



FROM RESEARCH TO INDUSTRY

NEW ADVANCED IN SENSITIVITY ANALYSIS BASED ON **HSIC** DEPENDENCE MEASURES

**Hilbert Schmidt Independence Criterion*

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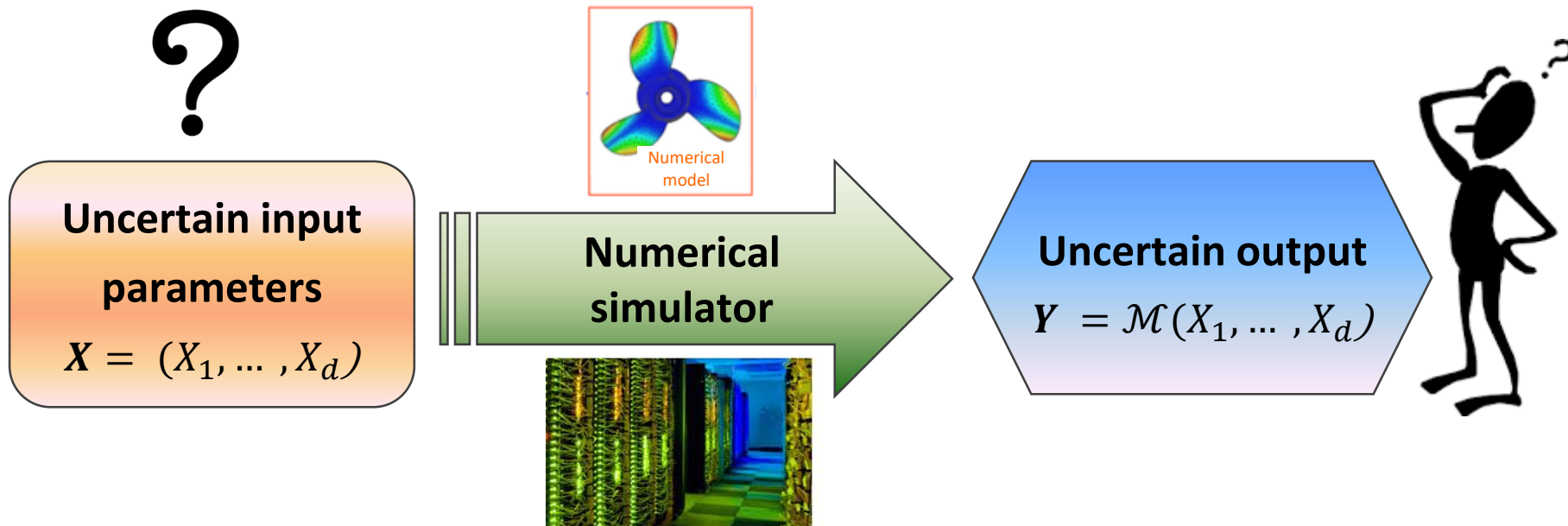
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Sensitivity Analysis in Uncertainty treatment

⇒ In numerical simulation

- **Numerical simulators:** fundamental tools to model & predict physical phenomena.
- **Large number of input parameters**, characterizing the studied phenomenon or related to its physical and numerical modelling.
- **Uncertainty on some input parameters** → impacts the **uncertainty on the output**
- **Black-box and time-expensive simulators** → limited number of simulations



⇒ Quantify how the variability of the input parameters influences the output
→ Aim of **Sensitivity Analysis (SA)** *Saltelli et al. [2000]*

cea Sensitivity Analysis (SA) in Uncertainty treatment ⇒ *In numerical simulation*

➤ Quantitative SA and Ranking purpose:

- Quantify the impact of each uncertain input and interaction → Ranking
→ Identify the variables to be fixed or further characterized in order to obtain the largest reduction of the output uncertainty

➤ Screening purpose. Separate the inputs into two groups: influential and non-influential

- Non-influential variables fixed without consequences on the output uncertainty
- In support of model reduction
- To build a simplified model, a metamodel

