

# Various generalizations of the Upper-Lower Index and their critical values

—  
Psychoco 2019

Lubomír Štěpánek<sup>1</sup>

Patrícia Martinková<sup>2, 3</sup>



<sup>1</sup>Institute of Biophysics and Informatics  
First Faculty of Medicine  
Charles University, Prague



FACULTY  
OF EDUCATION  
Charles University

<sup>2</sup>Faculty of Education  
Charles University, Prague



<sup>3</sup>Institute of Computer Science  
Czech Academy of Sciences

# Content

- 1 Introduction
- 2 Revisiting the Upper-Lower Indices
- 3 Critical values for Upper-Lower Indices
- 4 Future work
- 5 Summary

# Concept of item discrimination

- one of the item parameters characterizing item properties
  - besides item difficulty, pseudo-guessing, inattention
- describes how accurately an item can discriminate between individuals with high vs. with low ability
- item discrimination can be evaluated using both Classical Test Theory (CTT) and Item-Response Theory (IRT)

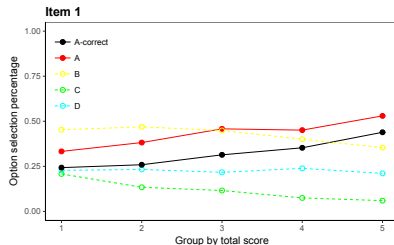
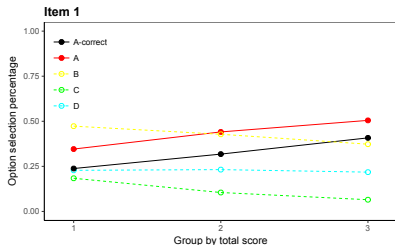
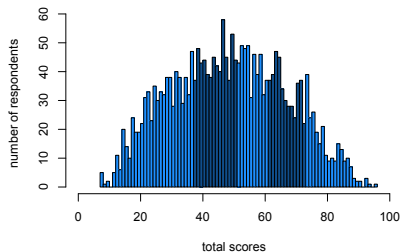
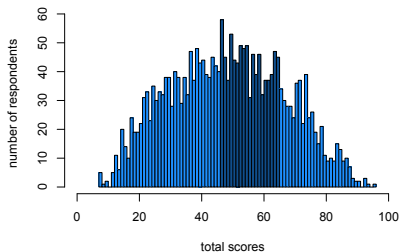
# Indicators of item discrimination

- based on (Pearson's) correlation
  - item-test correlation (RIT)
  - item-rest correlation (RIR)
  - tetrachoric coefficient of correlation,  $\chi^2$  statistics
- based on difference of proportions of correct answers
  - Kelley's discrimination index (classical Upper-Lower Index, ULI)
  - its variants and generalizations (generalized Upper-Lower Index, gULI)

# Motivation

- traditional methods
  - easily explainable to practitioners
  - easily calculated
- methods based on IRT
  - sophisticated models using  $p$ -values, information criteria etc.

# Motivation (con't)



# Aims of the study

- to do literature review
- to look into properties of generalizations of ULI (gULI)
- to find reasoning behind suggested critical values

# Kelley's item discrimination index

- let's assume a test with  $n$  test-takers and  $k$  binary items
- using ordered total scores of the  $n$  test-takers, the group is divided into three parts considering only highest  $m$  and lowest  $m$  of the test-takers
- assuming that  $p_U$ , or  $p_L$  is a proportion of correct answers in the highest, or lowest  $m$  test-takers, then<sup>1</sup>

$$\text{ULI} \equiv p_U - p_L$$

- Kelley suggests to choose  $m$  such that  $\frac{m}{n} \doteq 0.27$ , but did not offer any critical values for the calculated ULI

---

<sup>1</sup>T. L. Kelley. "The selection of upper and lower groups for the validation of test items." In: *Journal of Educational Psychology* 30.1 (1939), pp. 17–24. DOI: 10.1037/h0057123. URL: <https://doi.org/10.1037/h0057123>



