### Lecture 7: Generating Learning Algorithm

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- Review of Last Week
  - Exponential Families
  - Generalized Linear Models
- 2 Generative Learning Algorithms
  - Discriminative vs. Generative Learning Algorithms
  - Gaussian Discriminant Analysis (GDA)
  - Linear Discriminant Analysis (LDA)
  - LDA and Logistic Regression
  - MLE for GDA
  - Naïve Bayes

#### Reading materials & Reference

#### Reading materials:

- Chapter 4, Stanford CS 229 Lecture Notes, https://cs229.stanford.edu/notes2022fall/main\_notes.pdf
- Chapter 9, Probabilistic Machine Learning: an introduction.
- Learning From Data Lecture 4: Generative Learning Algorithms, http://yangli-feasibility.com/

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  - MLE for GDA
  - Naïve Bayes

Logistic Regression models the probability of the two classes  $\{0,1\}$  by

$$p(1|\eta) = \sigma(\eta) \text{ and } p(0|\eta) = 1 - \sigma(\eta) \,, \tag{1} \label{eq:posterior}$$

where  $\eta = \mathbf{x}^{\mathsf{T}}\mathbf{w}$ .

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It is called the link function.

$$p(y|\boldsymbol{\eta}) = \underbrace{h(y)}_{>0} \exp\left[\boldsymbol{\eta}^{\top} \boldsymbol{\phi}(y) - A(\boldsymbol{\eta})\right]$$
(3)

<sup>&</sup>lt;sup>1</sup>Assume that we are given independent samples from this distribution. We do know  $\phi(y)$  and h(y) but not  $\eta$ . In order to optimally estimate  $\eta$  given these samples, all we need is the empirical average of the  $\phi(y)$ .

A distribution belongs to the exponential family if it can be written in the form

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$$\int p(y|\boldsymbol{\eta})dy = \int h(y) \exp\left[\boldsymbol{\eta}^{\top} \boldsymbol{\phi}(y) - A(\boldsymbol{\eta})\right] dy = 1$$
 (4)

$$\Longrightarrow \int h(y) \exp\left[\boldsymbol{\eta}^{\top} \boldsymbol{\phi}(y)\right] dy = \int h(y) \exp\left[A(\boldsymbol{\eta})\right] dy = \exp\left[A(\boldsymbol{\eta})\right]$$
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- There is a 1-1 relationship between the "mean"  $\mu := \mathbb{E}\left[\phi(y)\right]$  and natural parameter  $\eta$ , defined using a so-called *link function*  $\mathbf{g}$ :

$$\eta = \mathbf{g}(\boldsymbol{\mu} := \mathbb{E}[\phi(y)]) \Longleftrightarrow \boldsymbol{\mu} = \mathbf{g}^{-1}(\boldsymbol{\eta}) = \nabla A(\boldsymbol{\eta})$$
(6)

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- The GLM framework extends these ideas to the general exponential family.

#### **Constructing GLMs**

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The condition probability is thus modeled as:

$$p(y|\mathbf{x}; \mathbf{w}) = h(y_n) \exp\left(\eta \phi(y) - A(\eta)\right) \qquad \text{for} \quad \eta = g \circ f(\mathbf{x}^\top \mathbf{w})$$
 (7)

Note that:

$$\mathcal{L}(\mathbf{w}) = -\frac{1}{N} \sum_{n=1}^{N} \ln p(y_n | \mathbf{x}_n^{\top} \mathbf{w}) = -\frac{1}{N} \sum_{n=1}^{N} (\ln(h(y_n)) + \eta_n \phi(y_n) - A(\eta_n))$$
(8)

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If we rewrite this sum by using the matrix notation, we get

$$\nabla_{\mathbf{w}} \mathcal{L}(\mathbf{w}) = -\frac{1}{N} \sum_{i=1}^{N} (\mathbf{x}_{n} \phi(y_{n}) - \nabla_{\mathbf{w}} A(\eta_{n}))$$

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(9)

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(8)

(9)

(10)

(11)

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$$= -\frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_n \phi(y_n) - \mathbb{E} \left[ \phi(Y_n) \right] \mathbf{x}_n) = -\frac{1}{N} \sum_{n=1}^{N} \left( \mathbf{x}_n \phi(y_n) - g^{-1}(\mathbf{x}_n^{\top} \mathbf{w}) \mathbf{x}_n \right)$$