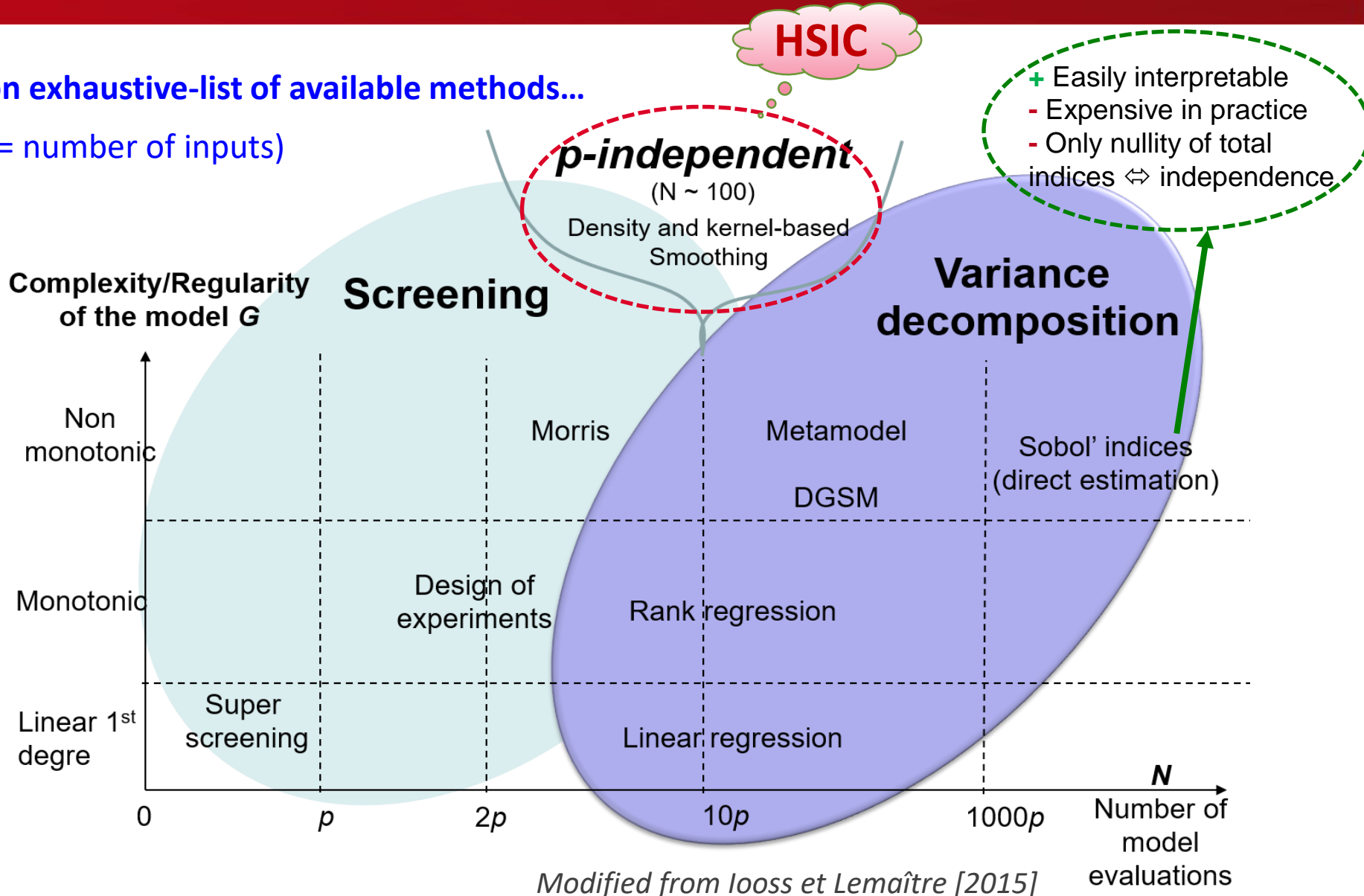


Global SA within a probabilistic framework

→ Valuable information to understand \mathcal{M} and underlying phenomenon

Non exhaustive-list of available methods...

(p = number of inputs)



HSIC* Review

**Hilbert Schmidt Independence Criterion*

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► Black-box model

$$Y = \mathcal{M}(X_1, \dots, X_d)$$

- X_1, \dots, X_d are d independent inputs, evolving in domain $\mathcal{X}_1, \dots, \mathcal{X}_d$
- Y evolves in domain \mathcal{Y}
- P_X denotes the probability distribution of X and p_X its density if X is continuous
- $P_{Y|X}$: the conditional distribution of Y given X
- $P_{X,Y}$: the joint probability measure and $P_Y \otimes P_X$ the product of marginal distributions

► \mathcal{M} unknown, only a **n -sample of simulations** $(X^{(j)}, Y^{(j)})_{1 \leq j \leq n}$ where $Y^{(j)} = \mathcal{M}(X^{(j)})$ for $j = 1, \dots, n$

GSA built upon RKHS framework

► How to evaluate the sensitivity in a probabilistic way? \Leftrightarrow independence

→ By comparing $P_{X,Y}$ with $P_X \otimes P_Y$

$$S_i = d(P_{X_i,Y}, P_{X_i} \otimes P_Y)$$

where d a dissimilarity measure between two probability distributions

d can be based on **Maximum Mean Discrepancy**:

$$\text{MMD}(\mathbb{P}, \mathbb{Q}) = \sup_{f \in \mathcal{F}} [\mathbb{E}_{\mathbb{P}} f(Y) - \mathbb{E}_{\mathbb{Q}} f(Y)]$$

with \mathcal{F} = **unit ball in a (characteristic) Reproducing Kernel Hilbert Space (RKHS)**

(Sriperumbudur et al. [2008])

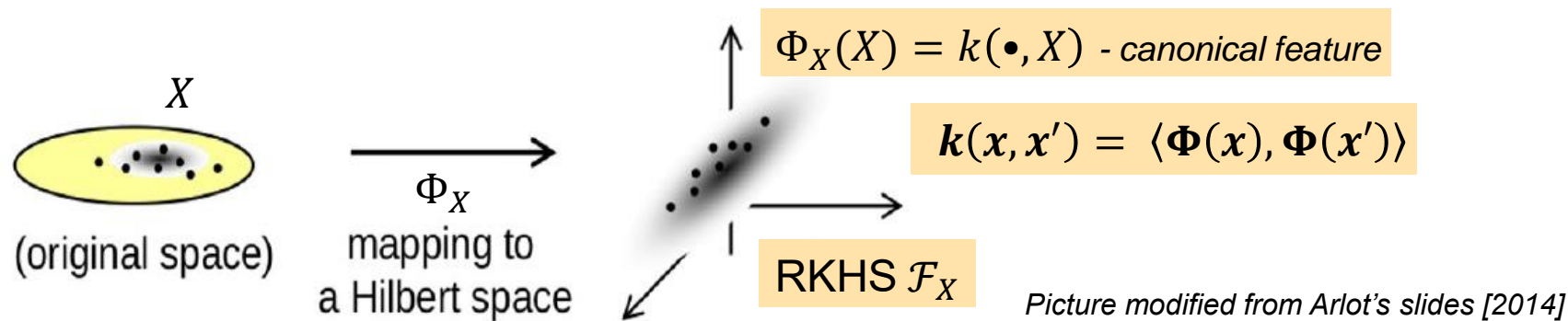
$$\Rightarrow S_i = \text{MMD}^2(P_{X,Y}, P_Y \otimes P_X) = \text{HSIC}(X, Y)$$

Hilbert-Schmidt Independence Criterion

GSA built upon RKHS framework

► Link between embeddings in RKHS and MMD

Association of RKHS \mathcal{F}_X to X : with Φ_X mapping function from \mathcal{X} to \mathcal{F}_X where the inner product is defined by kernel k (s.d.f. function).



