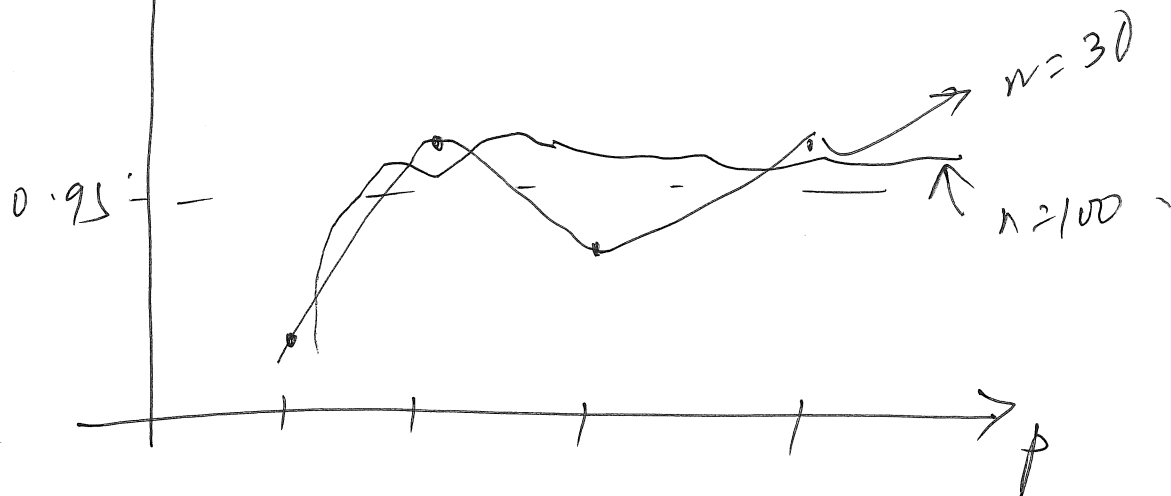


Project #2

#2

Estimated
coverage
prob.



\hat{p}_n R:
plot(

, xlim=c(-1, 1),
ylim=c(-1, 1), type="l")

lines(, lty=2)

lines(, lty=3)

legend(locator(1), legend=c("n=30", "-"),
lty=1:3)

Summary of steps in a hypothesis test:

- Formulate H_0 and H_1
- Find a test statistic T and get its null distribution
- Compute T_{obs}
- Use the null distribution to compute either the critical point or the p -value for the test.
- State your conclusion. (in layman terms also).

↓
don't just say H_0 is accepted or rejected.
 $T_{\text{obs}} \mid H_0 \text{ is true}$
 T is as extreme or more extreme than

Recall:

$p\text{-value} = P[T \text{ is as extreme or more extreme than } T_{\text{obs}} \mid H_0 \text{ is true}]$
Level- α test:

If $p\text{-value} \leq \alpha$, reject H_0 , but accept it.
if $p\text{-value} > \alpha$.

Some specific tests

One-sample tests for μ where $X \sim N(\mu, \sigma^2)$

Case 1: z-test (known σ^2): $H_0: \mu = \mu_0$

Test statistic: $T = \frac{\hat{\theta} - \theta_0}{\hat{\sigma}_c(\hat{\theta}_0)}$

Here: $Z =$

$$\frac{\bar{X} - \mu_0}{\sigma / \sqrt{n}}$$

Note: When H_0 is true:
 $\bar{X} \sim N\left[\mu = \mu_0, \frac{\sigma^2}{n}\right]$

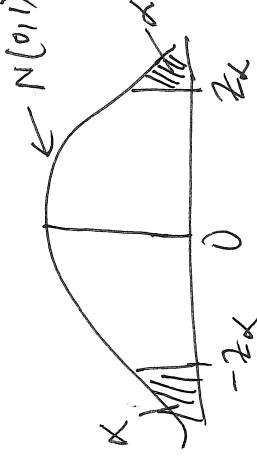
$\sim N(0,1)$ when H_0 is true.

