

Inference in Topological Data Analysis

Wolfgang Polonik
Department of Statistics, UC Davis

joint with

Johannes Krebs and Benjamin Roycraft

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Synopsis

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- (i) brief intro to TDA (persistence homology);
- (ii) large sample distribution of persistent Betti numbers (and Euler characteristic process);
- (iii) statistical inference via bootstrap for persistent Betti numbers (and Euler characteristics);
- (iv) discuss statistical insights

TDA - persistent homology

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- ▶ set of feature extraction methods;
- ▶ features are topological/geometric in nature;

Introductory texts with different foci:

Edelsbrunner and Harer (2010), Otter et al. (2017), Wasserman (2018),
Boissonnat et al. (2018), Rabadan and Blumberg (2019), Chazal and Michel
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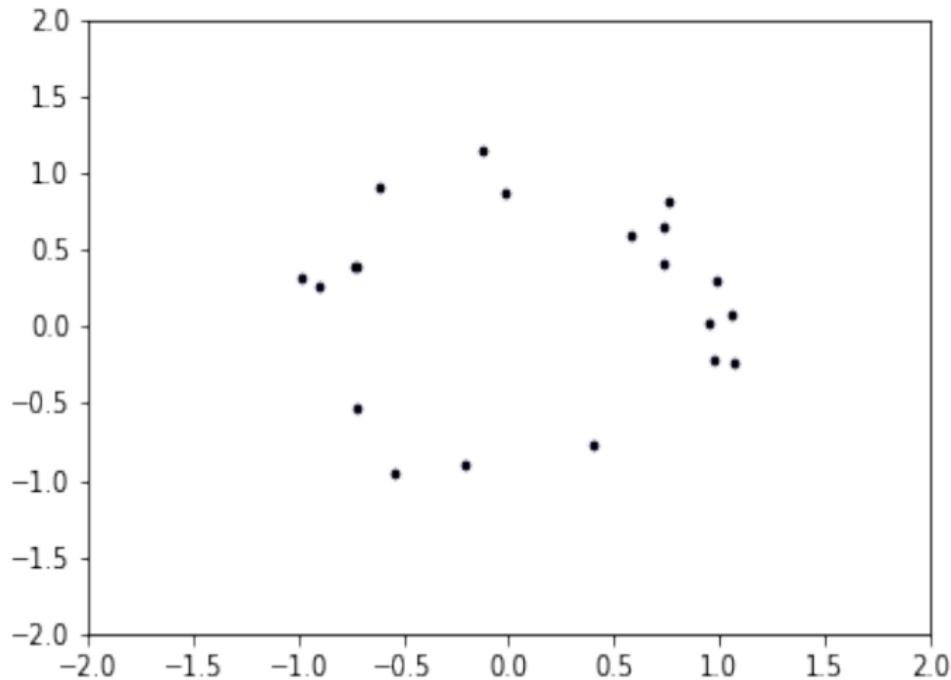
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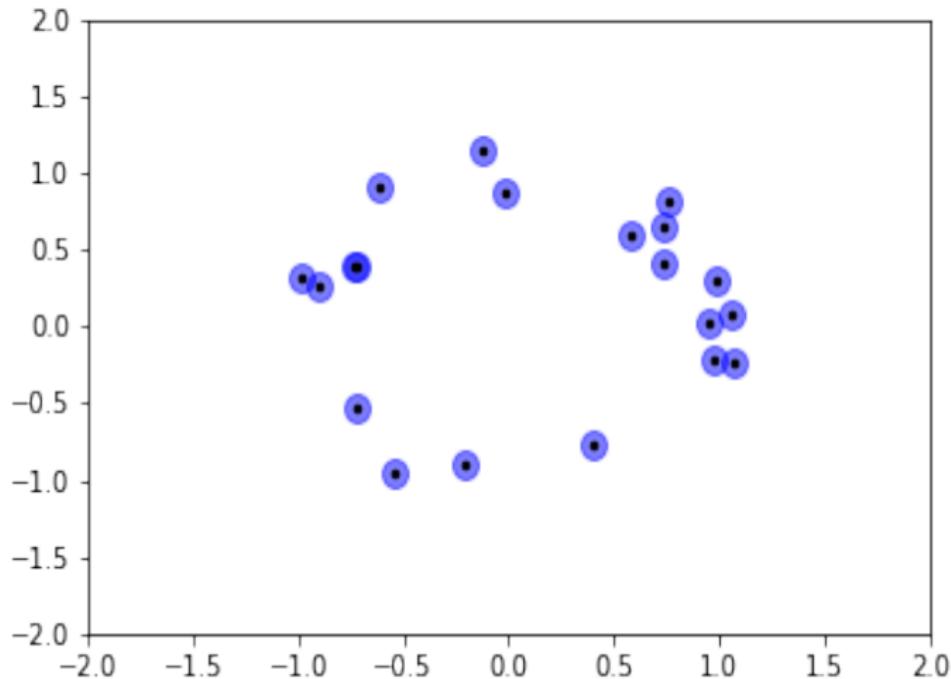
Recent surveys of applications in specific fields:

- Bukkuri et al. (2021) in oncology,
- Dłotko et al. (2019) in financial time series,
- Rabadan and Blumberg (2019) in genomics and evolution,
- Davies (2022) in cyber security,
- Amézquita et al. (2020) in biology,
- Smith et al. (2021) in chemical engineering,
- Joshi and Joshi (2019) in big data in health care, and
- Salch et al. (2021) discuss TDA methods in biomedical imaging.

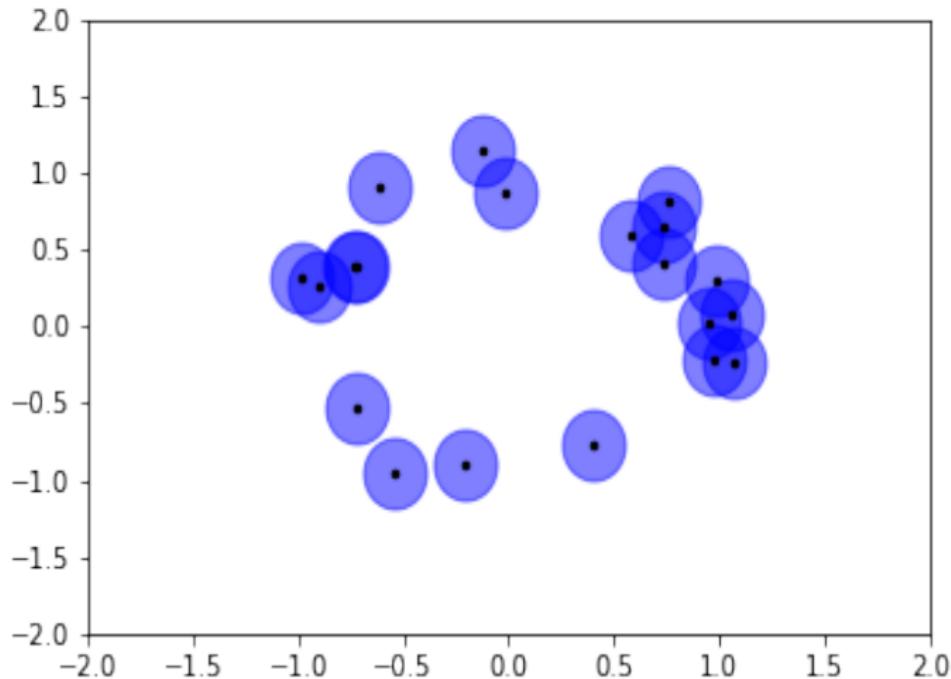
The persistence diagram based on the Čech filtration



