14. (a) Eq. 2.9 & 2.10: p(8/y) BDA HWI Ziyue Wang Jan. 27, 2-21

$$\propto \exp\left\{-\frac{1}{2}\left[\frac{(y+\theta)^{2}}{\sigma^{2}} + \frac{(\theta+b)^{2}}{\tau_{0}^{2}}\right]\right\}$$

$$\propto \exp\left\{-\frac{1}{2}\left[\frac{(\sigma^{2}+\tau_{0}^{2})\theta^{2}-2(y\tau_{0}^{2}+\mu_{0}\sigma^{2})\theta+y^{2}\tau_{0}^{2}+\mu_{0}^{2}\sigma^{2}}{\sigma^{2}\tau_{0}^{2}}\right]\right\}$$

$$\sim \exp\left\{-\frac{1}{2}\left[\frac{(\sigma^{2}+\tau_{0}^{2})\theta^{2}-2(y\tau_{0}^{2}+\mu_{0}\sigma^{2})\theta+y^{2}\tau_{0}^{2}+\mu_{0}^{2}\sigma^{2}}{\sigma^{2}\tau_{0}^{2}}\right]\right\}$$

$$\sim \exp\left\{-\frac{1}{2}\left[\frac{\sigma^{2}+\tau_{0}^{2}}{\sigma^{2}\tau_{0}^{2}}\right]\left[\theta^{2}-2\frac{y\tau_{0}^{4}\sigma^{2}+\mu_{0}\sigma^{4}\tau_{0}^{2}}{\sigma^{2}+\tau_{0}^{2}}\theta+C\right]\right\}$$

$$\sim \exp\left\{-\frac{1}{2}\left[\frac{\sigma^{2}+\tau_{0}^{2}}{\sigma^{2}\tau_{0}^{2}}\right]\left[\theta^{2}-2\frac{y\tau_{0}^{4}\sigma^{2}+\mu_{0}\sigma^{4}\tau_{0}^{2}}{\sigma^{2}+\tau_{0}^{2}}\right]+\left(-\left(\frac{y\tau_{0}^{4}\sigma^{2}+\mu_{0}\sigma^{4}\tau_{0}^{2}}{\sigma^{2}+\tau_{0}^{2}}\right)\right\}$$

$$\sim \exp\left\{-\frac{1}{2}\left[\frac{\sigma^{2}+\tau_{0}^{2}}{\sigma^{2}\tau_{0}^{2}}\right]\left(\theta-\frac{y\tau_{0}^{4}\sigma^{2}+\mu_{0}\sigma^{4}\tau_{0}^{2}}{\sigma^{2}+\tau_{0}^{2}}\right]+\left(-\left(\frac{y\tau_{0}^{4}\sigma^{2}+\mu_{0}\sigma^{4}\tau_{0}^{2}}{\sigma^{2}+\tau_{0}^{2}}\right)\right\}$$

$$\sim \exp\left\{-\frac{1}{2}\left[\frac{\sigma^{2}+\tau_{0}^{2}}{\sigma^{2}\tau_{0}^{2}}\right]\left(\theta-\frac{y\tau_{0}^{4}\sigma^{2}+\mu_{0}\sigma^{4}\tau_{0}^{2}}{\sigma^{2}+\tau_{0}^{2}}\right)\right\}$$

$$\sim \exp\left\{-\frac{1}{2}\left[\frac{\sigma^{2}+\tau_{0}^{2}}{\sigma^{2}\tau_{0}^{2}}\right]\left(\theta-\frac{y\tau_{0}^{4}\sigma^{2}+\mu_{0}\sigma^{4}\tau_{0}^{2}}{\sigma^{2}+\tau_{0}^{2}}\right)\right\}$$

$$\sim \exp\left\{-\frac{1}{2}\left[\frac{\sigma^{2}+\tau_{0}^{2}}{\sigma^{2}\tau_{0}^{2}}\right]\left(\theta-\frac{y\tau_{0}^{4}\sigma^{2}+\mu_{0}\sigma^{4}\tau_{0}^{2}}{\sigma^{2}+\tau_{0}^{2}}\right)\right\}$$