Critical point for the level α test:

One-sided alternative: Z_α or $-Z_\alpha$ right-tailed left-tailed

Two-sided alternative:

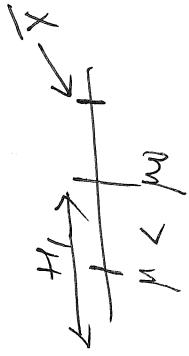
$\pm Z_{\alpha/2}$



Note: These critical points guarantee that

$$P[\text{reject } H_0 \mid H_0] = \alpha.$$

\uparrow
 $P[\text{Type I error}]$



p-value:

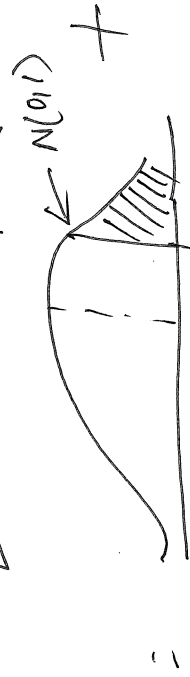
$$F(x) = \text{cdf of } N(0,1).$$

| H_1 | reject when | p-value | computing p-value |
|------------------|----------------------------|-------------------------------|-----------------------|
| $\mu \neq \mu_0$ | $ z_{obs} > z_{\alpha/2}$ | $P[Z \geq z_{obs} H_0]$ | $2(1 - F(z_{obs}))$ |
| $\mu > \mu_0$ | $z_{obs} > z_{\alpha}$ | $P[Z \geq z_{obs} H_0]$ | $1 - F(z_{obs})$ |
| $\mu < \mu_0$ | $z_{obs} < -z_{\alpha}$ | $P[Z \leq z_{obs} H_0]$ | $F(z_{obs})$ |

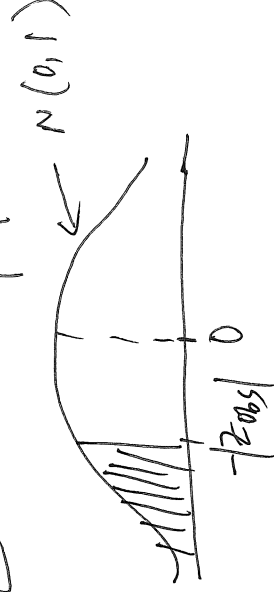
computing: $P[|Z| \geq |z_{obs}| | H_0]$

$$= P[Z \geq |z_{obs}| | H_0] + P[Z \leq -|z_{obs}| | H_0]$$

$$= P[Z \geq |z_{obs}| | H_0] + P[Z \leq -|z_{obs}| | H_0]$$



$$= 2(1 - F(|z_{obs}|))$$



Case 2: t-test (unknown σ^2): $H_0: \mu = \mu_0$

Test statistic:

$$t = \frac{\bar{X} - \mu_0}{\hat{se}(\bar{X})} = \frac{\bar{X} - \mu_0}{s/\sqrt{n}}$$

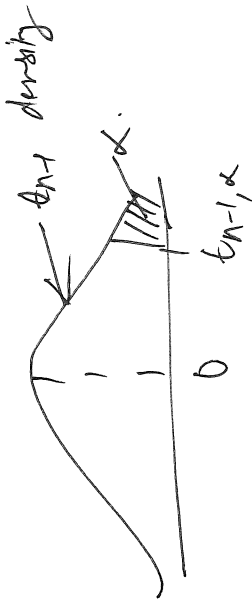
sample.
variance

$\sim t_{n-1}$ when H_0 is true.

Critical point for the level α test:

