

PGM foundations Part II

Representation in continuous domain, approximate inference

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Outline



- PGMs in continuous domain
- Generative approach
- Approximate inference



- Thus far, we've been using only discrete variables
- Conditional Probability Tables
- Extension to continuous domain is almost trivial...



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- Conditional Probability Tables
- Extension to continuous domain is almost trivial...
- But with it, some concepts become more relevant
 - Prior
 - Conjugate prior



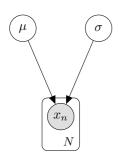
General form



- We use functions instead of tables
- Typically, each function is a well-known distribution (or combination of them)
- ullet Every distribution is parameterized by a set Θ



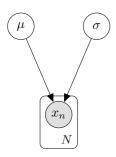
Guassian distribution



- A well-known example is the Gaussian (or Normal) distribution
- In this PGM, we assume to have observations x_n , that follow a Gaussian distribution
- It has two parameters (mean μ , variance σ^2)



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- It has two parameters (mean μ , variance σ^2)
- Inference
 - It has a well-known log likelihood function

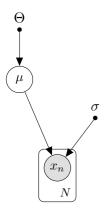


- A Graphical Model allows for a full Bayesian treatment:
 - We can assign *priors* to the parameters
 - We can use domain knowledge
 - Good to prevent overfitting



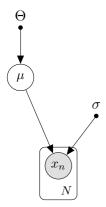
- A Graphical Model allows for a full Bayesian treatment:
 - We can assign *priors* to the parameters
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 - Good to prevent overfitting
 - What would be the form of those priors?





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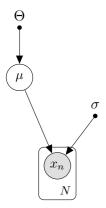




- \bullet To simplify, let's assume we know σ but not μ
- Can we pick *any* distribution, $D(\mu|\Theta)$?
- Our joint distribution would become:

$$p(\mu, \mathbf{X}|\Theta, \sigma) = D(\mu|\Theta) \prod_{n=1}^{N} p(x_n|\mu, \sigma)$$

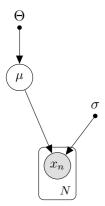




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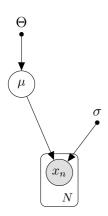




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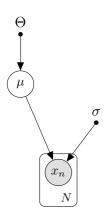


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- If $D(\mu|\Theta)$ is normal, then $p(\mu,\mathbf{X})$ is normal too!
- \bullet If $p(\mu, \mathbf{X})$ is not a known distribution, we may have trouble deriving it





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Conjugate priors



ullet For many known distributions, there is a corresponding *conjugate prior*, P, that preserves its form under multiplication. I.e., if we have distribution L and its conjugate prior P_0 , we should have

$$P_1 = L \times P_0$$

- ullet where P_1 has the same form as P_0
- For example, the Beta distribution is the conjugate prior of Bernoulli; and we've seen that the Normal is the conjugate for the mean of the Normal (when variance is known).
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- If we have a known closed form for model, inference is generally more efficient!
- This is great for online learning (why?)!

Conjugate priors



• We usually use a table

Likelihood	Model parameters	Conjugate prior distribution	Prior hyperparameters	Posterior hyperparameters	Interpretation of hyperparameters ^[note 1]	Posterior predictive[note 2
Bernoulli	p (probability)	Beta	α, β	$\alpha + \sum_{i=1}^n x_i, \ \beta + n - \sum_{i=1}^n x_i$	$\begin{array}{l} \alpha-1 \text{ successes, } \beta-1 \\ \text{failures}^{[\text{note 1}]} \end{array}$	$p(ilde{x}=1)=rac{lpha'}{lpha'+eta'}$
Binomial	p (probability)	Beta	α, β	$\alpha + \sum_{i=1}^n x_i, \ \beta + \sum_{i=1}^n N_i - \sum_{i=1}^n x_i$	$\begin{array}{l} \alpha-1 \text{ successes, } \beta-1 \\ \text{failures}^{[\text{note 1}]} \end{array}$	$\operatorname{BetaBin}(\tilde{x} lpha',eta')$ (beta-binomial)
Negative binomial with known failure number, r	p (probability)	Beta	α, β	$\alpha + \sum_{i=1}^n x_i, \beta + rn$	$\begin{array}{l} \alpha-1 \text{ total successes, } \beta-1 \\ \text{failures}^{\text{(note 1)}} \text{ (i.e., } \frac{\beta-1}{r} \\ \text{experiments, assuming } r \text{ stays} \\ \text{fixed)} \end{array}$	
Poisson	λ (rate)	Gamma	k, θ	$k+\sum_{i=1}^n x_i,\;rac{ heta}{n heta+1}$	k total occurrences in $\frac{1}{\theta}$ intervals	$\operatorname{NB}(\tilde{x} k', \theta')$ (negative binomial)
			α , $\beta^{[\text{note 3}]}$	$\alpha + \sum_{i=1}^{n} x_i, \ \beta + n$	α total occurrences in β intervals	$\operatorname{NB}\!\left(ilde{x} lpha',rac{1}{1+eta'} ight)$ (negative binomial)
Categorical	<pre>p (probability vector), k (number of categories; i.e., size of p)</pre>	Dirichlet	α	$oldsymbol{lpha} + (c_1, \dots, c_k),$ where c_i is the number of observations in category i	$lpha_i - 1$ occurrences of category $i^{[{ m note } 1]}$	$\begin{split} p(\tilde{x} = i) &= \frac{{\alpha_i}'}{\sum_i {\alpha_i}'} \\ &= \frac{{\alpha_i} + c_i}{\sum_i {\alpha_i} + n} \end{split}$