Define embedding of \mathbb{P} in RKHS: $\mu_{\mathbb{P}} \in \mathcal{F}_X$ such as $E_{\mathbb{P}}[f(X)] = \langle f, \mu_{\mathbb{P}} \rangle_{\mathcal{F}_X}$ for all $f \in \mathcal{F}_X$

Reproducing property $\Rightarrow \mu_{\mathbb{P}} = E_{\mathbb{P}}[k(\bullet, X)] = E_{\mathbb{P}}[\Phi_X(X)]$

$$\Rightarrow \text{MMD}(\mathbb{P}, \mathbb{Q}) = \sup_{f \in \mathcal{F}_X, \|f\|_{\mathcal{F}_X} \leq 1} [E_{\mathbb{P}}f(Y) - E_{\mathbb{Q}}f(Y)] = \|\mu_{\mathbb{P}} - \mu_{\mathbb{Q}}\|_{\mathcal{F}}$$

Kernel trick (reproducing prop.) $\Rightarrow \text{MMD}^2(\mathbb{P}, \mathbb{Q}) = E[k(X, X')] + E[k(Y, Y')] - 2E[k(X, Y)]$

where $X, X' \sim \mathbb{P}$ and $Y, Y' \sim \mathbb{Q}$

HSIC definition

► **MMD²** applied between $P_{X,Y}$ and $P_Y \otimes P_X \Rightarrow HSIC(X, Y)_{\mathcal{F}_{X_i}, \mathcal{F}_Y}$

With \mathcal{F}_{X_i} and \mathcal{F}_Y **RKHS** associated to X_i and Y , resp.

k_X, k_Y , kernels on \mathcal{X}, \mathcal{Y} associated with RKHS $\mathcal{F}_X, \mathcal{F}_Y$

$k_X k_Y$ kernel on $\mathcal{X} \times \mathcal{Y}$ associated with RKHS $\mathcal{F}_X \otimes \mathcal{F}_Y$

Steinwart and Christmann [2008]

$$\Rightarrow HSIC(X, Y)_{\mathcal{F}_{X_i}, \mathcal{F}_Y} = MMD_{\mathcal{F}_{X_i}, \mathcal{F}_Y}^2(P_{X,Y}, P_Y \otimes P_X) = \|\mu_{P_{X,Y}} - \mu_{P_Y \otimes P_X}\|_{\mathcal{F}_X, \mathcal{F}_Y}^2$$

Hilbert-Schmidt Independence Criterion

Gretton et al. [2005]

Kernel trick (reproducing property) again yields:

$$HSIC(X, Y) = \mathbb{E}[k_X(X, X')k_Y(Y, Y')] + \mathbb{E}[k_X(X, X')]\mathbb{E}[k_Y(Y, Y')] - 2\mathbb{E}[\mathbb{E}[k_X(X, X')|X]\mathbb{E}[k_Y(Y, Y')|Y]]$$

where (X', Y') is an independent and identically distributed copy of $(X, Y) \sim P_{X,Y}$.

➤ **HSIC is in fact a “super generalized” covariance**

(Gretton et al. [2005])

1. Cross-covariance operator $C_{X,Y}$ **between the feature maps**

$$C_{X,Y} = \mathbb{E}_{X,Y}[\Phi_X(X) \otimes \Phi_Y(Y)] - \mathbb{E}_X[\Phi_X(X)] \otimes \mathbb{E}_Y[\Phi_Y(Y)] = \mathbb{E}_{X,Y}[\Phi_X(X) \otimes \Phi_Y(Y)] - \mu_{P_X} \otimes \mu_{P_Y}$$

2. HSIC corresponds to the squared **Hilbert-Schmidt norm** of $C_{X,Y}$

$$\text{HSIC}(X, Y)_{\mathcal{F}_X, \mathcal{F}_Y} = \|C_{X,Y}\|_{HS}^2 = \|\mu_{P_{X,Y}} - \mu_{P_Y \otimes P_X}\|^2$$

⇒ **A larger panel of input-output dependency can be captured by this operator,**

HSIC somehow "summarizes" the cross-cov between feature maps applied to X and Y

► **Characteristic kernels and RKHS** ⇒ *Injective feature map*

⇒ **Equivalence to independence:** $\text{HSIC}(X, Y) = 0 \Leftrightarrow X \perp Y$

Ex: Gaussian Kernel

$$k(x_i, x'_i) = \exp\left(-\frac{(x_i - x'_i)^2}{2\lambda^2}\right)$$

➤ **HSIC** expression with only expectations of kernels

$$\text{HSIC}(X, Y) = \mathbb{E}[k_X(X, X')k_Y(Y, Y')] + \mathbb{E}[k_X(X, X')]\mathbb{E}[k_Y(Y, Y')] - 2\mathbb{E}[\mathbb{E}[k_X(X, X')|X]\mathbb{E}[k_Y(Y, Y')|Y]]$$

where (X', Y') is an independent and identically distributed copy of (X, Y) .

