

# Regression Subset Selection in R

For linear regression: `lm()` from stats package (built-in) For logistic regression: `glm()` from stats package

For lasso and ridge regression: `glmnet()` from glmnet package

Load packages

```
library(glmnet)
library(QuantPsyc) # for standardized regression coefficients
library(subselect)
#library(lars)
library(tabuSearch)
```

Additionally

```
library(lavaan) # for HolzingerSwineford1939 dataset
HS <- HolzingerSwineford1939

library(elasticnet) # regularized PCA
#library(fanc) # regularized FA
#library(FAiR) # semi-exploratory factor analysis; only works on Windows
library(GA) # genetic algorithm for subset selection
```

<http://www.jstatsoft.org/v53/i04/paper>

You can also embed plots, for example:

```
data(diabetes)
X <- diabetes$x
Y <- diabetes$y
```

```
lm(Y ~ X)
# note: equivalent to
lm(y ~ x,data=diabetes)
# which is equivalent to
lm(diabetes$y ~ diabetes$x)
```

Note, in this dataset, `x` is essentially a matrix within a dataframe. This is a little unusual, where the typical format would be:

```
diabetes2 <- data.frame(cbind(Y,X))
head(diabetes2)
```

```
##      Y      age      sex      bmi      map      tc
## 1 151  0.038075906  0.05068012  0.06169621  0.021872355 -0.044223498
## 2  75 -0.001882017 -0.04464164 -0.05147406 -0.026327835 -0.008448724
## 3 141  0.085298906  0.05068012  0.04445121 -0.005670611 -0.045599451
## 4 206 -0.089062939 -0.04464164 -0.01159501 -0.036656447  0.012190569
## 5 135  0.005383060 -0.04464164 -0.03638469  0.021872355  0.003934852
## 6  97 -0.092695478 -0.04464164 -0.04069594 -0.019442093 -0.068990650
```

```
##          ldl          hdl          tch          ltg          glu
## 1 -0.03482076 -0.043400846 -0.002592262  0.019908421 -0.017646125
## 2 -0.01916334  0.074411564 -0.039493383 -0.068329744 -0.092204050
## 3 -0.03419447 -0.032355932 -0.002592262  0.002863771 -0.025930339
## 4  0.02499059 -0.036037570  0.034308859  0.022692023 -0.009361911
## 5  0.01559614  0.008142084 -0.002592262 -0.031991445 -0.046640874
## 6 -0.07928784  0.041276824 -0.076394504 -0.041180385 -0.096346157
```

```
lm(Y ~ ., data=diabetes2)
```

```
##
## Call:
## lm(formula = Y ~ ., data = diabetes2)
##
## Coefficients:
## (Intercept)          age          sex          bmi          map
##      152.13      -10.01     -239.82      519.84      324.39
##          tc          ldl          hdl          tch          ltg
##     -792.18      476.75      101.04      177.06      751.28
##          glu
##       67.63
```

Using the “.” means we want to use all variables that aren’t the outcome as predictors  
Fun Fact: Can do the same thing with a SEM package

```
library(lavaan)
lm.mod <- '
Y ~ age + sex + bmi + map + tc + ldl + hdl + tch + ltg + glu
Y ~ 1 # intercept
'
lm.sem <- sem(lm.mod,diabetes2)
#summary(lm.sem)
#parameterEstimates(lm.sem)
coef(lm.sem)
```

```
##      Y~age      Y~sex      Y~bmi      Y~map      Y~tc      Y~ldl      Y~hdl      Y~tch
##    -10.012    -239.819    519.840    324.390   -792.184    476.746    101.045    177.064
##      Y~ltg      Y~glu      Y~1      Y~~Y
##     751.279     67.625    152.133    2859.690
```

Exact same answer.

```
lm.out <- lm(y ~ x,data=diabetes)
#summary(lm.out)

# check assumptions
#plot(lm.out)

# get standardized coefficients
lm.beta(lm.out) # from QuantPsyc
```

```
##           xage           xsex           xbmi           xmap           xtc
## -0.006178065 -0.147981279  0.320769113  0.200166344 -0.488820243
##           xldl           xhdl           xtch           xltg           xglu
##  0.294177828  0.062349936  0.109258122  0.463579756  0.041728501
```

So we are doing pretty good,  $R^2$  of 0.51, with only four significant predictors. So the question we are going to answer today is whether we can get rid of a few predictors and still do a good job of predicting the outcome. Now, there are two + reasons to do this:

1. In future studies, maybe time is of the essence, or each additional question costs a certain amount of money. By reducing the number of questions we have to ask participants, both money and time can be saved. The question is what is the tradeoff, can we reduce the number of scales/items/questionnaires, and still answer the questions we want?
2. Remember when using  $R^2$  as a criterion, by using more variables as predictors, these can only improve are within sample predictive power. But when we become concerned with generalizability, then in some cases, a reduced number of predictors, only important ones, can generalize better than a larger set of X's. This was somewhat demonstrated in the “preprocessing” lab.

Let's try #2 on the diabetes dataset

```
ids <- sample(1:nrow(diabetes2), .5*nrow(diabetes2),replace=FALSE)
diab.train <- diabetes2[ids,]
diab.test <- diabetes2[-ids,]

lm.trainFull <- lm(Y ~ ., data= diab.train)
summary(lm.trainFull)$r.squared
```

```
## [1] 0.4973194
```

```
lm.trainSub <- lm(Y ~ sex + bmi + map + ltg, data= diab.train)
summary(lm.trainSub)$r.squared
```

```
## [1] 0.4575152
```

```
pred.full <- predict(lm.trainFull,diab.test)
pred.sub <- predict(lm.trainSub,diab.test)

cor(pred.full,diab.test$Y)**2
```

```
## [1] 0.5233413
```

```
cor(pred.sub,diab.test$Y)**2
```

```
## [1] 0.5166219
```

Not in this case, but let's try some different methods specifically designed for subset selection.

## Subset Selection

First off, why don't we just try out all combination of predictors – entering them all separately into `lm()`? Problems:

1. How do we choose.  $R^2$  can only go up with added predictors (RSS can only go down).
2. This is usually computationally infeasible, as there are  $2^p$  possible models, where  $p$  is the number of predictors. In our case  $2^{10} = 1024$ , which is a lot but not too many.

