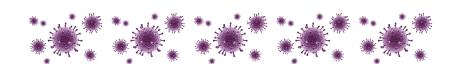
Modèles Linéaires Appliqués / Régression Analyse Discriminante & Courbe ROC

Arthur Charpentier

UQAM

Hiver 2020 - COVID-19 # 6





Supposons que le but soit *simplement* de prédire une classe, $\hat{y} \in \{0,1\}$ (ou plus généralement $\hat{y} \in \{a_1,a_2,\cdots,a_J\}$)

$$m^{\star}(\mathbf{x}) = \underset{y \in \{0,1\}}{\operatorname{argmin}} \{ \mathbb{P}[Y = y | \mathbf{X} = \mathbf{x}] \}$$

soit

$$m^{\star}(\mathbf{x}) = \underset{y \in \{0,1\}}{\operatorname{argmin}} \left\{ \frac{\mathbb{P}[\mathbf{X} = \mathbf{x} | Y = y]}{\mathbb{P}[\mathbf{X} = \mathbf{x}]} \right\}$$

(où $\mathbb{P}[X = x]$ devient f(x) dans le cas continu). Si y prend deux valeurs – i.e. $\{0, 1\}$

$$m^{\star}(\mathbf{x}) = \begin{cases} 1 \text{ si } \mathbb{E}(Y|\mathbf{X} = \mathbf{x}) > \frac{1}{2} \\ 0 \text{ sinon} \end{cases}$$



L'ensemble

$$\mathcal{D}_S = \left\{ \boldsymbol{x} : \ \mathbb{E}(Y|\boldsymbol{X} = \boldsymbol{x}) = \frac{1}{2} \right\}$$

est appelé frontière de décision.

Supposons que $\pmb{X}|Y=0\sim\mathcal{N}(\pmb{\mu}_0,\pmb{\Sigma}_0)$ et $\pmb{X}|Y=1\sim\mathcal{N}(\pmb{\mu}_1,\pmb{\Sigma}_1)$, alors, si r_y^2 est la distance de Manalahobis de \pmb{x} à $\pmb{\mu}_y$

$$r_y^2 = [\mathbf{x} - \boldsymbol{\mu}_y]^{\top} \boldsymbol{\Sigma}_y^{-1} [\mathbf{x} - \boldsymbol{\mu}_y] \text{ pour } y \in \{0, 1\},$$

$$m^{\star}(\mathbf{x}) = \begin{cases} 1 \text{ si } r_1^2 < r_0^2 + 2\log \frac{\mathbb{P}(Y=1)}{\mathbb{P}(Y=0)} + \log \frac{|\mathbf{\Sigma}_0|}{|\mathbf{\Sigma}_1|} \\ 0 \text{ sinon} \end{cases}$$



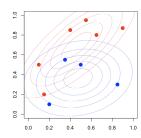
Soit δ_{v} la fonction définie (pour $y \in \{0,1\}$) par

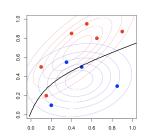
$$\delta_{y}(\mathbf{x}) = -\frac{1}{2}\log|\mathbf{\Sigma}_{y}| - \frac{1}{2}[\mathbf{x} - \boldsymbol{\mu}_{y}]^{\top}\mathbf{\Sigma}_{y}^{-1}[\mathbf{x} - \boldsymbol{\mu}_{y}] + \log\mathbb{P}(Y = y)$$

telle que la frontière de décision soit

$$\{ \boldsymbol{x} \text{ tel que } \delta_0(\boldsymbol{x}) = \delta_1(\boldsymbol{x}) \}$$

qui est quadratique en x

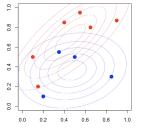


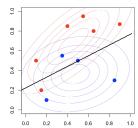


Fisher (1936) a rajouté hypothèse $\Sigma_0 = \Sigma_1$. Alors

$$\delta_y(\mathbf{x}) = \mathbf{x}^{\top} \mathbf{\Sigma}^{-1} \boldsymbol{\mu}_y - \frac{1}{2} \boldsymbol{\mu}_y^{\top} \mathbf{\Sigma}^{-1} \boldsymbol{\mu}_y + \log \mathbb{P}(Y = y)$$

et la frontière de décision est linéaire en x







Si
$$\pmb{X}|Y=0\sim\mathcal{N}(\pmb{\mu}_0,\pmb{\Sigma})$$
 et $\pmb{X}|Y=1\sim\mathcal{N}(\pmb{\mu}_1,\pmb{\Sigma})$ alors