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$$(10)$$

(8)

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(8)

(9)

(10)

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Classification: we observe some data  $S = \{x_n, y_n\}_{n=1}^N \in \mathcal{X} \times \underbrace{\mathcal{Y}}_{\text{Discrete se}}$ 

Goal: given a new observation x, we want to predict its label y

How: relates input to a categorical variable

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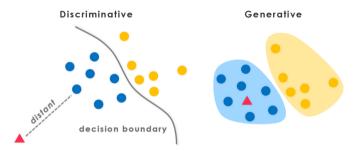
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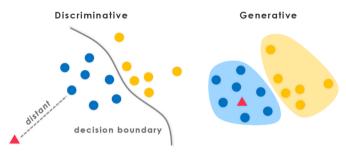
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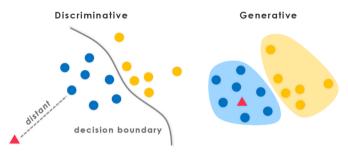


**Discriminative Learning Algorithm:** 



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Try to learn the conditional probability  $p(y|\mathbf{x}, \mathbf{w})$  directly or learn mappings directly from  $\mathcal{X}$  to  $\mathcal{Y}$ 

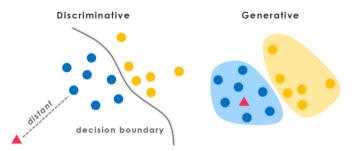


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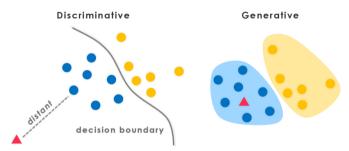
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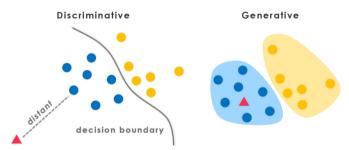
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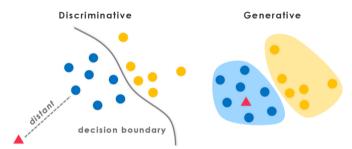
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- equivalently, it models  $p(\mathbf{x}|y,\mathbf{w})$  and  $p(y|\mathbf{w})$
- learned models are transformed to  $p(y|\mathbf{x}, \mathbf{w})$  later to classify data using Bayes' rule:

$$p(y|\mathbf{x}) = \frac{p(\mathbf{x}|y)p(y)}{p(\mathbf{x})}$$

The posterior distribution on y given x:

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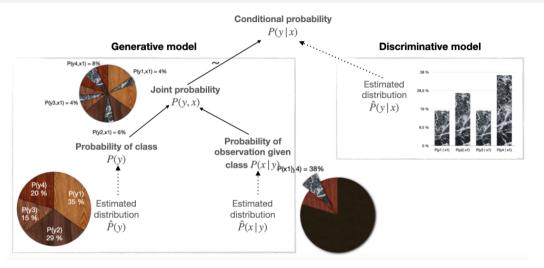
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Make predictions in a generative model (exercise):

$$\underset{y \in \mathcal{Y}}{\arg \max} p(y|\mathbf{x}, \mathbf{w}) = \underset{y \in \mathcal{Y}}{\arg \max} \frac{p(\mathbf{x}|y, \mathbf{w})p(y|\mathbf{w})}{p(\mathbf{x}|\mathbf{w})} = \underset{y \in \mathcal{Y}}{\arg \max} p(\mathbf{x}|y, \mathbf{w})p(y|\mathbf{w})$$

## Discriminative v/s Generative Learning Algorithms (an alternative view)



### Discriminative Models v.s. Generative Models (an alternative view)

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## Definition 1 (The density of a random vector from the multi-variate normal distribution)

The density of a random vector from the multi-variate normal/Gaussian distribution.

with mean 
$$\mu \in \mathbb{R}^d$$
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(16)

23/50

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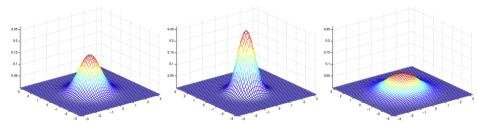
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23/50



