H_0)}{1 - P(H_0)} = \frac{1 - P(H_1)}{P(H_1)}, \text{ 判为 } H_1;$$

$$\frac{f(\vec{Y} \mid H_1)}{f(\vec{Y} \mid H_0)} < \frac{P(H_0)}{P(H_1)} = \frac{P(H_0)}{1 - P(H_0)} = \frac{1 - P(H_1)}{P(H_1)}, \text{ 判为 } H_0.$$



检验准则

- MAP

$$\frac{P(H_1 | \vec{Y})}{P(H_0 | \vec{Y})} \geq 1, \text{ 判为 } H_1;$$
$$\frac{P(H_1 | \vec{Y})}{P(H_0 | \vec{Y})} < 1, \text{ 判为 } H_0。$$

- ML准则

$$\frac{f(Y|H_1)}{f(Y|H_0)} \geq 1, \text{ 判为 } H_1;$$

否则, 判为 H_0 。



检验准则

• 极大极小 $\frac{f(\vec{Y} | H_1)}{f(\vec{Y} | H_0)} > \tau$, 判为 H_1 ;

$\frac{f(\vec{Y} | H_1)}{f(\vec{Y} | H_0)} = \tau$, 以概率 η 判为 H_1 ;

$\frac{f(\vec{Y} | H_1)}{f(\vec{Y} | H_0)} < \tau$, 判为 H_0 。

门限的确定：

$$C_{10}\alpha(q_0) + C_{00}[1 - \alpha(q_0)] = C_{01}\beta(q_0) + C_{11}[1 - \beta(q_0)]$$



检验准则

• NP

$$\frac{f(\vec{Y} \mid H_1)}{f(\vec{Y} \mid H_0)} > \tau, \text{ 判为 } H_1;$$
$$\frac{f(\vec{Y} \mid H_1)}{f(\vec{Y} \mid H_0)} = \tau, \text{ 以概率 } \eta \text{ 判为 } H_1;$$
$$\frac{f(\vec{Y} \mid H_1)}{f(\vec{Y} \mid H_0)} < \tau, \text{ 判为 } H_0.$$

门限的确定：

$$P_{fa} = \int_{\tau}^{\infty} f(L \mid H_0) dL = \alpha$$





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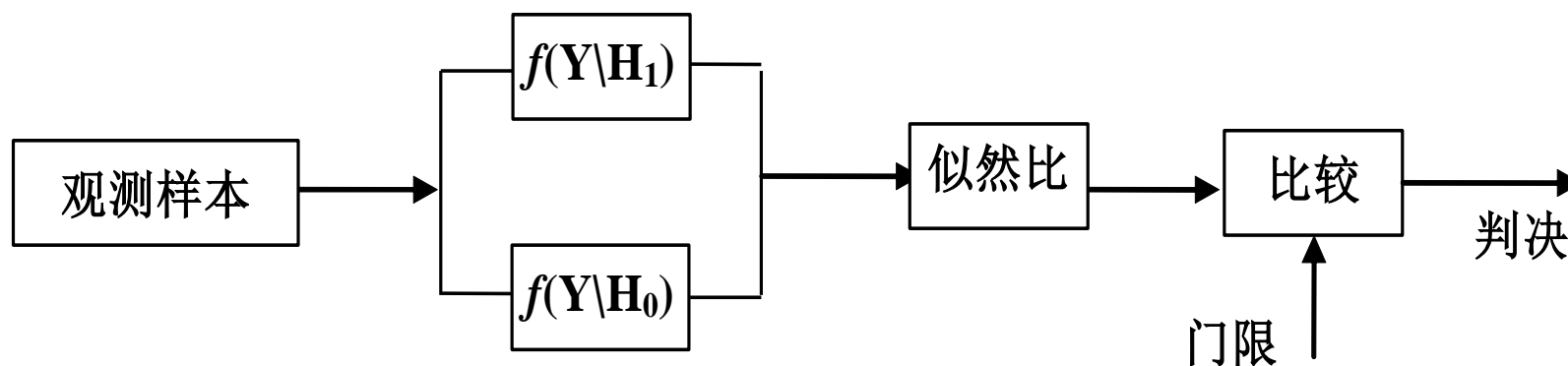
最佳检测

