

Mobility, Mixing and Ergodicity: A Physically-Motivated Measure for Economic Mobility

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Abstract

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^{||}We thank...

1 Introduction

Economic mobility describes “dynamic aspects of inequality” (Shorrocks, 1978). It quantifies how wealth (or income¹) ranks of individuals evolve over time. Intuitively, when mobility is high, ranks evolve quickly, and the chances of an individual to change her position in the wealth distribution over a given time period are high. When mobility is low, individuals are unlikely to change their rank in the distribution over time, or that it changes slowly.

Shorrocks (1978) described several required properties of statistical measures of mobility and set the standard for such measures. Mobility measures are assumed to be derived from a transition matrix, a bistochastic matrix describing the conditional probabilities of individuals to move between different ranks over some time period. An example for such a measure, used extensively in the mobility literature, is the rank correlation, the correlation between individual wealth ranks in two points in time. Another canonical measure of mobility, used most typically in studies of intergenerational income mobility, is the intergenerational earnings elasticity (IGE), defined as the regression coefficient between log-incomes of parents and children.²

The standard statistical measures of mobility have several limitations. First, they are generally incomparable. For example, a rank correlation of 0.3 would have a different meaning if it corresponds to a period of one year or ten years. Second, for a given time period, it is not possible to tell whether some value of rank correlation is high or low, as this is a dimensionless quantity. Thus, rank correlation over some time period can only be high or low in comparison to other economies over a similar time period, or to other time periods of the same length. In addition, the interpretation of the rank correlation depends on the underlying wealth distribution and its dynamics. The same rank correlation cannot be interpreted similarly when the underlying wealth distribution remains unchanged, and when it becomes more and more unequal. The IGE has similar limitations, but also others. Most notably, the IGE is sensitive by-design to the level of inequality and not only to the transition matrix, as discussed in detail in the mobility literature (*e.g.* Chetty et al. (2014)).

This paper introduces mixing time, a property of stochastic processes, as a measure of mobility. When wealth is an ergodic observable (Peters and Adamou, 2018), and assuming the wealth distribution approaches a steady state, if the wealths of an arbitrary group of individuals is followed over time, the distribution of wealth within this group will gradually become similar to the steady-state wealth distribution. The characteristic time of this process is the mixing time. Put simply, it is the time scale over which individuals mix into the wealth distribution. When mixing is rapid, *i.e.* the mixing time is short relative to the window of observation, we could interpret that as high wealth mobility. Slow mixing is interpreted as low mobility.

We then consider Reallocating Geometric Brownian Motion (RGBM (Marsili, Maslov and Zhang, 1998; Liu and Serota, 2017; Berman, Peters and Adamou, 2019)) as a model for wealth dynamics

¹We focus on wealth in this paper, but it applies also to income

²In fact, the rank correlation and the IGE are both measures of immobility, and to consider them as measures of mobility one has to subtract them from 1.

and study mixing in this model. In RGBM, individual wealth undergoes random multiplicative growth, modeled as Geometric Brownian Motion (GBM), and is reallocated among individuals by a simple pooling and sharing mechanism. RGBM is a null model of an exponentially growing economy with social structure. It has three parameters representing economic growth, random shocks to individual wealth, and economic interaction among agents, quantified by the reallocation rate. This model is known to reproduce several important stylized facts. In particular, when the reallocation rate is positive, the wealth distribution converges to a stationary distribution with a Pareto tail. The model has both ergodic and non-ergodic regimes, characterized by the sign of the reallocation rate parameter (Berman, Peters and Adamou, 2019).

We find that in RGBM the mixing time scales with the inverse of the reallocation rate. As the reallocation rate becomes higher, *i.e.* when a larger share of each individual’s wealth is pooled and then shared per unit time, mixing time becomes shorter proportionally, and mobility increases. As the reallocation rate approaches zero, mixing times get longer, and mobility lower. Since decreasing reallocation rates also lead to increasing inequality, this result is in line with the empirical observation that as inequality increases mobility decreases, and vice versa (Corak, 2013).