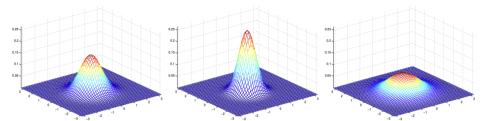
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(Left):  $\Sigma = I$ ; (Middle):  $\Sigma = 0.6I$ ; (Right):  $\Sigma = 2I$ . We can see that as

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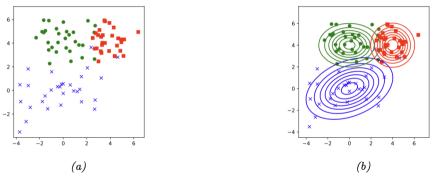
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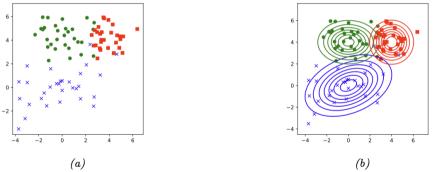
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The corresponding class posterior ((Gaussian Discriminant Analysis: GDA)) therefore has the form

$$p(y = c|\mathbf{x}, \mathbf{w}) \propto \pi_c \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_c, \boldsymbol{\Sigma}_c),$$
 (20)

where  $\pi_c = p(y = c | \mathbf{w})$ .

Write out the Probability Density Functions (PDFs):

$$p(y) = \phi^{y} (1 - \phi)^{1 - y} \tag{21}$$

$$p(\mathbf{x}|y=0) = \frac{1}{\sqrt{(2\pi)^d \det(\mathbf{\Sigma})}} \exp\left\{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_0)^{\top} \boldsymbol{\Sigma}_0^{-1}(\mathbf{x} - \boldsymbol{\mu}_0)\right\}$$
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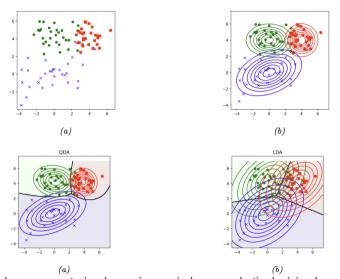
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The log posterior over class labels is given by

$$\log p(y=c|\mathbf{x},\mathbf{w}) = \log \pi_c - \frac{1}{2} \left| 2\pi \mathbf{\Sigma}_c \right| - \frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_c)^\top \mathbf{\Sigma}_c^{-1} (\mathbf{x} - \boldsymbol{\mu}_c) + \mathsf{const}$$

(22)



(a): GDA fits to data, where unconstrained covariances induce quadratic decision boundaries; (b): LDA fits to data, where tied covariances induce linear decision boundaries.

What is LDA?

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## The connection between LDA and logistic regression

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$$\log p(y = c | \mathbf{x}, \mathbf{w}) = \gamma_c + \mathbf{x}^{\top} \boldsymbol{\beta}_c + \kappa$$

Then,

$$p(y = c | \mathbf{x}, \mathbf{w}) = \frac{e^{\gamma_c + \mathbf{x}^{\top} \boldsymbol{\beta}_c}}{\sum_{c'} e^{\gamma_{c'} + \mathbf{x}^{\top} \boldsymbol{\beta}_{c'}}} = \frac{e^{\mathbf{w}_c^{\top}[1, \mathbf{x}]}}{\sum_{c'} e^{\mathbf{w}_{c'}^{\top}[1, \mathbf{x}]}},$$

where  $\mathbf{w}_c = [\gamma_c, \boldsymbol{\beta}_c]$ .

## The connection between LDA and logistic regression

Recall that in LDA, we have

$$\log p(y = c | \mathbf{x}, \mathbf{w}) = \gamma_c + \mathbf{x}^{\top} \boldsymbol{\beta}_c + \kappa$$

Then,

$$p(y = c | \mathbf{x}, \mathbf{w}) = \frac{e^{\gamma_c + \mathbf{x}^{\top} \boldsymbol{\beta}_c}}{\sum_{c'} e^{\gamma_{c'} + \mathbf{x}^{\top} \boldsymbol{\beta}_{c'}}} = \frac{e^{\mathbf{w}_c^{\top}[1, \mathbf{x}]}}{\sum_{c'} e^{\mathbf{w}_{c'}^{\top}[1, \mathbf{x}]}},$$

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To gain further insights, let's consider the binary case.