$$\log \frac{\mathbb{P}(Y=1|\boldsymbol{X}=\boldsymbol{x})}{\mathbb{P}(Y=0|\boldsymbol{X}=\boldsymbol{x})}$$

est égal à

$$\mathbf{x}^{\top}\mathbf{\Sigma}^{-1}[\boldsymbol{\mu}_{\scriptscriptstyle{\mathcal{Y}}}] - \frac{1}{2}[\boldsymbol{\mu}_{1} - \boldsymbol{\mu}_{0}]^{\top}\mathbf{\Sigma}^{-1}[\boldsymbol{\mu}_{1} - \boldsymbol{\mu}_{0}] + \log\frac{\mathbb{P}(Y=1)}{\mathbb{P}(Y=0)}$$

qui est linéaire en x, autrement dit

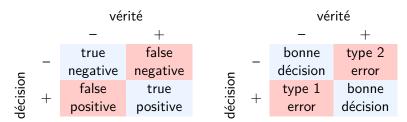
$$\log \frac{\mathbb{P}(Y=1|\boldsymbol{X}=\boldsymbol{x})}{\mathbb{P}(Y=0|\boldsymbol{X}=\boldsymbol{x})} = \boldsymbol{x}^{\top}\boldsymbol{\beta}$$

ce qui rappelle la régression logistique...



Test et Décision

$$y \in \{0,1\}$$
 et on va construire $\widehat{y} \in \{0,1\}$ – via $\widehat{y} = \mathbf{1}(\widehat{p} > s)$



On a un tradeoff entre les types d'erreurs, cf base rate fallacy.

Cf théorie des tests, où on teste H_0 – qui peut être valide (ou pas) - et on doit prendre une décision : rejeter H_0 ou accepter H_0 .





Test et Décision

disease 100

100

200

disease 100

100

200

9.851

149

10,000

9,410

590

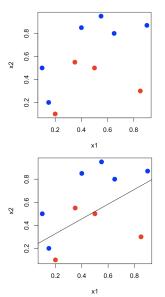
10,000

Prevalence $\frac{200}{10,000} = 2\%$		_
Specificity $\frac{9,751}{9,800} = 99.5\%$		well
9,800	_	9,751
Sensitivity $\frac{100}{200} = 50\%$	+	49
200		9,800
Positive Predictive Value $\frac{166}{149} \sim 67\%$		well
9 310	_	9,310
Specificity $\frac{9,310}{9,800} = 95\%$	+	490
100		9,800
Positive Predictive Value $\frac{100}{590} \sim 17\%$	'	

Cf Wainer & Savage (2008) (until proven guilty: False positives and the war on terror)

1 > x1 = c(.4,.55,.65,.9,.1,.35,
2 .5,.15,.2,.85)
3 > x2 = c(.85,.95,.8,.87,.5,.55,
4 .5,.2,.1,.3)
5 > y = c(1,1,1,1,1,0,0,1,0,0)
6 > df = data.frame(x1=x1,x2=x2,
7 y=as.factor(y))
8 > plot(x1,x2,col=1+y)
9 > reg = glm(y~x1+x2,data=df,
10 family=binomial(link = "logit"))
11 > b = coefficients(reg)
12 > abline(a=-b[1]/b[3],b=-b[2]/b[3])

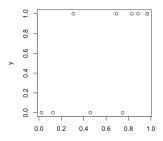
$$\mathbb{P}[Y = 1|X = x] = \mathbb{P}[Y = 0|X = x]$$
 si
 $e^{x^T\beta} = 1$ ou $x^T\beta = 0$
soit ici $\beta_0 + \beta_1 x_1 + \beta_2 x_2 = 0$.

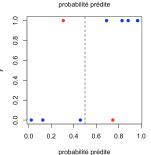


```
_1 > Y = df$y
2 > S = predict(reg,type="response")
3 > plot(S,y)
4 > seuil = .5
5 > Yhat = (S>seuil)*1
6 > plot(S,y,col=1+(y==Yhat))
7 > abline(v=seuil,lty=2)
```

$$\widehat{y}_i = \begin{cases} 1 & \text{si } \widehat{p}_i > \text{seuil} \\ 0 & \text{si } \widehat{p}_i \le \text{seuil} \end{cases}$$

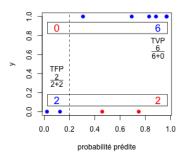
```
> table(Yhat,Y)
Yhat 0 1
  0 3 1
   1 1 5
```