$$\sim \exp\left\{-\frac{1}{2}\left[\frac{\sigma^{2}+\tau_{0}^{2}}{\sigma^{2}\tau_{0}^{2}}\right]\left(\theta-\frac{y\tau_{0}^{4}\sigma^{2}+\mu_{0}\sigma^{4}\tau_{0}^{2}}{\sigma^{2}+\tau_{0}^{2}}\right)\right\}$$

$$\sim \exp\left\{-\frac{1}{2}\left[\frac{\sigma^{2}+\tau_{0}^{2}}{\sigma^{2}\tau_{0}^{2}}\right]\left(\theta-\frac{y\tau_{0}^{4}\sigma^{2}+\mu_{0}\sigma^{2}\tau_{0}^{2}}{\sigma^{2}+\tau_{0}^{2}}\right)\right\}$$

$$\sim \exp\left\{-\frac{1}{2}\left[\frac{\sigma^{2}+\tau_{0}^{2}}{\sigma^{2}\tau_{0}^{2}}\right]\left(\theta-\frac{y\tau_{0}^{4}\sigma^{2}+\mu_{0}\sigma^{2}\tau_{0}^{2}}{\sigma^{2}+\tau_{0}^{2}}\right)\right\}$$

$$\sim \exp\left\{-\frac{1}{2}\left[\frac{\sigma^{2}+\tau_{0}^{2}}{\sigma^{2}+\tau_{0}^{2}}\right]\left(\theta-\frac{y\tau_{0}^{4}\sigma^{2}+\mu_{0}\sigma^{2}\tau_{0}^{2}}{\sigma^{2}+\tau_{0}^{2}}\right)\right\}$$

$$\sim \exp\left\{-\frac{1}{2}\left[\frac{\sigma^{2}+\tau_{0}^{2}}{\sigma^{2}+\tau_{0}^{2}}\right]\left(\theta-\frac{y\tau_{0}^{2}+\tau_{0}^{2}}{\sigma^{2}+\tau_{0}^{2}}\right)\right\}$$

$$\sim \exp\left\{-\frac{1}{2}\left[\frac{\sigma^{2}+\tau_{0}^{2}}{\sigma^{2}+\tau_{0}^{2}}\right]\left(\theta-\frac{\eta^{2}+\tau_{0}^{2}+\tau_{0}^{2}}{\sigma^{2}+\tau_{0}^{2}}\right)\right\}$$

where C, C' are constant w.r.t.  $\theta$ ,

drop  $\exp(-C')$  and define  $T' = \frac{\sigma^2 T_0^2}{\sigma^2 + T_0^2}$ ,  $\mu_1 = \frac{y t_0^4 \sigma^2 + \mu_0 \sigma^4 T_0^2}{\sigma^2 + T_0^2}$ we conclude that  $= \frac{1}{t_0^2 + t_0^2} = \frac{1}{t_0^2 + t_0^2}$   $= \frac{1}{t_0^2 + t_0^2} = \frac{1}{t_0^2 + t_0^2}$   $= \frac{1}{t_0^2 + t_0^2} = \frac{1}{t_0^2 + t_0^2}$ 

Eg. 2.11 & 2.12.

$$p(\theta|y) \propto \exp\left\{-\frac{1}{2}I\frac{\theta^2-2y\theta+c}{\sigma^2/n} + \frac{1}{t_0^2(\theta-\mu_0)^2}\right\}$$
where  $c = \frac{1}{n}\Sigma yi^2$  is constant w.r.t.  $\theta$ .

Thus it reduces to the same format as in Eq. 2.9 & 2.10 with  $\sigma^2$  replaced by  $\frac{n}{\sigma^2}$  and  $\gamma$  replaced by  $\bar{\gamma}$  and the result follows.

(b) We've show that with prior 
$$N(p_0, T_0^2)$$

$$\theta \mid y_1 \sim N(\theta \mid p_1, T_1^2) \qquad \text{(with exchangeability)}$$
where  $p_1 = \frac{\frac{1}{T_0^2} \cdot p_0 + \frac{1}{\sigma^2} y_1}{\frac{1}{T_0^2} + \frac{1}{\sigma^2}}$  and  $\frac{1}{T_1^2} = \frac{1}{T_0^2} + \frac{1}{\sigma^2}$ .