### The posterior is given by

$$p(y = 1|\mathbf{x}, \mathbf{w})$$

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$$p(y=1|\mathbf{x},\mathbf{w}) = \frac{e^{\gamma_1 + \boldsymbol{\beta}_1^{\mathsf{T}}\mathbf{x}}}{e^{\gamma_1 + \boldsymbol{\beta}_1^{\mathsf{T}}\mathbf{x}} + e^{\gamma_0 + \boldsymbol{\beta}_0^{\mathsf{T}}\mathbf{x}}}$$

$$p(y=1|\mathbf{x},\mathbf{w}) = \frac{e^{\gamma_1 + \boldsymbol{\beta}_1^\top \mathbf{x}}}{e^{\gamma_1 + \boldsymbol{\beta}_1^\top \mathbf{x}} + e^{\gamma_0 + \boldsymbol{\beta}_0^\top \mathbf{x}}} = \frac{1}{1 + e^{(\gamma_0 - \gamma_1) + (\boldsymbol{\beta}_0 - \boldsymbol{\beta}_1)^\top \mathbf{x}}}$$

$$p(y=1|\mathbf{x},\mathbf{w}) = \frac{e^{\gamma_1 + \boldsymbol{\beta}_1^{\top} \mathbf{x}}}{e^{\gamma_1 + \boldsymbol{\beta}_1^{\top} \mathbf{x}} + e^{\gamma_0 + \boldsymbol{\beta}_0^{\top} \mathbf{x}}} = \frac{1}{1 + e^{(\gamma_0 - \gamma_1) + (\boldsymbol{\beta}_0 - \boldsymbol{\beta}_1)^{\top} \mathbf{x}}} = \sigma \left( (\boldsymbol{\beta}_1 - \boldsymbol{\beta}_0)^{\top} \mathbf{x} + (\gamma_1 - \gamma_0) \right)$$

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#### Note that

$$\gamma_1 - \gamma_0$$

$$p(y=1|\mathbf{x},\mathbf{w}) = \frac{e^{\gamma_1 + \boldsymbol{\beta}_1^{\top}\mathbf{x}}}{e^{\gamma_1 + \boldsymbol{\beta}_1^{\top}\mathbf{x}} + e^{\gamma_0 + \boldsymbol{\beta}_0^{\top}\mathbf{x}}} = \frac{1}{1 + e^{(\gamma_0 - \gamma_1) + (\boldsymbol{\beta}_0 - \boldsymbol{\beta}_1)^{\top}\mathbf{x}}} = \sigma\left((\boldsymbol{\beta}_1 - \boldsymbol{\beta}_0)^{\top}\mathbf{x} + (\gamma_1 - \gamma_0)\right)$$

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$$\gamma_1 - \gamma_0 = (\log \pi_1 - \frac{1}{2}\boldsymbol{\mu}_1^{\top}\boldsymbol{\Sigma}^{-1}\boldsymbol{\mu}_1) - (\log \pi_0 - \frac{1}{2}\boldsymbol{\mu}_0^{\top}\boldsymbol{\Sigma}^{-1}\boldsymbol{\mu}_0)$$

$$p(y=1|\mathbf{x},\mathbf{w}) = \frac{e^{\gamma_1 + \boldsymbol{\beta}_1^{\top} \mathbf{x}}}{e^{\gamma_1 + \boldsymbol{\beta}_1^{\top} \mathbf{x}} + e^{\gamma_0 + \boldsymbol{\beta}_0^{\top} \mathbf{x}}} = \frac{1}{1 + e^{(\gamma_0 - \gamma_1) + (\boldsymbol{\beta}_0 - \boldsymbol{\beta}_1)^{\top} \mathbf{x}}} = \sigma \left( (\boldsymbol{\beta}_1 - \boldsymbol{\beta}_0)^{\top} \mathbf{x} + (\gamma_1 - \gamma_0) \right)$$