YB: Here we need to describe how mixing time and rank correlation or other standard measures are related in RGBM YB

In practice, many economic systems are best modeled as non-ergodic (Peters, 2019). In particular, Berman, Peters and Adamou (2019) argue that the US economy is best described in RGBM as one in which wealth is systematically reallocated from poorer to richer, *i.e.* the reallocation rate is negative. In such a case there is no mixing, so the mixing time is infinite. Thus, measuring mobility using standard measures under this regime may be misleading. The thorough study of RGBM in this regime is outside of the scope of this paper and left for future work.

The paper is organized as follows. Section 2 discusses the concept of mixing time and how it provides a physically-motivated measure for mobility. Section 4 studies mobility using mixing times in reallocating geometric Brownian motion as a model for wealth. We conclude in Section 6.

2 Standard measures of economic mobility

To motivate the implementation of mixing time, we begin by showing the characteristics of standard measures of economic mobility. These measures examine the properties of the bivariate joint distribution describing the wealth of the population in two different periods. Various methods have been developed for evaluating mobility under these circumstances, which, in general, exhibit similar characteristics. Therefore, in order to depict these characteristics we are going to utilize the three most widely used measures of economic mobility: the Spearman rank correlation, the intragenerational earnings elasticity and the wealth transition matrix.

2.1 Quantifying economic mobility

Spearman rank correlation: The Spearman rank correlation between the wealths corresponding to two different time periods, t_m and t_n ($t_m < t_n$) is defined as

$$r_{t_m, t_n} = 1 - \frac{6 \sum_i (rg(x_i(t_m)) - rg(x_i(t_n)))^2}{N(N^2 - 1)},$$

where $rg(x)$ is the rank transformation of x , $x_i(t)$ is the wealth of individual i in period t and N is the population size. This measure is bounded between -1 and 1 , with lower values suggesting greater economic mobility.

Intragenerational earnings elasticity: The second measure is the intragenerational earnings elasticity, quantified via the slope b_{t_m, t_n} of the regression

$$x_i(t_n) = b_0 + b_{t_m, t_n} x_i(t_m) + u_i,$$

where b_0 is the intercept and u_i is the error term. This is a simple linear regression and therefore,

$$b_{t_m, t_n} = \text{corr}(x(t_n), x(t_m)) \frac{\text{var}(x(t_n))}{\text{var}(x(t_m))}, \quad (2.1)$$

where $\text{corr}(x, y)$ is the correlation between the variables x and y and $\text{var}(x)$ is the variance of x . As with the rank correlation, lower intragenerational earnings elasticity also indicates greater mobility. However, this measure is unbounded and may take on any real values.

Wealth transition matrix: The last measure, disaggregates wealth rankings and summarizes economic mobility in a transition matrix \mathbf{A} in which cell entries A_{kl} show the probability that an individual in wealth quantile k in period t_m is found in wealth quantile l in period t_n . In a perfectly mobile economy, the entries of the transition matrix are all equal to each other. In an immobile economy, on the other hand, the largest values are concentrated in the diagonal entries.

2.2 Limitations

To visualize the limitations of the standard measures, in Fig 1 we construct an artificial example where we examine the economic mobility in a population of 9 individuals during 3 time periods. In Fig 1A, we show the dynamics of the wealth rankings. Besides this, we also highlight the amount of wealth owned by a person with a colored circle, with darker colored and larger circle implying a wealthier person.

In our artificial economy, between the three periods, only the four individuals with largest wealth are able to change their status. In particular, in the second period, t_2 the second wealthiest person