Similarly, with prior N(p1, T,2).

$$\Theta(y_2 \sim \mathcal{N}(\Theta \mid p_2, T_2^2)$$

where 
$$y_2 = \frac{\frac{1}{C_1^2} y_1 + \frac{1}{\sigma^2} y_2}{\frac{1}{C_1^2} + \frac{1}{\sigma^2}}$$
 and  $\frac{1}{C_2^2} = \frac{1}{C_1^2} + \frac{1}{\sigma^2}$ .

or equivalently, 
$$\frac{1}{T_{z}^{2}} = \frac{1}{T_{i}^{2}} + \frac{1}{\sigma^{2}} = \left(\frac{1}{T_{0}^{2}} + \frac{1}{\sigma^{2}}\right) + \frac{1}{\sigma^{2}}$$

$$\lambda_{2} = \frac{\left(\frac{1}{T_{0}^{2}} + \frac{1}{\sigma^{2}}\right) \mu_{1} + \frac{1}{\sigma^{2}} \mu_{2}}{\left(\frac{1}{T_{0}^{2}} + \frac{1}{\sigma^{2}}\right) + \frac{1}{\sigma^{2}}}$$

$$= \left(\frac{1}{T_{0}^{2}} + \frac{1}{\sigma^{2}}\right) \cdot \frac{\frac{1}{T_{0}^{2}} \cdot \mu_{0} + \frac{1}{\sigma^{2}} \mu_{1}}{T_{0}^{2} + \frac{1}{\sigma^{2}}} + \frac{1}{\sigma^{2}} \mu_{2} \right)$$

$$= \frac{1}{T_{0}^{2}} \cdot \mu_{0} + \frac{1}{T_{0}^{2}} \left(y_{1} + y_{2}\right) + \frac{1}{T_{0}^{2}}$$

By induction we have

$$y_n = \frac{\frac{1}{T_0}y_0 + \frac{n}{\sigma^2}\bar{y}}{\frac{1}{T_0^2} + \frac{n}{\sigma^2}} \quad \text{and} \quad \frac{1}{T_0^2} = \frac{1}{T_0^2} + \frac{n}{\sigma^2}$$

15. 
$$E[Z^{m}(1-2)^{n}]$$
:= 
$$\frac{1}{B(d,\beta)} \cdot \int Z^{m}(1-2)^{n} \cdot Z^{d-1} \cdot (1-2)^{\beta+1} dZ$$
= 
$$\frac{B(d+m,\beta+n)}{B(d,\beta)} \cdot \int \frac{1}{B(d+m,\beta+n)} \cdot Z^{d+m-1} (1-2)^{\beta+n-1} dZ$$
= 
$$\frac{B(d+m,\beta+n)}{B(d,\beta)} \cdot 1$$
where 
$$B(a,b) := \frac{\Gamma(d,\beta)}{\Gamma(d+\beta)}$$
(at m=1, n=0, we have
$$EZ = E[Z'(1-2)^{n}]$$
= 
$$\frac{B(d+1,\beta)}{B(d,\beta)}$$
= 
$$\frac{\Gamma(d+1)}{\Gamma(d)} \cdot \frac{\Gamma(d+\beta)}{\Gamma(d+\beta+1)}$$

$$= \frac{\Gamma(d+1)}{\Gamma(d)} \cdot \frac{\Gamma(d+\beta)}{\Gamma(d+\beta+1)}$$

$$= \frac{d}{d+\beta}$$

(et m=2, n=0. re have

$$E Z^{2} = E[Z^{2}(1-Z)^{\circ}]$$

$$= \frac{\Gamma(d+2)}{\Gamma(d)} \cdot \frac{\Gamma(d+\beta)}{\Gamma(d+\beta+2)}$$

$$= \frac{(d+1)d}{(d+\beta)(d+\beta+1)}$$

Thus, 
$$Var(2) = E2^2 - (E2)^2$$

$$= \frac{(d+1)d}{(d+\beta)(d+\beta+1)} - \frac{d^2}{(d+\beta)^2}$$

$$= \frac{d\beta}{(d+\beta)^2(d+\beta+1)}$$

$$(d > 0)$$

$$\beta > 0$$

2.19: (a) (support of y,0,\$\phi\$ omitted)
$$p(\theta|y) \propto p(\theta) p(y|\theta)$$

$$\propto \frac{\beta^{\alpha}}{r(\alpha)} \cdot \theta^{\alpha-1} e^{-\beta\theta} \cdot \theta e^{-\theta y}$$

$$\propto \theta^{(\alpha+1)-1} e^{-\theta(\beta+y)}$$
Thus  $\theta|y \sim Gamma(\alpha+1, \beta+y)$ 

(b) Let  $\theta \sim Gamma(x, \beta)$  and  $\phi := \frac{1}{\theta}$ .