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$$p(y = 1 | \mathbf{x}, \mathbf{w}) = \frac{e^{\gamma_1 + \boldsymbol{\beta}_1^\top \mathbf{x}}}{e^{\gamma_1 + \boldsymbol{\beta}_1^\top \mathbf{x}} + e^{\gamma_0 + \boldsymbol{\beta}_0^\top \mathbf{x}}} = \frac{1}{1 + e^{(\gamma_0 - \gamma_1) + (\boldsymbol{\beta}_0 - \boldsymbol{\beta}_1)^\top \mathbf{x}}} = \sigma \left( (\boldsymbol{\beta}_1 - \boldsymbol{\beta}_0)^\top \mathbf{x} + (\gamma_1 - \gamma_0) \right)$$

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If we define

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then we have  $\mathbf{w}^{\top}\mathbf{x}_0 = -(\gamma_1 - \gamma_0)$ ,

$$p(y=1|\mathbf{x},\mathbf{w}) = \frac{e^{\gamma_1 + \boldsymbol{\beta}_1^{\top}\mathbf{x}}}{e^{\gamma_1 + \boldsymbol{\beta}_1^{\top}\mathbf{x}} + e^{\gamma_0 + \boldsymbol{\beta}_0^{\top}\mathbf{x}}} = \frac{1}{1 + e^{(\gamma_0 - \gamma_1) + (\boldsymbol{\beta}_0 - \boldsymbol{\beta}_1)^{\top}\mathbf{x}}} = \sigma\left((\boldsymbol{\beta}_1 - \boldsymbol{\beta}_0)^{\top}\mathbf{x} + (\gamma_1 - \gamma_0)\right)$$

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then we have  $\mathbf{w}^{\top}\mathbf{x}_0 = -(\gamma_1 - \gamma_0)$ , and hence

$$p(y = 1 | \mathbf{x}, \mathbf{w}) = \sigma \left( \mathbf{w}^{\top} (\mathbf{x} - \mathbf{x}_0) \right) .$$

$$p(y=1|\mathbf{x},\mathbf{w}) = \frac{e^{\gamma_1 + \boldsymbol{\beta}_1^{\top}\mathbf{x}}}{e^{\gamma_1 + \boldsymbol{\beta}_1^{\top}\mathbf{x}} + e^{\gamma_0 + \boldsymbol{\beta}_0^{\top}\mathbf{x}}} = \frac{1}{1 + e^{(\gamma_0 - \gamma_1) + (\boldsymbol{\beta}_0 - \boldsymbol{\beta}_1)^{\top}\mathbf{x}}} = \sigma\left((\boldsymbol{\beta}_1 - \boldsymbol{\beta}_0)^{\top}\mathbf{x} + (\gamma_1 - \gamma_0)\right)$$

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This has the same form as binary logistic regression, and the MAP decision rule is

$$\hat{y}(\mathbf{x}) = 1$$
 iff  $\mathbf{w}^{\top}\mathbf{x} > c$ , where  $c = \mathbf{w}^{\top}\mathbf{x}_0$ 

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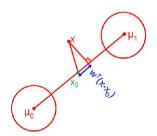
To interpret this equation geometrically, suppose  $\Sigma = \sigma^2 \mathbf{I}$   $\Longrightarrow \mathbf{w} = \sigma^{-2}(\boldsymbol{\mu}_1 - \boldsymbol{\mu}_0)$ , which is parallel to a line joining the two centroids,  $\boldsymbol{\mu}_0$  and  $\boldsymbol{\mu}_1$ .

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Geometry of LDA in the 2 class case where  $\Sigma_1 = \Sigma_2 = I$ .

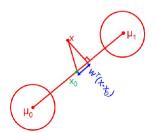
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 $\Longrightarrow$  we can classify a point by projecting it onto this line, and check its distance to  $\mu_0$  and  $\mu_1$ .



### Some remarks

#### **GDA**

- maximizes the joint likelihood  $\prod_{i=1}^n p(\mathbf{x}^{(i)}, y^{(i)})$
- modeling assumptions:

$$\mathbf{x}|y = b \sim \mathcal{N}(\boldsymbol{\mu}_b, \boldsymbol{\Sigma}), y \sim \mathsf{Bernoulli}(\phi)$$

 when modeling assumptions are correct, GDA is asymptotically efficient<sup>a</sup> and data efficient.