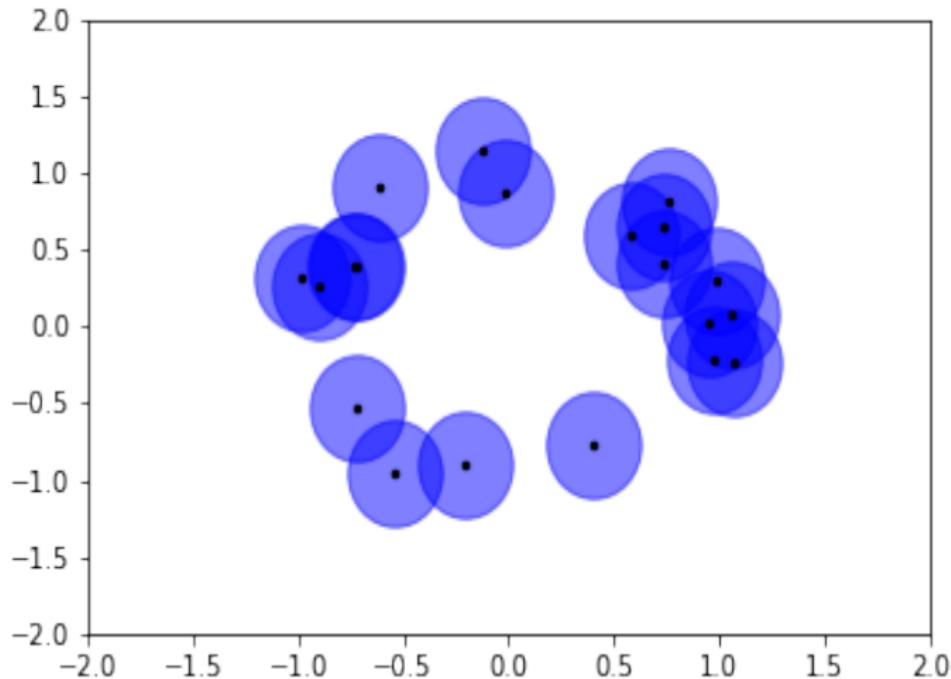
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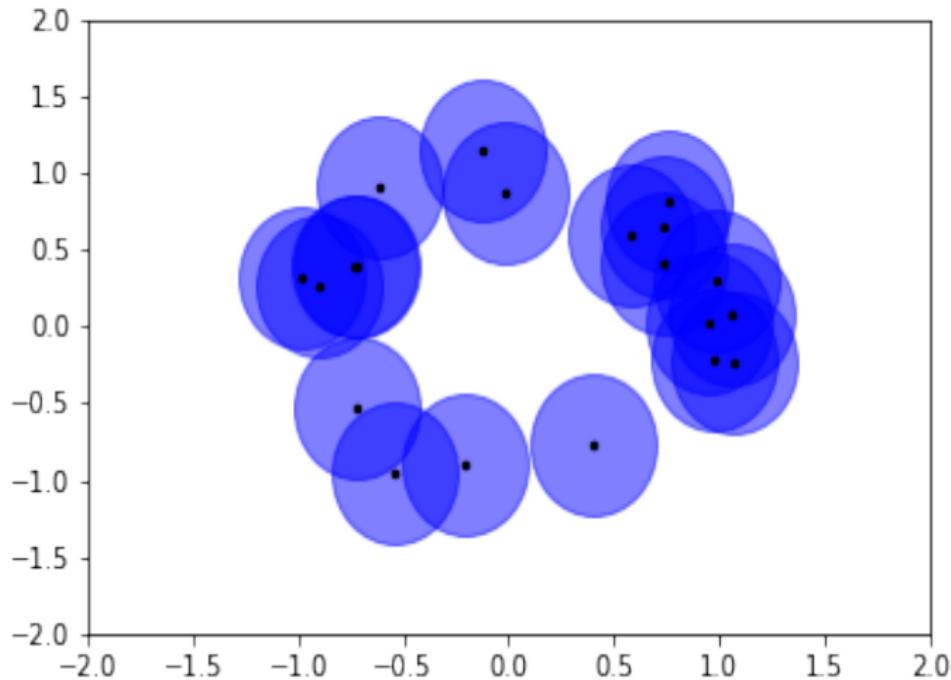
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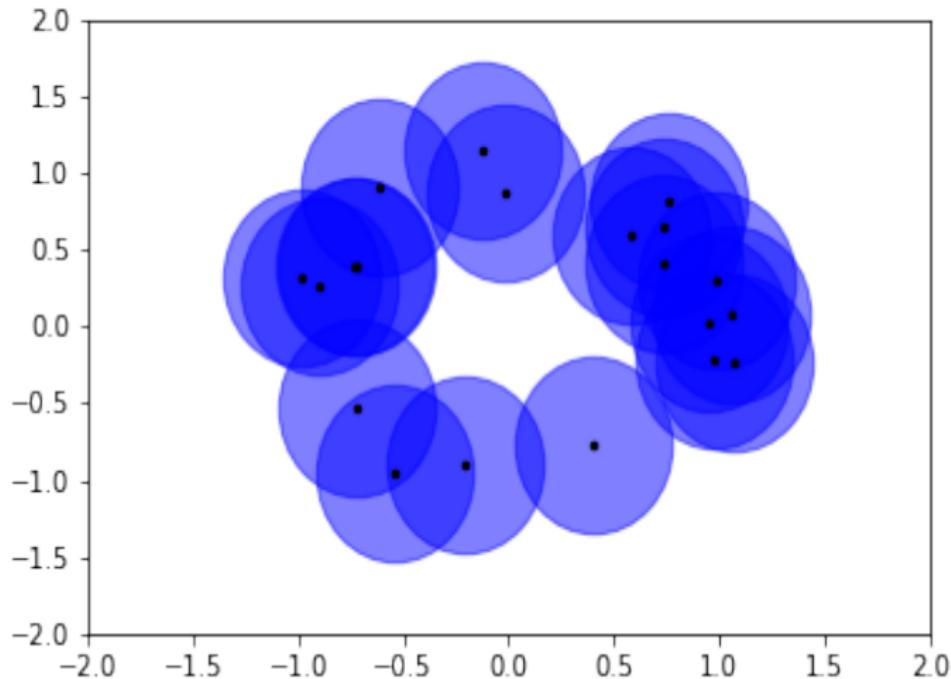
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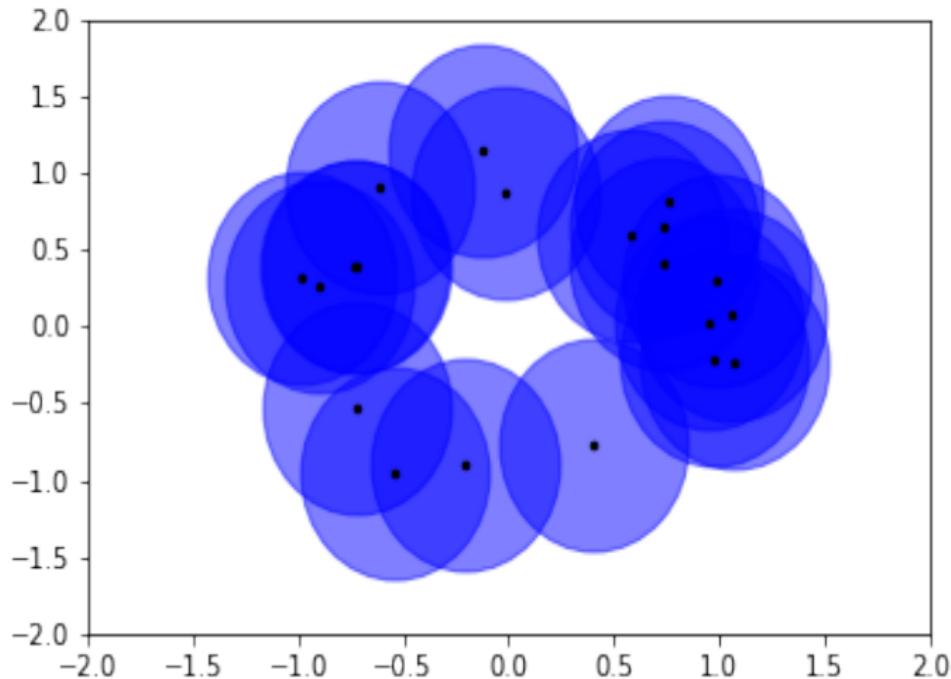
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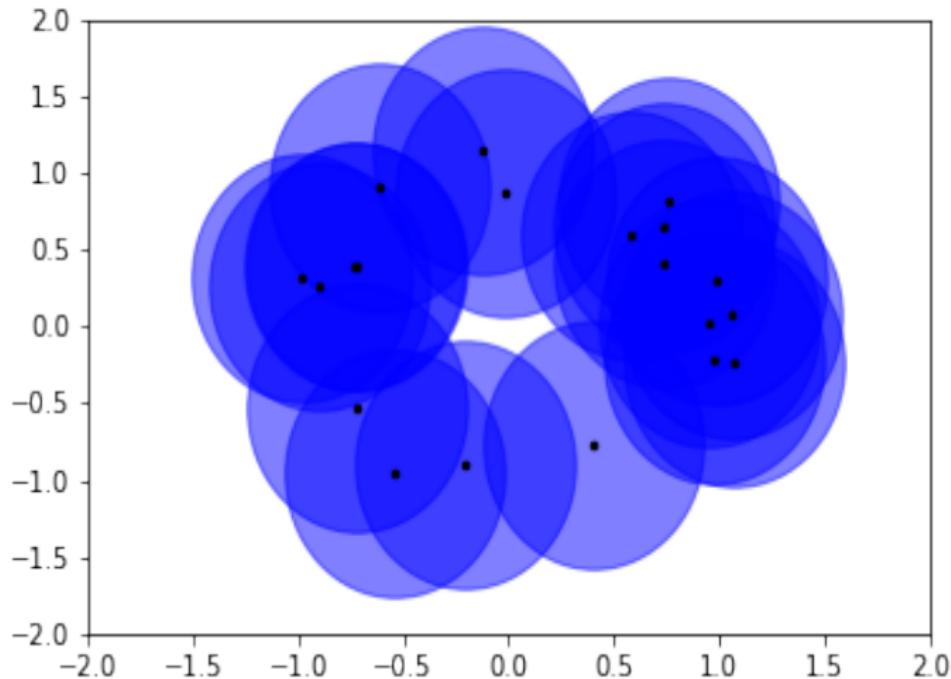
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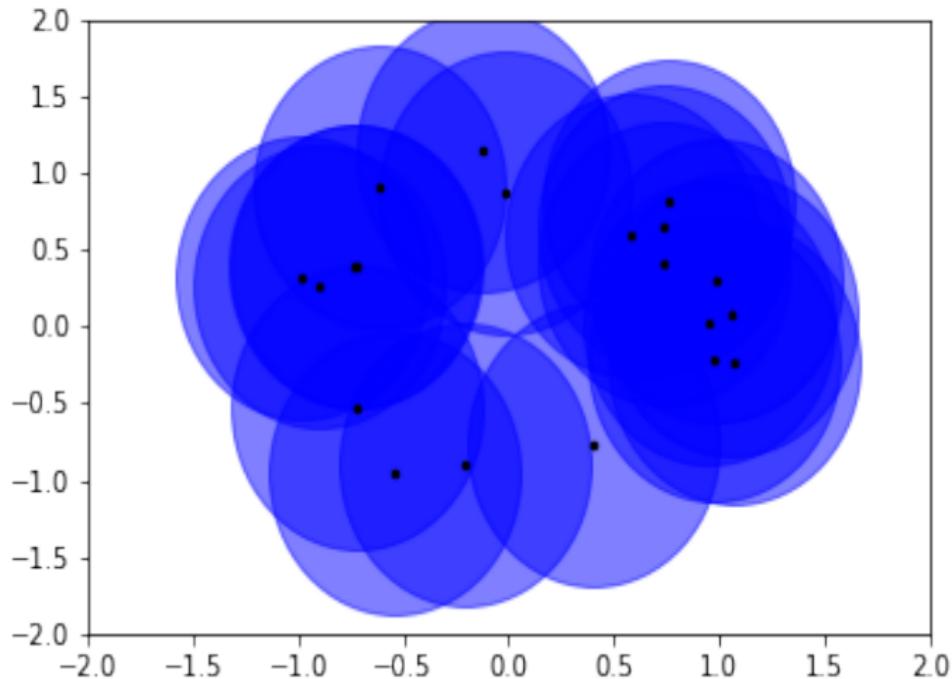
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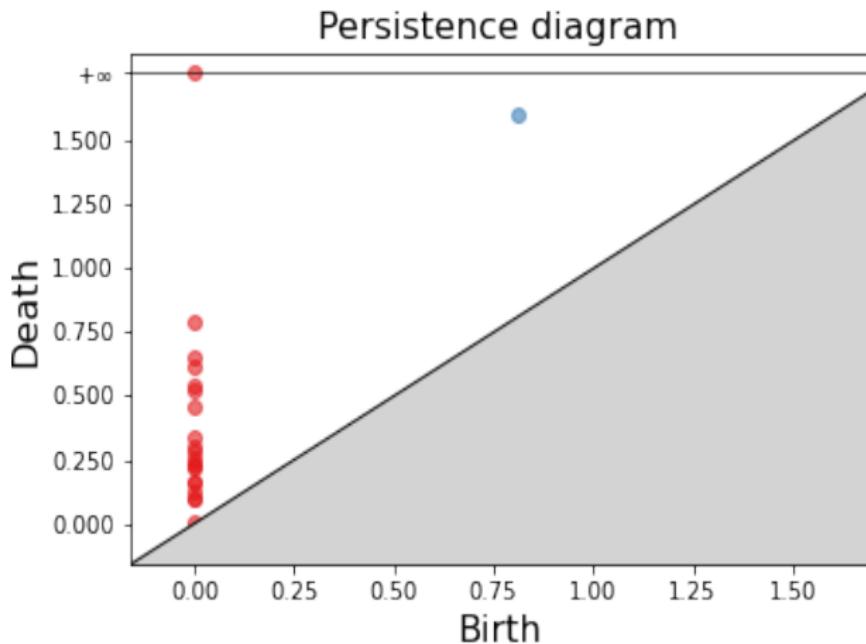
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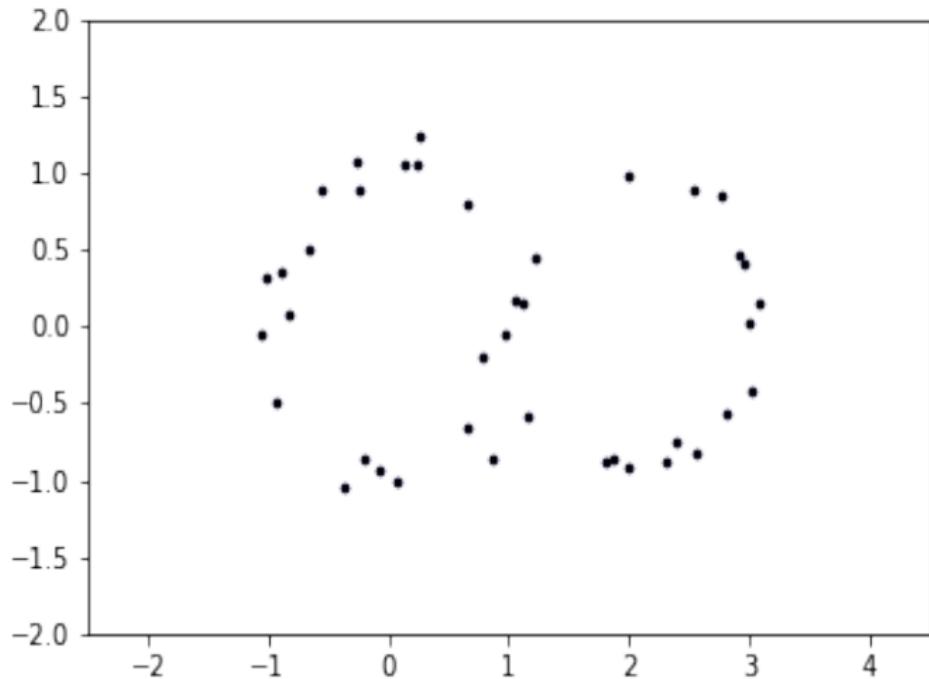
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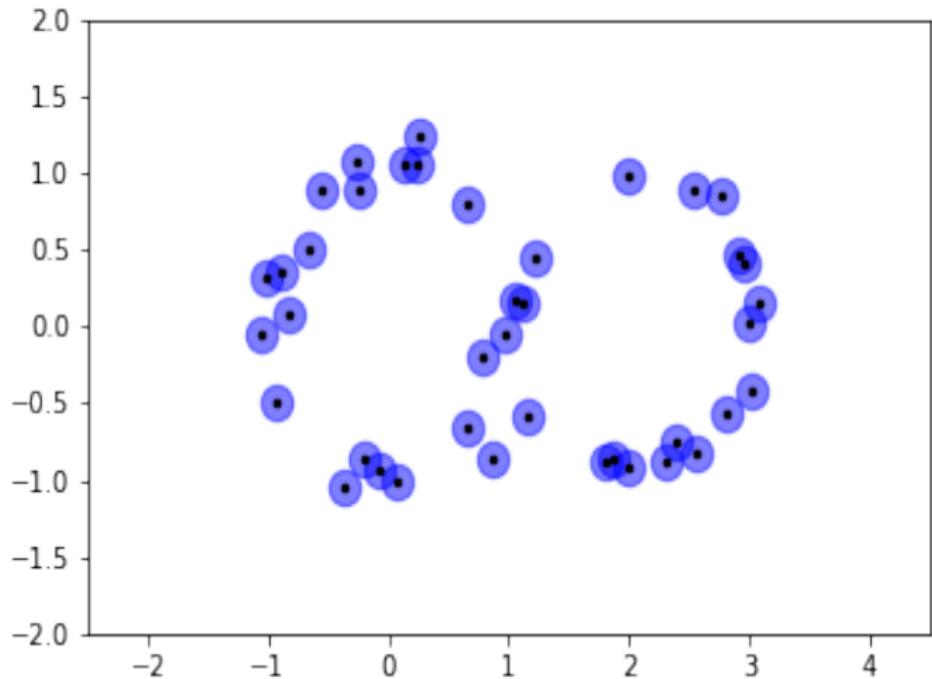
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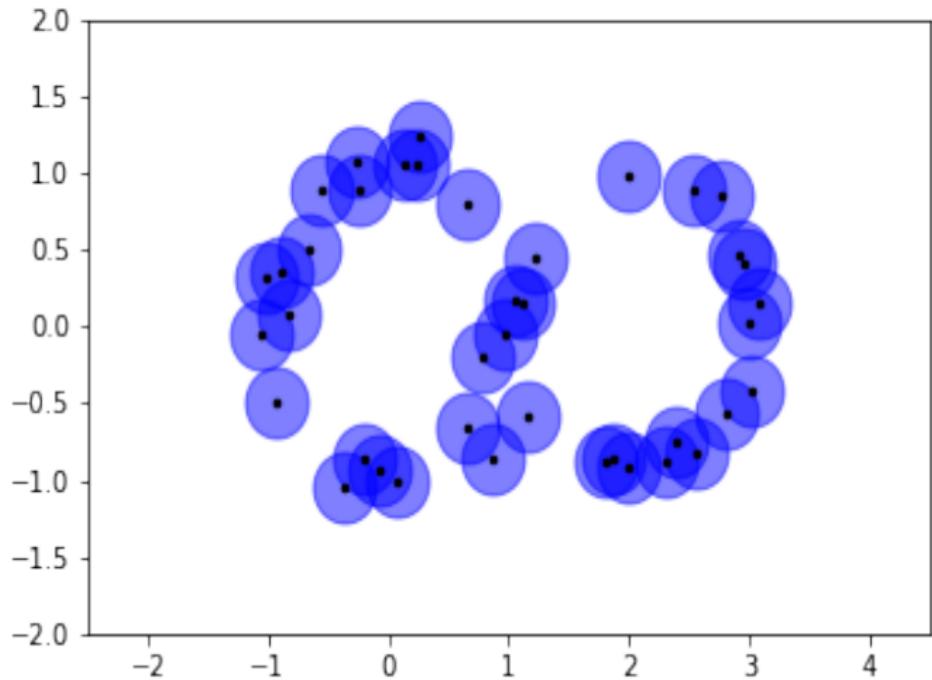
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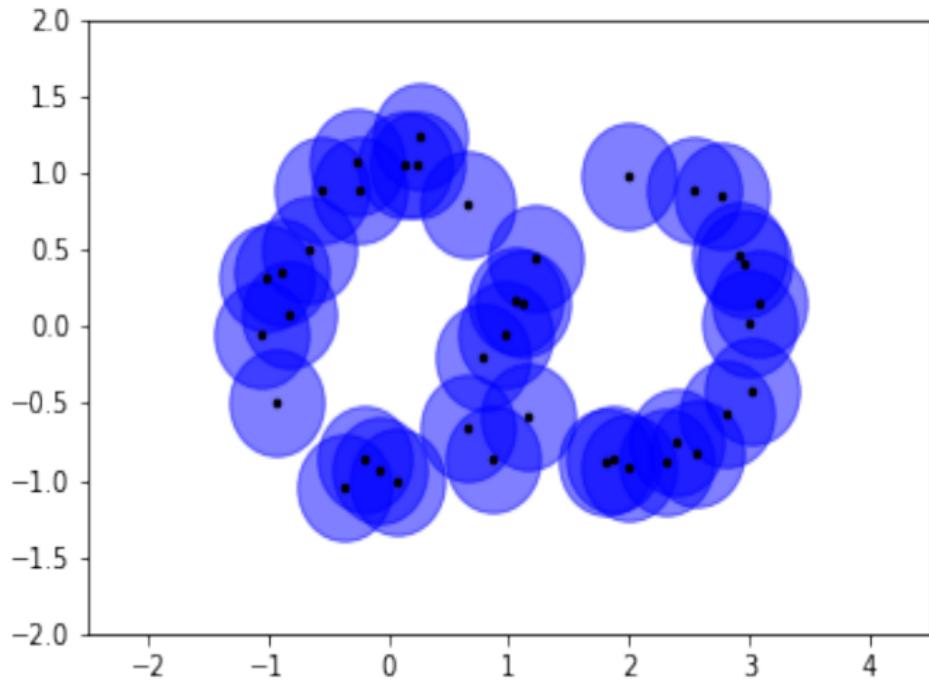