## Stepwise Selection

### Forward

Efficient, but not guaranteed to find best overall model.

```
library(MASS)
lmOut <- lm(Y ~ ., data=diab.train)
stepFor <- stepAIC(lmOut,direction="forward")
```

```
stepFor$anova
```

```
## Stepwise Model Path
## Analysis of Deviance Table
##
## Initial Model:
## Y ~ age + sex + bmi + map + tc + ldl + hdl + tch + ltg + glu
##
## Final Model:
## Y ~ age + sex + bmi + map + tc + ldl + hdl + tch + ltg + glu
##
##
##   Step Df Deviance Resid. Df Resid. Dev      AIC
## 1                210    646261.9 1785.756
```

```
pred.for<- predict(stepFor,diab.test)
```

```
cor(pred.for,diab.test$Y)**2
```

```
## [1] 0.5233413
```

AIC (Akaike Information Criterion) induces a penalty for complexity – meaning that it will try and choose a model that balances predictive accuracy with simplicity (less predictors).

In this example, forward stepwise doesn't suggest getting rid of any predictors

### Backward

```
library(MASS)
lmOut <- lm(Y ~ ., data=diab.train)
stepBack <- stepAIC(lmOut,direction="backward")
```

```
stepBack$anova
```

```
## Stepwise Model Path
## Analysis of Deviance Table
##
## Initial Model:
## Y ~ age + sex + bmi + map + tc + ldl + hdl + tch + ltg + glu
##
## Final Model:
## Y ~ sex + bmi + map + tc + ldl + ltg
##
##
```

##	Step	Df	Deviance	Resid. Df	Resid. Dev	AIC
##	1			210	646261.9	1785.756
##	2 - tch	1	98.53511	211	646360.4	1783.790
##	3 - glu	1	217.09857	212	646577.5	1781.864
##	4 - hdl	1	663.58536	213	647241.1	1780.091
##	5 - age	1	682.91551	214	647924.0	1778.324

```
pred.back<- predict(stepBack,diab.test)
```

```
cor(pred.back,diab.test$Y)**2
```

```
## [1] 0.5275474
```

Here, backwards suggests getting rid of 4 predictors.

## Ridge and Lasso Regression

- Including a penalty on the  $\beta$  parameters, and by varying the penalty we can shrink some of the  $\beta$ 's to zero, doing a form of “automatic” subset selection.

Although there are a number of packages to do this, maybe the best is *glmnet*

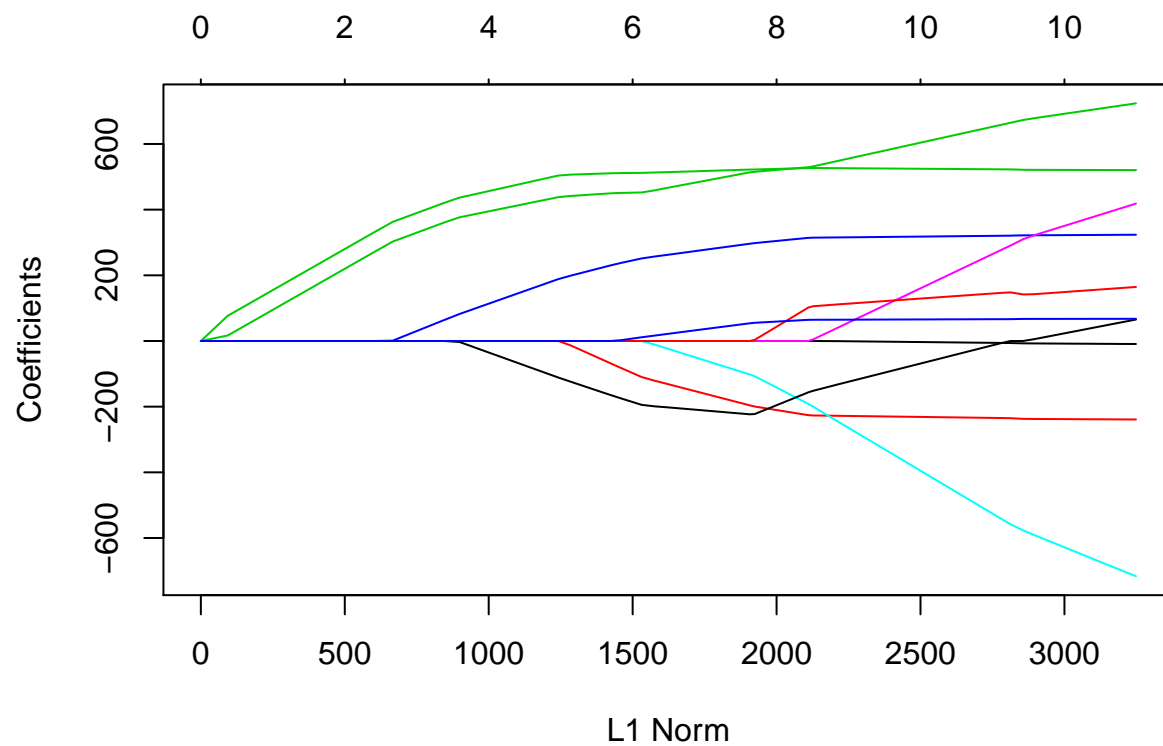
Note, for *glmnet*, your data has to be set up in two separate matrices. Doing this can be accomplished by:

```
YY <- as.matrix(diabetes2$Y)
XX <- as.matrix(diabetes2[,2:11])
# or
XX <- as.matrix(diabetes2[,c("age", "sex", "bmi", "map", "tc", "ldl", "hdl", "tch", "ltg", "glu")])
```

Two things to note: 1. Because we are doing regression with a continuous outcome, we specify the family(distribution) as “gaussian” 2. Shrinkage in lasso and ridge is sensitive to the scale of the variables, therefore, it is best to standardize the predictors before entering. *glmnet* does this by default (look at ?*glmnet*).