### **Logistic Regression**

- maximizes the conditional likelihood  $\prod_{i=1}^{n} p(y^{(i)}|\mathbf{x}^{(i)})$
- modeling assumptions:
   p(y|x) is a logistic function; no restriction on p(x)
- more robust and less sensitive to incorrect modeling assumptions.

 $<sup>^{</sup>a}$ in the limit of very large training sets (large n), there is no algorithm that is strictly better than GDA.

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#### 2 Generative Learning Algorithms

- Discriminative vs. Generative Learning Algorithms
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The likelihood function of a GDA model is as follows:

$$p(\mathcal{D}|\mathbf{w}) = \prod_{i=1}^{N} \mathsf{Cat}(y_i|\boldsymbol{\pi}) \prod_{c=1}^{C} \mathcal{N}(\mathbf{x}_i|\boldsymbol{\mu}_c, \boldsymbol{\Sigma}_c)^{1\{y_i=c\}}$$

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Hence the log-likelihood is given by

$$\log p(\mathcal{D}|\mathbf{w}) = \left[\sum_{i=1}^{n} \sum_{c=1}^{C} 1\{y_i = c\} \log \pi_c\right] + \sum_{c=1}^{C} \left[\sum_{i:y_i = c} \log \mathcal{N}\left(\mathbf{x}_i | \boldsymbol{\mu}_c, \boldsymbol{\Sigma}_c\right)\right]$$

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where we can optimize  $\pi$  and the  $(\mu_c, \Sigma_c)$  terms separately. MLF estimation:

- $\hat{\pi}_c = \frac{N_c}{N}$
- $\hat{\boldsymbol{\mu}}_c = \frac{1}{N_c} \sum_{i:y_i=c} \mathbf{x}_n$
- $\hat{\mathbf{\Sigma}}_c = \sum_{i:y_i=c} (\mathbf{x}_i \hat{\boldsymbol{\mu}}_c) (\mathbf{x}_i \hat{\boldsymbol{\mu}}_c)^{\top}$

$$\ell(\phi, \boldsymbol{\mu}_0, \boldsymbol{\mu}_1, \boldsymbol{\Sigma})$$

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$$= \log \prod_{i=1}^n p(\mathbf{x}^{(i)}, y^{(i)}; \phi, \boldsymbol{\mu}_0, \boldsymbol{\mu}_1, \boldsymbol{\Sigma})$$

$$\begin{split} &\ell(\phi, \boldsymbol{\mu}_0, \boldsymbol{\mu}_1, \boldsymbol{\Sigma}) \\ &= \log \prod_{i=1}^n p(\mathbf{x}^{(i)}, y^{(i)}; \phi, \boldsymbol{\mu}_0, \boldsymbol{\mu}_1, \boldsymbol{\Sigma}) \\ &= \log \prod_{i=1}^n p(\mathbf{x}^{(i)}|y^{(i)}; \boldsymbol{\mu}_0, \boldsymbol{\mu}_1, \boldsymbol{\Sigma}) p(y^{(i)}; \phi) \end{split}$$

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Log-likelihood of the data:

 $\ell(\phi, \boldsymbol{\mu}_0, \boldsymbol{\mu}_1, \boldsymbol{\Sigma})$ 

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By maximizing  $\ell$  w.r.t. the parameters, the maximum likelihood estimate of the parameters:

$$egin{aligned} \phi &= rac{1}{n} \sum_{i=1}^n 1\{y^{(i)} = 1\} \ egin{aligned} m{\mu}_b &= rac{\sum_{i=1}^n 1\{y^{(i)} = b\} \mathbf{x}^{(i)}}{\sum_{i=1}^n 1\{y^{(i)} = b\}} ext{ for } b = 0, 1 \ m{\Sigma} &= rac{1}{n} \sum_{i=1}^n (\mathbf{x}^{(i)} - m{\mu}_{y^{(i)}}) (\mathbf{x}^{(i)} - m{\mu}_{y^{(i)}})^{ op} \end{aligned}$$

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Let's talk about a different learning algorithm in which  $x_i$ 's are discrete-valued.