One-sided alternative: $t_{n-1, \alpha}$ or $-t_{n-1, \alpha}$

Two-sided alternative: $t_{n-1, \alpha/2}$

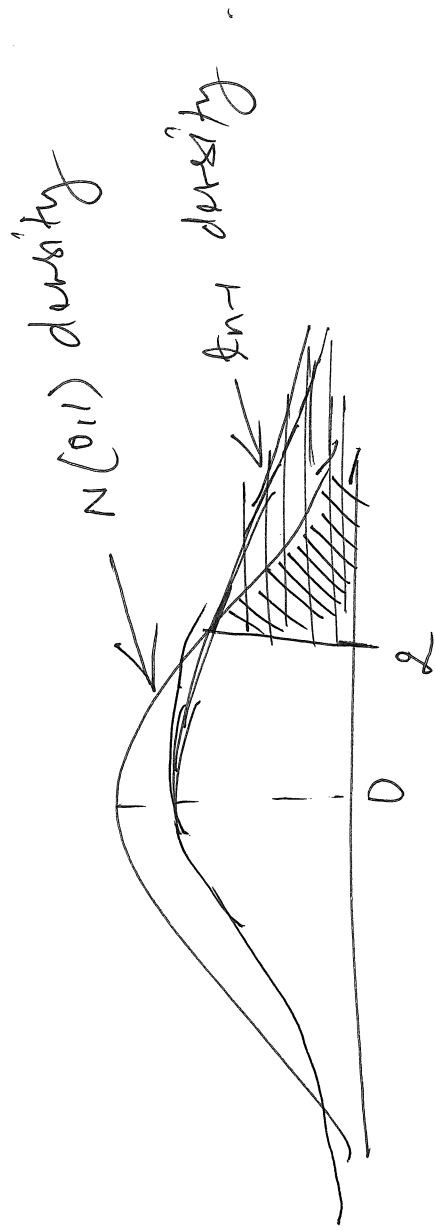


p-value:

here: $F(x) = \text{cdf of } t_{n-1}$

| H_1 | reject when | p-value | computing p-value |
|------------------|---|--------------------------------------|------------------------------|
| $\mu \neq \mu_0$ | $ t_{\text{obs}} \geq t_{n-1, \alpha/2}$ | $P(t \geq t_{\text{obs}} H_0)$ | $2(1 - F(t_{\text{obs}}))$ |
| $\mu > \mu_0$ | $t_{\text{obs}} \geq t_{n-1, \alpha}$ | $P(t \geq t_{\text{obs}} H_0)$ | $1 - F(t_{\text{obs}})$ |
| $\mu < \mu_0$ | $t_{\text{obs}} \leq -t_{n-1, \alpha}$ | $P(t \leq t_{\text{obs}} H_0)$ | $F(t_{\text{obs}})$ |

Recall:



~~probability assuming t_{n-1} > probability assuming~~

$$1 - F(2) \text{ assuming } N(0,1) \leq 1 - F(2) \text{ assuming } t_{n-1}.$$

One-sample test for μ when X is nonnormal

Large-sample z -test: $H_0 : \mu = \mu_0$

- Need large n but works for mean of any (non-normal) population

- Use the z -test with test statistic

$$z = \frac{\bar{X} - \mu_0}{s/\sqrt{n}} \quad n \quad t = \frac{\bar{X} - \mu_0}{s/\sqrt{n}} \quad \sim N(0,1) \text{ when } n \text{ is large.}$$

$\sim N(0,1)$

when n is large

- When n is large, the null distribution is approximately $N(0, 1)$ due to central limit theorem.
- This test has approximate level α .

One-sample test for population prop p

- The large-sample z -test works because in this case $X \sim \text{Bernoulli}(p)$ and $E(X) = p$. $H_0: p = p_0$.
- Use the z -test with test statistic

$$Z = \left\{ \frac{\hat{p} - p_0}{\sqrt{\frac{p_0(1-p_0)}{n}}} \right\} \sim N(0,1) \text{ when } n \text{ is large.}$$

\nearrow \rightarrow better uses.

α $\frac{\hat{p} - p_0}{\sqrt{\frac{\hat{p}(1-\hat{p})}{n}}}$ \rightarrow ok to use. \nearrow se. (\hat{p}) when H_0 is true.

- This test has approximate level α .

Ex 1: A long-time authorized user of a computer account takes 0.2 seconds on average between keystrokes. One day, when a user typed in the username and password, 15 times between \bar{x} keystrokes were recorded. These data had mean of 0.3 seconds and standard deviation of 0.12 seconds. Do these data given evidence of an unauthorized login attempt? Assume normality for the time b/w keystrokes, and 5% level of significance.

Recall:

$$X = \text{time b/w key strokes for the person who is trying to log in}$$

$$\theta = E[X]$$

$$H_0: \theta = 0.2 \text{ s} \quad \text{vs} \quad H_1: \theta \neq 0.2 \text{ s}$$

\uparrow
authorized attempt

\uparrow
unauthorized attempt

pop. is unknown.

t-test because population is normal and SD \neq pop.

$$t_{\text{obs}} = \frac{0.3 - 0.2}{0.12 / \sqrt{15}} = \frac{3.227}{0.031} = 104.1$$

compute:
 $t_{n-1, \alpha/2} = t_{14, 0.025, df=14} = 2.145$

p-value: $2[1 - \Phi(3.227)] = 0.006$
 level: 5% \uparrow int: since p-value $< 0.05 \Rightarrow$ reject H_0 .

\Rightarrow attempt is ~~unauthorized~~ unauthorized