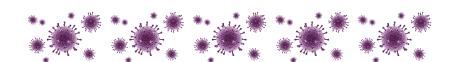
# Modèles Linéaires Appliqués / Régression Régression Logistique: Extensions

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## Régression Bernouilli $y = \mathbf{1}_A$ •••

```
1 > reg1 = glm((Survived==1)~Pclass+Sex+Age+I(Age^2)+I(
     Age^3)+SibSp, family=binomial, data=base)
2 > summary(reg1)
3
4 Coefficients:
               Estimate Std. Error z value Pr(>|z|)
5
6 (Intercept) 5.616e+00 6.565e-01 8.554 < 2e-16 ***
7 Pclass2 -1.360e+00 2.842e-01 -4.786 1.7e-06 ***
8 Pclass3 -2.557e+00 2.853e-01 -8.962 < 2e-16 ***
9 Sexmale -2.658e+00 2.176e-01 -12.216 < 2e-16 ***
10 Age
        -1.905e-01 5.528e-02 -3.446 0.000569 ***
11 I(Age^2) 4.290e-03 1.854e-03 2.314 0.020669 *
12 \text{ I}(Age^3) -3.520e-05 1.843e-05 -1.910 0.056188 .
13 SibSp
        -5.041e-01 1.317e-01 -3.828 0.000129 ***
14 > predict(reg1)[1]
15 -2.592995
16 > predict(reg1, type="response")[1]
17 0.06959063
```

#### Régression Bernouilli $y = \mathbf{1}_{A^C}$

```
1 > reg0 = glm((Survived==0)~Pclass+Sex+Age+I(Age^2)+I(
     Age^3)+SibSp, family=binomial, data=base)
2 > summary(reg0)
3
 Coefficients:
               Estimate Std. Error z value Pr(>|z|)
5
 (Intercept) -5.616e+00 6.565e-01 -8.554 < 2e-16 ***
7 Pclass2
           1.360e+00 2.842e-01 4.786 1.7e-06 ***
8 Pclass3 2.557e+00 2.853e-01 8.962 < 2e-16 ***
9 Sexmale 2.658e+00 2.176e-01 12.216 < 2e-16 ***
           1.905e-01 5.528e-02 3.446 0.000569 ***
10 Age
11 I(Age^2) -4.290e-03 1.854e-03 -2.314 0.020669 *
12 I(Age^3)
          3.520e-05 1.843e-05 1.910 0.056188 .
13 SibSp
            5.041e-01 1.317e-01 3.828 0.000129 ***
14
15 > predict(reg0)[1]
16 2.592995
17 > predict(reg0, type="response")[1]
18 0.9304094
```

#### Régression Binomiale

Au lieu de  $Y_i \sim \mathcal{B}(p_i)$ ,  $Y_i \sim \mathcal{B}(n_i, p_i)$  où  $n_i$  est connue.

$$\mathbb{E}\left(\frac{Y_i}{n_i}\right) = p_i = \frac{e^{\mathbf{x}_i^\top \beta}}{1 + e^{\mathbf{x}_i^\top \beta}}$$

1 > reg = glm(cbind(cbind(Y,n-Y) ~ X1+X2, data = base, family=binomial)

On pose  $z_i = y_i/n_i$ , dont la densité est

$$f(y_i, p_i) = \binom{n_i}{n_i y_i} \exp \left[ n_i y_i \log \left( \frac{p}{1-p} \right) + n_i \log(1-p) \right]$$

et on estime  $\beta$  par maximum de vraisemblance



Pour une loi de Bernoulli,  $y \in \{0, 1\}$ ,

$$\mathbb{P}(Y=1) = \frac{e^{\mathbf{x}^{\top}\boldsymbol{\beta}}}{1 + e^{\mathbf{x}^{\top}\boldsymbol{\beta}}} = \frac{\rho_1}{\rho_0 + \rho_1} \text{ et } \mathbb{P}(Y=0) = \frac{1}{1 + e^{\mathbf{x}^{\top}}} = \frac{\rho_0}{\rho_0 + \rho_1}$$

Pour une loi multinomiale,  $y \in \{A, B, C\}$ ,  $\mathbf{y} = (\mathbf{1}_A, \mathbf{1}_B, \mathbf{1}_C)$ 

$$\mathbb{P}(Y = A) = \frac{p_A}{p_A + p_B + p_C} \propto p_A \text{ i.e. } \mathbb{P}(Y = A) = \frac{e^{\mathbf{x}^\top \beta_A}}{e^{\mathbf{x}^\top \beta_B} + e^{\mathbf{x}^\top \beta_B} + 1}$$

$$\mathbb{P}(Y = B) = \frac{p_B}{p_A + p_B + p_C} \propto p_B \text{ i.e. } \mathbb{P}(Y = B) = \frac{e^{\mathbf{x}^\top \beta_B}}{e^{\mathbf{x}^\top \beta_A} + e^{\mathbf{x}^\top \beta_B} + 1}$$

$$\mathbb{P}(Y = C) = \frac{p_C}{p_A + p_B + p_C} \propto p_C \text{ i.e. } \mathbb{P}(Y = C) = \frac{1}{e^{\mathbf{x}^\top \beta_A} + e^{\mathbf{x}^\top \beta_B} + 1}$$