Lasso

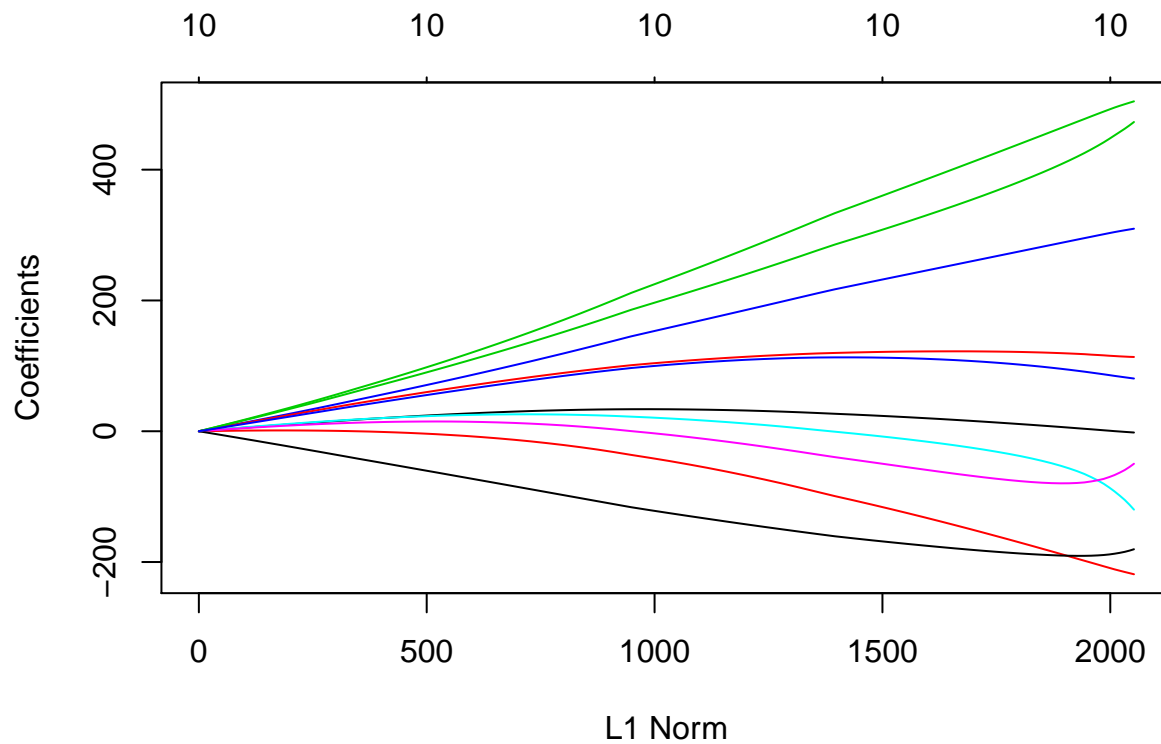
```
?glmnet
lasso.out <- glmnet(XX,YY,family="gaussian",alpha=1)
plot(lasso.out)
```



```
#gaussian for continuous outcomes, "binomial" for categorical
# alpha=1 is lasso, alpha=0 is ridge
```

Ridge

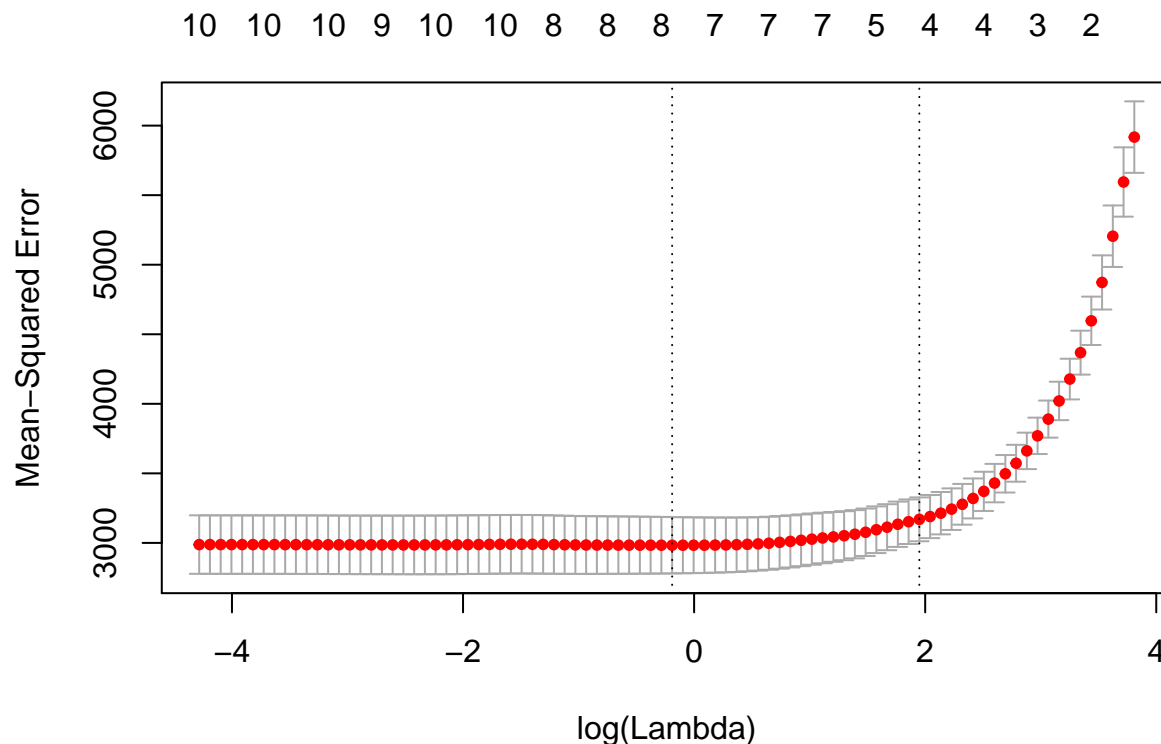
```
ridge.out <- glmnet(XX,YY,family="gaussian",alpha=0)
#plot(ridge.out,type.coef="2norm")
plot(ridge.out)
```



Since ridge regression does not shrink the  $\beta$  coefficients to 0 with increase penalization, it does not do an “automatic” form of subset selection

The problem now becomes, which value of  $\lambda$  (amount of shrinkage) do we choose? Using cross-validation is one of the better ways, and is implemented the glmnet package

```
cv.lasso <- cv.glmnet(XX,YY,family="gaussian",alpha=1)
plot(cv.lasso)
```



Two-strategies for selecting  $\lambda$ : either pick the lowest CV error, or the best solution within 1 standard error. I don't think that there is a clear best choice. The one advantage of using the 1SE rule is that you need fewer predictors. In our example 4 instead of 7.

```
#str(cv.lasso)
(lmin <- cv.lasso$lambda.min)
```

```
## [1] 0.826762
```

```
(lminSE <- cv.lasso$lambda.1se)
```

```
## [1] 7.025438
```

```
lasso.out2 = glmnet(XX,YY,family="gaussian",alpha=1,lambda=lminSE)
lasso.out2
```

```
##
## Call:  glmnet(x = XX, y = YY, family = "gaussian", alpha = 1, lambda = lminSE)
##
##      Df    %Dev Lambda
## [1,]  4 0.4753  7.025
```

Note that Df correspond to the number of non-zero  $\beta$ 's  
So how are we doing?

```
# ?predict.glmnet
pred.1se <- predict(lasso.out2,XX)
cor(pred.1se,YY)**2
```



```
##           [,1]
## s0 0.4891983
```

So with only 4 predictors entered into the model, we only lose 2-3% of our predicted variance( $R^2$ ). With the lasso, there are two recommended strategies for using the results. 1. Taking the predictors with non-zero  $\beta$ 's, and just using that subset in linear regression. 2. Or, bypass this all together and use least angle regression.

