Yassin_Eslam_Tuto1_correction

Question 1

Question 1.a

- When extracting the parameter θ from the product, θ should have been exponentiated to n.
- The indicator function is also impacted by the product operation resulting from computing the likelihood of several observations. The product of indicator functions $\forall i \in \{1,\dots,n\},\ 1_{x_i \geq \eta}$ is equal to the indicator function $1_{\min(X) \geq \eta}$ as any x below η implies that the whole product is null.

$$egin{aligned} f_i(heta,\eta) &= heta e^{- heta(x_i-\eta)} \mathbb{1}_{x_i \geq \eta} \quad ext{(with f the density of X_i)} \ \mathcal{L}(heta,\eta) &= \prod_{i=1}^n f_i(heta,\eta) \ &= \prod_{i=1}^n heta e^{- heta(x_i-\eta)} \mathbb{1}_{min(x_i) \geq \eta} \ &= heta^n e^{- heta \sum_{i=1}^n (x_i-\eta)} \mathbb{1}_{min(X) \geq \eta} \end{aligned}$$

Question 1.b

Let's recall the instruction:

Simulate n = 10 i.i.d ISI with this model. For fixed theta, plot the likelihood as a function of eta.

The exercise meant to have a fixed true value for both θ and η (but where the θ is actually known, η would not in theory).

Then the goal was to plot a 2-dimensional graph where the x-axis would represent a range of candidates for η and the y-axis the corresponding likelihood (for a given known θ).

In this situation, it is expected that the plot will display the maximal likelihood where the x-axis equates or gets closer to the true η .

The answer's main issue was to try generating the 10 observations using R's standard library normal distribution function rnorm. The actual process would have been to:

- 1. Generate 10 observations using the exponential distribution with true parameter θ , then adding the true delay η (The exponential function can be generated by R using the standard library function rexp.)
- 2. Generate a fine, uniform sequence of candidates η using the R standard library function seq
- 3. Map the likelihood function over that generated sequence (provided the fixed known θ) to yield a likelihood at each candidate η
- 4. Plot the likelihood given η graph

Notes

ullet As noted in Question 1.a , the likelihood function lik missed the exponentiation of heta in computing the variable out

```
lik <- function(f_eta){
  theta<- f_eta[1] #theta
  eta<- f_eta[2]
  #out <-theta*exp(-theta*(sum(x)-length(x)*eta))
  out <- theta^n*exp(-theta*sum(x-eta))
  return(out)
}</pre>
```

Question 1.c

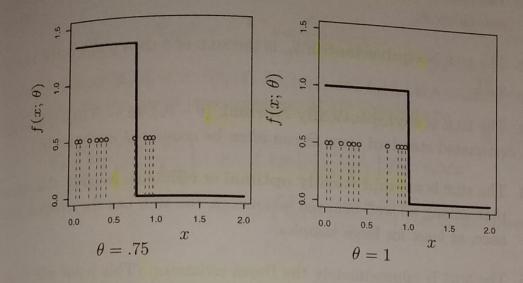
About the MLE of η

Overall, unless a MLE is an actual constant, the estimator $\hat{\eta}$ should have been represented as a function of the observations (Wasserman, L., All of Statistics, p107: "Let $T_n=g(X_1,\ldots,X_n)$ be a statistic, that is, T_n is any function of the data.")

The trick is to see that the model's maximum likelihood for η is very similar to the maximum likelihood definition of the lower bound parameter of a uniform distribution (https://en.wikipedia.org

/wiki/Continuous_uniform_distribution#Maximum_likelihood_estimator). The demonstration for the uniform distribution can be found on StackExchange (https://math.stackexchange.com/questions/411145/maximum-likelihood-estimation-of-a-b-for-a-uniform-distribution-on-a-b).