Figure: From Wikipedia

Some conjugate priors to remember...

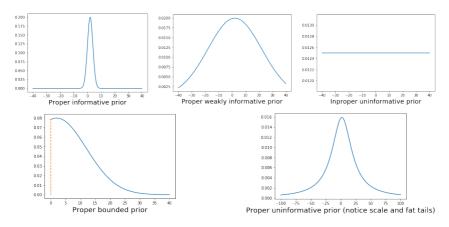


Likelihood	Prior
Normal with known variance	Normal
Normal with known mean	Inverse Gamma
Multivariate normal, known	Inverse Wishart
mean	
Multivariate normal, unknown	Normal-inverse-Wishart
mean and variance	
Exponential	Gamma
Bernoulli	Beta
Mulitnomial	Dirichlet
Poisson	Gamma

Last note on priors



• Depending on what you know of the problem (or the constraints you want to impose...):



Playtime!



- Open notebook "3-PGM fundamentals.ipynb"
- Do part 1 (est. duration=30 min)

Generative approach



- By now, you understand that you can combine variables in multiple ways in your graphical model
- On the other hand, you may be overwhelmed about where to start doing your own
 - Small models, with few variables, are simple
 - What if you have a lot of variables, assumptions, domain knowledge?...

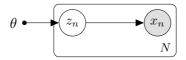
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- You need to think from a generative perspective...

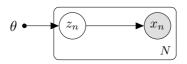


• How is a data point generated?





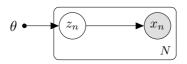
• How is a data point generated?



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- \bullet For n=1..N, do
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• How is a data point generated?



- ullet Given a parameter heta
- \bullet For n=1..N, do
 - **1** Draw a random latent variable, $z_n \sim p(z|\theta)$
 - **2** Given z_n , generate x_n such that $x_n \sim p(x|\theta,z_n)$
- In fact, this resembles a program structure!

A more complex example - Dwell time prediction



For a given bus stop, that serves a single line, can we predict the amount of time the next bus will be stopped there to load/unload passengers (the dwell time)?

- Our dataset contains $\{x_n=\{0,1\}$ -representing peak/non-peak hour, dt_n dwell time $\}$.
- Notice that, sometimes, the bus does not stop at all!
- ullet When it stops, we measure the duration as dt



Given N, σ_{β} , σ_{ϵ} and π

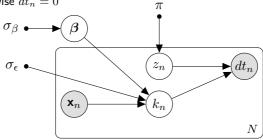
- **1** Draw a pair of parameters¹, $\boldsymbol{\beta} \sim \mathcal{N}(\mathbf{0}, I\sigma_{\beta})$
- **2** For n = 1..N
 - **1** Draw one value for z_n , such that $z_n \sim Bern(\pi)$.
 - If $z_n = 1$, the bus has stopped ($z_n = 0$ otherwise).
 - ullet Distributed as Bernoulli, with parameter π
 - **2** Draw one value for k_n , such that $k_n \sim \mathcal{N}(\mathbf{x}_n^T \boldsymbol{\beta}, \sigma_{\epsilon})$
 - **3** If $z_n = 1$, $dt_n = k_n$,
 - otherwise $dt_n = 0$

 $^{^1}$ We need two values for eta, one for the intercept, another for the peak/non-peak information.