“One approach for reducing this bias is to run the lasso to identify the set of non-zero coefficients, and then fit an un-restricted linear model to the selected set of features.” p. 91 Hastie et al., 2009

In our case, we will take the predictors with non-zero  $\beta$ 's and use them with `lm()` to get our final model. This will probably be our most realistic estimate of  $R^2$  when caring about generalization, as we are using the test dataset to derive the estimate.

```
coef(lasso.out2)
```

```
## 11 x 1 sparse Matrix of class "dgCMatrix"
##              s0
## (Intercept) 152.1335
## age          .
## sex          .
## bmi          498.9887
## map          180.6668
## tc           .
## ldl          .
## hdl         -103.2937
## tch          .
## ltg          433.5503
## glu         .
```

```
lm.lasso <- lm(Y ~ bmi + map + hdl + ltg,diab.train)
lmLas.pred <- predict(lm.lasso,diab.test)

cor(lmLas.pred,diab.test$Y)**2
```

```
## [1] 0.5127864
```

Try Elastic Net – from caret

This optimizes both the  $\lambda$  and mixing percentage

```
library(caret)
XX <- as.matrix(diab.train[,-1])
YY <- diab.train$Y # important to change class of variable
enet.out <- train(XX,YY,method="glmnet",tuneLength=8)

row.result <- best(enet.out$results,"Rsquared",maximize=T)
enet.out$results[row.result,]
```

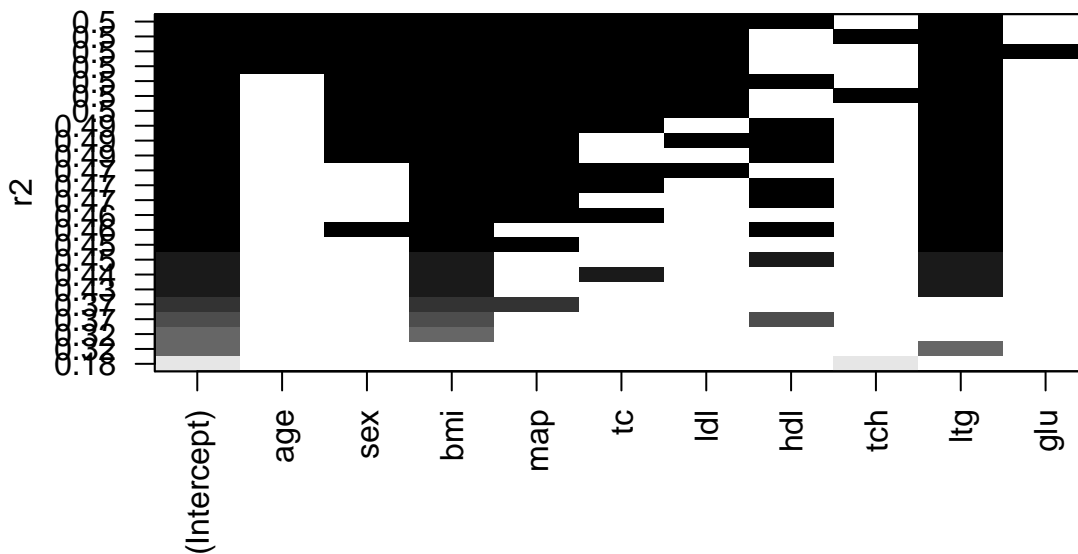
## Stochastic Search

Use 3 packages: “tabuSearch” “leaps” “GA” (genetic algorithm, note there is also “genalg”)

```
library(leaps)
```

```
##
## Attaching package: 'leaps'
##
## The following object is masked from 'package:subselect':
##
## leaps
```

```
# leaps and bounds
leaps <- regsubsets(Y ~.,,data=diab.train,nbest=3)
#summary(leaps)
# plot a table of models showing variables in each model.
# models are ordered by the selection statistic.
plot(leaps,scale="r2")
```



```
# plot statistic by subset size
library(car)
```

```
##
## Attaching package: 'car'
##
## The following object is masked from 'package:boot':
##
## logit
```

```
subsets(leaps, statistic="rsq",legend=F)
```