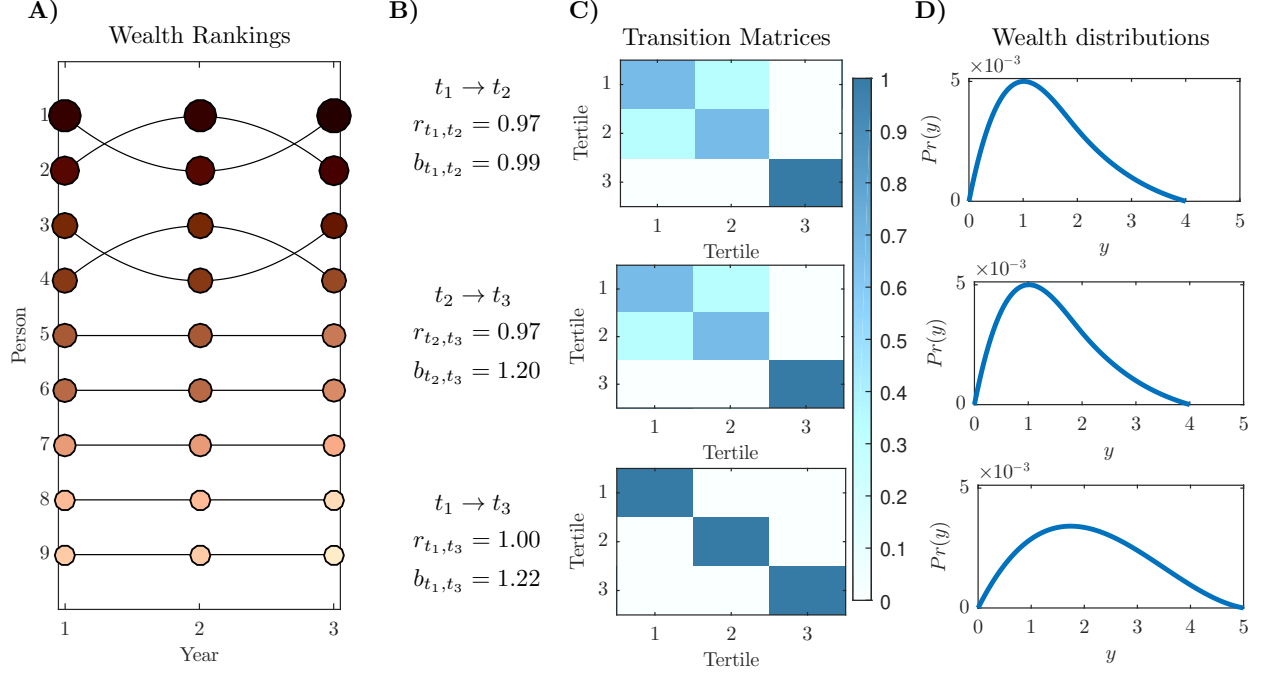


Figure 1: **Performance of mobility measures.** **A)** Dynamics of the wealth rankings in an artificial economy consisting of 9 people, in a span of 3 years. **B)** Estimated Spearman correlation and Intragenerational earnings elasticity during the studied periods. **C)** The corresponding mobility transition matrices. **D)** The distribution of wealth for the 3 time periods.

becomes the one with the most wealth, while the wealthiest person in the first period t_1 fails to second position. In addition, the third and fourth wealthiest persons change position. In period t_3 , the dynamics are reversed, thus ending up with the same wealth rankings as in t_1 . The only difference is that in this period, the wealth inequality increased, *i.e.*, the rich got richer and the poor got poorer.

As a result, the rank correlation tells us that the mobility was the same between t_1 and t_2 , and t_1 and t_3 , but overall between t_1 and t_3 there was no mobility ($r_{t_1, t_3} = 1$) (Fig 1B). The intragenerational earnings elasticity gives us similar results. In fact, since inequality increased in the last period, it further suggests that the economy became less mobile (Fig 1D).

The same conclusions hold when looking at the transition matrices estimated by dividing the wealth rankings in tertiles, as shown in Fig 1C. An astonishing observation of the last measure is that the transition matrices between t_1 and t_2 and t_2 and t_3 suggest that there is a 1/3 chance for a person belonging in the second tertile to climb to the richer tertile and vice versa. However, we structured our economy in a way that allows movement only between the persons belonging on the edge of the tertile. Hence, the transition matrices fail to adequately represent the movement of the typical person in the tertile.

Obviously, standard measures of economic mobility represent aggregate values of the changes in the wealth rankings of the individuals which constitute the population between two time periods.

Therefore, they i) rely on a relevant time period for which the wealth rankings are compared, and ii) do not quantify the mobility of the typical representative of the population.

3 Mixing time

3.1 What is a mixing time?

Mixing time is able to overcome the limitations of standard economic mobility measures by providing an estimate for the characteristic time scale over which individuals mix into the wealth distribution. Think of the economy as a cup of coffee and the typical person's wealth as the milk you pour in the coffee. Mixing time quantifies the time required for the milk to blend with the coffee. This enables the measure to be used for appropriate comparison between different time periods and economies.