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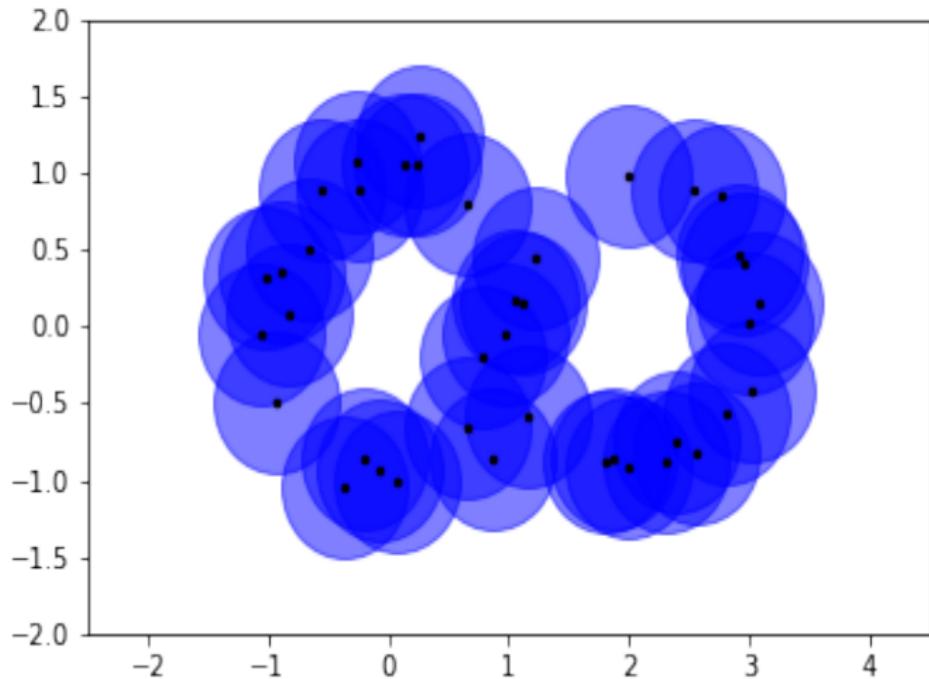
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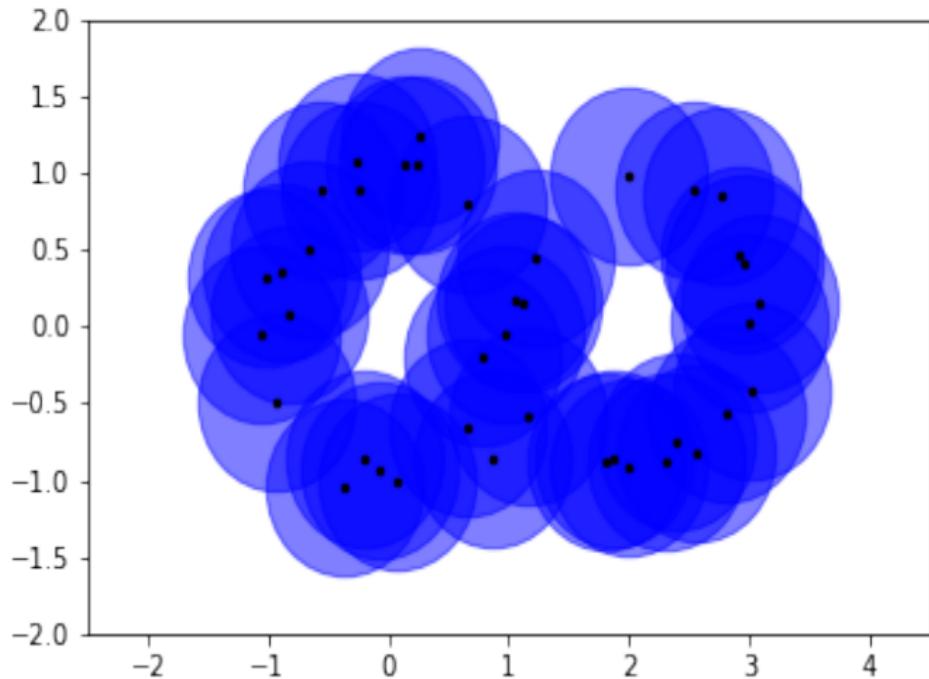
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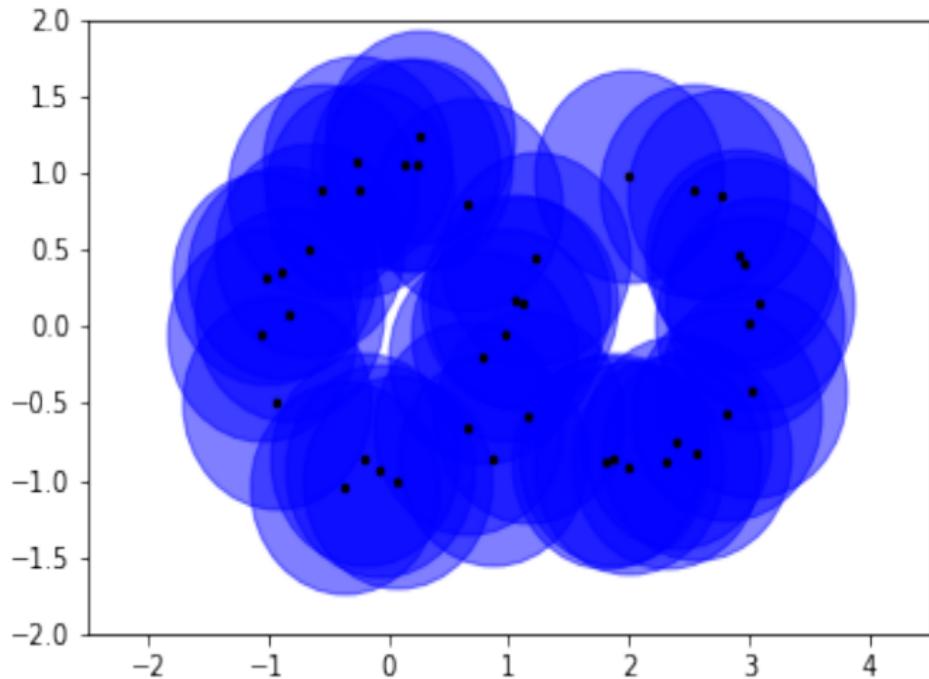
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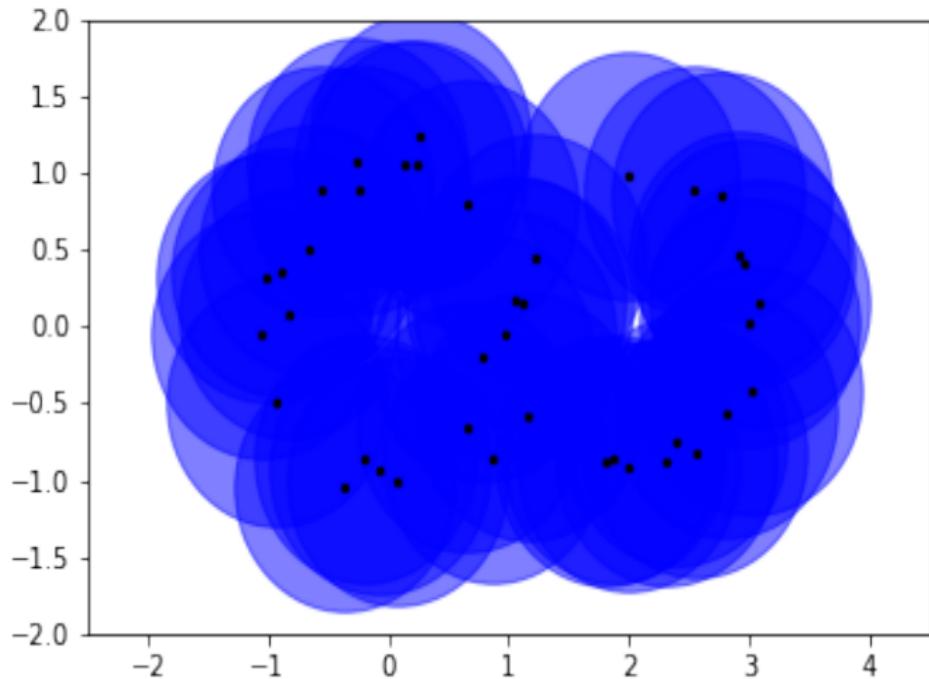
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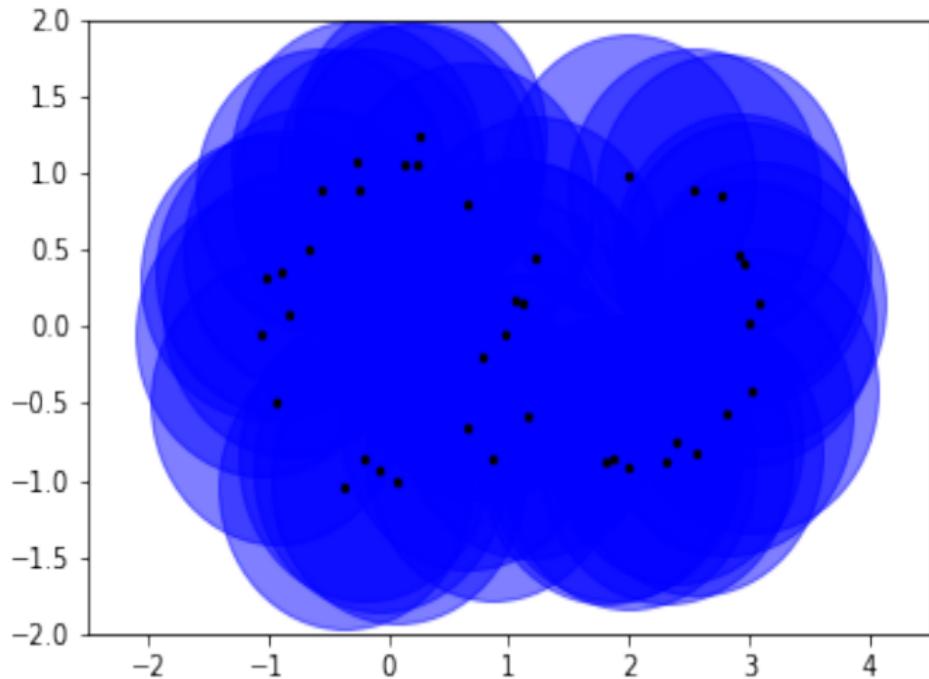
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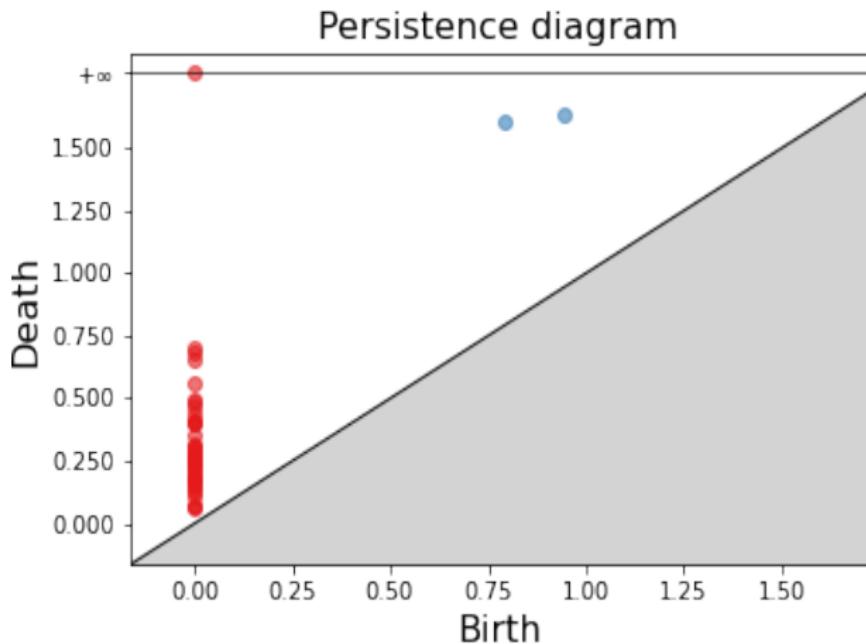
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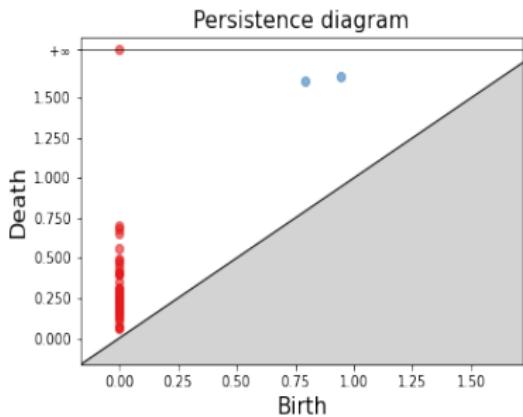
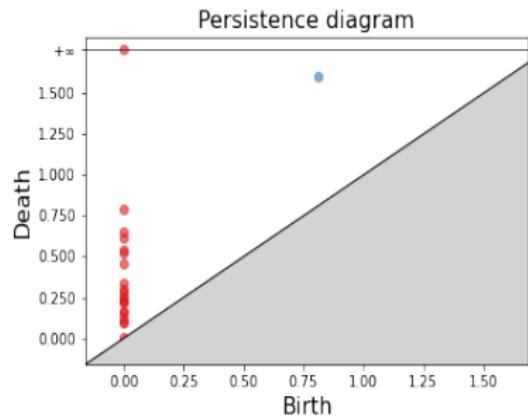
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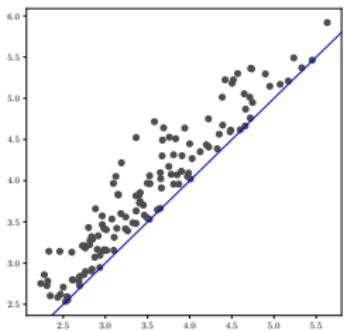
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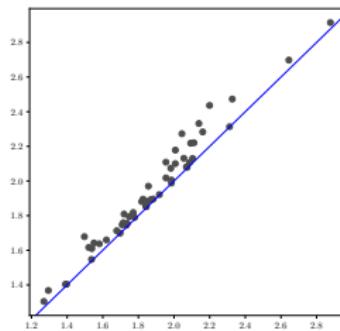
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Persistence diagram for multivariate normal