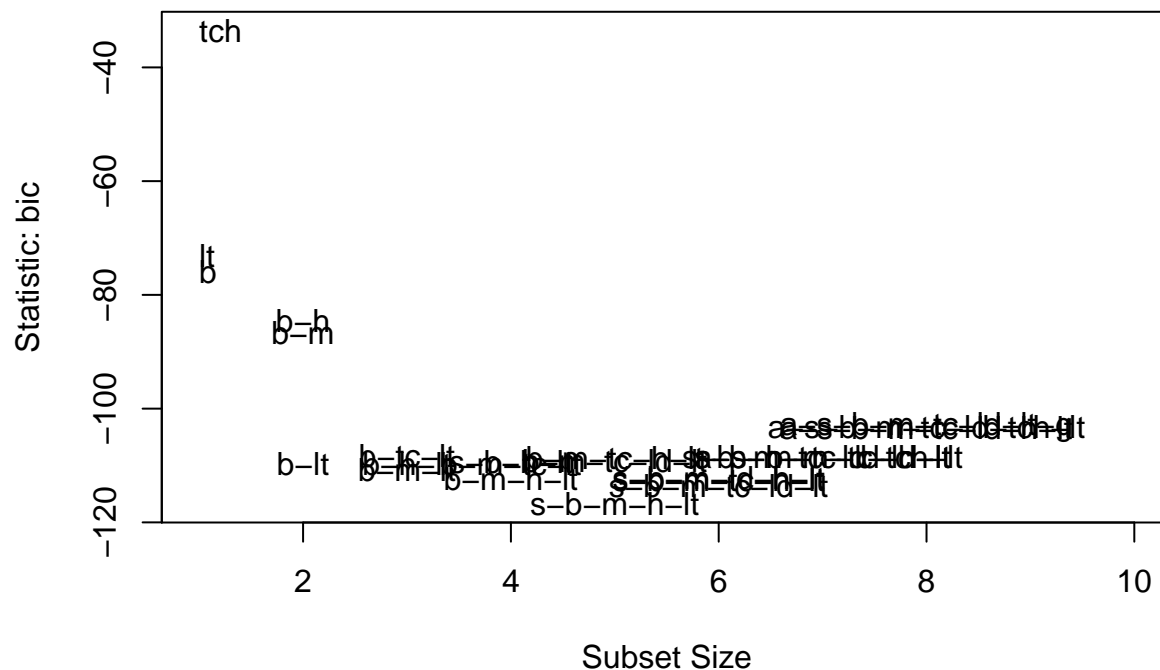


Based on this, looks like a subset size of 5 might be best, as we get pretty close to an  $R^2$  of 0.5

```
summary(leaps)$bic
```

```
## [1] -75.96156 -73.12543 -33.54266 -109.62248 -86.61304 -84.54964
## [7] -110.67974 -109.71542 -108.31248 -112.38244 -109.63451 -109.39965
## [13] -116.69919 -109.11901 -108.57735 -113.64905 -112.48698 -112.10288
## [19] -108.48394 -108.40795 -108.38271 -103.31248 -103.26582 -103.16310
```

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##	Abbreviation
## age	a
## sex	s
## bmi	b
## map	m
## tc	tc
## ldl	ld
## hdl	h
## tch	tch
## ltg	lt
## glu	g

Looks like we get a similar answer, but seem to also have a “clearer” best model The

## Genetic Algorithm

An example using the “GA” package to do this: <http://www.jstatsoft.org/v53/i04/paper>

But there is an easier way:

```
library(glmulti)
```

```
## Loading required package: rJava
```

Using “GA”

```
mod <- lm(Y ~ ., diab.train)

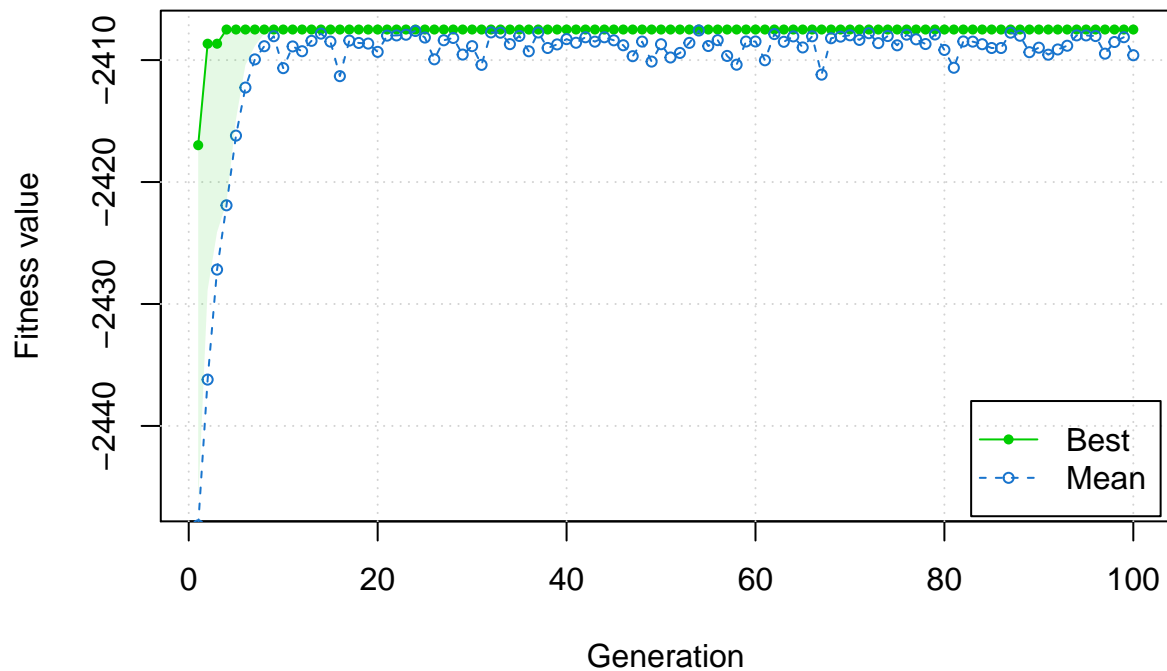
x <- model.matrix(mod)[, -1]
y <- model.response(model.frame(mod))
```

```

fitness <- function(string) {
  inc <- which(string == 1)
  X <- cbind(1, x[,inc])
  mod <- lm.fit(X, y)
  class(mod) <- "lm"
  -AIC(mod)
}

GA <- ga("binary", fitness = fitness, nBits = ncol(x),
        names = colnames(x), monitor=F)
plot(GA)

```



```
summary(GA)
```

```

## +-----+
## |      Genetic Algorithm      |
## +-----+
##
## GA settings:
## Type                = binary
## Population size      = 50
## Number of generations = 100
## Elitism              = 2
## Crossover probability = 0.8
## Mutation probability = 0.1
##
## GA results:
## Iterations           = 100
## Fitness function value = -2407.495
## Solution              =
##   age sex bmi map tc ldl hdl tch ltg glu

```

```
## [1,] 0 1 1 1 1 1 0 0 1 0
```

```
summary(GA)$solution
```

```
##      age sex bmi map tc ldl hdl tch ltg glu
## [1,] 0  1  1  1  1  1  0  0  1  0
```