# Kelley's optimal splitting of a group of test takers

- firstly let's assume there are  $n$  test-takers with scores sorted increasingly,  $(x_{(1)}, x_{(2)}, \dots, x_{(n)})^T$  such that  $\forall i \in \{1, 2, \dots, n\}$  is  $x_{(i)} \in \mathbb{R}$  and the scores follow  $\mathcal{N}(0, 1^2)$
- Kelley searched for  $q = \frac{m}{n}$  in order to maximize  $t$  statistics of two sample  $t$ -test
- assuming that  $\mu_U$  and  $\mu_L$  are mean scores of upper and lower  $m$  test-takers,  $\sigma$  is a pooled standard deviation of test scores,

$$t = \frac{\mu_U - \mu_L}{\frac{\sigma}{\sqrt{m}}}$$

- given the assumed  $\mathcal{N}(0, 1^2)$  distribution, therefore and equal size we get  $\mu_U = -\mu_L$  and

$$t = \frac{2\mu_U \cdot \sqrt{m}}{\sigma} \propto \mu_U \cdot \sqrt{m}$$

# Kelley's optimal splitting of a group of test takers

- since mean of the truncated normal distribution is  $\mu_U = \frac{z_q}{q}$  where  $z_q$  is  $q$ -th quantile and  $q = \frac{m}{n}$  we get

$$t \propto \mu_U \cdot \sqrt{m} \propto \frac{z_q}{q} \cdot \sqrt{m} \propto \frac{z_q}{q} \cdot \sqrt{qn} \propto \frac{z_q}{\sqrt{q}} \cdot \sqrt{n} \propto \frac{z_q}{\sqrt{q}}$$

- the first derivative with respect to  $x$  where  $x$  and  $-x$  are points of division in a unit normal distribution is

$$\frac{dt}{dx} \propto \frac{z}{\sqrt{q}} \left( -x + \frac{z}{2q} \right)$$

- setting the derivation equal to zero we get

$$q = \frac{z}{2x}$$

and using tables for the truncated normal distribution we could find  
 $q \doteq 0.270$

# Generalized Upper-Lower Index for binary items by Brennan

- Brennan generalized the ULI (here assigned  $ULI_B$ ) by relaxing the assumption of equal sizes of the upper/lower group of test-takers<sup>2</sup>
- let's assume a test with  $n$  test-takers and  $k$  binary items
- using ordered total scores of the  $n$  test-takers, the group is divided into three parts considering only highest  $n_U$  and lowest  $n_L$  of the test-takers
- assuming that  $m_U$ , or  $m_L$  is a number of correct answers in the highest  $n_U$ , or lowest  $n_L$  test-takers, then

$$ULI_B \equiv \frac{m_U}{n_U} - \frac{m_L}{n_L}$$

---

<sup>2</sup>Robert L. Brennan. "A Generalized Upper-Lower Item Discrimination Index". In: *Educational and Psychological Measurement* 32.2 (July 1972), pp. 289–303. DOI: 10.1177/001316447203200206. URL: <https://doi.org/10.1177/001316447203200206>

# Critical values for Upper-Lower Index by Brennan

- Brennan also offered critical values for the  $ULI_B$  searching for the minimal values of  $ULI_B$  for which  $z$  statistic of  $z$ -test for two proportions becomes significant
- under null hypothesis  $H_0 : p = p_U = p_L$  where  $p_U$  and  $p_L$  is a proportion of correct answers in both upper and lower part

$$z = \frac{\left( \frac{m_U}{n_U} - \frac{m_L}{n_L} \right) - 0}{\sqrt{\frac{n_U + n_L}{n_U n_L} p(1 - p)}}$$

follows standard normal distribution

- for  $n_U = n_L = 30$  and  $p = p_U = p_L = 0.5$  such that  $\mathbb{E}\left(\frac{m_U}{n_U}\right) = p_U$  and  $\mathbb{E}\left(\frac{m_L}{n_L}\right) = p_L$  he got the well-known critical value for the ULI (or  $ULI_B$ ) equal to 0.3 (approx.)

# Generalized Upper-Lower Index for polytomous items by Metsämuuronen

- Metsämuuronen generalized the ULI in order to make it applicable to polytomous items

$$\text{GDI}_m = \frac{\sum_{i=1}^m O_i^U - \sum_{i=1}^m O_i^L}{m(\max(O_i) - \min(O_i))}$$

where  $O_i^U$  and  $O_i^L$  refer to the score of the  $i$ -th highest and lowest test-taker given  $m$  highest and lowest observations<sup>3</sup>

---

<sup>3</sup>Jari Metsämuuronen. "Generalized Discrimination Index". en. In: (2018). DOI: 10.13140/rg.2.2.30933.88804/1. URL:

<http://rgdoi.net/10.13140/RG.2.2.30933.88804/1>

# Simulated optimal splitting of a group of test takers

- based on Kelley, let's assume a test with  $n$  test-takers and  $k$  binary items
- using ordered total scores of the  $n$  test-takers, the group is divided into three parts considering only highest  $m$  and lowest  $m$  of the test-takers
- assuming that  $p_U$ , or  $p_L$  is a proportion of correct answers in the highest, or lowest  $m$  test-takers, then