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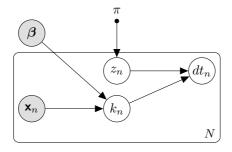
 $^{^{1}}$ We need two values for β , one for the intercept, another for the peak/non-peak information.



- After you define your model, you need to estimate it. I.e. infer the following:
 - Distribution of β
 - Optimal values of σ_{ϵ} , σ_{β} , and π (we defined them as constants!)
- Of course, when you have them, you can make your predictions!
- Your model will look different:



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- Set up the building blocks, as per available knowledge
- Easy to change data distributions inside the model
- Can be used to actually generate data!
 - Ancestral sampling

Playtime!



- Open notebook "3-PGM fundamentals.ipynb"
- Do part 2 (est. duration=30 min)

Mixture models

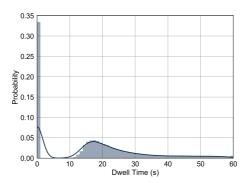


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- Mixture models are pervasive in data modelling in general

Mixture models



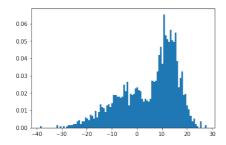
- A PGM is composed of observed and latent variables, parameters, constants.
- In this course, we'll approach some examples from this very large family
- Mixture models are pervasive in data modelling in general
- Problem:
 - Sub-populations of data
 - Data generated from combination/competition of multiple sources
 - Number of sources usually discrete and finite



The canonical example: Gaussian Mixture



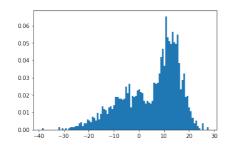
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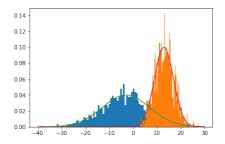
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• What we observe



• What really happens



Generative story



Given:

- A dataset with N points (or vectors) $(\mathbf{x}_1, \mathbf{x}_2, ..., \mathbf{x}_n)$ and a value K
- $oldsymbol{0}$ Draw $oldsymbol{\pi}$, and $(oldsymbol{\mu_k}, oldsymbol{\Sigma_k})$ for all K gaussians
- **2** For n = 1, 2, ..., N
 - **1** Draw $z_n \sim Multinomial(\pi)$
 - ullet π is a vector $(1 \times K)$ with the probabilities of each class
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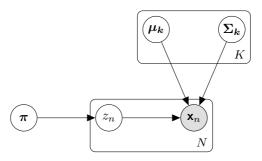
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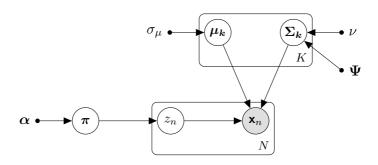


Note: in practice we need to be exhaustive



...particularly in probabilistic programming (e.g. STAN)

- $\pi \sim Dir(\alpha)$
- $\mu_k \sim \mathcal{N}(\mathbf{0}, I\sigma_\mu)$
- ullet $oldsymbol{\Sigma}_{oldsymbol{k}} \sim \mathcal{W}^{-1}(oldsymbol{\Psi},
 u)$
 - ullet Typically, u= number of dimensions, and $oldsymbol{\Psi}=I$



Playtime!



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The problem of inference



- ...your last exercise should show that we need efficient inference methods
 - Complex distribution (e.g. involving log of sum; an unknown form; etc.)
 - High dimensionality (e.g. more than a couple of parameters is often too many!)
 - Continuous dimensions instead of discrete

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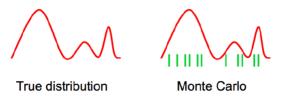
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- Two general approaches:
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- Before we get practical (i.e. STAN), we need to understand a bit how inference can be done
 - Important to manipulate STAN and understand its output
- STAN uses Approximate Inference (we'll talk about it today)
- In a later class, we'll get more detailed (in both Exact and Approx.).