▪ **Monte-Carlo estimator from a n -sample of simulations $(X_i^{(j)}, Y^{(j)})_{1 \leq j \leq n}$**

$$\widehat{\text{HSIC}}(X, Y) = \frac{1}{n^2} \text{Tr}(K_X H L_Y H) \quad (\text{Gretton et al. [2005]})$$

where $H = I_n - \frac{1}{n}$, $K_X = \left(k_X(X^{(j)}, X^{(j')}) \right)_{1 \leq j, j' \leq n}$ and $L_Y = \left(k(Y^{(j)}, Y^{(j')}) \right)_{1 \leq j, j' \leq n}$

▪ **Statistical properties of $\widehat{\text{HSIC}}$:**

- Asymptotically unbiased, variance of order $O(1/n)$
- If $\mathbf{X} \perp \mathbf{Y}$, $n\widehat{\text{HSIC}}(\mathbf{X}, \mathbf{Y})$ converges asymptotically to a Gamma distribution

HSIC for sensitivity analysis

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■ Normalization for sensitivity analysis:

(Da Veiga [2015])

$$R^2_{HSIC,i} = \frac{HSIC(X_i, Y)}{\sqrt{HSIC(X_i, X_i) HSIC(Y, Y)}}$$

$\Rightarrow R^2_{HSIC} \in [0,1]$ for easier interpretation

\Rightarrow Use for ranking of inputs

$$\text{Influence}(X_{[1]}) > \text{Influence}(X_{[2]}) > \dots > \text{Influence}(X_{[d]})$$

Where order $[\cdot]$ is such that $\widehat{R^2_{H,X_{[1]}}} > \widehat{R^2_{H,X_{[2]}}} > \dots > \widehat{R^2_{H,X_{[d]}}}$

► Several illustrations on analytical examples and industrial applications

- HSIC detect non-influential factors easily and robustly, even with small sample size
- HSIC indices can capture a large spectrum of dependence
 - Good ranking on usual GSA functions from sample size $n \sim 100$
 - Efficiency for screening → Even better: HSIC independence test

HSIC-based Independence test

► Use HSIC for screening → with Independence test

$$HSIC(X_i, Y) = 0 \Leftrightarrow X_i \perp Y \quad (\text{with } \underline{\text{characteristic kernels!}})$$

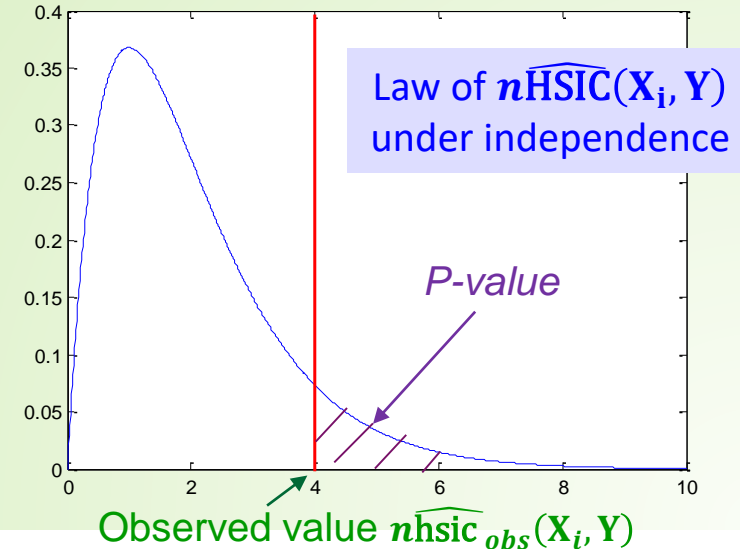
- Null hypothesis: $\mathcal{H}_0 : X_i \perp Y_i$ against $\mathcal{H}_1 : X_i \not\perp Y_i$
- Test statistics: $n\widehat{HSIC}(X_i, Y)$
- Decision rule to obtain a test of level $\alpha = \mathbb{P}_{\mathcal{H}_0} [\text{reject } \mathcal{H}_0]$ (α fixed at 5% or 10%)

$\mathcal{H}_{0,i}$ rejected iff $n\widehat{HSIC}(X_i, Y) > q_{1-\alpha}$ where $q_{1-\alpha}$ is the $(1 - \alpha)$ quantile under $\mathcal{H}_{0,i}$

→ Test function: $\Delta_\alpha = 1_{n\widehat{HSIC}(X_i, Y) > q_{1-\alpha}}$

- In practice, computation of p-value:

$$p\text{-value} = \mathbb{P}[\widehat{HSIC}(X_i, Y) > \widehat{hsic}_{obs}(X_i, Y)]$$



⇒ How to have the distribution of $n\widehat{HSIC}(X_i, Y)$ under \mathcal{H}_0 ?

⇒ How to have the distribution $n\widehat{\text{HSIC}}(X_i, Y)$ under \mathcal{H}_0 to compute p -value?

- **Asymptotic computation** with Gamma approximation for large n (Gretton et al. (2008])
- **Permutation-based approximation** for smaller sample size n (De Lozzo & Marrel (2016a], Meynaoui et al. [2019])

Algorithm 1 – Permutation-based independence test (for each X_i)

Require: The learning sample (X_i, Y) of n inputs/outputs $\{(X_i^{(1)}, Y^{(1)}), \dots, (X_i^{(n)}, Y^{(n)})\}$, B and α