1-dim. features



2-dim. features

- PD based on sample of size 200 from a 6-dimensional standard normal.
- Each point corresponds to a 1-dim. hole or a 2-dim. hole, respectively.
- There might be higher-dimensional features (holes)

Filtrations

- Let $\mathcal{B}_r = \bigcup_{i=1}^n B_d(x_i, r) \subset \mathbb{R}^d$; then

$$\mathcal{B} = \{\mathcal{B}_r : r \geq 0\}$$

defines a **filtration** of a topological space (in this case of \mathbb{R}^d): For any $r_1 \leq r_2 \dots \leq r_N$,

$$\mathcal{B}_{r_1} \subseteq \mathcal{B}_{r_2} \subseteq \dots \subseteq \mathcal{B}_{r_N} \subseteq \dots \subseteq \mathbb{R}^d$$

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Filter function for union of balls (called Čech filtration), is

$$f(x) = D_{\mathbb{X}}(x) = \min_{i=1, \dots, n} d(x, X_i) \quad \text{distance function,}$$

for

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Question is: Which filter function f is useful? (Depends on problem at hand.)

Persistence homology

Tracking the 'dynamics' of the topological features (holes, homology) given by the filtration

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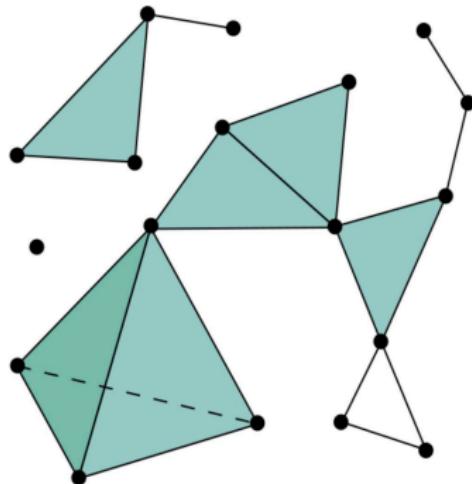
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- Persistence diagram 'summarizes' this dynamics
- Each point in the persistence diagram corresponds to an **equivalence classes of cycles**, each of them 'encircling' the same 'hole'; (k -dimensional holes are encircled by cycles of k -dimensional simplices)

Simplicial complexes

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Definition (abstract simplicial complex)

Given a finite set V , an abstract simplicial complex \mathcal{C} with vertex set V is a collection of subsets of V such that

- (i) each element of V lies in \mathcal{C} ;
- (ii) $D \in \mathcal{C}$ and $C \subset D \Rightarrow C \in \mathcal{C}$.

Each $D \in \mathcal{C}$ is called a *simplex*, and its dimension is $|D| - 1$. The *dimension of \mathcal{C}* is the maximum dimension of the simplices in \mathcal{C} .

The Čech and the Vietoris-Rips filtration

Definition (Čech-complex)

Let (\mathbb{X}, d) be a metric space, and let $\mathbb{X}_n = \{x_1, \dots, x_n\} \subset \mathbb{X}$. For $r \geq 0$, the Čech-complex $\mathcal{C}_r(\mathbb{X}_n)$ at scale r over \mathbb{X}_n is the abstract simplicial complex given by:

$$[x_0, \dots, x_k] \in \mathcal{C}_r(\mathbb{X}_n) \Leftrightarrow \bigcap_{i=0}^k \overline{B}_r(x_i) \neq \emptyset.$$

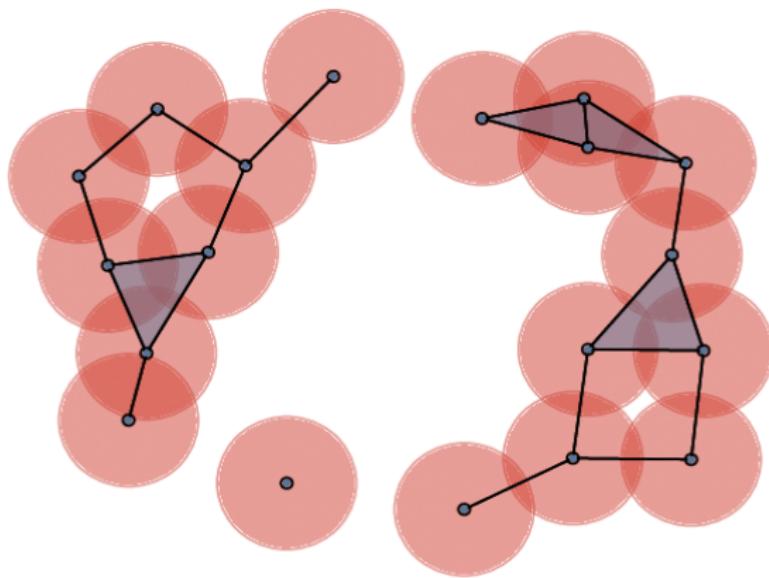
Definition (VR-complex)

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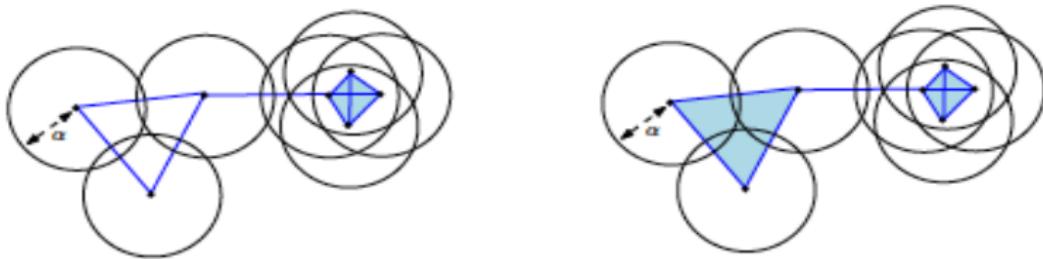
$$[x_0, \dots, x_k] \in \text{VR}_r(\mathbb{X}_n) \Leftrightarrow d(x_i, x_j) \leq r \quad \forall 0 \leq i \leq j \leq k.$$

The Čech complex

Note: The Čech complex $\mathcal{C}_r(\mathbb{X})$ is homotopy equivalent to the union of balls $\bigcup_{i=1}^n \overline{B}_r(x_i)$. (Nerve Theorem)

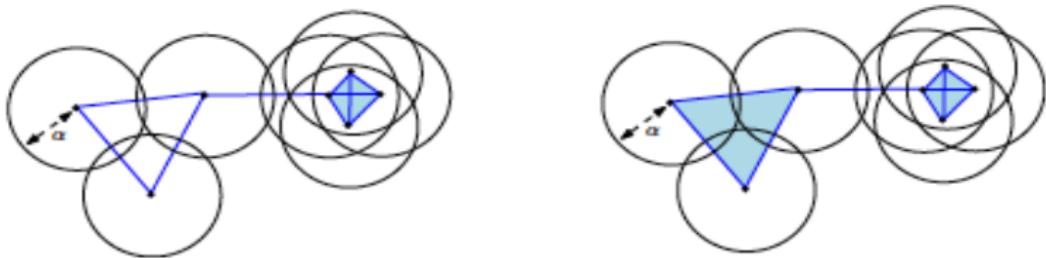


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Čech complex at scale r (left) and VR-complex at scale $2r$ (right).

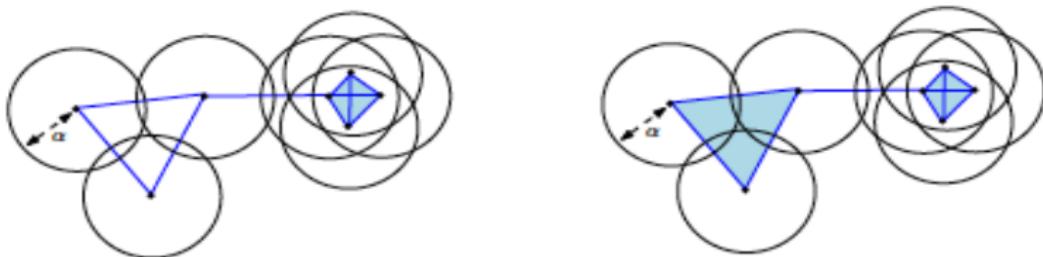
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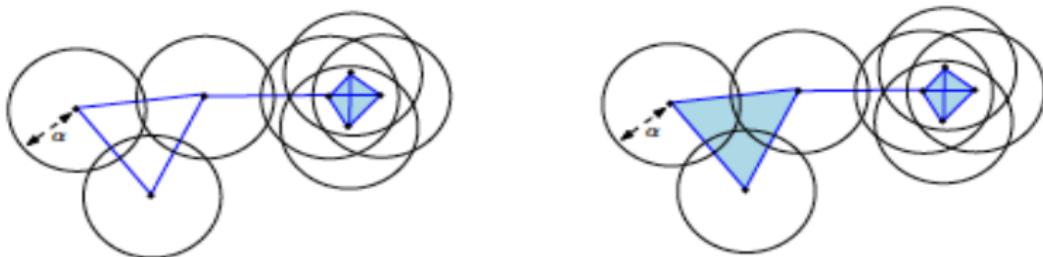
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Proposition: Let \mathbb{X}_n be a finite set of points in \mathbb{R}^d . For any $\alpha \geq 0$,

$$\text{VR}_r(\mathbb{X}_n) \subset \mathcal{C}_r(\mathbb{X}_n) \subset \text{VR}_{2r}(\mathbb{X}_n).$$

Filtrations: The collections $\{\mathcal{C}_r(\mathbb{X}_n) : r \geq 0\}$ and $\{\text{VR}_r(\mathbb{X}_n) : r \geq 0\}$ are called Čech and VR-filtration, respectively.

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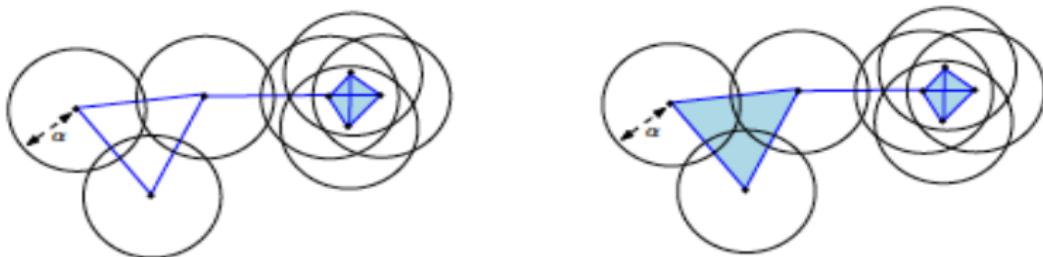
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General filtration of a simplicial complex C : Increasing sequence

$$C_1 \subseteq C_2 \subseteq \cdots \subseteq C_N = C.$$

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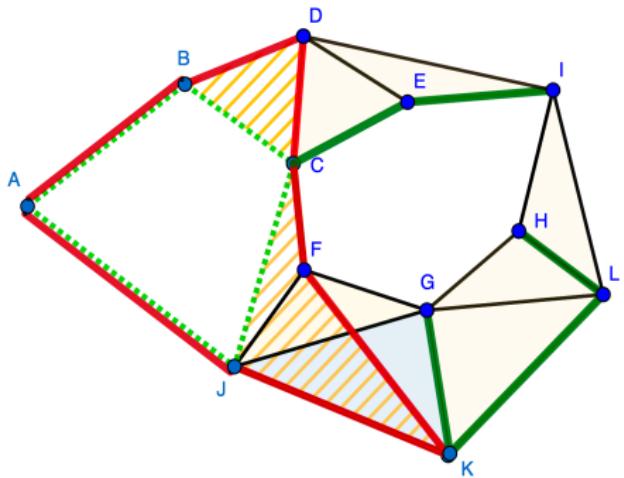
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Add one simplex at a time \rightsquigarrow algorithm.