- 二元假设检验:

$$H_1: Y = S_1 + N$$

$$H_0: Y = S_0 + N$$

- $Y = [y_1, y_2, \dots, y_M]^T$; $S_j = [s_{j1}, s_{j2}, \dots, s_{jM}]^T$
- $N = [n_1, n_2, \dots, n_M]^T$, 白噪声, 均值为零



白噪声

- 噪声各分量独立同分布 (IID)
- 两种假设下观测值 Y 为高斯分布

- 均值:

$$H_1: E(Y|H_1) = S_1$$

$$H_0: E(Y|H_0) = S_0$$

- 方差:

$$E \left\{ (y_i - s_{1i})^2 \mid H_1 \right\} = E \left\{ (y_i - s_{0i})^2 \mid H_0 \right\} = \sigma^2$$



似然表达

- 似然比

$$L(Y) = \frac{f(Y \setminus H_1)}{f(Y \setminus H_0)} = \frac{\prod_{i=1}^M f(y_i \setminus H_1)}{\prod_{i=1}^M f(y_i \setminus H_0)} = \prod_{i=1}^M \frac{f(y_i \setminus H_1)}{f(y_i \setminus H_0)}$$

$$= \exp \left\{ -\frac{1}{2\sigma^2} \sum_{i=1}^M \left[2y_i s_{0i} - 2y_i s_{1i} - (s_{0i}^2 - s_{1i}^2) \right] \right\}$$

$$\begin{matrix} H_1 \\ \geq th \end{matrix}$$

- 对数似然比

$$l(Y) = -\frac{1}{2\sigma^2} \sum_{i=1}^M \left[2y_i s_{0i} - 2y_i s_{1i} - (s_{0i}^2 - s_{1i}^2) \right]$$

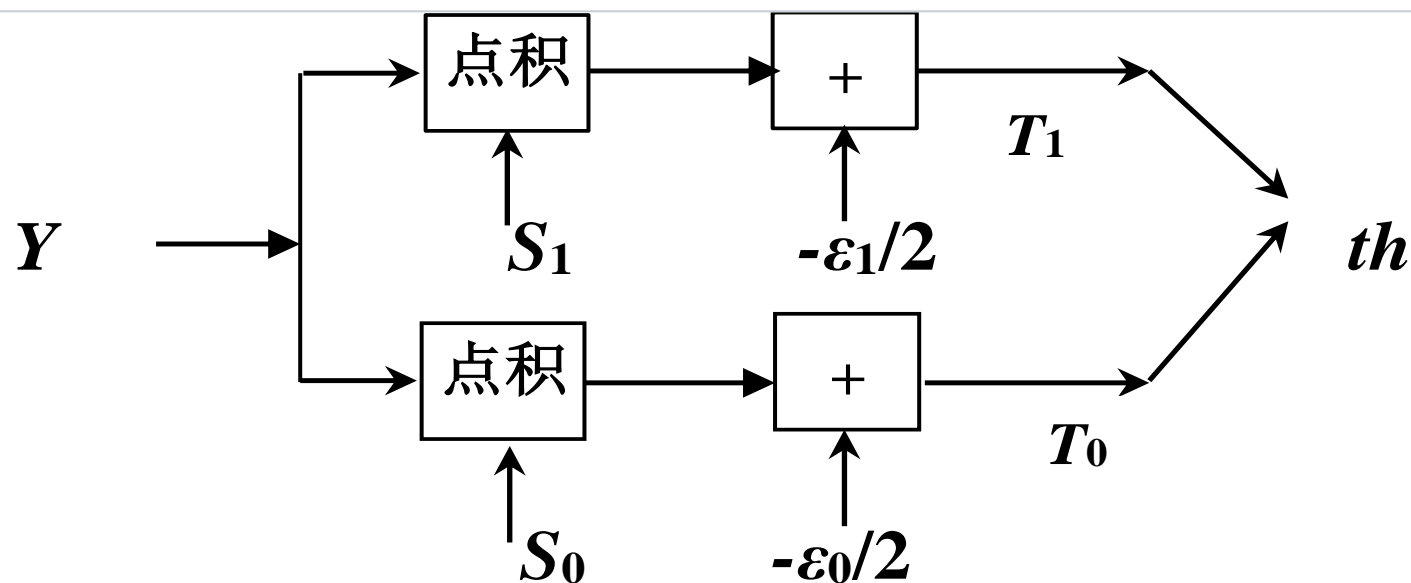


相关检测器

$$T(Y) = \sigma^2 \cdot l(Y) = \sum_{i=1}^M y_i (s_{1i} - s_{0i}) - \frac{1}{2}(\varepsilon_1 - \varepsilon_0) = T_1(Y) - T_0(Y)$$