- Stochastic
- Variational

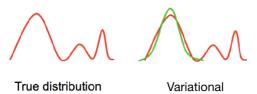


- Stochastic
 - We sample from the distribution
 - Markov Chain Monte Carlo (MCMC)
- Variational





- Stochastic
- Variational
 - We look for a simpler but similar distribution
 - Becomes an optimization problem (of minimizing the difference between *true* and *approximate* distribution)





- Stochastic
- Variational
- STAN uses
 - MCMC
 - Automatic Differentiation Variational Inference (ADVI) a combination of variational and stochastic...

Intuition on Markov Chain Monte Carlo (MCMC)



- Major challenge: how do we choose the sample points?
- Don't forget that "one sample" means:
 - Assignment of a value to each variable of the joint distribution
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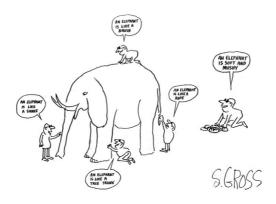
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- With a good number of points, we can:
 - Obtain approximate statistics for the distribution
 - Obtain estimates for individual parameters

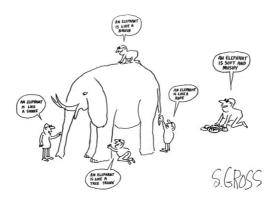


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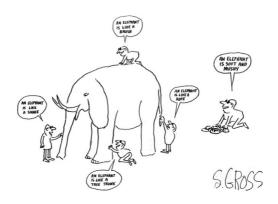
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• Major challenge: how do we choose the sample points?



- Option 1: Just uniformly.
- Option 2: Using the true distribution (cleverly... ;-))



- Our mixture model exercise. We want to estimate values for parameters π , $\mu = [\mu_1, \mu_2]^T$ having the expression for $p(\pi, \mu | \sigma, \sigma_{\mu}, \alpha)$
- Gibbs sampling for our Gaussian Mixture exercise².
 - Initialize π , $\pmb{\mu}$ with random values (from their priors). Let's call them $\pi^{(0)}$, $\pmb{\mu}^{(0)}$
 - **2** For t = 1...T, do
 - **1** Choose new value $\pi^t \sim p(\pi | \boldsymbol{\mu}^{t-1})$.
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we're dropping the full notation to better give the intuition. Notice all variables other than π and μ are fixed anyway. In other words, we're estimating $p(\pi, \mu)$, but we mean $p(\pi, \mu|\sigma, \sigma_{\mu}, \alpha)$.



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- With T sufficiently large, we get enough points to estimate what we want! :-)

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- We want to approximate $p(\mathbf{x})$, where $\mathbf{x} = [x_1, x_2, ..., x_k]$ and $q(\mathbf{x})$ is a prior distribution for \mathbf{x}
- Generic Gibbs sampling algorithm:
 - $\begin{array}{l} \textbf{1} \text{ Initialize } \mathbf{x} \sim q(\mathbf{x}) \\ \textbf{2} \text{ For } t = 1...T \text{, do} \\ \textbf{1} \ x_1^t \sim p(x_1|x_2^{t-1}, x_3^{t-1}, ..., x_k^{t-1}) \\ \textbf{2} \ x_2^t \sim p(x_2|x_1^t, \ , x_3^{t-1}, ..., x_k^{t-1}) \\ \vdots \\ \textbf{3} \ x_k^t \sim p(x_k|x_1^t, x_2^t, x_3^t, ..., x_k^t) \end{array}$

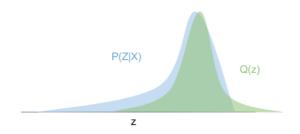


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- In general MCMC algorithms all have this flavor
- STAN uses Hamiltonian Monte Carlo
- We'll get back to this later... ;-)

Intuition on Variational Inference



• Key idea: approximate intractable distribution with a simpler, tractable one.



- We use a method to compare the two distributions, called Kullback-Leibler (KL) divergence
- We turn into an optimization problem, of minimizing KL divergence
- We later use the simpler distribution, to make our inference in the model

Conclusions



- PGMs are extremely flexible. They can combine:
 - Discrete and continuous variables
 - Parametric and non-parametric models
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 - Partial and complete data
- Think in a generative way helps design a model

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 - Parametric and non-parametric models
 - Informative and non-informative priors
 - Online learning with conjugate priors
 - Partial and complete data
- Think in a generative way helps design a model
- The more complex the model is, the harder inference may be
- Markov Chain Monte Carlo and Variational Inference (exact inference later...)

References



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