## Tabu Search

An example: <http://www.r-bloggers.com/finding-the-best-subset-of-a-gam-using-tabu-search-and-visualizing-it-in-r/>

```
library(tabuSearch)

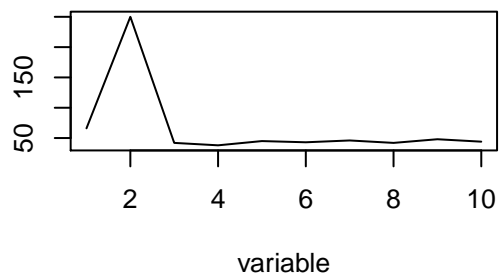
mod <- lm(Y ~ ., diab.train)

x <- model.matrix(mod)[, -1]
y <- model.response(model.frame(mod))

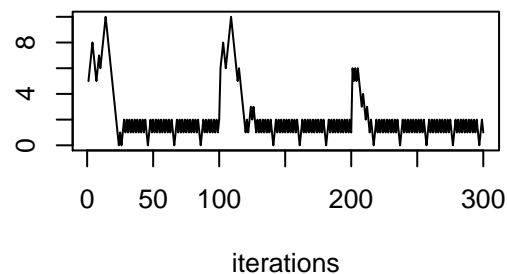
fitness2 <- function(string) {
  inc <- which(string == 1)
  X <- cbind(1, x[,inc])
  mod <- lm.fit(X, y)
  class(mod) <- "lm"
  -AIC(mod) + 100000 # won't take negative
}

result <- tabuSearch(size = 10, iters = 100, objFunc = fitness2)
plot(result) #fit margins too large
```

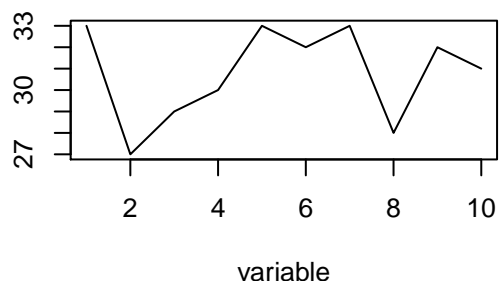
**No of times selected**



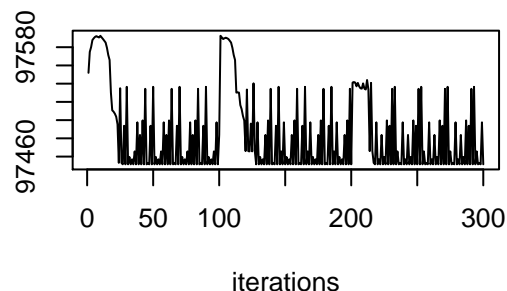
**Sum of included variables**



**Most frequent moves**



**Objective Function**



```
summary(result,verbose=T) # 6 predictors
```

```
## Tabu Settings
##   Type                               = binary configuration
##   No of algorithm repeats             = 1
##   No of iterations at each prelim search = 100
##   Total no of iterations              = 300
##   No of unique best configurations     = 53
##   Tabu list size                      = 9
##   Configuration length                = 10
##   No of neighbours visited at each iteration = 10
## Results:
##   Highest value of objective fn       = 97592.50529
##   Occurs # of times                   = 2
##   Optimum number of variables         = c(6, 6)
## Optimum configuration:
##      [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8] [,9] [,10]
## [1,]    0    1    1    1    1    1    0    0    1    0
## [2,]    0    1    1    1    1    1    0    0    1    0
```

## glmulti

Looks for interactions as well

```
library(glmulti)
# method = "g" for genetic algorithm
# default fitfunction = "glm"
multi.out = glmulti(Y ~., data=diab.train,method="g",plotty=F,
                    report=F,fitfunction="lm",crit="aic")
```

```
## TASK: Genetic algorithm in the candidate set.
## Initialization...
## Algorithm started...
## Improvements in best and average IC have been going on below the specified goals.
## Algorithm is declared to have converged.
## Completed.
```

```
#summary(multi.out)
summary(multi.out)$bestmodel
```

```
## [1] "Y ~ 1 + sex + bmi + map + tc + ldl + hdl + ltg + map:age + map:bmi + "
## [2] "      tc:bmi + tc:map + hdl:map + tch:sex + tch:map + ltg:hdl + "
## [3] "      glu:ldl + glu:hdl + glu:tch"
```

Now we can take the output and test it out in lm()

```
eq = summary(multi.out)$bestmodel
lm.multi = lm(eq,data=diab.train)
summary(lm.multi)
```

```
##
## Call:
## lm(formula = eq, data = diab.train)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -115.651  -35.666   -0.265   31.718  133.029
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    140.605      4.629  30.372 < 2e-16 ***
## sex           -211.729     84.639  -2.502 0.013159 *
## bmi            410.735     99.186   4.141 5.08e-05 ***
## map            258.410     90.414   2.858 0.004709 **
## tc           -2261.529    704.476  -3.210 0.001543 **
## ldl            1948.791    609.747   3.196 0.001617 **
## hdl             507.116    275.523   1.841 0.067153 .
## ltg            1216.043    269.027   4.520 1.05e-05 ***
## map:age        3134.578   1855.303   1.690 0.092662 .
## bmi:map        5120.912   1846.306   2.774 0.006063 **
## bmi:tc        -3667.190   2055.736  -1.784 0.075944 .
## map:tc        12725.555   3612.883   3.522 0.000529 ***
## map:hdl     -10510.047   4107.781  -2.559 0.011243 *
## sex:tch       -2946.263   1687.626  -1.746 0.082365 .
## map:tch      -14902.962   5002.341  -2.979 0.003244 **
## hdl:ltg       -3985.408   2361.781  -1.687 0.093059 .
## ldl:glu       -7539.661   3394.820  -2.221 0.027466 *
## hdl:glu       10460.555   3976.401   2.631 0.009179 **
## tch:glu       14119.843   4724.852   2.988 0.003152 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 51.88 on 202 degrees of freedom
## Multiple R-squared:  0.5772, Adjusted R-squared:  0.5395
## F-statistic: 15.32 on 18 and 202 DF,  p-value: < 2.2e-16
```

The addition of interactions increases our  $R^2$  even though our criterion for glmulti was AIC.

## simulated annealing

<http://topepo.github.io/caret/SA.html>

part of this, other options including GA available: <http://topepo.github.io/caret/featureselection.html>

To do, have to install most current version of caret from github. CRAN version doesn't include functions

The great thing about this function is that we can use it for all of the methods in caret (100+).

```
library(devtools)
devtools::install_github("cran/caret")
library(caret)

ctrl <- safsControl(functions = caretSA)
```



```
obj <- safes(x = diab.train[,-1],
            y = diab.train$Y,
            iters = 50,
            safesControl = ctrl,
            method = "lm")
#quartz()
plot(obj) + theme_bw()
# should increase the iterations
```

Use with genetic algorithm: <http://topepo.github.io/caret/GA.html>

Pretty slow

```
ctrl <- gafsControl(functions = caretGA)
obj <- gafs(x = diab.train[,-1],
            y = diab.train$Y,
            iters = 50,
            gafsControl = ctrl,
            method = "lm")
```

## Classification

All of the methods used previously will also work in the classification context in using forms of logistic regression.