As we will see in the following, there is a direct relation between mixing time and standard mobility measures, whenever mixing time is a finite quantity. However, the standard measures may still indicate that there is mobility even when mixing time does not exist. This is because, mixing time depends on the existence of mobility between *every* quantile in the wealth distribution, whereas for standard measures it is enough for to have transitions only between two quantiles for mobility to exist. In other words, mixing time not only describes the capabilities of poor individuals to rise and become leading social actors, but also evaluates the time required for a rich person to fall in the social ladder. Therefore, we believe that our measure embodies the true concept of mobility.

3.2 Measuring mixing time

In physical terms, mixing describes the property of a dynamical system of being strongly intertwined. That is, any set of particles moving according to the laws of the dynamical system and satisfying a small spread of uncertainty in their initial conditions, follow paths that enter into any region of the phase space, and in a relatively "uniform" way. After a sufficient period of time, the mixing time, the percentage of the particles found within a particular region, is proportional to the volume of that region.

As defined, mixing is strongly related to the concept of ergodicity. However, the latter is a much broader concept: A dynamical system is said to be ergodic if time-averages are equal to ensemble averages. Put differently, ergodicity only implies that every trajectory spreads around the phase space. Hence, every dynamical system that is mixing is also ergodic, but the opposite is not necessarily true.

The ergodicity indicates that in a mixing economy there is always a stationary distribution to which some rescaled transformation of a person's wealth converges. This characteristic can be utilized to develop an estimation procedure for mixing time. Fig 2 gives an example for how to estimate

mixing time, whereas the formal procedure is as follows:

1. **Select stationary distribution:** Set a target distribution to which the rescaled transformation of wealth converges (black line in the inset plots);
2. **Select typical representatives:** Select a subsample of the population constituted of people close to the typical representative (red dashed line in the inset plots). For example, this can be the persons whose wealth is closest to either the mean/median/mode wealth, depending on the properties of the stationary distribution;
3. **Track wealth dynamics:** Track the wealth dynamics of each person in the subsample for a sufficient amount of time. Snapshots for the wealth distribution at several time periods are given in the inset plots of Fig 2.
3. **Quantify the differences between the subsample and the stationary wealth distributions:** At certain time periods quantify the differences between the subsample wealth distribution and the stationary distribution. This can be done by utilizing usual statistical distribution distance measures such as the Kullback-Leibler divergence, the Hellinger distance or the Kolmogorov-Smirnov statistic. In Fig 2, with a blue line we plot the the log of one such statistic as a function of time. In a mixing economy the statistic will exhibit three states. First, there will be a transient state in which the differences between the subsample and stationary wealth distributions will be large and stable. Next, there will be a mixing time state during which the two distributions will slowly converge towards each other, and the statistic will slowly decrease. Finally, there will be a stable state with no significant differences between the stationary and subsample wealth distributions;
1. **Estimate mixing time:** Estimate the slope of the regression in which the dependent variable is the log of the statistic and the independent is time (black dash-dotted line in the figure). Take the data only for the mixing time state. Mixing time is the additive inverse of the reciprocal value of the slope ($-1/slope$).

4 Mixing in a simple model of an economy

4.1 Reallocating geometric Brownian motion

To illustrate the application of mixing time in economic systems, we utilize the RGBM model. Under RGBM, the dynamics of the wealth of person i is specified as

$$dx_i = x_i (\mu dt + \sigma dW_i) - \tau (x_i - \langle x \rangle_N) dt, \quad (4.1)$$

with $\mu > 0$ being the drift term, $\sigma > 0$ the noise amplitude, and dW_i is an independent Wiener increment, $W_i(t) = \int_0^t dW_i$. In the equation τ is a parameter that quantifies reallocation of wealth.