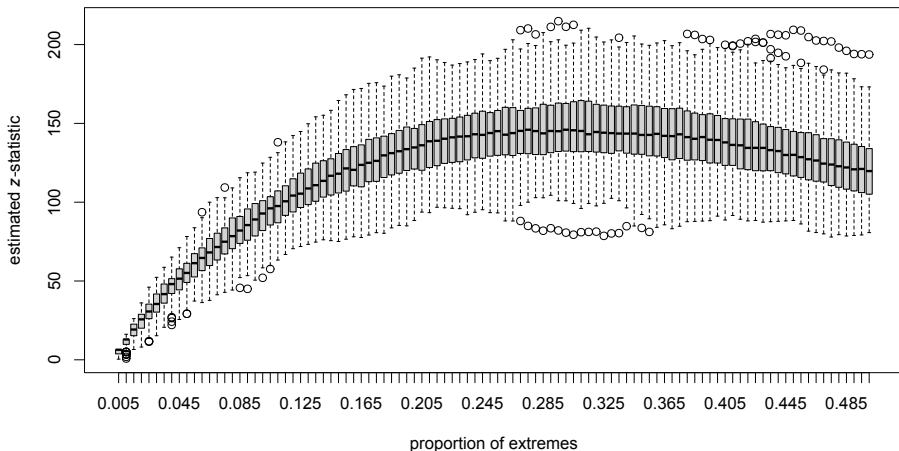
$$\text{ULI} \equiv p_U - p_L$$

- let's suppose a probability of correct answer is given by  $\frac{\exp(\theta)}{1+\exp(\theta)}$  where  $\theta \sim \mathcal{N}(0, 1^2)$  and  $p_L \sim \text{Bernoulli}\left(\frac{\exp(\theta)}{1+\exp(\theta)}\right)$  as well as  $p_U \sim \text{Bernoulli}\left(\frac{\exp(\theta)}{1+\exp(\theta)}\right)$

# Simulated optimal splitting of a group of test takers

```
n := 1000      // size of a sample of test-takers (or their scores);
k := 200       // number of repetitions per one given proportion i;
z := {}        // a list of z-statistics;
for i = 1 : ⌊ $\frac{n}{2}$ ⌋ do
  z-vector := {}      // vector of z-statistics for given i;
  for j = 1 : k do
    abilities := vector of n numbers following  $\sim \mathcal{N}(0, 1^2)$ ;
    scores := {};
    for  $\theta$  in abilities do
      scores = {scores, Bernoulli( $\frac{\exp(\theta)}{1+\exp(\theta)}$ )};
    end
    z-statistic := z-statistic of test of two proportions based on lower i and upper
      i number of the vector scores;
    z-vector = {z-vector, z-statistic};
  end
  z = {z, z-vector};
end
```

# Simulated optimal splitting of a group of test takers





# Critical values for gULI

- based on Brennan, let's assume a test with  $n$  test-takers and  $k$  binary items
- using ordered total scores of the  $n$  test-takers, the group is divided into three parts considering only highest  $n_U$  and lowest  $n_L$  of the test-takers
- assuming that  $m_U$ , or  $m_L$  is a number of correct answers in the highest  $n_U$ , or lowest  $n_L$  test-takers, then

$$\text{ULI}_B \equiv \frac{m_U}{n_U} - \frac{m_L}{n_L}$$

- assuming  $n$ ,  $n_U = \frac{1}{5}n$ ,  $n_L = \frac{1}{5}n$  is given in our case, there is  $|\{0, 1, \dots, n_L\} \times \{0, 1, \dots, n_U\}|$  combinations of possible  $\text{ULI}_B$
- we search for the lowest value of  $\text{ULI}_B$  such that  $z$  statistic of  $z$ -test of two proportions is significant given  $n$ ,  $n_U$ ,  $n_L$