其中

$$\begin{cases} \varepsilon_1 = \sum_{i=1}^M s_{1i}^2 \\ \varepsilon_0 = \sum_{i=1}^M s_{0i}^2 \end{cases}, \begin{cases} T_1(Y) = \sum_{i=1}^M y_i s_{1i} - \frac{1}{2} \varepsilon_1 \\ T_0(Y) = \sum_{i=1}^M y_i s_{0i} - \frac{1}{2} \varepsilon_0 \end{cases}$$



相关接收

- 最大似然—差异信号能量

$$Y^T (S_1 - S_0) \stackrel{H_1}{\geq} \frac{1}{2} (S_1^T \cdot S_1 - S_0^T \cdot S_0)$$

\Leftrightarrow

$$(Y - S_0)^T (Y - S_0) \stackrel{H_1}{\geq} (Y - S_1)^T (Y - S_1)$$





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n 维Euclidean实空间

- 矢量的点积

$$(X, Y) = \sum_{i=1}^M x_i y_i$$

- 范数

$$\|X\| = \sqrt{(X, X)} = \sqrt{\sum_{i=1}^M x_i^2}$$

- 距离

$$\|X - Y\| = \sqrt{\sum_{i=1}^M (x_i - y_i)^2}$$



似然表达

似然函数 $f(Y|H_j) = \prod f(y_i | H_j)$

$$= \frac{1}{(\sqrt{2\pi}\sigma)^M} \exp \left\{ -\frac{1}{2\sigma^2} \sum (y_i - s_{ji})^2 \right\}$$

$$= \frac{1}{(\sqrt{2\pi}\sigma)^M} \exp \left\{ -\frac{1}{2\sigma^2} \|Y - S_j\|^2 \right\}$$

似然比

$$\ln L = \ln \frac{f(Y|H_1)}{f(Y|H_0)} = -\frac{1}{2\sigma^2} \left\{ \|Y - S_1\|^2 - \|Y - S_0\|^2 \right\}$$

$$= \frac{1}{\sigma^2} [Y^T \cdot (S_1 - S_0)] - \frac{1}{2\sigma^2} \|S_1\|^2 + \frac{1}{2\sigma^2} \|S_0\|^2$$



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检验统计量分布

$$T(Y) = \sigma^2 \cdot l(Y) = \sum_{i=1}^M y_i (s_{1i} - s_{0i}) - \frac{1}{2}(\varepsilon_1 - \varepsilon_0) \geq \sigma^2 \cdot \ln th$$

$$E(T \mid H_0) = E \left\{ \sum_{i=1}^M y_i (s_{1i} - s_{0i}) \right\} - \frac{1}{2}(\varepsilon_1 - \varepsilon_0)$$

$$= \sum_{i=1}^M s_{0i} s_{1i} - \frac{1}{2} \sum_{i=1}^M s_{0i}^2 - \frac{1}{2} \sum_{i=1}^M s_{1i}^2$$

$$= -\frac{1}{2} \sum_{i=1}^M (s_{1i} - s_{0i})^2$$

$$= -\frac{1}{2} \|S_1 - S_0\|^2$$



检验统计量分布

$$\begin{aligned} \text{Var}(T \mid H_0) &= \text{Var}\left(\sum_{i=1}^M y_i (s_{1i} - s_{0i}) \mid H_0\right) \\ &= \sum_{i=1}^M \text{Var}(y_i) (s_{1i} - s_{0i})^2 \\ &= \sigma^2 \|S_1 - S_0\|^2 \end{aligned}$$

$$H_0 : T \in N\left(-\frac{1}{2}\|S_1 - S_0\|^2, \sigma^2 \|S_1 - S_0\|^2\right)$$

$$\text{同理 } H_1 : T \in N\left(\frac{1}{2}\|S_1 - S_0\|^2, \sigma^2 \|S_1 - S_0\|^2\right)$$