Note: Check the bottom right graph in the attached picture (Wasserman, L., All of Statistics, p125)



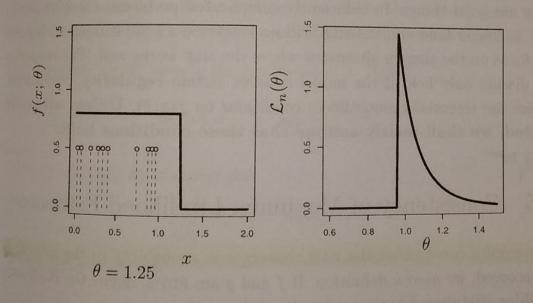


FIGURE 9.2. Likelihood function for Uniform $(0,\theta)$. The vertical lines show the Observed data. The first three plots show $f(x;\theta)$ for three different values of θ .

When θ When $\theta < X_{(n)} = \max\{X_1, \dots, X_n\}$, as in the first plot, $f(X_{(n)}; \theta) = 0$ and hence $f(X_n) = \max\{X_1, \dots, X_n\}$, as in the first plot, $f(X_n) = 0$ and hence $\lim_{n \to \infty} \frac{1}{n} \int_{n}^{\infty} X_{(n)} = \max\{X_1, \dots, X_n\},$ as in the first plot, $f(X_{(n)}, x)$ and hence $f(X_i; \theta) = \prod_{i=1}^n f(X_i; \theta) = 0$. Otherwise $f(X_i; \theta) = 1/\theta$ for each i and hence $f(X_i; \theta) = \prod_{i=1}^n f(X_i; \theta) = (1/\theta)^n$. The last plot shows the likelihood function.



The idea is that it is the reversed problem: the likelihood is correlated to the value of the exponential distribution, and thus η is upper-bounded by the lowest value among the observations x_i .

About the MLE of θ

Besides the erroneous $\hat{\eta}$ (mentioned above) and the missing exponentiation of the first instance of the parameter θ (yielding $n \log \theta$ in the log-likelihood formula), the derivation process is correct.

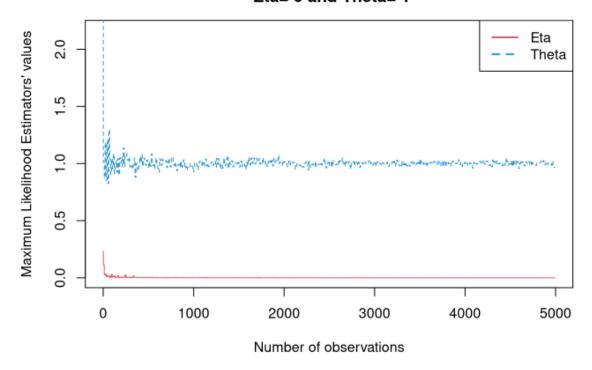
Question 1.d

Similar observations as Question 1.b with regards to modeling ISIs: the use of a normal distribution breaks the process. One correct way to broach the question would be the following:

- 1. Generate a sequence of observations numbers (e.g. 10, 100, 1000, etc.) and for each number n:
 - a. Generate n observations using the exponential distribution with the true parameter θ , then adding the true delay η
 - b. Compute the MLE $\bar{\theta}$ and $\hat{\eta}$ and records them in an array/matrix/data.frame
- 2. Plot the array/matrix/data.frame of MLEs (y-axis) given the number of observations n (x-axis)

This can be repeated for various true values for both θ and η . Example:

MLE convergence plot for the ISI model with Eta= 0 and Theta= 1



Question 1.e

A similar correction to the likelihood function (as mentioned in Question 1.b) should be applied here.

When the concept of time window was mentioned in the instructions, it indicated, for instance, that a MLE for heta could

have been computed on a. possibly rolling, window for instance (e.g. computing a MLE at every half-seconds; computing a MLE before and after an apparent spike, etc.)

Question 1.f

Not completed

Question 2

Question 2.a

Everything's good. No comment.

Question 2.b

Not completed

Question 2.c

Not completed

Question 2.d

Not completed

Question 2.e

Not completed