- 1: Compute $\widehat{\text{HSIC}}_{\text{obs}}(X_i, Y)$ from Eq. (2)
 - 2: Generate B permutation-based samples $(X_i, Y_{[b]})_{1 \leq b \leq B}$
 - 3: Compute the B permutation-based estimators $(\widehat{\text{HSIC}}_b(X_i, Y))_{1 \leq b \leq B}$ by replacing Y by $Y_{[b]}$ in Eq. (2)
 - 4: Estimate the p -value by Monte-Carlo estimator $\hat{p}_{\text{val},i}^B = \frac{1}{B} \sum_{b=1}^B \mathbb{1}_{\widehat{\text{HSIC}}_b(X_i, Y) > \widehat{\text{HSIC}}_{\text{obs}}(X_i, Y)}$
 - 5: **if** $\hat{p}_{\text{val},i}^B < \alpha$ **then**
 - 6: **return** reject (\mathcal{H}_0^i)
 - 7: **else**
 - 8: **return** accept (\mathcal{H}_0^i)
 - 9: **end if**
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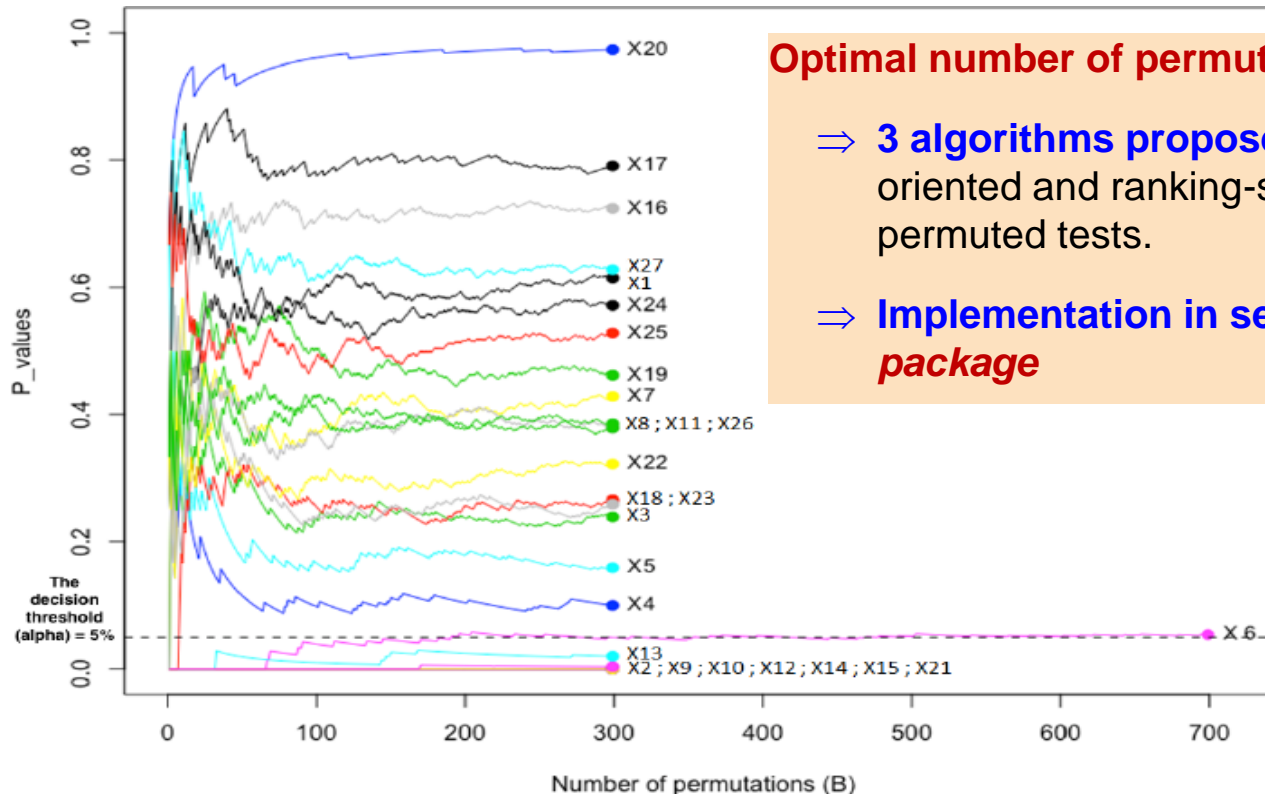
⇒ How to have the distribution $n\widehat{\text{HSIC}}(X_i, Y)$ under \mathcal{H}_0 to compute p -value?

- **Asymptotic computation** with Gamma approximation for large n (Gretton et al. (2008))
 - **Permutation-based approximation** for smaller sample size n (De Lozzo & Marrel (2016a), Meynaoui et al. [2019])
- Theoretical demonstration: **permuted-test is of level α** (Meynaoui et al. [2019])
- Empirically observed: **power of asymptotic and permutation test equivalent** for a sufficient number B of permutations
- Extension for **non independent** identically distributed samples: (El Amri & Marrel [2021b])
- ✓ **ok** for Latin Hypercube sample (LHS)
 - ✓ Correction with **Conditional Randomization Test** for scrambled low-discrepancy sequence (El Amri & Marrel [2021b])
 - ✗ Not possible for space-filling LHS

HSIC-based Independence test

► Permutation-based approximation (*El Amri & Marrel [2021a]*)

In practice, which number B of permutations to accurately estimate p-value?



Optimal number of permutations according to the final goal

⇒ **3 algorithms proposed**: screening-oriented, ranking-oriented and ranking-screening-oriented sequential permuted tests.

⇒ **Implementation in sensiHSIC of *R Sensitivity* package**

In practice reduction of B to few hundreds.

⇒ Sensitivity studies and convergence studies more tractable

HSIC-based Independence test

- Estimation of p-value with Pearson III-approximation (El Amri & Marrel [2021b])

$$\widehat{\text{HSIC}}(X, Y) = \frac{1}{n^2} \text{Tr}(K_X H L_Y H) = \frac{1}{n^2} \text{Tr}(\tilde{K}_X \tilde{L}_Y)$$

With $\tilde{K}_X = (H K_X H)$ and $\tilde{L}_Y = (H L_Y H)$ double centered matrices

$A \in \mathbb{R}^{n \times n}$ symmetric and with centered rows
 $W = U U^T$ with $U \in \mathbb{R}^{n \times p}$ and centered columns
 $S = \text{Tr}(AW)$

⇒ Distribution of S when rows of U are randomly permuted can be approximated by a Pearson III + expression of 3 first moments

(Kazi-Aoual et al. [1994])