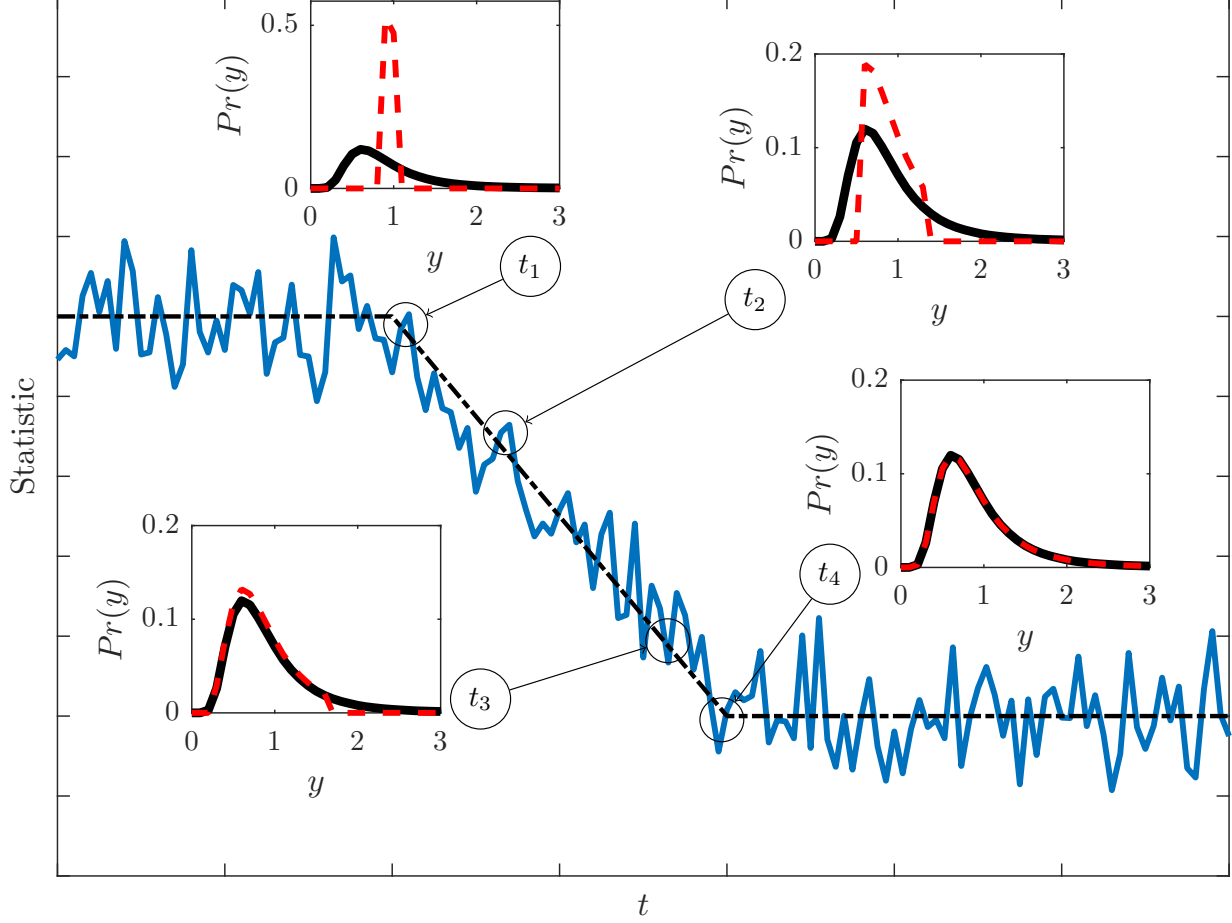


Figure 2: **Estimating mixing time.** Blue line: The log of the statistical distance between the subsample and population wealth distributions. **Black dash-dotted line**: Estimated slope of the regression of the log of the statistic on time. Inset plots: Snapshots of the empirical population wealth distribution (black line) and the subsample wealth distribution (red dashed line) at time periods t_1 , t_2, t_3 and t_4 .

It implies that every year, everyone in the economy contributes a proportion τ of their wealth to a central pool, and then the pool is shared evenly across the population ($\langle x \rangle_N$ is per-capita wealth). The parameter aggregates a multitude of effects: collective investment in infrastructure, education, social programs, taxation, rents paid, private profits made etc.

The dynamical behavior of RGBM is strictly dependent on the relation between τ and σ , and can be both ergodic and non-ergodic. Since mixing time is predicated on ergodic dynamics, we are going to focus on the ergodic regime which occurs when $\tau > \frac{\sigma^2}{2}$. Then, the model exhibits mean-reversion as each x_i reverts to the population average $\langle x \rangle_N$. The large population approximation $\langle x(t) \rangle_N = \exp[\mu t]$ is valid, and it can be used to write a mean field equation for the dynamics of the rescaled wealth $y_i = x_i / \langle x \rangle_N$ as