虚警概率

$$\begin{aligned}\alpha &= P(D_1 \setminus H_0) = \Pr\{T(Y) > th \setminus H_0\} \\ &= \int_{\sigma^2 \ln th}^{\infty} f(T \setminus H_0) dT \\ &= Q\left[\frac{\ln th}{d} + \frac{d}{2}\right]\end{aligned}$$

- 偏移系数

$$d^2 = \frac{[E(Y \setminus H_1) - E(Y \setminus H_0)]^2}{\text{Var}(Y \setminus H_0)} = \frac{\|S_1 - S_0\|^2}{\sigma^2}$$

注： $Q(x) = \int_x^{\infty} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{u^2}{2}\right) du$



发现概率和漏警概率

$$\begin{aligned} P_d &= P(D_1 \setminus H_1) = \Pr\{T(Y) \geq th' \setminus H_1\} \\ &= \int_{\sigma^2 \ln th}^{\infty} f(T \setminus H_1) dT = Q\left[\frac{\ln th}{d} - \frac{d}{2}\right] = Q\left[Q^{-1}(\alpha) - d\right] \end{aligned}$$

虚警概率的逆

$$\begin{aligned} \beta &= P(D_0 \setminus H_1) = \Pr\{T(Y) < th' \setminus H_1\} \\ &= \int_{-\infty}^{\sigma^2 \ln th} f(T \setminus H_1) dT = 1 - Q\left[\frac{\ln th}{d} - \frac{d}{2}\right] \end{aligned}$$



错误概率

$$P_e = P(H_0)P(D_1 | H_0) + P(H_1)P(D_0 | H_1)$$

- 若先验概率相等, 则

$$P_e = P(D_1 | H_0) = \Pr\{T(Y) > 0 | H_0\}$$

$$= \int_0^\infty f(T | H_0) dT$$

$$= Q\left(\frac{\frac{1}{2}\|S_1 - S_0\|^2}{\sqrt{\sigma^2\|S_1 - S_0\|^2}}\right) = Q\left(\frac{1}{2}\sqrt{\frac{\|S_1 - S_0\|^2}{\sigma^2}}\right)$$

偏移系数



错误概率

- 约束:平均能量有限

$$\bar{\varepsilon} = \frac{1}{2}(\varepsilon_1 + \varepsilon_0)$$

- 归一化

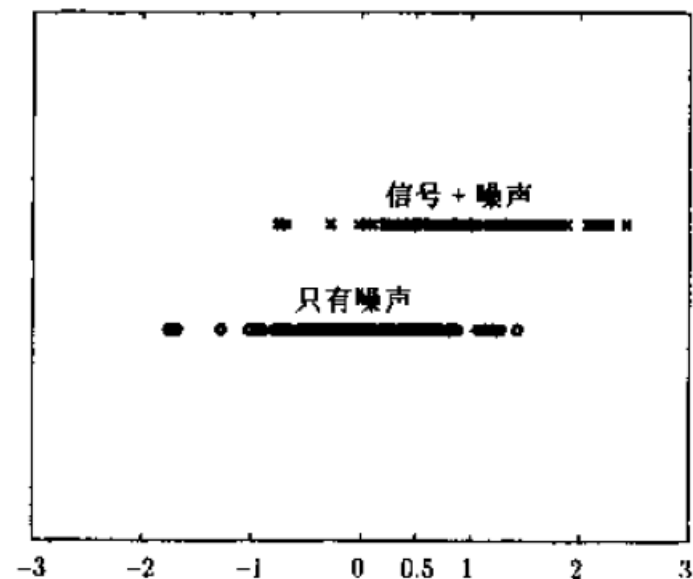
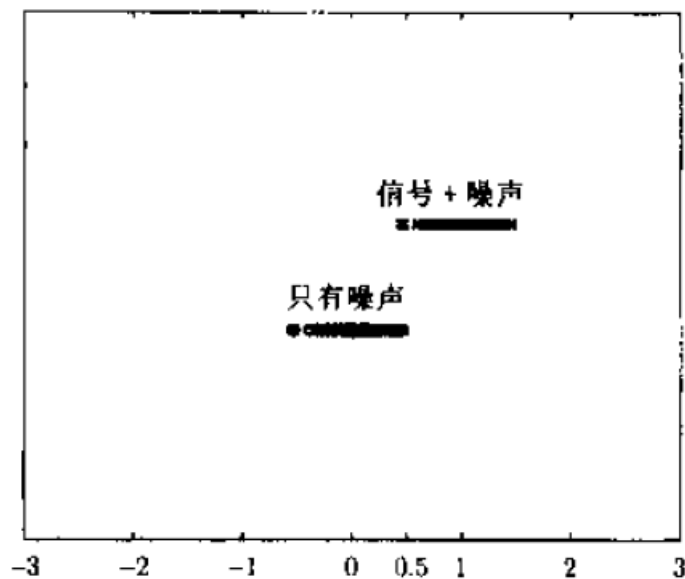
$$\rho_s = \frac{S_1^T S_0}{\frac{1}{2}(S_1^T S_1 + S_0^T S_0)}$$