- Application to permuted $\widehat{\text{HSIC}}_b(X, Y)$ statistics with $A = \tilde{K}_X$ and $W = \tilde{L}_Y$

- ⇒ Direct approximation, no permutation required
- ⇒ Fast and very efficient
- ⇒ **Best method for non-asymptotic framework**

Solutions for more powerful HSIC tests

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Solutions for more powerful HSIC-based test

► **Limitation: as mentionned before, usual kernels often depend on parameter**

For 1-D Gaussian kernel: **bandwidth parameter** λ ($\lambda > 0$), estimated in practice either with empirical standard deviation, or empirical median

$$k_G(x_i, x'_i) = \exp\left(-\frac{(x_i - x'_i)^2}{2\lambda^2}\right)$$

→ Heuristic choices without theoretical justifications

→ **Impact on the power of the HSIC-based test** (and quantitative interpretation of HSIC)

► **Solution 1: Aggregated HSIC-tests to take into account several kernels**

→ Improve the power of tests: theoretical demonstration on level and power

→ Improve the robustness of HSIC-based screening

Albert et al. [2021] and Meynaoui's PhD [2019]

Solution 1 for more powerful test: aggregation

► Methodology of aggregated testing procedure of level α ?

1. Consider a countable **collection of positive bandwidths** $\Lambda = (\lambda_i)_i$ and $\mathcal{U} = (\mu_i)_i$ and a **collection of positive weights** $\{\omega_{\lambda,\mu} / (\lambda,\mu) \in \Lambda \times \mathcal{U}\}$ such that $\sum_{(\lambda,\mu) \in \Lambda \times \mathcal{U}} e^{-\omega_{\lambda,\mu}} \leq 1$
2. Define a **level for each single test** of the collection $(\lambda,\mu) \in \Lambda \times \mathcal{U} : \mathbf{u}_\alpha e^{-\omega_{\lambda,\mu}}$

$$\text{Single test rejects } \mathcal{H}_0 \iff \widehat{HSIC}_{\lambda,\mu} > q_{1-\mathbf{u}_\alpha e^{-\omega_{\lambda,\mu}}}^{\lambda,\mu}$$

where \mathbf{u}_α is the less conservative value such that the aggregated test is of level α :

$$\mathbf{u}_\alpha = \sup \left\{ u > 0 ; \mathbb{P}_{f_X \otimes f_Y} \left[\sup_{(\lambda,\mu) \in \Lambda \times \mathcal{U}} \left(\widehat{HSIC}_{\lambda,\mu} > q_{1-u e^{-\omega_{\lambda,\mu}}}^{\lambda,\mu} \right) > 0 \right] \leq \alpha \right\}$$

\Rightarrow In practice \mathbf{u}_α estimated by permutation-based approach

3. Reject independence if there is **at least one single test** that rejects \mathcal{H}_0

$$\Delta_\alpha = 1 \iff \sup_{(\lambda,\mu) \in \Lambda \times \mathcal{U}} \left(\widehat{HSIC}_{\lambda,\mu} > q_{1-\mathbf{u}_\alpha e^{-\omega_{\lambda,\mu}}}^{\lambda,\mu} \right) > 0$$

► **Solution 2: HSIC-tests with optimal bandwidths** (*El Amri & Marrel [2021b]*)

$$(\lambda, \mu)_n^* = \operatorname{argmax}_{\lambda, \mu} \widehat{HSIC}_{\lambda, \mu}$$
$$\widehat{HSIC}^* = \widehat{HSIC}_{(\lambda, \mu)_n^*}$$

► **Methodology of optimal-bandwidth test:**

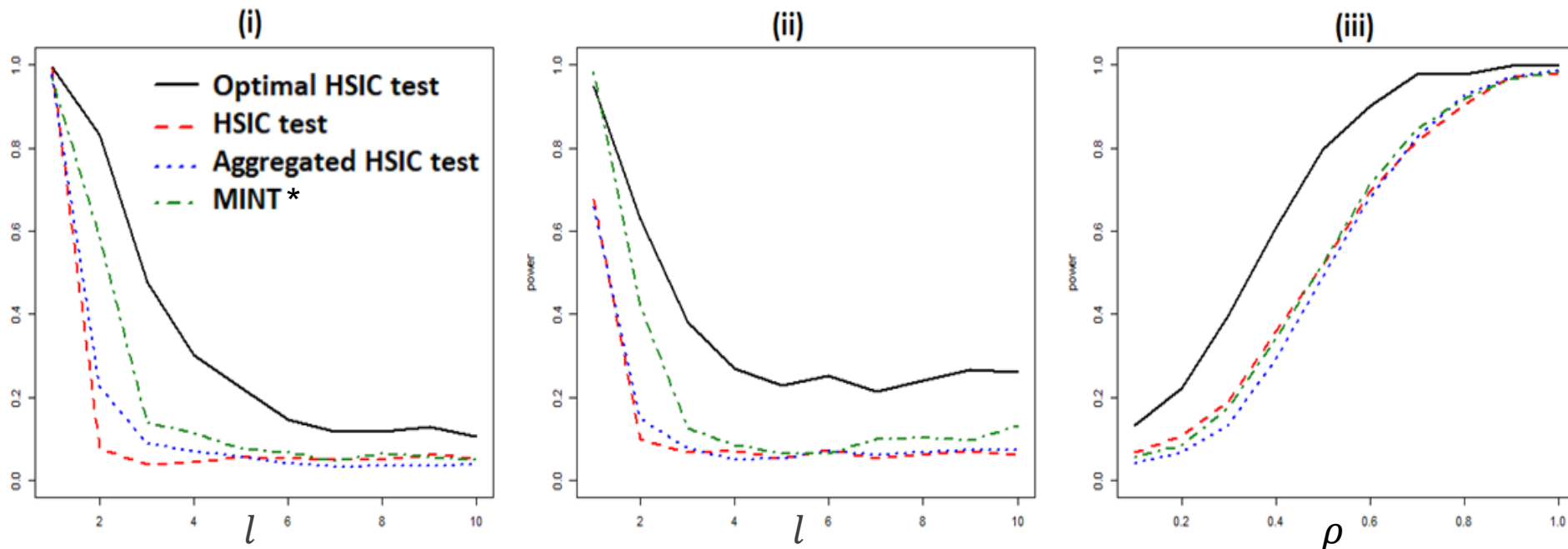
- Optimization solved using HSIC gradient and kernel derivatives
- Adaptation of permutation-based method by re-estimating the optimal bandwidths for each permuted sample
- Use of sequential permutation to optimize the number of permutation