Cycles and boundaries

Cycles and boundaries

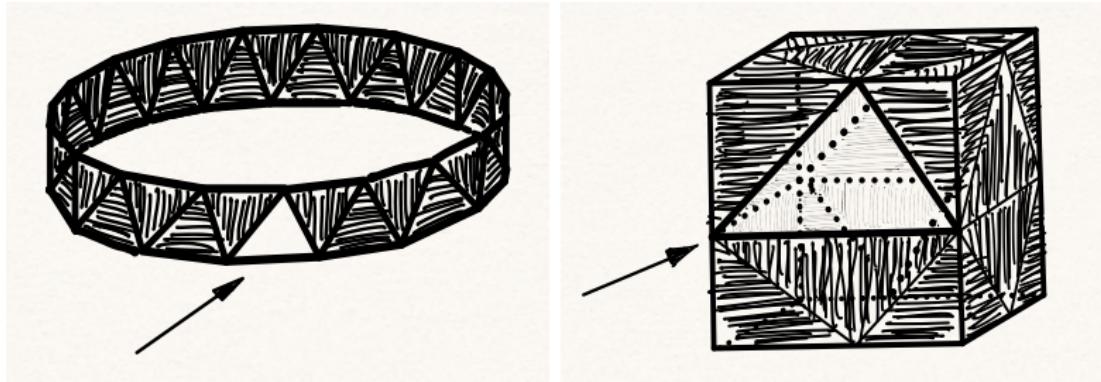
- $Z_k(K) \subset C_k$ (kernel of $\partial_k : C_k \rightarrow C_{k-1}$); cycles.
- $B_k(K) \subset C_k$ (image of $\partial_{k+1} : C_{k+1} \rightarrow C_k$); boundaries of $(k+1)$ -chains



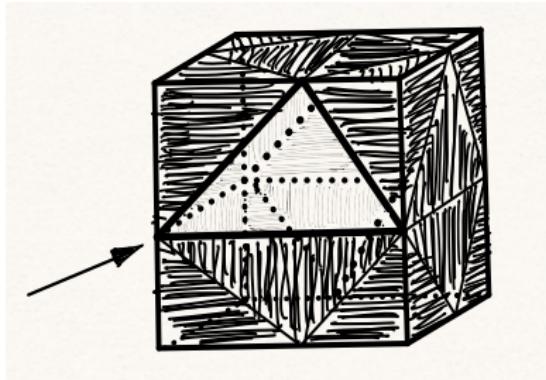
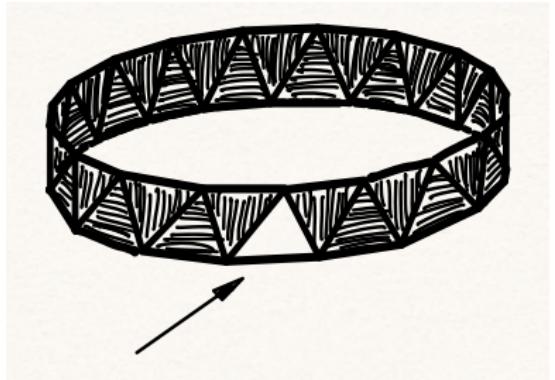
$H_k = Z_k / B_k$
(k -th homology group);
 $\text{rank}(H_k) = k\text{-th Betti}$
number

Both the solid red and the dashed green 1-chains are cycles. Their sum (or difference) is a boundary of a 2-chain (hatched). The 1-chain in bold green is not a cycle.

Example



Example



- How many 'holes' are here?
- What happens when add simplices closing the 'holes' ?

Inference for persistent homology I

Biscio, C.A.N. and Møller, J. (2019): The accumulated persistence function, a new useful functional summary statistic for topological data analysis, with a view to brain artery trees and spatial point process applications, *J. Comput. Graphical Statist.*, **28**, 671 - 681.

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Inference for persistent homology II

Garside, K., Gjoka, R., Henderson, H., Johson, H. and Makarenko, I. (2021): Event history and topological data analysis. *Biometrika*, **108**, 757 - 773.

Krebs, J. and WP (2019): On the asymptotic normality of persistent Betti numbers. <https://arxiv.org/abs/1903.03280>

Robinson, A. and Turner, K. (2017): Hypothesis testing for topological data analysis, *Journal of Applied and Computational Topology*, **1**, 241 - 261.

Krebs, J., Roycroft, B. and WP (2021): On approximation theorems for the Euler characteristic with applications to the bootstrap. *Electronic J. Stat.*, **15**, 4462-4509.

Roycroft, B., Krebs, J., and WP (2022): Bootstrapping Persistent Betti Numbers and Other Stabilizing Statistics. To appear in *Ann. Statist.*
<https://arxiv.org/abs/2005.01417>

Vejdemo-Johansson, M. and Mukherjee, S. (2022): Multiple hypothesis testing with persistent homology. *Foundations of Data Science*, **4**, 667 - 705.

Analyses of TDA methodologies: Challenges

Conceptual:

1. What is the information contained in a persistence diagram for a given filtration? (Topology of support? Shape of filter function (e.g. density)? Dependence of observed data? Which one is important?...)
2. Different filtrations result in different persistence diagrams with different behavior and different information; different tools are needed for their respective analysis;
3. How to compare persistence diagrams?

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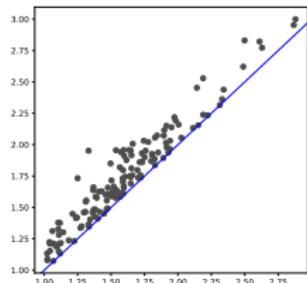
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Technical:

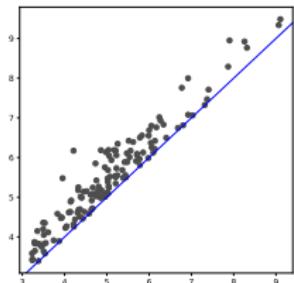
1. How to control *dependence* of points in persistence diagram?
2. What is the population counterpart of quantities of interest?
3. What is the 'right' asymptotic?
4. Can we conduct (asymptotically) valid statistical inference?

PDs for multivariate normals

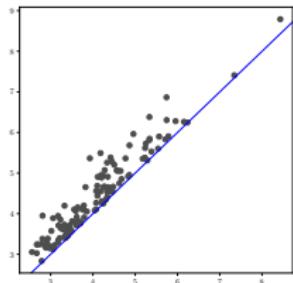
PDs (1-dim holes) for samples of size 200 from a 6-dimensional normal; diagonal covariance matrix; diagonal entries (variances) in captions



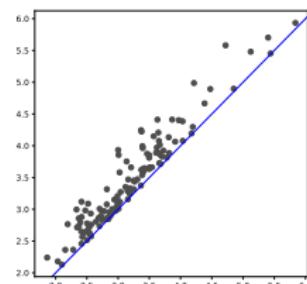
1, 1, 1, 1, 1, 1



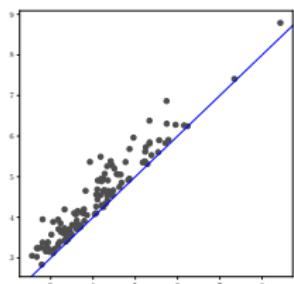
10, 10, 10, 10, 10, 10



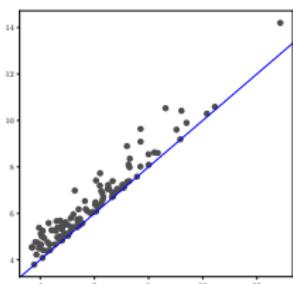
20, 1, 10, 10, 10, 2



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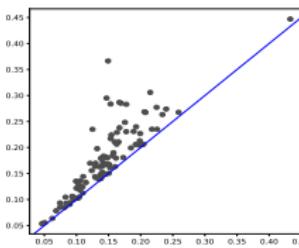
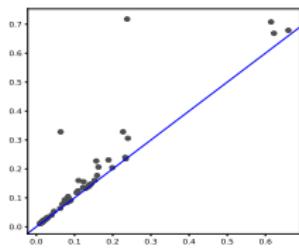
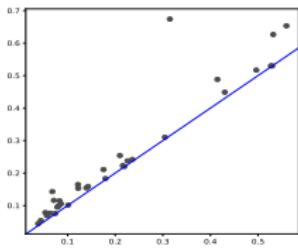
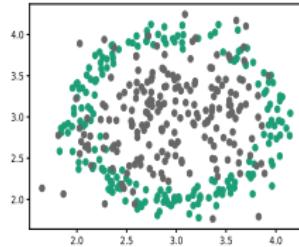
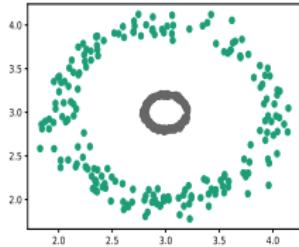
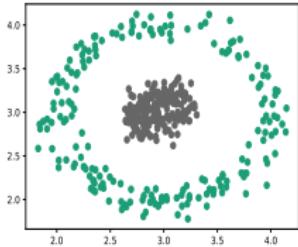
2, 1, 10, 10, 10, 20



1, 10, 10, 10, 10, 20

Nested circles

PDs (1-dim holes); samples of size 200



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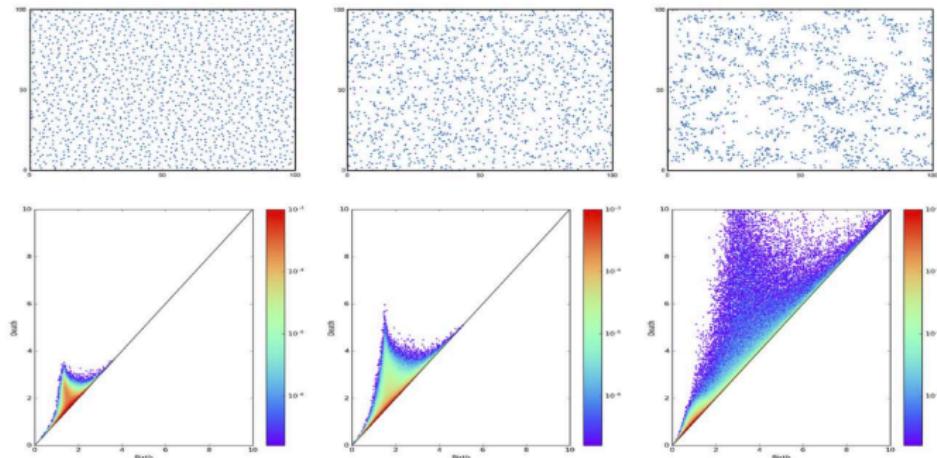


FIG. 1. Top: Point processes with negative (Ginibre), zero (Poisson), and positive (Poisson cluster) correlations. In these three point processes, the number of points and the density are set to be 1,000,000 and $1/2\pi$, respectively. Bottom: The normalized persistence diagrams $\xi_{1,L}/L^2$ of the above.

figure from Hiraoka et al. (2018)

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- Gaining insight into how the sampling distribution influences the shape of the PD \rightsquigarrow statistical inference.

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See also Aaromi et al. (2021), who study the dependence of shape of PD (using persistence landscapes) on the trace of the covariance matrix of d -dimensional observations.

Our set-up

- persistence diagrams are based on either the Čech or the VR-filtration
- $X_1, X_2, \dots, X_n, \dots$ iid from F in \mathbb{R}^d .

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We analyze topological noise based on iid data!

think of a null-model

Signatures of PDs - or, extracting features from features

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- ▶ Betti-curves (Bubenik and Kim, 2007)
- ▶ persistent Betti functions (Edelsbrunner et al., 2010)
- ▶ persistent homology transform (Curry et al., 2021)
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- ▶ persistence landscapes (Bubenik, 2015)
- ▶ persistence image (Adams et al., 2016)
- ▶ persistence surface/ persistence intensity (Chen et al., 2014)
- ▶ persistence terrace (Moon et al., 2018)
- ▶ methods based on kernel distance (Reininghaus et al., 2016)
- ▶ accumulated persistence function (Biscio and Møller, 2019)
- ▶ envelope embedding (Chevyrev et al. 2018)
- ▶ total persistence
- ▶

We will consider (persistent) Betti functions (and Euler characteristics) in the one-sample set-up.

Persistent Betti curves

$\mathcal{K} = \{K_t, t \in \mathbb{R}\}$ filtration of a simplicial complex, i.e. $K_t \subset K_{t'} \subset K$ for $t \leq t'$.

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Definition (k -th Persistent Betti curve)

For $-\infty < s \leq t < \infty$

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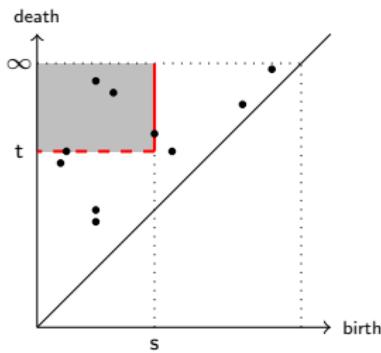
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We also have

$$\beta_k(s, t) = \sum_{(b, d) \in \text{Dgm}_k} 1(b \leq s, d > t).$$



- $\beta_k(t) = \beta_k(t, t)$ is **Betti curve**.