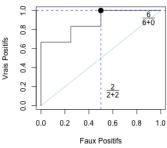




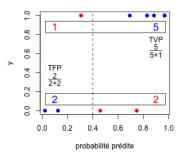


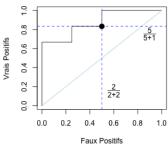
$$\textit{FPR} = \frac{\mathbb{P}[y=0, \widehat{y}=1]}{\mathbb{P}[y=0]} \text{ et } \textit{TPR} = \frac{\mathbb{P}[y=1, \widehat{y}=1]}{\mathbb{P}[y=1]}$$



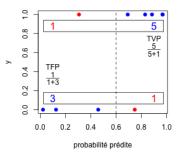


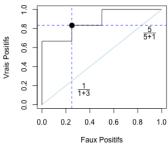
$$\textit{FPR} = \frac{\mathbb{P}[y=0, \widehat{y}=1]}{\mathbb{P}[y=0]} \text{ et } \textit{TPR} = \frac{\mathbb{P}[y=1, \widehat{y}=1]}{\mathbb{P}[y=1]}$$



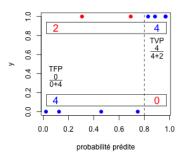


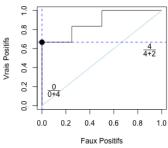
$$\textit{FPR} = \frac{\mathbb{P}[y=0, \widehat{y}=1]}{\mathbb{P}[y=0]} \text{ et } \textit{TPR} = \frac{\mathbb{P}[y=1, \widehat{y}=1]}{\mathbb{P}[y=1]}$$



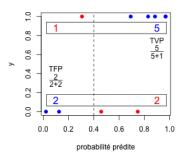


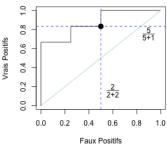
$$\textit{FPR} = \frac{\mathbb{P}[y=0, \widehat{y}=1]}{\mathbb{P}[y=0]} \text{ et } \textit{TPR} = \frac{\mathbb{P}[y=1, \widehat{y}=1]}{\mathbb{P}[y=1]}$$



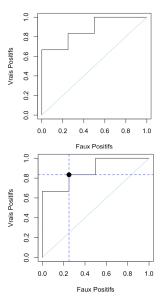


$$\textit{FPR} = \frac{\mathbb{P}[y=0, \widehat{y}=1]}{\mathbb{P}[y=0]} \text{ et } \textit{TPR} = \frac{\mathbb{P}[y=1, \widehat{y}=1]}{\mathbb{P}[y=1]}$$





```
roc.curve=function(s){
    Ps=(S>s)*1
2
    FP=sum((Ps==1)*(Y==0))/sum(Y==0)
3
    TP=sum((Ps==1)*(Y==1))/sum(Y==1)
4
   vect=c(FP,TP)
    names(vect)=c("FPR","TPR")
    return(vect) }
8 > u = seq(0,1,length=251)
 > V = Vectorize(roc.curve)(u)
 > plot(t(V),type="s")
  > table(Yhat,Y)
12
  Yhat 0 1
     0 3 1
14
16 > sum((Yhat)*(Y==0))/sum(Y==0)
  [1] 0.25
18 > sum((Yhat==1)*(Y==1))/sum(Y==1)
  [1] 0.8333333
```



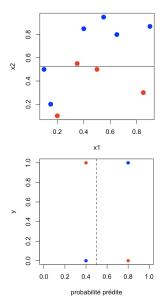
De nombreux packages permettent de tracer des courbes ROC, dont ROCR (ou pROC, plotROC)

```
> library(ROCR)
                                              0.8
pred = prediction(S,Y)
                                            Vrais Positifs
3 > plot(performance(pred, "tpr", "fpr
      "))
 > auc.perf = performance(pred,
                                              0.2
     measure = "auc")
                                              0.0
 > auc.perf@y.values[[1]]
 [1] 0.875
```

L'AUC – aire sous la courbe – donne une idée de la qualité de la classification.