Naïve Baves

⇒ Naïve Bayes

# Naïve Bayes: Motivating Example

#### Example 2 (Spam filter (document classification))

Classify email message x to spam (y = 1) and non-spam (y = 0) classes.

#### Hello

We need to confirm your info...

(1) FINAL MESSAGE: Payout Verification - \$3000 PAYOUT is ready to be addressed in your Name and we want to be sure it gets to the right place. Click below to start the confirmation process. The sooner you act, the sooner it can be in your hands!

Raging Bull Casino

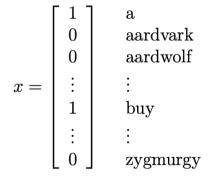
A sample spam email

# Example: Spam Filter

We would like to represent an email via a feature vector.

**Binary text features:** Given a dictionary of size n, represent a message composed of dictionary words as  $\mathbf{x} \in \{0, 1\}^n$ :

$$x_i = \begin{cases} 1 & i\text{-th dictionary word is in message} \\ 0 & \text{otherwise} \end{cases}$$



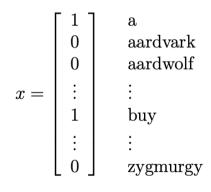
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To build a generative model, we have to model  $p(\mathbf{x}|y)$ 



# Naïve Bayes assumption and Naïve Bayes classifier

Probability of observing email  $x_1, \ldots, x_n$  given spam class y:

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As a result,

$$p(x_1,...,x_n|y) = p(x_1|y)p(x_2|y)...p(x_n|y) = \prod_{i=1}^{n} p(x_i|y)$$

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#### Model parameters:

- $\phi_{\nu} \in \mathbb{R}$
- $\phi_{i|y=1}$ ,  $\phi_{i|y=0}$  for i = 1, ..., n

# Naïve Bayes Parameter Learning

Likelihood of i.i.d. training data  $(\mathbf{x}^{(1)}, y^{(1)}), \dots, (\mathbf{x}^{(n)}, y^{(n)})$ :

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Maximum likelihood estimation of parameters:

$$\phi_y = \frac{1}{n} \sum_{i=1}^n 1\{y^{(i)} = 1\} \tag{\% of spam emails}$$
 
$$\phi_{i|y=b} = \frac{\sum_{i=1}^n 1\{x_j^{(i)} = 1, y^{(i)} = b\}}{\sum_{i=1}^n 1\{y^{(i)} = b\}} \text{ for } b = 1, 0$$
 (% of spam (non-spam) emails containing  $j$ -th dictionary word)

$$p(y=1|\mathbf{x})$$

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Choose label y = 1 (spam) if p(y = 1|x) > T, where  $T \in [0, 1]$  is a threshold, e.g., T = 0.5.

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In case of taking the  $x_j$  value in  $\{1, \ldots, k_j\}$ , we can model  $p(x_j|y)$  as multi-nomial distribution, rather than as Bernoulli.

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$$\phi_{j|y=1} = \phi_{j|y=0} = \frac{\sum_{i=1}^n 1\{x_j^{(i)} = 1, y^{(i)} = b\}}{\sum_{i=1}^n 1\{y^{(i)} = b\}} = 0 \,, \quad \text{ for } b = 1, 0$$

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We cannot compute class posterior

$$p(y=1|\mathbf{x}) = \frac{\prod_{j=1}^{d} p(x_j|y=1) p(y=1)}{\left(\prod_{j=1}^{d} p(x_j|y=1)\right) p(y=1) + \left(\prod_{j=1}^{d} p(x_j|y=0)\right) p(y=0)} = \frac{0}{0}$$

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Naïve Baves

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Laplace smoothing can be used to eliminate this issue:

$$\phi_j = \frac{1 + \sum_{i=1}^n 1\{z^{(i)} = j\}}{k+n}$$
 where

- $\phi_j \neq 0$  for all j
- $\sum_{i=1}^{k} \phi_i = 1$

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• In practice we don't apply Laplace smoothing to  $\phi_y = p(y=1)$ , which is greater than 0.

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#### This lecture:

- Discriminative vs. Generative learning algorithms
- Gaussian Discriminant Analysis (GDA)
- Linear Discriminant Analysis (LDA)
- LDA and Logistic regression
- Naïve Bayes