Since  $\frac{1}{x}$  is 1-to-1 transformation of x,  $(\frac{1}{x})^{7} = x$ , we have

$$P_{\theta}(\theta) = \left| \frac{d\theta}{d(y_{\theta})} \right| P_{\theta}(\theta)$$

$$= \left| -\frac{1}{(1/\theta)^{2}} \right| \cdot \frac{\beta^{2}}{\Gamma(\alpha)} \cdot \theta^{2} \cdot \theta^{2} \cdot \theta^{2}$$

$$= \frac{\beta^{2}}{\Gamma(\alpha)} \cdot \theta^{2} \cdot \theta^{2} \cdot \theta^{2} \cdot \theta^{2}$$

$$:= \frac{\beta^{2}}{\Gamma(\alpha)} \cdot \theta^{2} \cdot \theta^{2} \cdot \theta^{2} \cdot \theta^{2}$$

#

Also, we can show 
$$\beta | y \sim inv$$
-Gamma
$$p(\phi | y) \propto p(\phi) p(y | \phi)$$

$$\perp \frac{\beta \lambda}{\rho(\lambda)} \phi^{-\lambda - 1} e^{-\frac{\beta}{\phi}} \cdot \frac{1}{\phi} \cdot e^{-\frac{y}{\phi}}$$

$$\propto \phi^{-\lambda - 2} \rho^{-\frac{\beta + y}{\phi}}$$

Thus, \$14 ~ inv-Gamma (x+1, B+y).

$$SD(\theta) = \sqrt{\frac{\beta^2}{\beta^2}}, \quad E(\theta) = \frac{\alpha}{\beta}$$

Thus, coeff. of variation = 
$$\frac{SD(\theta)}{E(\theta)} = \frac{1}{\sqrt{\lambda}} = \frac{1}{2}$$

Given size n sample 
$$y=(y_1,...,y_n)$$
, the posterior is  $p(\theta|y) \propto \frac{1}{11} (\theta e^{-\theta y_i}) \cdot \theta^{d-1} e^{-\beta \theta}$ 

$$\propto \theta^{\lambda+\eta-1} e^{-\theta(\beta+\frac{\eta}{1-1}y_i)}$$

Thus, 
$$\theta \mid y \sim Gamma \left( d+n, \beta + \sum_{i=1}^{n} y_{i} \right)$$

and the coeff. of variation = 
$$\frac{1}{\sqrt{d+n}} = \frac{1}{10}$$
.

This implies 
$$n = 100 - \lambda = 96$$

(d) Similarly, for 
$$\phi \sim \text{inv-Gamma}(\alpha, \beta)$$
,

the coeff. of variation =  $\frac{SD(\phi)}{E(\phi)} = \sqrt{\frac{\beta^2}{(64)^3(64)^2}} / \frac{\beta}{(64)^3(64)^2}$  (for  $\alpha > 2$ )

$$= \sqrt{\frac{1}{d-2}} = \frac{1}{2}$$

This implies  $\lambda = 6$ .

The posterior of  $\beta$  is  $p(\phi|y) \propto \frac{1}{11} \left( \frac{1}{p} e^{-\frac{y_i}{p}} \right) \cdot \phi^{-\frac{1}{p}} e^{-\frac{\beta}{p}}$   $\lambda = \frac{\beta + \sum y_i}{p}$ Thus,  $\phi|y \sim \text{inv-Gamma}(\lambda + n, \beta + \sum y_i)$ The coeff. of variation  $= \int \frac{1}{\lambda + n^2} = \frac{1}{10}$ This implies  $n = 100 + 2 - \lambda = \frac{96}{2}$ 

The answer does not change.