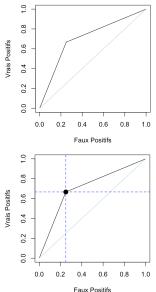
Régression de y sur $\mathbf{1}_{[.525,\infty)}(x_2)$

```
> reg = glm(y~I(x2> .525), data=df,
2 family=binomial(link = "logit"))
3 > abline(h=.525)
4 > Y = df$v
5 > S = predict(reg,type="response")
6 > plot(S,y,xlim=0:1)
7 > seuil = .5
8 > Yhat = (S>seuil)*1
 > table(Yhat,Y)
  Yhat 0 1
     0 3 2
     1 4 5
13
```



La courbe ROC est linéaire par morceaux

```
> pred = prediction(S,Y)
plot(performance(pred,"tpr","fpr"))
 > table(Yhat,Y)
 Yhat 0 1
    0 3 2
    1 1 4
 > auc.perf = performance(pred,
     measure = "auc")
 > auc.perf@y.values[[1]]
 [1] 0.7083333
```

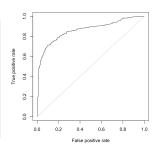


Survie des Passagers du Titanic

y : indicatrice de survie d'un passager du Titanic

```
1 > loc = "http://freakonometrics.free.fr/titanic.RData"
2 > download.file(loc, "titanic.RData")
3 > load("titanic.RData")
4 > base = base[!is.na(base$Age),1:7]
5 > reg = glm(Survived ~ Sex+poly(Age,3)+Pclass+SibSp,
6 family = "binomial", data = base)
```

```
1 > library(ROCR)
2 > Y = base$Survived
3 > S = predict(reg,type="response")
4 > pred = prediction(S,Y)
5 > plot(performance(pred,"tpr","fpr"))
6 > performance(pred, measure = "auc")
7     @y.values[[1]]
8  [1] 0.8627358
```



Survie des Passagers du Titanic

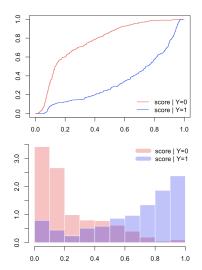
Kolmogorov-Smirnov (KS): Comparer les distributions de (S|Y = 0) et (S|Y = 1)

$$d = \sup_{x \in [0,1]} \{ |\widehat{F}_1(x) - \widehat{F}_0(x)| \}$$

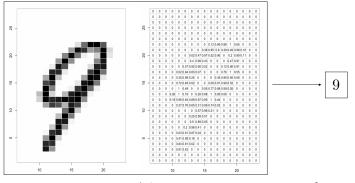
οù

$$\widehat{F}_1(x) = \frac{1}{n_1} \sum_{i:y_i=1} \mathbf{1}(s_i \leq x)$$

$$\widehat{F}_0(x) = \frac{1}{n_0} \sum_{i: v_i = 0} \mathbf{1}(s_i \le x)$$



$$(y_i, \mathbf{x}_i)$$
, où $y \in \{0, 1, 2, 3, \cdots, 9\}$ et $\mathbf{x}_i \in \mathcal{M}_{28,28} = [0, 1]^{28 \times 28}$



$$oldsymbol{x}_i \in \mathcal{M}_{28,28}$$

$$y_i \in \{0, 1, \cdots, 9\}$$

```
1 > library(keras)
2 > mnist = dataset_mnist()
```

Here $\{(y_i, x_i)\}$ with $y_i = "3"$ and $x_i \in [0, 1]^{28 \times 28}$



```
1 > library(keras)
2 > mnist = dataset_mnist()
3 > n = 1000
4 > V = mnist$train$x[1:n,,]
5 > MV = NULL
6 for(i in 1:n) MV=cbind(MV,as.vector(V[i,,]))
7 > MV = t(MV)
8 > df = data.frame(y=mnist$train$y[1:n],x=MV)
```

Peut-on reconnaître les '1'?

Problème numérique!

On a k = 784 variables explicatives... il faut réduire la dimension !

```
1 > library(factoextra)
2 > pca=prcomp(MV)
3 > res.ind = get_pca_ind(pca)
4 > PTS = res.ind$coord
5 > k=3
6 > dfpca = data.frame(y=mnist$train$y[1:n],x=PTS[,1:k])
7 > reg1 = glm((y==1)~.,data=dfpca,family=binomial)
8 > reg8 = glm((y==8)~.,data=dfpca,family=binomial)
```

On essaye de reconnaître les '1' et les '8'

```
1 > library(ROCR)
2 > Y1 = as.numeric((df$y == 1)*1)
3 > S1 = predict(reg1,type="response")
4 > pred1 = prediction(S1,Y1)
5 > plot(performance(pred1,"tpr","fpr"))
6 > Y8 = as.numeric((df$y == 8)*1)
7 > S8 = predict(reg8,type="response")
8 > pred8 = prediction(S8,Y8)
9 > plot(performance(pred8,"tpr","fpr"))
```

On utilise ici seulement les 3 premières composantes principales