Persistent Betti function

Remarks:

- persistent Betti function can be interpreted as multivariate survival function in particular, it characterizes the persistence diagram
- Betti curves are determined by the marginal distributions of the persistence diagram (distributions of birth and death times, respectively)

Asymptotics for (persistent) Betti curves for VR and Čech filtrations

One sample of size n and $n \rightarrow \infty$

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- increasing n \Rightarrow distances between observations decrease
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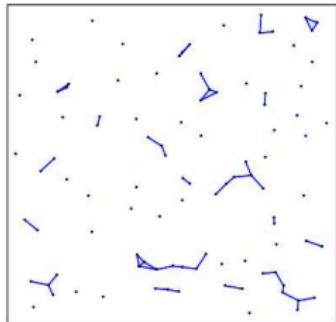
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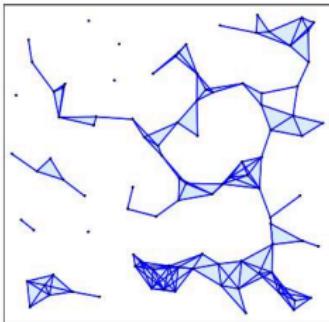
Motivation: behavior of nearest neighbor distance

Three different regimes



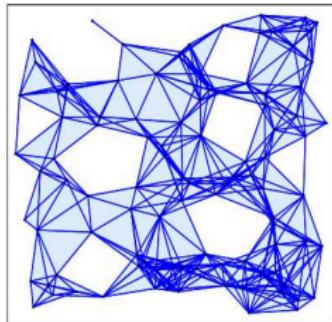
(a) Sparse regime

$$r \ll tn^{-1/d}$$



(b) Thermodynamic regime

$$r = tn^{-1/d}$$



(c) Dense regime

$$r \gg tn^{-1/d}$$

taken from Goel et al., (2019)

Theorem (Krebs and WP, 2019)

Let f be a bounded Lebesgue density on $[0, 1]^d$.

- (i) Let \mathcal{P}_n be a Poisson process on $[0, 1]^d$ with intensity nf . For $p = (s, t) \in \Delta$ let

$$Z_{n,k}(p) = n^{-1/2} (\beta_k(\mathcal{C}_s(n^{1/d}\mathcal{P}_n), \mathcal{C}_t(n^{1/d}\mathcal{P}_n)) - \mathbb{E}\beta_k(\mathcal{C}_s(n^{1/d}\mathcal{P}_n), \mathcal{C}_t(n^{1/d}\mathcal{P}_n)))$$

denote the centered and scaled persistent Betti numbers. For $k = 0, 1, \dots, d-1$, there exist functions $\sigma_k : \Delta \times \Delta \rightarrow [0, \infty)$, such that for any choice $p_1 = (s_1, t_1), \dots, p_m = (s_m, t_m) \in \Delta$, as $n \rightarrow \infty$

$$(Z_{n,k}(p_1), \dots, Z_{n,k}(p_m))' \rightarrow N(0, \Sigma(p_1, \dots, p_m)) \quad \text{in distribution,}$$

where with $X \sim f$, the covariance matrix $\Sigma(p_1, \dots, p_m) \in \mathbb{R}^{m \times m}$ is given by

$$\Sigma_{i,j}(p_1, \dots, p_m) = \mathbb{E} \left[\sigma_k \left[f^{\frac{1}{d}}(X)(s_i, t_i), f^{\frac{1}{d}}(X)(s_j, t_j) \right] \right].$$

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- (ii) Replacing \mathcal{P}_n in part (i) by a Binomial process \mathbb{X}_n with density f gives a similar asymptotic normality result as in (i), but with a covariance matrix $\tilde{\Sigma} \in \mathbb{R}^{m \times m}$ of the form

$$\tilde{\Sigma}_{i,j}(p_1, \dots, p_m) = \Sigma_{i,j}(p_1, \dots, p_m) - \mathbb{E} \left[\alpha \left[f^{\frac{1}{d}}(X)(s_i, t_i) \right] \right] \mathbb{E} \left[\alpha \left[f^{\frac{1}{d}}(X)(s_j, t_j) \right] \right],$$

for some function $\alpha : \Delta \rightarrow \mathbb{R}$, and with $\Sigma_{i,j}$ and X as in part (i).

Putting $s_i = t_i$ for all $i = 1, \dots, m$ gives the joint asymptotic normality for Betti numbers.

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5. Centering: expected Betti number

Convergence of expected Betti number in critical regime

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Theorem (convergence of expected values in the critical regime)

- (i) Let \mathcal{P}_n be a **homogeneous Poisson process** on $[0, 1]^d$ with intensity n . Then, for $k = 0, 1, 2, \dots, d - 1$ there exist functions $\gamma_k : \Delta \rightarrow [0, \infty)$, such that, as $n \rightarrow \infty$,

$$\frac{1}{n} \mathbb{E} \beta_k(\mathcal{C}_s(n^{1/d} \mathcal{P}_n), \mathcal{C}_t(n^{1/d} \mathcal{P}_n)) \rightarrow \gamma_k(s, t).$$

- (ii) Let f be a bounded probability density on $[0, 1]^d$ with compact support. Furthermore, let \mathbb{X}_n be a **Binomial process** on $[0, 1]^d$ with density f , and let \mathcal{P}_n be an **inhomogenous Poisson process** on $[0, 1]^d$ with intensity nf . Then, for $k = 0, 1, \dots, d - 1$ and with $\gamma_k(s, t)$ from part (i), as $n \rightarrow \infty$,

$$\frac{1}{n} \mathbb{E} \beta_k(\mathcal{C}_s(n^{1/d} \mathbb{X}_n), \mathcal{C}_t(n^{1/d} \mathbb{X}_n)) \rightarrow \mathbb{E}[\gamma_k(s f^{1/d}(X), t f^{1/d}(X))]$$

and

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with $X \sim f$. Setting $s = t$ in either (i) or (ii) gives results for the Betti numbers.

See Trinh (2017), Hiraoka et al. (2018), Trinh (2019), Goel et al. (2019), Owada and Thomas (2020).

Discussion of asymptotic means and variances

- In general, the form of the functions γ_k, σ_k and α_k is unknown (only for $k = 0$ more is known).
- Note, however, that these functions are determined by the behavior under a **homogeneous** Poisson sampling \rightsquigarrow they don't depend on the sampling density;

Consequences of this observation: The dependence of the limit on the sampling density f is through quantities of the form

$$\mathbb{E}_f [\Psi_{k,s,t}(f^{1/d}(X))].$$

This will be discussed further below.

For conducting statistical inference, one needs to know the (asymptotic) distribution. So we need to

estimate the limit variance of asymptotic normal.

We will do this using a **bootstrap procedure**.

The bootstrap

A computational device to estimate sampling distributions.

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Basic idea:

Statistical Model

- \mathbb{X}_n is drawn from F
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Bootstrap World

- \mathbb{X}_n^* is drawn from $\hat{F}_n = F_n(\mathbb{X}_n)$
- \hat{F}_n is known
- draw as many bootstrap samples as desired
- estimate sampling distribution of $T_n(F)$ by $T_n(\hat{F}_n)$
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Standard bootstrap: $F_n = \text{empirical distribution given by } \mathbb{X}_n$

Bootstrap confidence regions for Betti numbers for VR and Čech filtrations

Smooth bootstrap: Draw samples from a KDE $f_{n,h}(x)$ based on \mathbb{X}_n

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Recall, with \mathcal{F} either VR or \mathcal{C} ,

$$Z_{n,k}(p, \mathcal{F}) = n^{-1/2} \left(\beta_k(\mathcal{F}_s(n^{\frac{1}{d}} \mathbb{X}_n), \mathcal{F}_t(n^{\frac{1}{d}} \mathbb{X}_n)) - \mathbb{E}[\beta_k(\mathcal{F}_s(n^{\frac{1}{d}} \mathbb{X}_n), \mathcal{F}_t(n^{\frac{1}{d}} \mathbb{X}_n))] \right).$$

The corresponding bootstrap version is

$$Z_{n,k}^*(p, \mathcal{F}) = n^{-1/2} \left(\beta_k(\mathcal{F}_s(n^{\frac{1}{d}} \mathbb{X}_n^*), \mathcal{F}_t(n^{\frac{1}{d}} \mathbb{X}_n^*)) - \mathbb{E}[\beta_k(\mathcal{F}_s(n^{\frac{1}{d}} \mathbb{X}_n^*), \mathcal{F}_t(n^{\frac{1}{d}} \mathbb{X}_n^*)) | \mathbb{X}_n] \right)$$

where $\mathbb{X}_n^* \sim \hat{f}_{n,h}$.

Theorem (bootstrap for persistent Betti numbers)

Let \mathbb{X}_n be a Binomial process in \mathbb{R}^d with intensity f . Fix $k \geq 0$, and let $m \geq 1$ and $p_1, \dots, p_m \in \Delta$. If

- (i) $\|f\|_{2k+3} < \infty$,
 - (ii) $\|\widehat{f}_{n,h} - f\|_q \rightarrow 0$ in probability (or a.s.) as $n \rightarrow \infty$ for some $q > 2k + 3$,
- then, for $\mathcal{F} = \text{VR}$ and $\mathcal{F} = \mathcal{C}$,

$$(\mathcal{Z}_{n,k}(p_1, \mathcal{F}), \dots, \mathcal{Z}_{n,k}(p_m, \mathcal{F}))' \xrightarrow{\mathcal{D}} N(0, \Sigma_m) \quad \text{as } n \rightarrow \infty,$$

if and only if

$$(\mathcal{Z}_{n,k}^*(p_1, \mathcal{F}), \dots, \mathcal{Z}_{n,k}^*(p_m, \mathcal{F}))' \xrightarrow{\mathcal{D}} N(0, \Sigma_m) \quad \text{in probab. (or a.s.) as } n \rightarrow \infty.$$

Sufficient conditions for convergence in probability (for all $q > 0$):

f and kernel K bounded, and $nh^{2d} \rightarrow \infty$.

Real Data Example

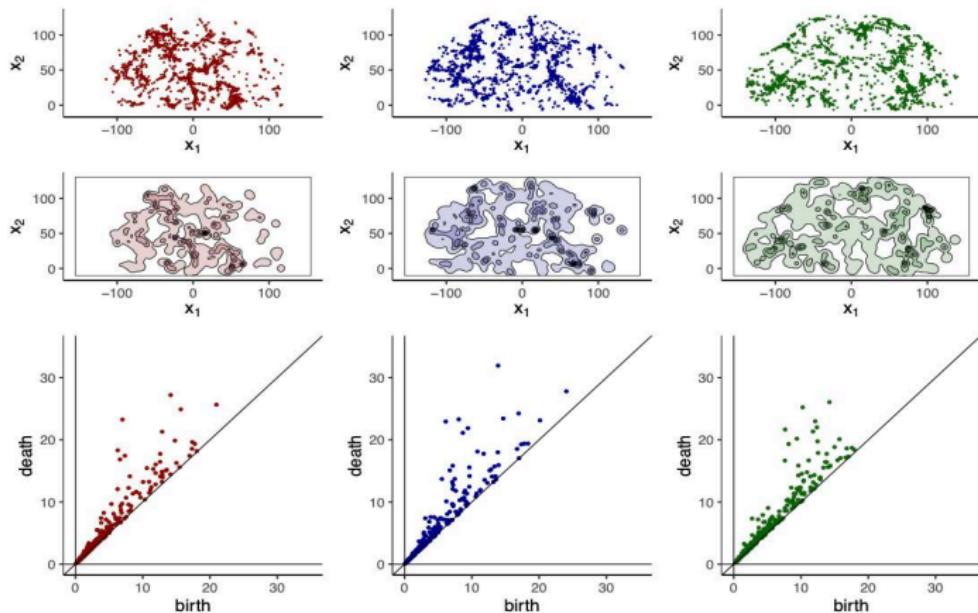
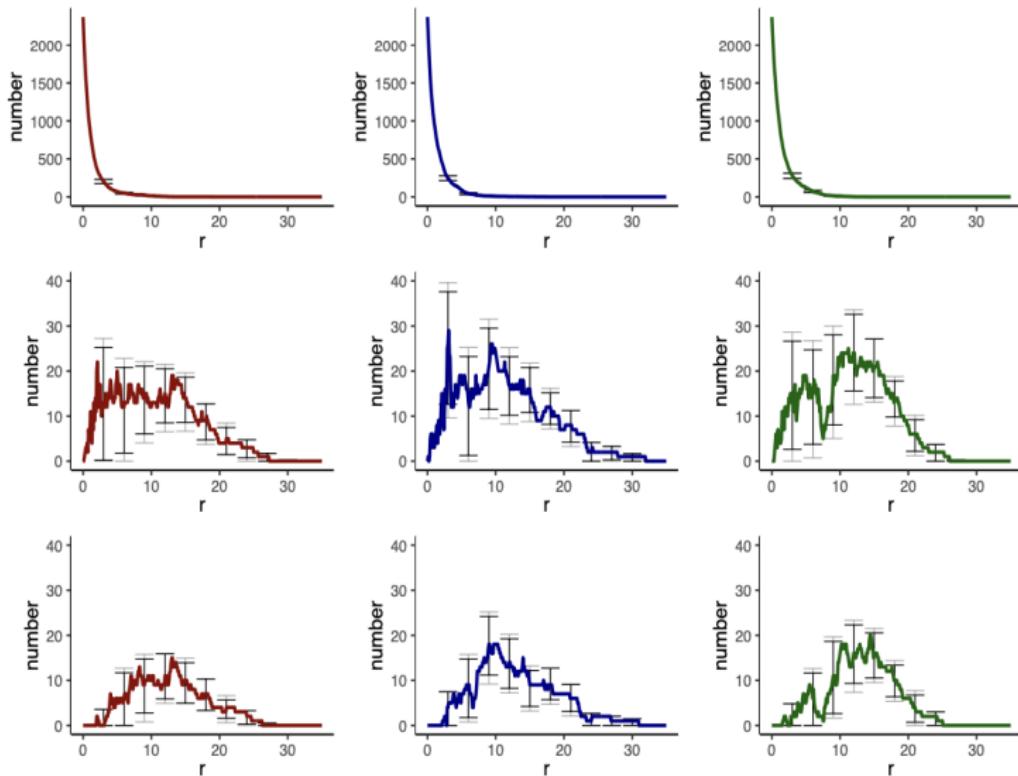


FIG 2. Top row: Transformed point clouds. Middle row: Density estimates using adaptive bandwidth. Bottom row: Persistence diagrams in dimension $q = 1$ for the Vietoris-Rips complex. Columns from left to right: Galaxies with redshifts within $(0.025, 0.026)$, $(0.027, 0.028)$, and $(0.029, 0.030)$, respectively. Axis units are given in Megaparsecs (Mpc).

