

## HW#5 - Solutions

ESE-520

Total : 4 op.

Pr. 1. [5p]

a) Given :  $X \sim \text{exp.}(\lambda) \Rightarrow f_X(x) = \lambda e^{-\lambda x}, x > 0$

$$\Rightarrow F_X(x) = \int_0^x \lambda e^{-\lambda y} dy = -e^{-\lambda y} \Big|_0^x = 1 - e^{-\lambda x}, x > 0 \text{ (0 otherwise)}$$

$$\Rightarrow P(X > t+s) = 1 - F_X(t+s) = e^{-\lambda(t+s)} = e^{-\lambda t} \cdot e^{-\lambda s} =$$

$$= P(X > t) \cdot P(X > s) \text{ for all } s, t \geq 0,$$

b) The equivalency of statements

$$P(X > t+s) = P(X > t) \cdot P(X > s)$$

and

$$P(X > t+s | X > s) = P(X > t)$$

follows from the fact that

$$\frac{P(X > t+s)}{P(X > s)} = \frac{P(X > t+s, X > s)}{P(X > s)} = P(X > t+s | X > s)$$

Since  $\{X > t+s\} \subseteq \{X > s\}$  for all  $s, t \geq 0$ .

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Pr.2 [5P] Given:  $(N_t), t \geq 0$  is a Poisson process

Calculate:  $P(N_s=0, N_t=1)$  for any  $0 \leq s \leq t$ .

$$\rightarrow P(N_s=0, N_t=1) = P(0 \leq s < T_1 \leq t < T_1 + T_2)$$

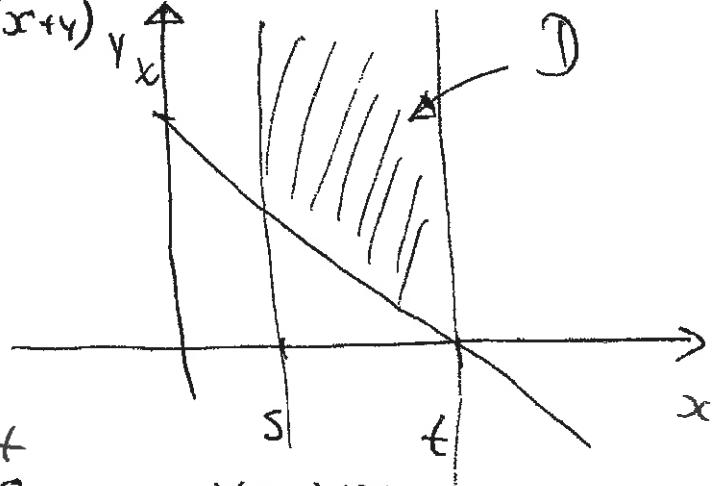
where  $T_1, T_2 \sim \exp(\lambda)$ ,  $\lambda > 0$  & independent.

$$f_{T_1}(x) = \lambda \cdot e^{-\lambda x}, f_{T_2}(y) = \lambda \cdot e^{-\lambda y}, x > 0, y > 0 \Rightarrow$$

$$f_{T_1, T_2}(x, y) = \lambda^2 \cdot e^{-\lambda(x+y)}, x > 0, y > 0$$

$$\Rightarrow P(0 \leq s < T_1 \leq t < T_1 + T_2) =$$

$$= \iint \lambda^2 \cdot e^{-\lambda(x+y)} dx dy =$$



$$= \int_s^t \int_{t-x}^{\infty} \lambda^2 e^{-\lambda(x+y)} dy dx = \int_s^t -\lambda [e^{-\lambda(x+y)}]_{t-x}^{\infty} dx =$$

$$= \int_s^t \lambda e^{-\lambda(x+t-x)} dx = \int_s^t \lambda e^{-\lambda t} dx = \boxed{\lambda(t-s) \cdot e^{-\lambda t}}$$

Pr.3 [5P] Given:  $(N_t), t \geq 0$  is a Poisson process

with parameter  $\lambda > 0$ .