Numerical comparisons

► Illustration on analytical examples (from Berreth & Samworth [2019], details in Appendix)

Results obtained from 1000 i.i.d. Monte-Carlo samples of size $n=100$

Power curves of independence tests according to shape parameters l and ρ



*MINT : Mutual Information-based test

Extension for Target and Functional Sensitivity Analysis

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► Goal-oriented SA for safety studies

(Marrel & Chabridon [2021])

- \Rightarrow To measure the input influence in a restricted output domain: $Y \in \mathcal{C}$
- \Rightarrow Numerous applications for **safety and risk assessment** ($\mathcal{C} = \{Y | Y > q_{0.9}\}$, e.g.)

► Technical point: choose the **characteristic kernel** according to the **type of data**:

- Output = « Is Y in domain \mathcal{C} ? »
 - **Target SA**: measures the influence of X over **the occurrence** of $Y \in \mathcal{C}$
 \rightarrow Bernoulli output: $\mathbf{1}_{Y \in \mathcal{C}}(Y) \sim \mathcal{B}(p_{\mathcal{C}})$ with $p_{\mathcal{C}} = \mathbb{P}[Y \in \mathcal{C}] \Rightarrow$ Dirac Kernel
 - **Conditional SA**: performed **within \mathcal{C}** only, ignoring what happens outside
 \rightarrow Real output: $Y | Y \in \mathcal{C}$ with $\mathbb{P}_{|Y \in \mathcal{C}}[\mathcal{A}] = \frac{\mathbb{P}[\mathcal{A} \cap Y \in \mathcal{C}]}{p_{\mathcal{C}}} \Rightarrow$ Gaussian kernel

► Goal-oriented HSIC for safety studies

(Marrel & Chabridon [2021])

\Rightarrow Brute versions:

- Target SA: $HSIC(X, \mathbf{1}_{Y \in \mathcal{C}}(Y))$ with Dirac Kernel
- Conditional SA: $HSIC(X, Y | Y \in \mathcal{C})$ with Gaussian Kernel

\Rightarrow Smoother versions to cope with the loss of information and take into account some information outside $\mathcal{C} \rightarrow$ Use of weight function $W_{\mathcal{C}}$ for relaxation

$$W_{\mathcal{C}} : \mathcal{Y} \rightarrow [0,1]$$

$$W_{\mathcal{C}}(y) = e^{-d_{\mathcal{C}}(y)/s} \quad \text{and} \quad d_{\mathcal{C}}(y) = \inf_{y' \in \mathcal{C}} \|y - y'\|$$

$\rightarrow HSIC(X, W_{\mathcal{C}}(Y))$ and $HSIC(X, W_{\mathcal{C}}(Y) | Y \in \mathcal{C})$

Similar use for optimization purpose in Spagnol et al. [2019]

✓ Implementation in sensiHSIC of *R Sensitivity package*:

Estimators of $HSIC(X, W_{\mathcal{C}}(Y))$ + asymptotic and permuted-based tests

► SA for functional data

(El Amri & Marrel [2021b])

Output is a random function of time or space \Rightarrow which kernel?

One solution: combine

1. Functional Principal Component Analysis (FPCA)
2. Truncation with the q first terms: $Y(t) - \mu(t) \approx \sum_{k=1}^q U_k \varphi_k(t)$
3. Weighed kernel based on the q FPCA (random) coefficients $(U_k)_{k=1,\dots,q}$

$$k(\mathbf{Y}^{(l)}, \mathbf{Y}^{(m)}) = \sum_{h=1}^q w_h k(\|U_h^{(l)} - U_h^{(m)}\|_2^2)$$

Where k is a usual kernel for real variables (Gaussian e.g.) and weights w_h correspond to the percentage of variance explained by each component U_h (cf. eigenvalues from FPCA)