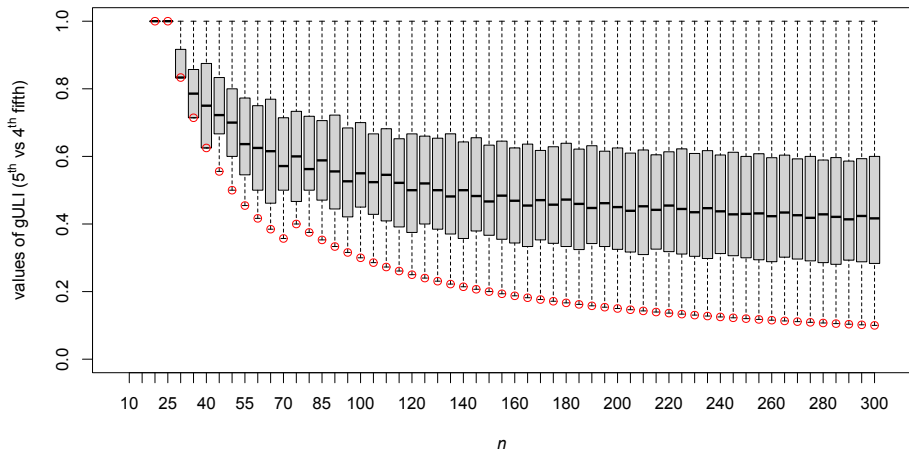
# Critical values for gULI using the 5<sup>th</sup> and 4<sup>th</sup> fifth of scores

```

 $n_{\min} := 10$  // minimal size of a sample of test-takers or their scores;
 $n_{\max} := 300$  // maximal size of a sample of test-takers or their scores;
 $ULI := \{\emptyset\}$  // a list of ULI's;
for  $n = n_{\min} : n_{\max}$  do
     $n_{\text{lower}} := 0.2n$  // a size of lower part of test-takers' sample;
     $n_{\text{upper}} := 0.2n$  // a size of upper part of test-takers' sample;
     $\text{significant\_ULI} := \{\emptyset\}$  // a vector of significant ULI's for given  $n$ ;
    for  $i = 0 : n_{\text{lower}}$  do
        for  $j = 0 : n_{\text{upper}}$  do
            if  $\frac{j}{n_{\text{upper}}} \geq \frac{i}{n_{\text{lower}}}$  and  $z$ -statistic for  $\frac{i}{n_{\text{lower}}}$  vs.  $\frac{j}{n_{\text{upper}}}$  is significant then
                 $\text{significant\_ULI} = \left\{ \text{significant\_ULI}, \frac{j}{n_{\text{upper}}} - \frac{i}{n_{\text{lower}}} \right\}$ ;
            end
        end
    end
     $ULI = \{ULI, \text{significant\_ULI}\}$ ;
end

```

# Critical values for gULI using the 5<sup>th</sup> and 4<sup>th</sup> fifth of scores



# Critical values for gULI using the 5<sup>th</sup> and 4<sup>th</sup> fifth of scores

- values for gULI between the 5<sup>th</sup> and 4<sup>th</sup> fifth of scores for given  $n$  guaranteed the  $z$  statistics using the gULI would be significant,  $\alpha = 0.05$ , regardless of  $p_L$ ,  $p_U$

$n$	gULI (5 <sup>th</sup> , 4 <sup>th</sup> fifth)
25	0.833
30	0.714
40	0.556
50	0.455
80	0.352
100	0.286
150	0.194




# Going further

- deriving analytically (optimally) properties of distribution of minimal values of significant ULI's
- more IRT-based approach
  - using the approach based on searching for lowest values of discrimination coefficient  $a$  of a selected IRT model such that  $H_0 : \hat{a} = 0$  is rejected
- taking account into the fact that  $z$  statistic does not respect an order or “closeness” of the upper and lower part of scores, e. g. the 5<sup>th</sup> and 4<sup>th</sup> fifth of scores vs. the 5<sup>th</sup> and 1<sup>st</sup> fifth of scores and their gULI's

# Conclusion

- there are still good reasons to use traditional item analysis approaches
- Upper-Lower index could be generalized to tailor this descriptive statistics to the needs of the individual test
- critical values for generalized ULI's are not usually provided or published
- relatively simple concepts and ideas such as basic inference hypotheses tests could help to search for the critical values
- some older concepts were revisited analytically and new one were derived using simulations in R language

# References

-  Brennan, Robert L. "A Generalized Upper-Lower Item Discrimination Index". In: *Educational and Psychological Measurement* 32.2 (July 1972), pp. 289–303. DOI: 10.1177/001316447203200206. URL: <https://doi.org/10.1177/001316447203200206>.
-  Kelley, T. L. "The selection of upper and lower groups for the validation of test items." In: *Journal of Educational Psychology* 30.1 (1939), pp. 17–24. DOI: 10.1037/h0057123. URL: <https://doi.org/10.1037/h0057123>.
-  Metsämuuronen, Jari. "Generalized Discrimination Index". en. In: (2018). DOI: 10.13140/rg.2.2.30933.88804/1. URL: <http://rgdoi.net/10.13140/RG.2.2.30933.88804/1>.

Thank you for your attention!

lubomir.stepanek@lf1.cuni.cz