$$\rightarrow \text{We recall: } \sinh(x) = \frac{e^x - e^{-x}}{2} = \sum_{n=0}^{\infty} \frac{x^{2n+1}}{(2n+1)!}$$

$$\cosh(x) = \frac{e^x + e^{-x}}{2} = \sum_{n=0}^{\infty} \frac{x^{2n}}{(2n)!}$$

• By definition & properties of  $(N_t)$ :

$$P(N_t \text{ is odd}) = \sum_{n=0}^{\infty} P(N_t = z_{n+1}) = \\ = \sum_{n=0}^{\infty} e^{-\lambda t} \frac{(\lambda t)^{2n+1}}{(2n+1)!} = e^{-\lambda t} \cdot \sinh(\lambda t),$$

$$P(N_t \text{ is even}) = \sum_{n=0}^{\infty} P(N_t = z_n) = \\ = \sum_{n=0}^{\infty} e^{-\lambda t} \frac{(\lambda t)^{2n}}{(2n)!} = e^{-\lambda t} \cdot \cosh(\lambda t). \quad \square,$$

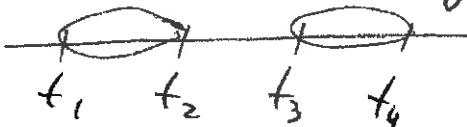
Pr. 5 [5p] Given:  $(W_t)$ ,  $t \geq 0$  is a Wiener process.

Show:  $V_t := W_{T+t} - W_T$ ,  $t \geq 0$  is a Wiener process

We have:

i)  $V_0 = W_T - W_T = 0$  ; non-overlapping

ii) For any  $t_1 < t_2 < t_3 < t_4$ ,

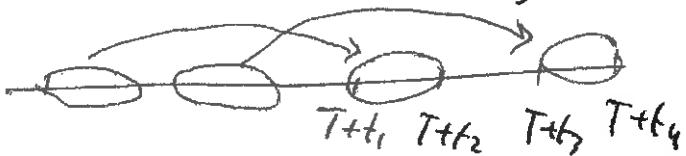


consider  $VV'$ :

$$V_{t_2} - V_{t_1} = W_{T+t_2} - W_T - (W_{T+t_1} - W_T) = W_{T+t_2} - W_{T+t_1}$$

$$V_{t_4} - V_{t_3} = W_{T+t_4} - W_T - (W_{T+t_3} - W_T) = W_{T+t_4} - W_{T+t_3}$$

It is clear that  $(V_t)$  has also independent increments.



iii) For any  $s < t$ :

$$V_t - V_s = W_{T+t} - W_T - (W_{T+s} - W_T) = W_{T+t} - W_{T+s}$$



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$V_t - V_s$  has normal distribution  $\mathbb{R}$

$$E(V_t - V_s) = E(W_{T+t}) - E(W_{T+s}) = 0,$$

$$\begin{aligned} \text{Var}(V_t - V_s) &= \text{Var}(W_{T+t} - W_{T+s}) = (T+t) - (T+s) \\ &= t-s \end{aligned}$$

$$\Rightarrow V_t - V_s \sim N(0, t-s).$$

$\Rightarrow (V_t)$  is a Wiener process.

$$\text{Pr. 6 [5p]} \text{ Given: } h(t, x, y) = \frac{1}{\sqrt{2\pi t}} \cdot e^{-\frac{(y-x)^2}{2t}}$$

$\rightarrow$  Pure differentiation exercise:

$$\frac{\partial h}{\partial t} = \frac{y^2 - 2yx + x^2 - t}{2t} \cdot h(t, x, y)$$

$$\frac{\partial h}{\partial y} = \frac{x-y}{t} \cdot h(t, x, y)$$

$$\frac{\partial^2 h}{\partial y^2} = \frac{y^2 - 2yx + x^2 - t}{t} \cdot h(t, x, y)$$

$\Rightarrow$

$$\frac{\partial h}{\partial t} = \frac{1}{2} \frac{\partial^2 h}{\partial y^2} \quad \square.$$

$\text{Pr. 7 [5p]} \text{ Given: } W_t^1 \text{ & } W_t^2 \text{ are independent BM's}$

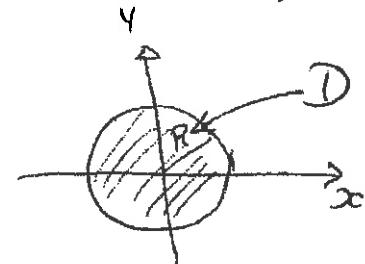
$$\text{and } |W_t| = \sqrt{(W_t^1)^2 + (W_t^2)^2}.$$