$$dy = y\sigma dW - \tau(y - 1)dt. \quad (4.2)$$

Notice that the mean field representation leads to decoupling of the dynamics of each individual i , and hence we can neglect the subscript i in equation (4.2). This is a convenient form of RGBM which allows for deriving powerful analytical results regarding the properties of the model. For instance, we can derive the stationary distribution for the rescaled wealth, which reads

$$p(y) = \frac{(\theta - 1)^\theta}{\Gamma(\theta)} \exp\left(-\frac{\theta - 1}{y}\right) y^{-(1+\theta)}, \quad (4.3)$$

where $\theta = 1 + \frac{2\tau}{\sigma^2}$ and $\Gamma(\cdot)$ is the Gamma function. The distribution has a power-law tail as $y \rightarrow \infty$, resembling Pareto's often confirmed observation that the frequency of large wealths tends to decay as a power law. The exponent of the power law, θ , is called the Pareto parameter and is one measure of economic inequality. More importantly, the usual stylized facts are recovered: the larger σ (more randomness in the dynamics) and the smaller τ (less reallocation), the smaller the tail index and the fatter the tail of the distribution (higher inequality).

4.2 Mixing time in RGBM

In RGBM, mixing time can be analytically estimated from its correlation function,

$$\text{corr}(y(t), y(t + \delta)) = \exp[-\tau\delta]. \quad (4.4)$$

This function measures the statistical relation between observations of rescaled wealth at two different time points t and $t - \delta$.

Let us define the time period t^* as the transition at which the distribution of the subsample converges to the stationary distribution. Then, we can reformulate the stationary distribution in terms of this period as,

$$\text{corr}(y(t^*), y(t)) = \exp[-\tau(t^* - t)],$$

for any time period t that belongs to the mixing time state. Obviously, the magnitude of any statistical measure $S(t)$ will be proportional to the value of this correlation, and its log will be of the form

$$\log(S(t)) = -\tau(t^* - t), \quad (4.5)$$

which implies that the mixing time in RGBM is $1/\tau$. The interpretation behind this result is fairly intuitive – in an economy in which reallocation from the rich to the poor is larger, there will be also more mixing. On the other hand, as the reallocation rate approaches zero, mixing times get longer, and mobility lower. Since decreasing reallocation rates also lead to increasing inequality, this result is in line with the empirical observations which suggests that as inequality increases mobility decreases, and vice versa ([Corak, 2013](#)).

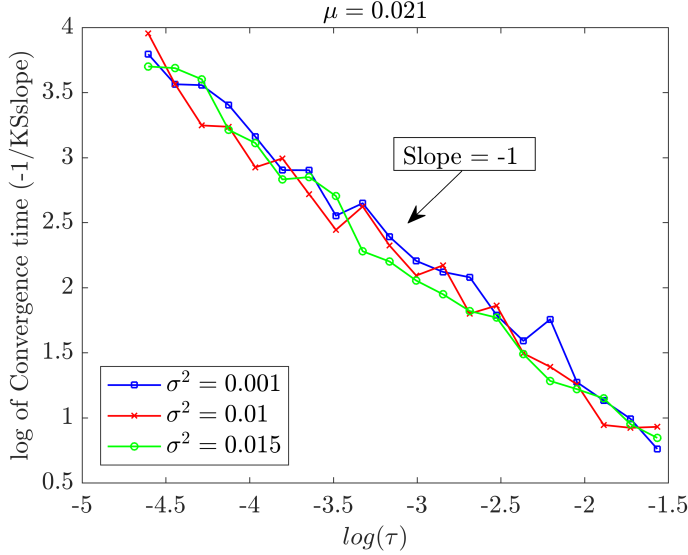


Figure 3: Mixing time in RGBM. Describe the figure.

Fig 3 compares our analytical expression with numerical simulations where mixing time was estimated through the procedure described in Section 3.2. In the estimation procedure we chose Kolmogorov-Smirnov statistic as a statistical distance measure.

4.3 Mixing time and standard measures of mobility

Differently from the mixing time, in RGBM some of the standard measures mobility, besides being dependent on the reallocation parameter, are also determined by the noise amplitude σ . This is a consequence of the randomness playing a significant role in the wealth dynamics when we consider time frames that are shorter than the mixing time.

In what follows we evaluate the relationship between mixing time and the standard measures of mobility in RGBM.