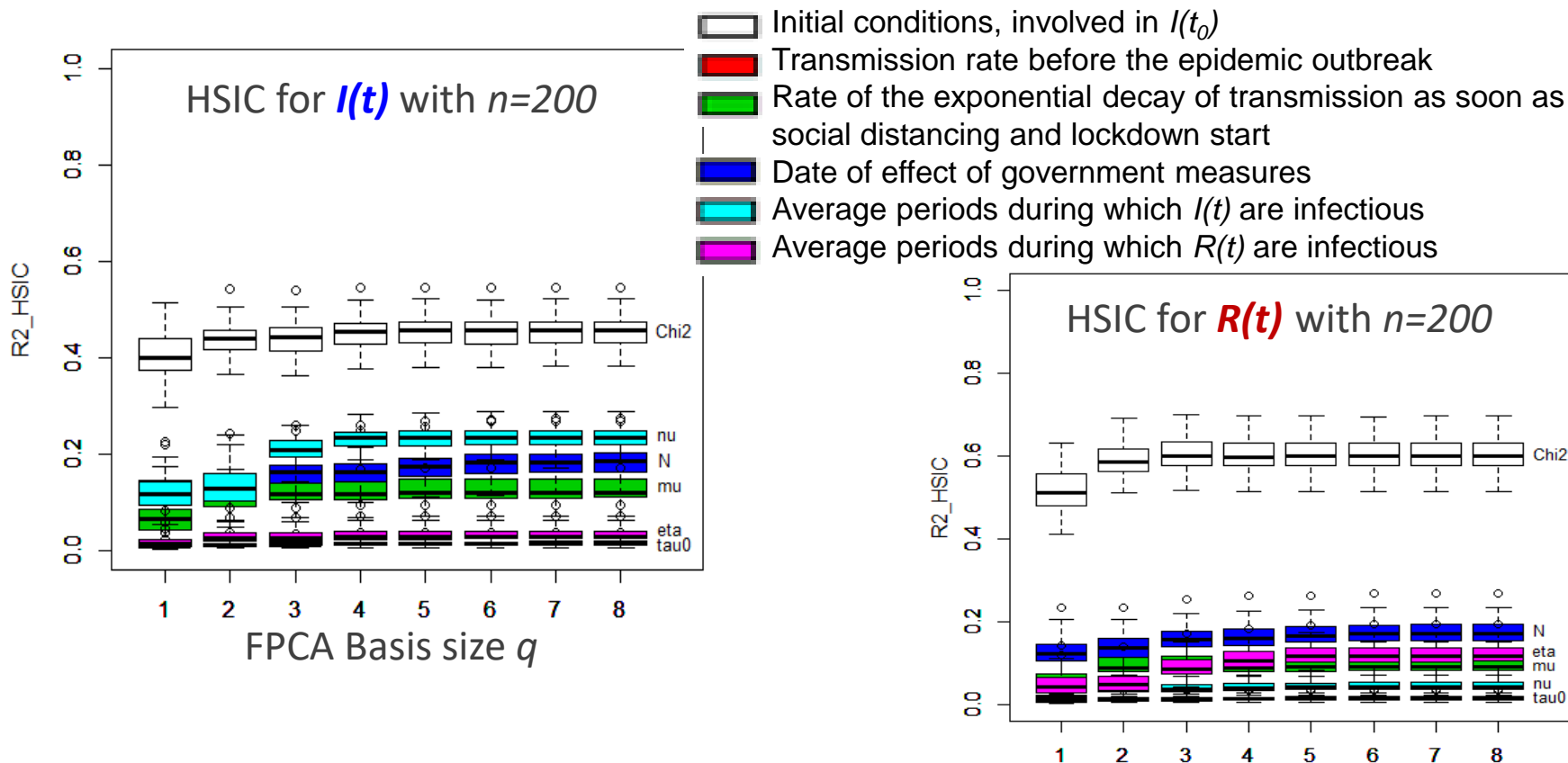
► SA for functional data

(El Amri & Marrel [2021b])

Illustration on compartmental epidemiologic model on COVID-19

Modified SIR model (Susceptible – Infected – Recovered) with 6 uncertain inputs

$I(t)$ and $R(t)$: number of **asymptomatic** and **reported symptomatic** infectious individuals at time t



Conclusion and prospects

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► HSIC as indices of Sensitivity Analysis

- Focus the SA analysis on the difference between $P_{X,Y}$ with $P_X \otimes P_Y$
- Power of RKHS → HSIC=one of the most successful non-parametric dependence measure
- Capture a large spectrum of relationships
- Able to deal with many factors and purposes (ranking, screening, goal-oriented SA)
- **Characterize independence** → efficient for screening and building independence tests !

► HSIC-tests of independence for screening

- Rigorous statistical framework, control of 1st and 2nd kind error
- Asymptotic and several non-asymptotic versions
- **P-value** of test → Really efficient for screening and for quantitative SA



**Efficiency demonstrated in numerous industrial applications,
especially with small sample size and large dimension**

► Limitations and prospects remain in HSIC SA indices

- Decomposition into main effects & interactions must be investigated
 - ⇒ Assess the use of **HSIC** with **ANOVA-like kernels** and **Shapley-HSIC** for **dependent inputs** (*Da Veiga [2021]*)
 - ⇒ Build associated independence tests
 - ⇒ Assess power for screening and relevancy for ranking



Simulation **A**nalytics and **M**eta-model-based solutions
for **O**ptimization, **U**ncertainty and **R**eliability **A**nalys**I**s

- Invariance properties → Preliminary isoprobabilistic transformation? (*Poczos et al. (2018)*)
- Extension “functional” HSIC-tests with other reduction techniques like Dynamic Time Warping?

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Thank you for your attention !



DE LA RECHERCHE À L'INDUSTRIE

APPENDIX

► **Illustration on analytical examples** (*usually used in indep. test comparison*)

3 different mechanisms of dependence, varying by a **shape parameter** (l or ρ)

Berreth & Samworth [2019]

(i). Defining the joint density f_l , $l = 1, \dots, 10$ of (X, Y) on $[-\pi, \pi]$ by

$$f_l(x, y) = \frac{1}{4\pi^2} \{1 + \sin(lx) \sin(ly)\}$$

(ii). Considering X and Y as $X = L \cos \Theta + \frac{\varepsilon_1}{4}$, $Y = L \sin \Theta + \frac{\varepsilon_2}{4}$,

where L , Θ , ε_1 and ε_2 are independent, with $L \sim \mathcal{U}\{1, \dots, l\}$

for $l = 1, \dots, 10$, $\Theta \sim \mathcal{U}[0, 2\pi]$ and $\varepsilon_1, \varepsilon_2 \sim \mathcal{N}(0, 1)$.

(iii). Defining $X \sim \mathcal{U}[-1, 1]$. For a given $\rho = 0.1, 0.2, \dots, 1$, Y is defined as $Y = |X|^\rho \varepsilon$, where $\varepsilon \sim \mathcal{N}(0, 1)$ independent with X .