- 平均错误概率

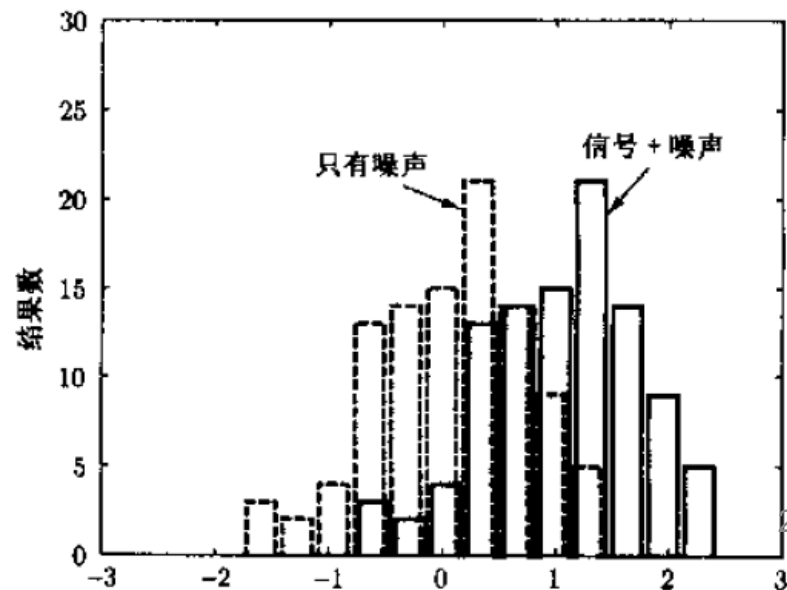
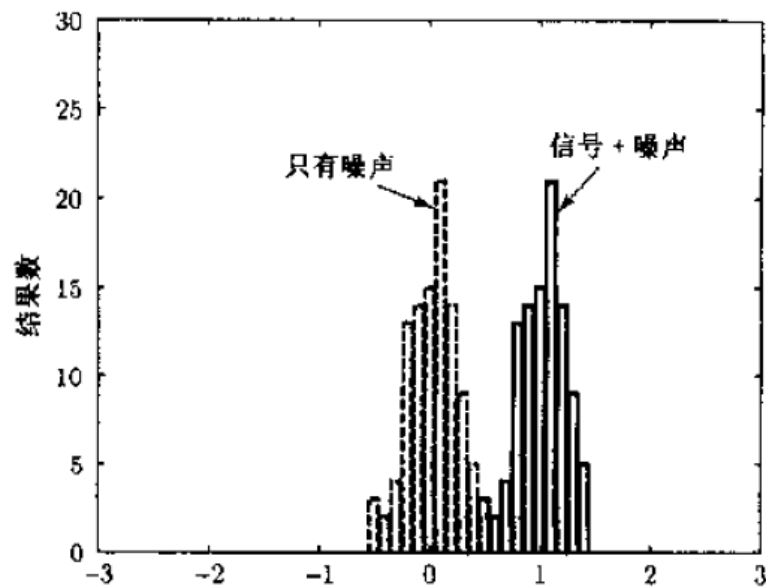
$$P_e = Q\left(\sqrt{\frac{\bar{\varepsilon}(1 - \rho_s)}{2\sigma^2}}\right)$$