Spearman rank correlation: In the ergodic regime of RGBM the Spearman rank correlation is inversely related to the mixing time. However, the rank correlation is also dependent on the noise amplitude σ and the temporal difference δ between the two periods that are being compared. Larger values for both parameters lead to a greater economic mobility. This can be seen in Fig. 4 where we plot the additive inverse of log of the rank correlation divided by δ as a function of τ for various noise amplitudes.

Intragenerational earnings elasticity: In RGBM, once the dynamics reaches a stationary state, the variance of $y(t)$ is invariant on time, *i.e.*, the variances term in equation (2.1) cancel out. Hence, the intergenerational earnings elasticity reduces simply to the correlation between

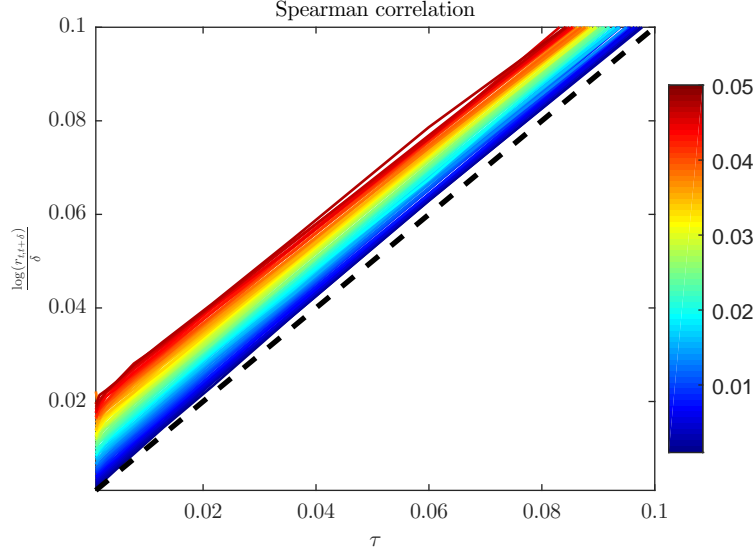


Figure 4: **Spearman rank correlation in RGBM.** The colored lines correspond to numerical estimations of the stationary Spearman rank correlation in RGBM, color-mapped according to the magnitude of the noise amplitude. In the simulation, $\sigma^2 \in [0.001, 0.05]$, $\delta = 20$, $N = 10^4$.

observations corresponding to two different time periods as given by equation (4.4). In other words, the only difference between mixing time and the intragenerational earnings elasticity is the shift in the magnitude of the later due to the dependence on the time period over which it is estimated. This finding is confirmed in Fig. 5, where we show the log of the stationary intragenerational earnings elasticity normalized for the time frame δ as a function of the parameter τ and vary σ^2 .

Transition matrices: As the transition matrix divides the wealth distribution into Q quantiles, it gives us a total of Q^2 measures for mobility, each quantifying the relationship between pairs of quantiles. Instead of studying each of those measures, we will focus on one aggregate quantity called the Global Mean First Passage Time (GMFPT). Formally, the mean first passage time (MFPT) between quantiles k and l gives an estimate of the time required for an individual whose wealth belongs to quantile k to move to quantile l , and the GMFPT is the average across all possible MFPT pairs. By construction, this quantity is measured in the same units as mixing time, and as such it can give us a direct relation between the properties of the standard measures and mixing time in RGBM.

As evidenced in Fig. 6, similarly to the rank correlation, the behavior of the transition matrices is dependent on both the noise amplitude and the temporal differences between the periods for which it is estimated. However, differently from the previous measures it appears that there is no linear transformation for the relationship between GMFPT and the mixing time. *Maybe elaborate...*