### Logistic Regression

```
Ybin <- ifelse(diab.train$Y > mean(diab.train$Y),1,0)
diab.train2 <- diab.train
diab.train2$Y <- Ybin
# logistic
glm.out <- glm(Y ~ ., diab.train2,family="binomial")
#summary(glm.out)
```

So how well did we do? I like using receive operating characteristic (ROC) curves and the area under the curve (AUC) to evaluate results in classification.

For binary outcomes, my favorite way to assess how well I did is using the confusionMatrix() function from caret

sensitivity = true positive – with event AND predicted to have event / having the event

people who respondend AND people predicted to respond / people who respondend

specificity = samples w/o event AND predicted as non-event / samples w/o event

actual non-response AND predicted non-response/ actual non-response

false-positive = 1-specificity

```
library(pROC)
library(caret)
```

```
glm.probs=predict(glm.out,type="response")
glm.pred=ifelse(glm.probs>0.5,1,0)
```

```
table(diab.train2$Y,glm.pred)
```

```
##      glm.pred
##      0  1
## 0 89 32
## 1 30 70
```

```
confusionMatrix(diab.train2$Y,glm.pred,positive="1") # from caret package
```

```
## Confusion Matrix and Statistics
##
##              Reference
## Prediction  0  1
##           0 89 32
##           1 30 70
##
##              Accuracy : 0.7195
##              95% CI : (0.6553, 0.7776)
##      No Information Rate : 0.5385
##      P-Value [Acc > NIR] : 2.664e-08
##
##              Kappa : 0.4348
##  Mcnemar's Test P-Value : 0.8989
##
##              Sensitivity : 0.6863
##              Specificity : 0.7479
##              Pos Pred Value : 0.7000
##              Neg Pred Value : 0.7355
##              Prevalence : 0.4615
##              Detection Rate : 0.3167
##      Detection Prevalence : 0.4525
##              Balanced Accuracy : 0.7171
##
##              'Positive' Class : 1
##
```

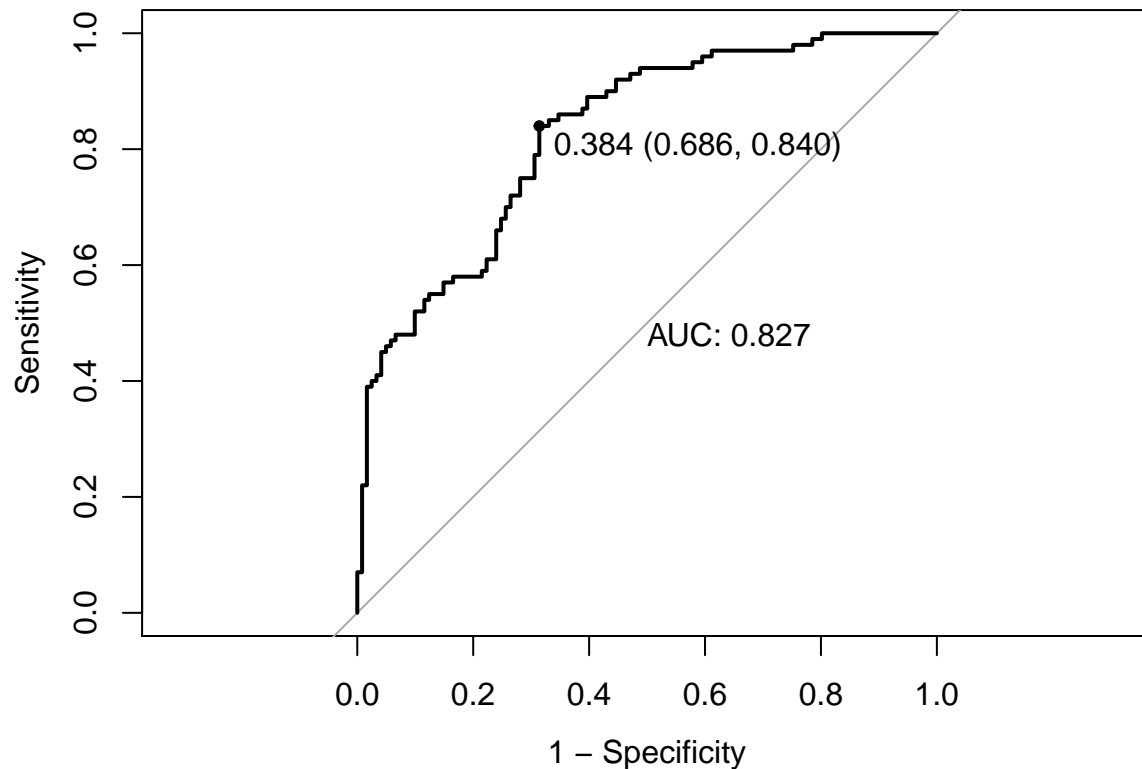
```
rocCurve <- roc(diab.train2$Y,glm.probs)
auc(rocCurve)
```

```
## Area under the curve: 0.8269
```

```
ci.roc(rocCurve)
```

```
## 95% CI: 0.7739-0.8798 (DeLong)
```

```
plot(rocCurve, legacy.axes = TRUE, print.thres=T, print.auc=T)
```