检测性能



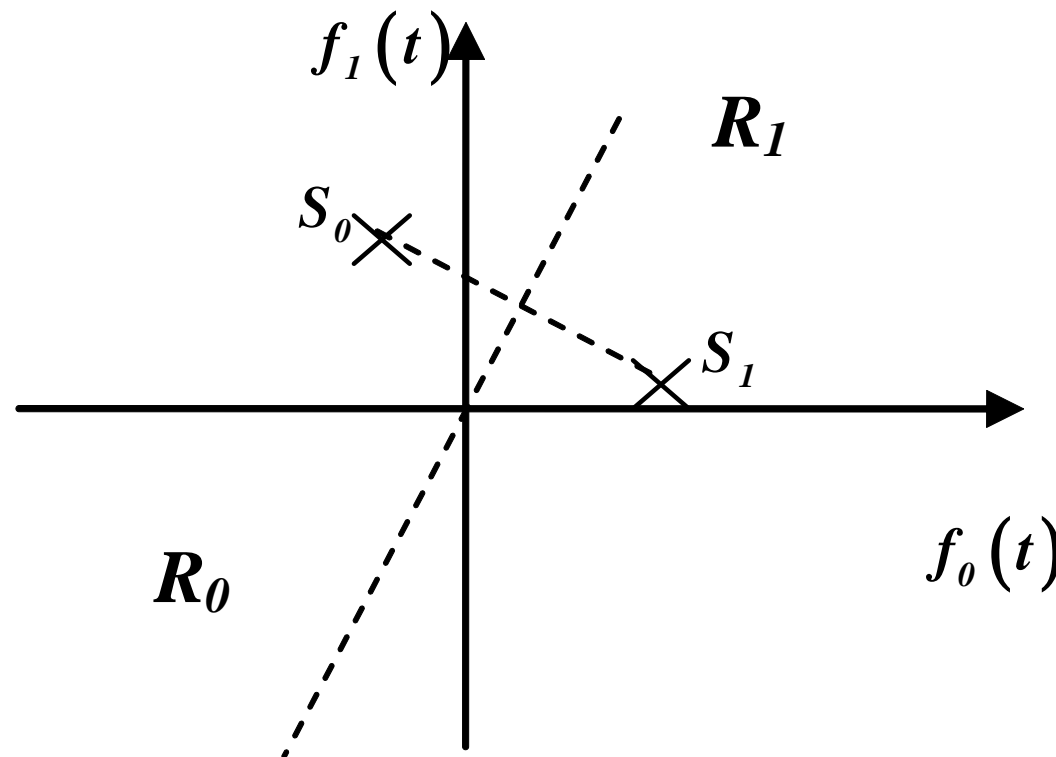
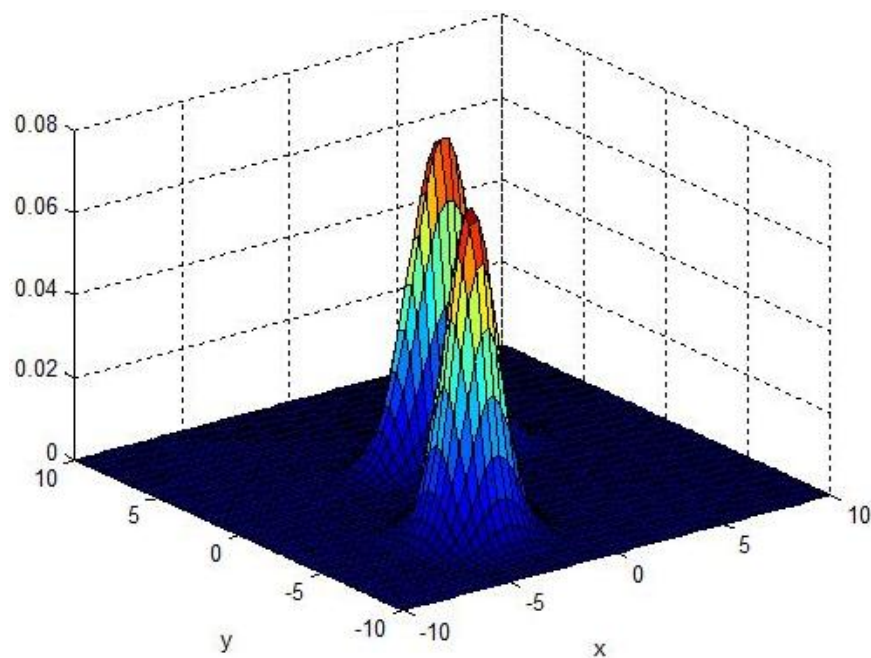
$$\begin{cases} x=1+n; \\ x=n \end{cases}$$