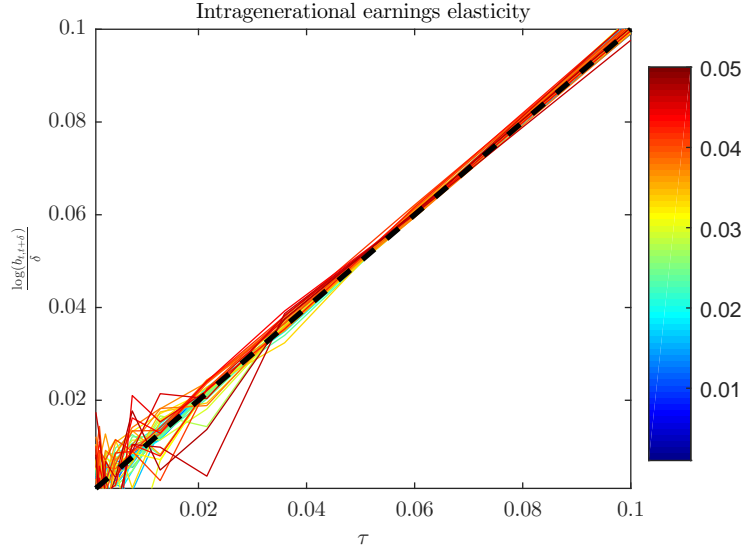


Figure 5: **Intragenerational earnings elasticity in RGBM.** The colored lines correspond to numerical estimations of the stationary Intragenerational earnings elasticity in RGBM, color-mapped according to the magnitude of the noise amplitude. In the simulation, $\sigma^2 \in [0.001, 0.05]$, $\delta = 20$, $N = 10^4$.

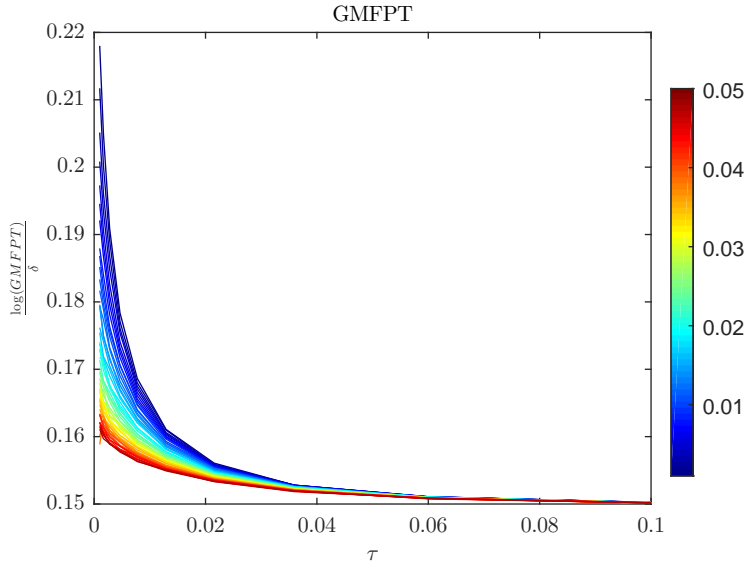


Figure 6: **GMFPT in RGBM.** The colored lines correspond to numerical estimations of the stationary GMFPT in RGBM, color-mapped according to the magnitude of the noise amplitude. In the simulation, $\sigma^2 \in [0.001, 0.05]$, $\delta = 20$, $N = 10^4$.

5 Discussion

- maybe merge figs 4-6 into one figure
- comment on non-ergodicity

- comment relationship with inequality
- comment on real data?

6 Conclusion

References

- Berman, Yonatan, Ole Peters, and Alexander Adamou.** 2019. “Wealth Inequality and the Ergodic Hypothesis: Evidence from the United States.” Available at SSRN.
- Chetty, Raj, Nathaniel Hendren, Patrick Kline, and Emmanuel Saez.** 2014. “Where is the Land of Opportunity? The Geography of Intergenerational Mobility in the United States.” Quarterly Journal of Economics, 129(4): 1553–1623.
- Corak, Miles.** 2013. “Income Inequality, Equality of Opportunity, and Intergenerational Mobility.” Journal of Economic Perspectives, 27(3): 79–102.
- Liu, Z., and Rostislav A. Serota.** 2017. “Correlation and Relaxation Times for a Stochastic Process with a Fat-tailed Steady-state Distribution.” Physica A, 474: 301–311.
- Marsili, Matteo, Sergei Maslov, and Yi-Cheng Zhang.** 1998. “Dynamical Optimization Theory of a Diversified Portfolio.” Physica A, 253: 403–418.
- Peters, Ole.** 2019. “The Ergodicity Problem in Economics.” Nature Physics, 15: 1216–1221.
- Peters, Ole, and Alexander Adamou.** 2018. “Ergodicity Economics.” 5.0, 2018/06/30.
- Shorrocks, Anthony F.** 1978. “The Measurement of Mobility.” Econometrica, 5: 1013–1024.