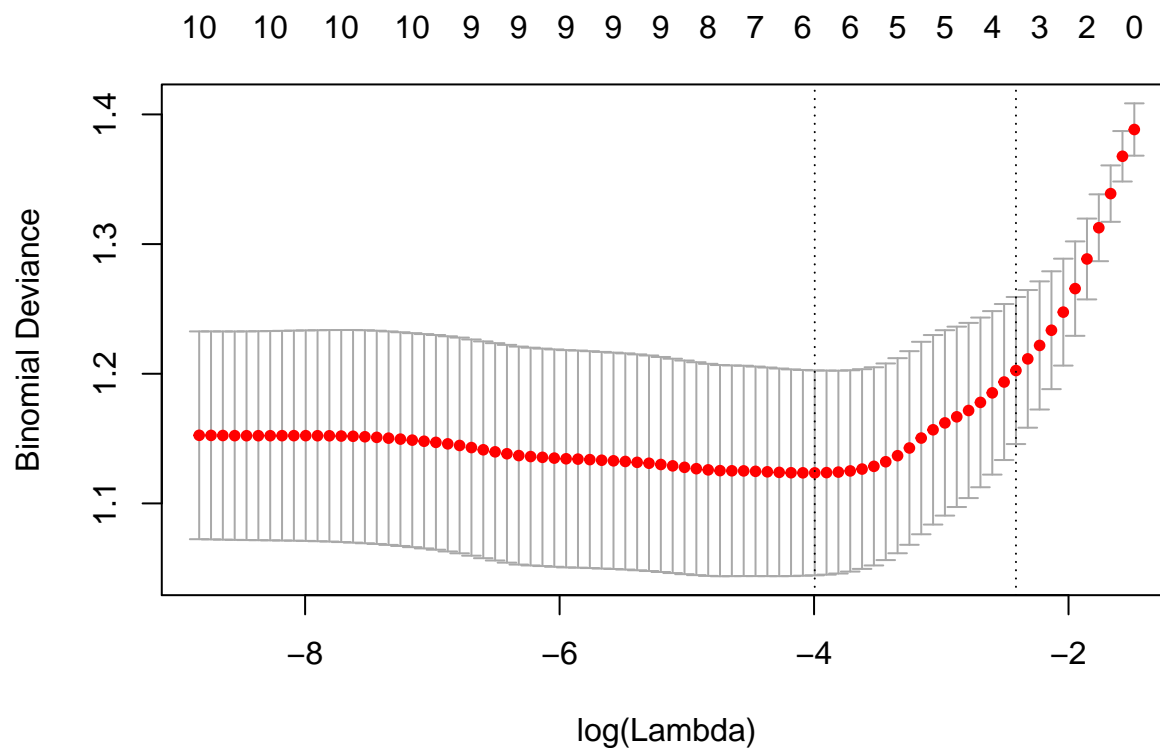


```
##  
## Call:  
## roc.default(response = diab.train2$Y, predictor = glm.probs)  
##  
## Data: glm.probs in 121 controls (diab.train2$Y 0) < 100 cases (diab.train2$Y 1).  
## Area under the curve: 0.8269
```

Try also with elastic net

Regularization in caret package: [http://topepo.github.io/caret/L1\\_Regularization.html](http://topepo.github.io/caret/L1_Regularization.html)

```
XX <- as.matrix(diab.train2[,-1])  
YY <- diab.train2$Y  
lasLog <- cv.glmnet(XX,YY,family="binomial")  
plot(lasLog)
```



```
pred.lasLog <- predict(lasLog, XX, s="lambda.1se",type="response")
```

How did we do?

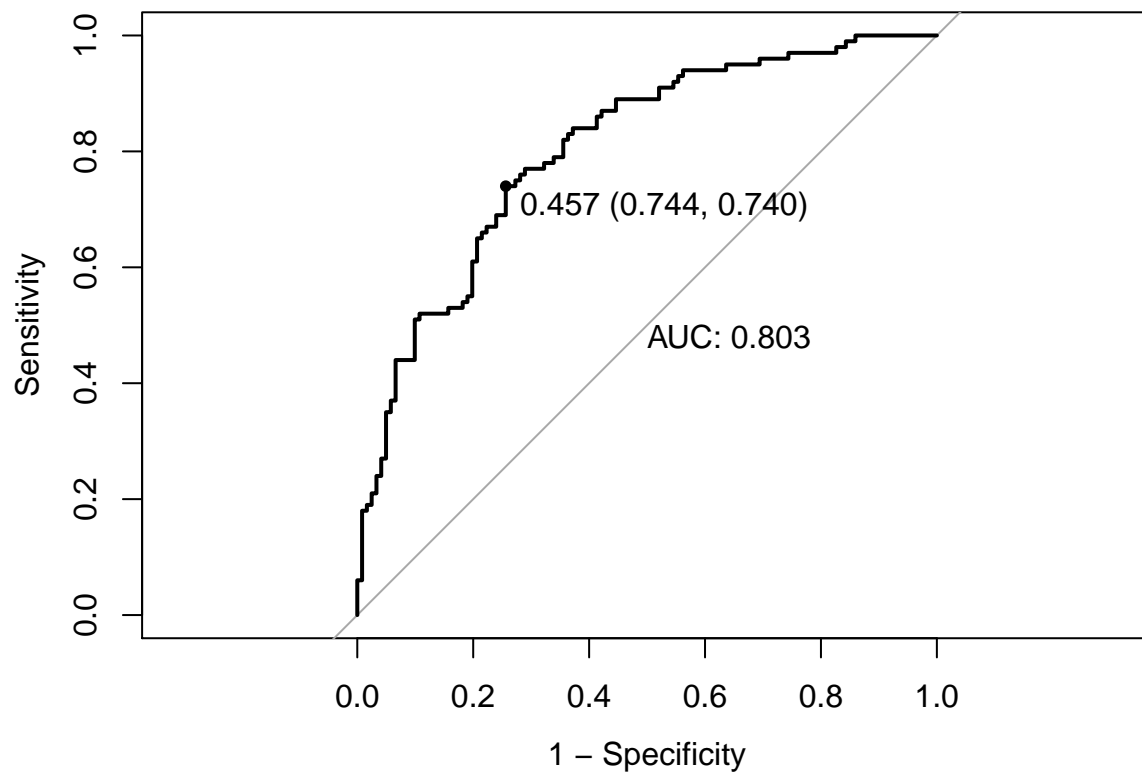
```
rocCurve2 <- roc(diab.train2$Y,pred.lasLog)
auc(rocCurve2)
```

```
## Area under the curve: 0.8034
```

```
ci.roc(rocCurve2)
```

```
## 95% CI: 0.7461-0.8606 (DeLong)
```

```
plot(rocCurve2, legacy.axes = TRUE,print.thres=T,print.auc=T)
```



```
##
## Call:
## roc.default(response = diab.train2$Y, predictor = pred.lasLog)
##
## Data: pred.lasLog in 121 controls (diab.train2$Y 0) < 100 cases (diab.train2$Y 1).
## Area under the curve: 0.8034
```

We got rid of 6 predictors and only lost 0.01 in the AUC