最小欧氏距离判决准则 (ML)

$$R_1 = \{Y : \|Y - S_1\|^2 < \|Y - S_0\|^2\}$$

$$R_0 = \{Y : \|Y - S_1\|^2 \geq \|Y - S_0\|^2\}$$





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似然函数

$$Y=[Y_1, Y_2, \dots, Y_n]^T$$

$$\text{均值: } E(Y \setminus H_i) = [E\{Y_1\}, E\{Y_2\}, \dots, E\{Y_n\}]^T = [m_{Y_{i1}}, m_{Y_{i2}}, \dots, m_{Y_{in}}]^T$$

$$\text{协方差矩阵: } Cov(Y \setminus H_i) = \begin{bmatrix} E\left\{\left(Y_1 - m_{Y_{i1}}\right)^2\right\} & \dots & \dots & E\left\{\left(Y_1 - m_{Y_{i1}}\right)\left(Y_n - m_{Y_{in}}\right)\right\} \\ \vdots & \dots & \dots & \vdots \\ \vdots & \dots & \dots & \vdots \\ E\left\{\left(Y_n - m_{Y_{in}}\right)\left(Y_1 - m_{Y_{i1}}\right)\right\} & \dots & \dots & E\left\{\left(Y_n - m_{Y_{in}}\right)^2\right\} \end{bmatrix}$$

$$\text{似然函数: } f(Y \setminus H_i) = \frac{1}{(2\pi)^{n/2} |C_i|^{1/2}} \exp\left\{-\frac{1}{2}[(Y - m_i)^T C_i^{-1} (Y - m_i)]\right\}$$



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检验统计量

- 似然比 $L(Y) = \frac{f(Y \mid H_1)}{f(Y \mid H_0)} \geq th$
- 对数似然比 $l(Y) = (Y - m_0)^T C_0^{-1} (Y - m_0) - (Y - m_1)^T C_1^{-1} (Y - m_1) \geq 2 \ln th + \ln |C_1| - \ln |C_0|$

- 等协方差

$$\Rightarrow 2\Delta m^T C^{-1} Y + m_0^T C^{-1} m_0 - m_1^T C^{-1} m_1 \geq 2 \ln th$$

其中 $\Delta m = m_1 - m_0$

- 判决准则

$$T(Y) = \Delta m^T C^{-1} Y \geq \ln th + \frac{1}{2} (m_0^T C^{-1} m_0 - m_1^T C^{-1} m_1)$$



检验统计量的分布

$$E\{T(Y \setminus H_0)\} = \Delta m^T C^{-1} E(Y \setminus H_0) = \Delta m^T C^{-1} m_0$$

$$\begin{aligned} C\{T(Y \setminus H_0)\} &= E\left\{\left[\Delta m^T C^{-1} Y - \Delta m^T C^{-1} m_0\right]\left[\Delta m^T C^{-1} Y - \Delta m^T C^{-1} m_0\right]^T\right\} \\ &= \Delta m^T C^{-1} E\left\{\left[Y - m_0\right]\left[Y - m_0\right]^T\right\} C^{-1} \Delta m = \Delta m^T C^{-1} \Delta m \end{aligned}$$

同理 $E\{T(Y \setminus H_1)\} = \Delta m^T C^{-1} m_1$

$$C\{T(Y \setminus H_0)\} = \Delta m^T C^{-1} \Delta m$$

$$d^2 = \frac{\left[E(T \setminus H_1) - E(T \setminus H_0)\right]^2}{C(T \setminus H_0)} = \Delta m^T C^{-1} \Delta m$$

Q: 白? 有色?





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检验统计量

$$\begin{aligned} l(Y) &= (Y - m)^T C_0^{-1} (Y - m) - (Y - m)^T C_1^{-1} (Y - m) \\ &= (Y - m)^T (C_0^{-1} - C_1^{-1}) (Y - m) \\ &\geq 2 \ln th + \ln |C_1| - \ln |C_0| \end{aligned}$$

$$\text{令 } m_0 = m_1 = m = 0$$

$$\Delta C^{-1} = C_1^{-1} - C_0^{-1}$$

$$\text{二次型函数 } T(Y) = Y^T \Delta C^{-1} Y \geq 2 \ln th + \ln |C_1| - \ln |C_0|$$

Q: 白? 有色?