Real Data Example



Galaxy data (three slices): top row: $\beta_0(r)$; middle row: $\beta_1(r)$; bottom row: $\beta_1(r, r + 1)$

Weak convergence of the Euler characteristic process

Theorem (Krebs & WP, 2021, Thomas and Owada, 2021)

Suppose that f can be approximated uniformly by a sequence of blocked density functions on $[0, 1]^d$ of the form $f_m = \sum_{i=1}^{m^d} b_i 1_{B_i}$ with blocks B_i of the form $B_i = [a_1, b_1] \times \cdots \times [a_d, b_d]$ such that for all $i = 1, \dots, d$ one has $c \frac{1}{m} \leq b_i - a_i \leq C \frac{1}{m}$ for some $c, C > 0$. Then, for a Poisson sampling scheme with intensity nf , as $n \rightarrow \infty$,

$$\nu_n \Rightarrow G \quad \text{weakly, in } D[0, T],$$

where $D[0, T]$ denotes Skorohod space, and G is a mean zero Gaussian process on $[0, T]$ with covariance of the form

$$\text{Cov}(G(s), G(t)) = \mathbb{E}[\gamma(f(X)^{\frac{1}{d}}(s, t))].$$

where $X \sim f$. A similar result holds for a Binomial sampling scheme (with density f), where the covariance function changes to

$$\text{Cov}(G(s), G(t)) = \mathbb{E}[\gamma(f(X)^{\frac{1}{d}}(s, t))] - \mathbb{E}[\alpha(f(X)^{\frac{1}{d}} s)] \mathbb{E}[\alpha(f(X)^{\frac{1}{d}} t)]$$

for some functions $\alpha : \Delta \times \Delta \rightarrow \mathbb{R}$, $\gamma : \Delta \times \Delta \rightarrow [0, \infty)$.

How does persistence diagram depend on sampling density?

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Recall:

- Limits of expected values and limit variances all depend on f through quantities of the form

$$\int_{\mathbb{X}} \Psi(f^{1/d}(x)) f(x) dx$$

for some functions Ψ not depending on f .

Writing

$$\int_{\mathbb{X}} \Psi(f^{1/d}(x)) f(x) dx = \mathbb{E} \Psi(f^{1/d}(X)),$$

we see that this value

depends on f only through the distribution of $f(X)$, where $X \sim f$.

See Vishwanath et al. (2020)

How does persistence diagram depend on sampling density?

What is this distribution? The survival function of $Y = f(X)$ with $X \sim f$ is

$$S_f(t) = P_f(f(X) \geq t) = \int_{\Gamma(\lambda)} f(x)dx = F(\Gamma(t)),$$

where $\Gamma(t) = \{x : f(x) \geq t\}$ (superlevel set of f); and F is distribution with density f .

$$S_f = S_g$$

\Rightarrow

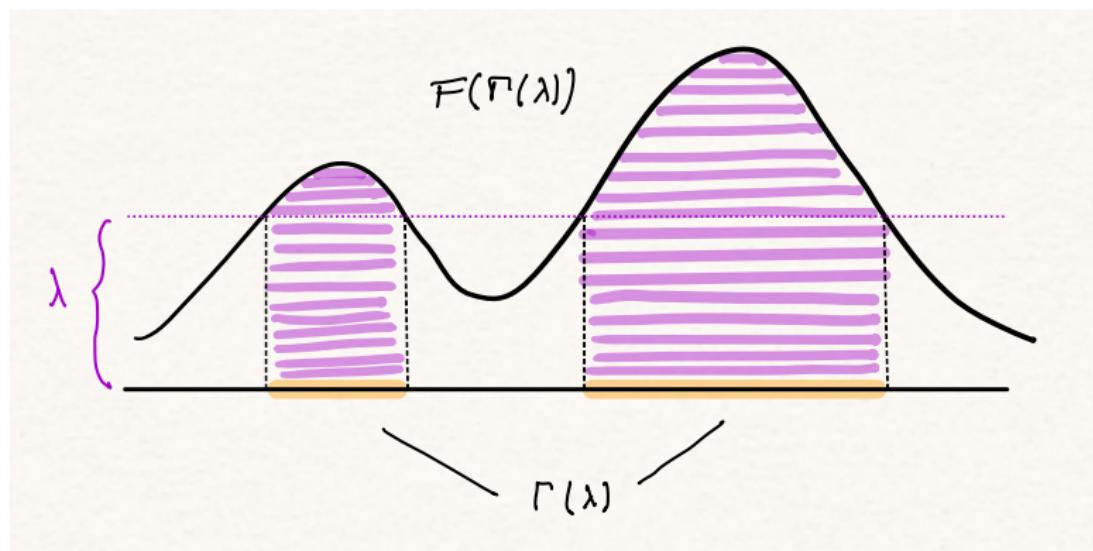
same asymptotic behavior of persistent Betti function (and Euler characteristic)

How does persistence diagram depend on sampling density?

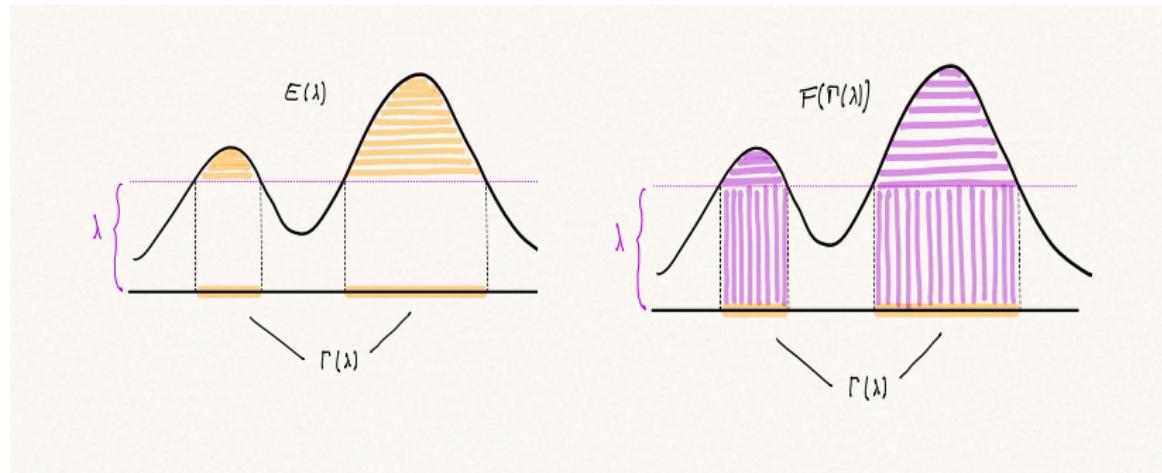
When is $S_f = S_g$?

How does persistence diagram depend on sampling density?

When is $S_f = S_g$?



How does persistence diagram depend on sampling density?



Excess mass function

We have: $S_f = S_g \Leftrightarrow E_f = E_g \Leftrightarrow \text{Leb}(\Gamma_f(\lambda)) = \text{Leb}(\Gamma_g(\lambda)) \forall \lambda$

E.g. f and g its monotonically decreasing rearrangement satisfy this!

Bandwidth selection for TDA

Bandwidth selection for TDA

Label	Description						
F_1	Rotationally symmetric in \mathbb{R}^2 , finite L_8 norm						
F_2	Rotationally symmetric in \mathbb{R}^2 , finite L_2 norm, infinite L_8 norm						
F_3	\mathbb{S}^1 embedded in \mathbb{R}^2 , additive Gaussian noise						
F_4	Uniformly distributed over $B_0(1)$ in \mathbb{R}^3 , additive Gaussian noise						
F_5	5 clusters in \mathbb{R}^3 , additive exponential noise						
F_6	\mathbb{S}^2 embedded in \mathbb{R}^5 , additive Cauchy noise						
F_7	Flat figure-8 embedded in \mathbb{R}^{10} , additive Gaussian noise						

Distr.	F_1	F_2	F_3	F_4	F_5	F_6	F_7	F_4	F_5	F_6	F_7
	$q = 1$							$q = 2$			
r	4.94	5.20	3.03	1.92	0.30	1.78	1.28	2.96	0.39	2.71	1.46
s	5.36	5.60	3.28	2.12	0.31	1.91	1.32	3.04	0.40	2.80	1.47
$n = 100$	0.896	0.965	0.921	0.859	0.954	0.19		0.908	0.705	0.038	
	0.931	0.959	0.914	0.809	0.941	0.133		0.903	0.604	0.045	
	0.903	0.97	0.91	0.859	0.927	0.049		0.902	0.363	0.002	
	0.359	0.931	0.942	0.864	0	0	0.656	0.902	0	0	0.045
$n = 200$	0.908	0.971	0.94	0.898	0.942	0.159		0.878	0.795	0.125	
	0.92	0.972	0.946	0.891	0.923	0.106		0.872	0.707	0.074	
	0.888	0.975	0.959	0.906	0.892	0.06		0.908	0.277	0.031	
	0.299	0.954	0.903	0.899	0	0	0.766	0.882	0	0	0.537
$n = 300$	0.9	0.971	0.926	0.921	0.94	0.183		0.854	0.906	0.225	
	0.94	0.971	0.938	0.896	0.94	0.087		0.854	0.917	0.072	
	0.913	0.971	0.94	0.896	0.922	0.054		0.855	0.964	0.074	
	0.283	0.956	0.925	0.906	0	0	0.835	0.856	0	0	0.508
$n = 400$	0.918	0.961	0.947	0.934	0.96	0.175		0.851	0.883	0.259	
	0.927	0.951	0.938	0.92	0.955	0.063		0.839	0.88	0.076	
	0.908	0.976	0.933	0.924	0.939	0.062		0.863	0.958	0.099	
	0.266	0.961	0.909	0.922	0.114	0	0.891	0.859	0	0	0.584

TABLE 2

Coverage proportions for 95% smoothed bootstrap confidence intervals on the mean persistent Betti numbers; coverage is estimated using $N = 1,000$ independent base samples with $B = 500$ bootstrap samples each. True mean persistent Betti numbers are estimated using a large ($N = 100,000$) number of independent samples from the true distribution. For each case, the values from top to bottom: Coverage proportions using “Hpi.diag”, “Hlscv.diag”, “Hscv.diag”, and “bw.silv” bandwidth selectors, respectively (see Section 5).

Some comments on the proof of the above results

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Proofs heavily rely on the notion of **stabilization** of the **add-one cost function**.

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Proofs heavily rely on the notion of **stabilization** of the **add-one cost function**.

- $\mathcal{X} = \text{set of all finite (multi) sets of } \mathbb{R}^d$
- $H : \mathcal{X} \rightarrow \mathbb{R}$ (such as Betti number or Euler characteristic)

Definition (add-one cost)

For $z \in \mathbb{R}^d$, the *add-one cost* function (for H) is

$$D_z(\mathbb{X}) = H(\mathbb{X} \cup \{z\}) - H(\mathbb{X}).$$

Examples: add-one cost

Example 1: Let

- $H(\mathbb{X}) = |S_k(\mathbb{X})| = \text{number of } k\text{-simplices in } \text{VR}_r(\mathbb{X})$ (for some fixed $r > 0$);

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Then

$$D_0(\mathbb{X}) = \sum_{\{x_{i_1}, \dots, x_{i_k}\} \subset \mathbb{X}} \mathbf{1}(\psi(0, x_{i_1}, \dots, x_{i_k}) \leq r).$$

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Then,

$$D_0(\mathbb{X}) = \sum_{x \in \mathbb{X}} d(0, x_j) \mathbf{1}(0 \in \text{kNN}(x, \mathbb{X}); x \in \text{kNN}(0, \mathbb{X})).$$

Stabilization

Definition (weak stabilization)

A functional $H : \mathcal{X} \rightarrow \mathbb{R}$ is *weakly stabilizing* on a locally finite point process \mathbb{Y} in \mathbb{R}^d , if there exists a random variable Δ_∞ , such that, for any sequence $\{B_n\}_{n \geq 1}$ of (measurable) sets satisfying $B_n \rightarrow \mathbb{R}^d$ as $n \rightarrow \infty$, we have

$$D_z(\mathbb{Y} \cap B_n) \rightarrow \Delta_\infty \quad \text{a.s. as } n \rightarrow \infty.$$

Definition (strong stabilization)

A functional $H : \mathcal{X} \rightarrow \mathbb{R}$ is *strongly stabilizing* on a locally finite point process \mathbb{Y} in \mathbb{R}^d , if there exists random variables S (*radius of stabilization*) and Δ_∞ such that, for $z \in \mathbb{R}^d$, with probability 1,

$$D_z((\mathbb{Y} \cap B_S(z)) \cup A) = \Delta_\infty \quad \text{for all } A \text{ finite with } A \subset \mathbb{R}^d \setminus B_S(z).$$

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- Clearly: strong stabilization \Rightarrow weak stabilization;

Comments

- ‘stabilization’ formalized in Penrose and Yukich (2001); ideas go back to Kesten and Lee (1996)
- H stabilizes \rightsquigarrow difference operator determined by ‘local’ information
 \rightsquigarrow control of dependence
- adding moment conditions \rightsquigarrow limit theorems

Examples

Example 1: Simplex counts:

$$D_0(\mathbb{Y} \cup B_n) = \sum_{\{X_{i_1}, \dots, X_{i_k}\} \subset B_n} \mathbf{1}(\psi(0, X_{i_1}, \dots, X_{i_k}) \leq r).$$

(Recall : $\Psi(\sigma)$ is filtration time, and we consider filtration VR_r .)

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Consider strong stabilization criterion.