On va essayer de comprendre la classe  $y \in \{1, 2, 3\}$  sur les données du Titanic

```
1 > loc = "http://freakonometrics.free.fr/titanic.RData"
2 > download.file(loc_fichier, "titanic.RData")
3 > load("titanic.RData")
4 > regclass = multinom(Pclass ~ Sex+Age+SibSp, base)
5 > regclass
6
 Coefficients:
    (Intercept) Sexmale
                                           SibSp
                                  Age
9 2 1.416426 0.2662196 -0.04526865 -0.2150871
10 3 2.420469 1.0330840 -0.07541502 -0.1149161
11
12 Residual Deviance: 1347.672
13 AIC: 1363.672
```

Avec ici  $\beta_2$  et  $\beta_3$  (la class 1 est la référence)

Idée : comme la class 1 est la référence,

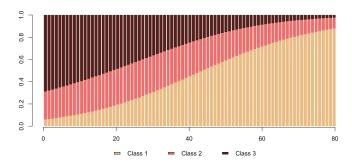
$$\mathbb{P}(Y=1) \propto 1, \ \mathbb{P}(Y=2) \propto e^{\mathbf{x}^{\top} \boldsymbol{\beta}_2} \text{ et } \mathbb{P}(Y=3) \propto e^{\mathbf{x}^{\top} \boldsymbol{\beta}_3}$$

$$\mathbb{P}(Y=1) \propto 1, \ \mathbb{P}(Y=2) \propto e^{\mathbf{x}^{\top} \boldsymbol{\beta}_2} \ \text{et} \ \mathbb{P}(Y=3) \propto e^{\mathbf{x}^{\top} \boldsymbol{\beta}_3}$$

$$\mathbb{P}(Y=1)\frac{1}{e^{\mathbf{x}^{\top}\boldsymbol{\beta}_{2}}+e^{\mathbf{x}^{\top}\boldsymbol{\beta}_{3}}+1},\ \mathbb{P}(Y=2)=\frac{e^{\mathbf{x}^{\top}\boldsymbol{\beta}_{A}}}{e^{\mathbf{x}^{\top}\boldsymbol{\beta}_{B}}+e^{\mathbf{x}^{\top}\boldsymbol{\beta}_{B}}+1},\cdots$$

```
1 > t(exp(b%*%x))/sum(exp(b%*%x))
3 [1.] 0.7170873 0.1954892 0.08742357
```

```
x = cbind(1,0,0:80,0)
_{2} > p2 = exp(apply((x%*%b[2,]),1,sum))
3 > p3 = exp(apply((x%*%b[3,]),1,sum))
4 > pp2 = p2/(1+p2+p3)
5 > pp3 = p3/(1+p2+p3)
6 > p = rbind(1-pp2-pp3,pp2,pp3)
7 > barplot(p)
```



Considérons une approche alternative : régressions Bernoulli itérées considérons un premier modèle de Bernoulli  $y_1 = \mathbf{1}_A$ 

```
1 > reg1 = glm((Pclass==1) ~ Sex+Age+SibSp, base, family
     =binomial)
```

considérons un premier modèle de Bernoulli  $y_2 = \mathbf{1}_B$ , entre les classes B et C

```
1 > reg2 = glm((Pclass==2) ~ Sex+Age+SibSp, base, family
     =binomial, subset = (Pclass!=1))
```

Idée : 
$$\mathbb{P}(y = B) = \mathbb{P}(y = B|y \neq A) \cdot \mathbb{P}(y \neq A)$$

- 1 > p11 = predict (reg1, newdata=base, type="response")
- 2 > p12 = predict (reg2, newdata=base, type="response")
- 3 > itp = cbind(p11, (1-p11)\*p12, (1-p11)\*(1-p12))

On peut comparer les modèles logit itérés (à gauche) et le modèle multinomial (à droite)

```
> mmp = predict(regclass, newdata=base, "probs")
 > head(cbind(itp,mmp))
4 1 0.129 0.204 0.668 0.126 0.201 0.673
5 2 0.459 0.274 0.267 0.462 0.275 0.264
6 3 0.256 0.328 0.416 0.259 0.330 0.411
7 4 0.412 0.285 0.303 0.417 0.284 0.299
8 5 0.227 0.253 0.521 0.229 0.253 0.518
```