$H_0 : y \sim N(m, \sigma_0^2)$, 若 $\sigma_1^2 > \sigma_0^2$ 且先验概率相等, 求检测准则。
 $H_1 : y \sim N(m, \sigma_1^2)$

似然函数

$$f(y | H_0) = \frac{1}{\sqrt{2\pi\sigma_0^2}} \exp \left\{ -\frac{(y-m)^2}{2\sigma_0^2} \right\}$$

$$f(y | H_1) = \frac{1}{\sqrt{2\pi\sigma_1^2}} \exp \left\{ -\frac{(y-m)^2}{2\sigma_1^2} \right\}$$

$$\Rightarrow \frac{\sigma_1^2 - \sigma_0^2}{2\sigma_1^2 \sigma_0^2} (y-m)^2 \stackrel{H_1}{\geq} \ln \frac{\sigma_1}{\sigma_0}$$





当 $y \geq th' + m$ 或 $y \leq -th' + m$ 时，判决 H_1 为真

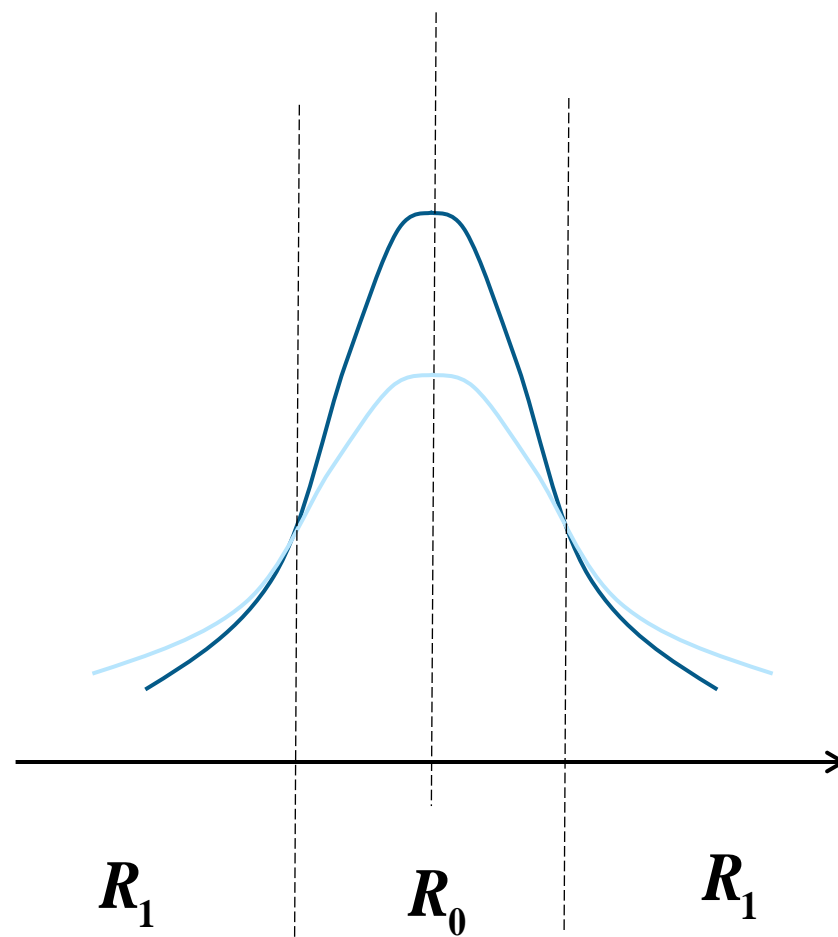
当 $-th' + m < y < th' + m$ 时，判决 H_0 为真

$$\text{其中 } th' = \frac{2\sigma_1^2\sigma_0^2}{\sigma_1^2 - \sigma_0^2} \ln \frac{\sigma_1}{\sigma_0}$$

Q:此时的虚警、漏警概率?

均值为0时有区别吗?

噪声 IID时多样本?



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summary

- ✓ 高斯噪声、参量检测
- ✓ 最佳检测等效于相关接收
- ✓ 检测性能由偏移系数确定

Ref: §3.1-§3.5(赵版)or §3.1-§3.7、 §4.5(KAY版)



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Q

多样本的必要性？

样本数如何影响检测性能？





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