$\rightarrow$  The joint pdf of the vector  $(W_t)$  is  $-\frac{(x^2+y^2)}{2t}$

$$f_{W_t}(x, y) = f_{W_t^1}(x) \cdot f_{W_t^2}(y) = \frac{1}{2\pi t} \cdot e^{-\frac{(x^2+y^2)}{2t}}, x, y \in \mathbb{R}$$

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$$\Rightarrow P(|W_t| < R) = \iint_{\{(x,y) : \sqrt{x^2+y^2} < R\}} \frac{1}{2\pi t} e^{-\frac{(x^2+y^2)}{2t}} dx dy$$



= we use polar coordinates =

$$= \frac{1}{2\pi t} \cdot \int_0^R \int_0^{2\pi} r \cdot e^{-\frac{r^2}{2t}} d\theta dr =$$

$$= \frac{1}{t} \int_0^R r \cdot e^{-\frac{r^2}{2t}} dr = -e^{-\frac{r^2}{2t}} \Big|_0^R = 1 - e^{-\frac{R^2}{2t}}$$

Pr. 4 [ 5p ]

We find the characteristic function of  $M_t$ :

$$\begin{aligned} \varphi_{M_t}(s) &= E\left[e^{isM_t}\right] = E\left[e^{is\sum_{k=1}^{N_t} Y_k}\right] = \\ &= \sum_{n=0}^{\infty} E\left[e^{is\sum_{k=1}^{N_t} Y_k} \mid N_t = n\right] \cdot P(N_t = n) = \\ &= \sum_{n=0}^{\infty} E\left[e^{is\sum_{k=1}^n Y_k} \mid N_t = n\right] \cdot P(N_t = n) \\ &= \text{use independence} = \sum_{n=0}^{\infty} E\left[e^{is\sum_{k=1}^n Y_k}\right] \cdot P(N_t = n) = \\ &= \sum_{n=0}^{\infty} \prod_{k=1}^n E\left(e^{isY_k}\right) \cdot P(N_t = n) \\ &= \sum_{n=0}^{\infty} \left(e^{is} p + (1-p)\right)^n \cdot e^{-nt} \cdot \frac{(it)^n}{n!} = \end{aligned}$$

$$= e^{-pt} \cdot e^{[e^{is}p + (-p)]t} = e^{ispt - pt}$$

$\Rightarrow M_f \sim \text{Poisson}(pt)$

Prob. 8 [5pt]

$$\text{Let } X = \begin{pmatrix} X_1 \\ X_2 \\ \vdots \\ X_n \end{pmatrix} = \begin{pmatrix} W_{t_1} \\ W_{t_2} - W_{t_1} \\ \vdots \\ W_{t_n} - W_{t_{n-1}} \end{pmatrix} \quad \& \quad Y = \begin{pmatrix} Y_1 \\ \vdots \\ Y_n \end{pmatrix} = \begin{pmatrix} W_{t_1} \\ W_{t_2} \\ \vdots \\ W_{t_n} \end{pmatrix}$$

By independence of increments,

$$f_X(x_1, x_2, \dots, x_n) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi(t_i - t_{i-1})}} e^{-\frac{x_i^2}{2(t_i - t_{i-1})}}$$

$$\text{Let } A = \begin{pmatrix} 1 & 0 & \cdots & 0 \\ 1 & 1 & 0 & \cdots & 0 \\ 1 & 1 & 1 & \ddots & \vdots \\ 1 & 1 & \ddots & \ddots & 1 \end{pmatrix} \Rightarrow Y = A \cdot X \quad \text{with } f_0 = 0.$$

By transformation formula,

$$f_Y(y_1, \dots, y_n) = f_X(A^{-1}y) \cdot \frac{1}{|\det A|}$$

$$\text{Since } f_X(A^{-1}y) = f_X(A^{-1}y)$$

$$X = \begin{pmatrix} X_1 \\ \vdots \\ X_n \end{pmatrix} = \begin{pmatrix} W_1 - w_0 \\ W_2 - W_1 \\ \vdots \\ W_n - W_{n-1} \end{pmatrix}, \Rightarrow -\frac{(W_i - W_{i-1})^2}{2(t_i - t_{i-1})}$$

$$f_X(w_1, \dots, w_n) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi(t_i - t_{i-1})}} e^{-\frac{(w_i - w_{i-1})^2}{2(t_i - t_{i-1})}}$$