Question to answer: Are there random variables S and $\Delta_\infty = \Delta_\infty(r)$, such that for all finite $A \subset (B_S(z))^\complement$,

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Observe:

$d(0, A) > 2r \Rightarrow$ points in A cannot be part of a simplex containing 0

$\Rightarrow S = 2r$ is a radius of stabilization.

Note: S is not random; holds for both VR and Čech complex!

- this translates to Euler characteristic (as alternating sum of simplex counts)

Examples

Example 2: Length of k-NN graph; $d = 2$ (for simplicity)

Again, we show strong stabilization; here stabilization radius is random.

Better control of the tail behavior of the stabilization radius

⇒ stronger results

A key technical result

A key technical result

A key ingredient for the proofs of bootstrap validity:

Proposition

Let \mathbb{X}_n and \mathbb{Y}_n be Binomial processes with densities f and g , respectively, and suppose that $\|f\|_p < \infty$ for $p \geq 2$. Further let ψ be *strongly stabilizing* over \mathbb{X}_n . Then there exists a coupling between \mathbb{X}_n and \mathbb{Y}_n such that

$$\sup_{n \in \mathbb{N}} \text{Var} \left[\frac{1}{\sqrt{n}} [(\psi(n^{1/d}\mathbb{Y}_n) - \psi(n^{1/d}\mathbb{X}_n)] \right] \leq \gamma(\|f - g\|_p),$$

where the rate function $\gamma: \mathbb{R}_+ \rightarrow \mathbb{R}$ is increasing with $\lim_{\delta \rightarrow 0} \gamma(\delta) = 0$ and depends only on f and p .

This applies to both persistent Betti numbers and Euler characteristic (pointwise)

Another key ingredient - the Geometric Lemma

Lemma (Corollary of Hiraoka et al. (2018), Lemma 2.11)

Let $A \subseteq B$ be two finite point sets of \mathbb{R}^d . Then, with $s \leq t$,

$$\begin{aligned} & \left| \beta_k(\mathcal{F}_s(n^{1/d}B), \mathcal{F}_t(n^{1/d}B)) - \beta_k(\mathcal{F}_s(n^{1/d}A), \mathcal{F}_t(n^{1/d}A)) \right| \\ & \leq \sum_{j=q}^{q+1} |S_j(\mathcal{F}_t(B) \setminus \mathcal{F}_t(A))|. \end{aligned}$$

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BACK-UP SLIDES

The Euler characteristic

Recall: The Euler characteristic $\chi(K)$ of a simplicial complex K is defined as

$$\chi(K) = \sum_{k=0}^{\infty} (-1)^k n_k(K),$$

where $n_k(K)$ denotes the number of k -simplices in K .

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In general: $\dim(K) = d \Rightarrow$ sum in the definition of the Euler characteristic is finite.

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$$\chi(K) = \sum_{k=0}^{\infty} (-1)^k n_k(K),$$

where $n_k(K)$ denotes the number of k -simplices in K . In particular, for a polyhedron in \mathbb{R}^d ,

$$\chi(P) = \#\{\text{vertices}\} - \#\{\text{edges}\} + \#\{\text{faces}\}.$$

In general: $\dim(K) = d \Rightarrow$ sum in the definition of the Euler characteristic is finite.

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Relation to Betti numbers:

$$\chi(K) = \sum_{k=0}^{\infty} (-1)^k \beta_k(K).$$

(This follows from observing that $\beta_k(K) = |S_k^+(K)| - |S_{k+1}^-(K)|$, with $S_k^\pm(K)$ positive/negative simplices in filtration.)

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Advantages: Easier to compute - no need to find persistence diagram, just count number of simplices.

Euler characteristic and the Euler characteristic curve

Applications:

Kerscher (2000); Kilner et al. (2005), Scholz et al. (2012); Khanamiri et al. (2018); Amézquita et al. (2020), Yen and Cheong (2021)

Various contexts: porous matter; astronomy; expected Euler characteristic heuristic; VR-complex and biological shapes; . . .

Theory and methodology related to our work:

Decreusefond et al. (2014), Bobrowski and Mukherjee (2015), Thomas and Owada (2021)

Weak convergence of the Euler characteristic process

Again, we consider

- VR and Čech filtrations
- \mathbb{X}_n either Binomial or Poisson point process
- critical regime

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The Euler characteristic process: With \mathcal{F} either VR or \mathcal{C} :

$$\nu_n(t) = \frac{1}{\sqrt{n}} (\chi(\mathcal{F}_t(n^{\frac{1}{d}} \mathbb{X}_n)) - \mathbb{E}[\chi(\mathcal{F}_t(n^{\frac{1}{d}} \mathbb{X}_n))]), \quad t \in [0, T].$$

Weak convergence of the Euler characteristic process

Theorem (Krebs & WP, 2021, Thomas and Owada, 2021)

Suppose that f can be approximated uniformly by a sequence of blocked density functions on $[0, 1]^d$ of the form $f_m = \sum_{i=1}^{m^d} b_i 1_{B_i}$ with blocks B_i of the form $B_i = [a_1, b_1] \times \cdots \times [a_d, b_d]$ such that for all $i = 1, \dots, d$ one has $c \frac{1}{m} \leq b_i - a_i \leq C \frac{1}{m}$ for some $c, C > 0$. Then, for a Poisson sampling scheme with intensity nf , as $n \rightarrow \infty$,

$$\nu_n \Rightarrow G \quad \text{weakly, in } D[0, T],$$

where $D[0, T]$ denotes Skorohod space, and G is a mean zero Gaussian process on $[0, T]$ with covariance of the form

$$\text{Cov}(G(s), G(t)) = \mathbb{E}[\gamma(f(X)^{\frac{1}{d}}(s, t))].$$

where $X \sim f$. A similar result holds for a Binomial sampling scheme (with density f), where the covariance function changes to

$$\text{Cov}(G(s), G(t)) = \mathbb{E}[\gamma(f(X)^{\frac{1}{d}}(s, t))] - \mathbb{E}[\alpha(f(X)^{\frac{1}{d}} s)] \mathbb{E}[\alpha(f(X)^{\frac{1}{d}} t)]$$

for some functions $\alpha : \Delta \times \Delta \rightarrow \mathbb{R}$, $\gamma : \Delta \times \Delta \rightarrow [0, \infty)$.

Discussion of weak convergence of Euler characteristic process

Discussion of weak convergence of Euler characteristic process

- The form of the functions γ_k, σ_k and α_k is not well understood.
- These functions are determined by the behavior under a homogeneous Poisson sampling \rightsquigarrow they don't depend on the sampling density;
- The dependence of the limit on the sampling density f is through quantities of the form

$$\mathbb{E}_f [\Psi_{k,s,t} f^{1/d}(X)].$$

This will be discussed further below.

- We do not have a process-level result for the persistence Betti numbers!
- As for persistent Betti numbers, process is centered by its expected value.
Does the expected value converge to a limit?

Expected value of the Euler characteristic

Proposition (Bobrowski and Mukherjee, 2015; Thomas and Owada, 2021)

Let f be a density w.r.t. uniform measure on an m -dimensional manifold \mathbb{X} embedded in \mathbb{R}^d , and assume f to be bounded. Furthermore, let \mathcal{P}_n be a Poisson process with intensity nf . Then

(i)

$$\frac{1}{n} \mathbb{E}\chi(\mathcal{C}_t(n^{1/m}\mathcal{P}_n)) \rightarrow \sum_{k=1}^m (-1)^k c_k(t) \quad \text{as } n \rightarrow \infty,$$

where the non-negative $c_k(t)$ are of the form $c_k(t) = \int_{\mathbb{X}} \Psi_{k,t}(f(x)) f(x) dx$ for some functions $\Psi_{k,t}$, $k = 0, 1, \dots, m$ not depending on f . A similar result holds for \mathcal{P}_n replaced by a Binomial process \mathbb{X}_n with density f .

(ii) If f is a density on \mathbb{R}^d , then

$$\frac{1}{n} \mathbb{E}\chi(\text{VR}_t(n^{1/d}\mathcal{P}_n)) \rightarrow \sum_{k=1}^{\infty} (-1)^k d_k(t) \quad \text{as } n \rightarrow \infty,$$

where $d_k(t)$ are of the form $d_k(t) = \int_{\mathbb{X}} \tilde{\Psi}_{k,t}(f(x)) f(x) dx$ for some functions $\tilde{\Psi}_{k,t}$, $k = 0, 1, \dots$ not depending on f .

Bootstrapping the Euler characteristic process

Again, smooth bootstrap:

- $N \sim \text{Poiss}(n)$
- $\mathbb{X}_n^* = (X_1^*, \dots, X_n^*), X_i^* \sim_{iid} \hat{f}_n$ (density estimator - not necessarily KDE)
- $\mathcal{P}^* = (X_1^*, \dots, X_N^*)$

With \mathcal{F} being either VR or \mathcal{C} , let

$$\mathcal{F}_{n,t}^* \text{ being either } \mathcal{F}_t(n^{1/d}(\mathbb{X}_n^*)) \text{ or } \mathcal{F}_t(n^{1/d}\mathcal{P}_n^*).$$

Define bootstrap version of Euler characteristic process:

$$\nu_n^*(t) = \frac{1}{\sqrt{n}} (\chi(\mathcal{K}_{n,t}^*) - \mathbb{E}^* [\chi(\mathcal{K}_{n,t}^*)]), \quad t \in [0, T].$$

Pointwise bootstrap for the Euler characteristic curve in the critical regime

First we discuss point-wise results:

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Theorem (Krebs & WP, 2021)

Let f be a density on $[0, 1]^d$ and let $(\hat{f}_n : n \in \mathbb{N})$ be a sequence of density estimators with the property that $\lim_{n \rightarrow \infty} \|\hat{f}_n - f\|_\infty = 0$ a.s. (in probability). Then, for ν_n^* based on bootstrap samples drawn from \hat{f}_n ,

$$\|\hat{f}_n - f\|_\infty^{-1/2} \cdot \sup_{t \in [0, T]} W_1(\nu_n^*(t), \nu_n(t)) = O(1) \quad \text{a.s. (in probability)},$$

where W_1 is the 1-Wasserstein distance.

Furthermore, for each $t \in [0, T]$

$$\{\|\hat{f}_n - f\|_\infty^{1/2} + n^{-1/2}\}^{-1} \cdot d_K(\nu_n^*(t), \nu_n(t)) = O(1) \quad \text{a.s. (in probability)},$$

where d_K denotes Kolmogorov-distance.

E.g.: For $\hat{f}_n = \hat{f}_{n,h}$ a KDE with p -th order kernel:

$$\|\hat{f}_{n,h} - f\|_\infty = O\left(\left(\frac{\log n}{n}\right)^{p/(d+2p)}\right) \quad \text{a.s.}$$

Bootstrap for the Euler characteristic process in the critical regime

Now we consider a process-level result. For simplicity, smooth bootstrap based on the KDE:

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Theorem (Krebs & WP, 2021)

For either the VR or the Čech filtration constructed over either a (rescaled) Binomial process with density f on $[0, 1]^d$, or a (rescaled) Poisson process with intensity nf . Suppose that for some $p > d$,

- f is p times continuously differentiable;
- bootstrap samples are based on a KDE $\hat{f}_{n,h}$, based on a p -th order kernel;

Then, for any fixed $T > 0$,

$$W_1^{D[0,T]}(\nu_n^*, \nu_n) = O((\log n)^\alpha n^{-\beta}),$$

where $\frac{1}{3} < \alpha = \frac{2p}{4d+2p} < 1$ and $0 < \beta = \frac{3p}{4d+8p} - \frac{1}{4} < \frac{1}{8}$, and where $W_1^{D[0,T]}$ denotes the 1-Wasserstein distance on the Skorohod space $D[0, T]$.

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- Above results come with rates of approximation;
- Bootstrap results are based on bounds for

$$W_1(\nu_{n,f}, \nu_{n,g}) \quad \text{and} \quad d_K(\nu_{n,f}, \nu_{n,g})$$

in terms of distances between f and g , where $\nu_{n,f}$ and $\nu_{n,g}$ denote the Euler characteristic process for samples from f and g , respectively. (See below.)

A key technical result

A key technical result

In case of the Euler characteristic (with finite stabilization radius(!)), we have a stronger result:

Theorem (Approximating property in the Wasserstein-Kantorovich distance)

Let f be an essentially bounded density on $[0, 1]^d$, $r \in \mathbb{R}_+$, and $g \in B_\infty(f, r)$. There are coupled Poisson processes $(\mathcal{P}_n, \mathcal{Q}_n)$ with intensities (nf, ng) and coupled binomial processes $(\mathbb{X}_n, \mathbb{Y}_n)$ with densities f, g , respectively, and a constant $C_{0,f} \in \mathbb{R}_+$ depending on f and r but not on g (as long as $g \in B_\infty(f, r)$), such that

$$\sup_n \sup_{t \in [0, T]} \text{Var}(\nu_{f,n}(t) - \nu_{g,n}(t)) \leq C_{0,\kappa} \|f - g\|_\infty.$$

In particular,

$$\sup_n \sup_{t \in [0, T]} W_1(\nu_{f,n}(t), \nu_{g,n}(t)) \leq C_{0,\kappa} \|f - g\|_\infty^{1/2}.$$

Filter functions

Union of balls with midpoints at the data points $\mathbb{X}_n = \{X_1, \dots, X_n\}$ are sublevel set of the distance function $d_{\mathbb{X}_n}(x)$:

$$\bigcup_{i=1}^n \overline{B}(r; X_i) = \{d_{\mathbb{X}_n}(x) \leq r\}$$

where $\overline{B}(r; X_i) = \text{closed ball of radius } r \text{ and midpoint } X_i$, and

$$d_{\mathbb{X}_n}(x) = \inf \{\|x - X_i\|, i = 1, \dots, n\}.$$

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More generally: If a persistence diagram is based on the sublevel set filtration of the form $\mathcal{F}_f = \{f^{-1}((-\infty, t]), t \in \mathbb{R}\}$, then we call f a **filter function**.