

ESE 520 Probability and Stochastic Processes

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EXAM 2-Sample

Name (Please print):

Solutions

Instructions:

1. You must show all work to completely justify your answers in order to receive any credit.
2. You can use One one-sided sheet of paper with your own formulas.

1. a) Assume X has a normal distribution with mean μ and variance σ^2 . Find the moment-generating function of X .

- $X \sim N(\mu, \sigma^2) \Rightarrow Y = \delta \cdot X + \mu$ with $Y \sim N(0, 1)$
- $\psi_Y(t) = E e^{tY} = \int_{\mathbb{R}} e^{ty} \frac{1}{\sqrt{2\pi}} e^{-y^2/2} dy = e^{\mu t} \underbrace{\int_{\mathbb{R}} \frac{1}{\sqrt{2\pi}} e^{-\frac{(y-t)^2}{2}} dy}_{=1}$
 $= e^{\mu t + \frac{t^2}{2}}, t \in \mathbb{R}$
- $\psi_X(t) = E(e^{t(X+\delta Y)}) = e^{\mu t} \cdot E e^{t\delta Y} = e^{\mu t} e^{\frac{\delta^2 t^2}{2}}$
 $= e^{\mu t + \frac{\delta^2 t^2}{2}}$

b) Use part a) to find $E(X)$;

$$\psi'_X(t) = (\mu + t\delta^2) \cdot e^{\mu t + \frac{\delta^2 t^2}{2}}$$

$$\therefore E(X) = \psi'_X(0) = \mu.$$

2. Let X and Y be two random variables having jointly continuous distribution with the joint pdf given as

$$f(x, y) = \begin{cases} 8xy, & 0 < y < x < 1 \\ 0, & \text{otherwise.} \end{cases}$$

Find the best linear mean-square estimate of Y by X .

- $E(\bar{Y}|X) = \mu_Y + \frac{\text{cov}(X, Y)}{\text{Var}(X)} \cdot (X - \mu_X)$,
- $f_X(x) = \int_0^x 8xy dy = 4x^2 y |_0^x = 4x^3$;
- $\mu_X = E(X) = \int_0^1 4x^4 dx = \frac{4}{5} x^5 |_0^1 = \frac{4}{5}$;
- $f_Y(y) = \int 8xy dx = 4x^2 y |_y^1 = 4(y-y^2)$;
- $\mu_Y = E(Y) = \int_0^1 4(y^2 - y^4) dy = 4\left(\frac{y^3}{3} - \frac{y^5}{5}\right) |_0^1 = \frac{8}{15}$;
- $E(X^2) = \int_0^1 4x^5 dx = \frac{2}{3}$;
- $\text{Var}(X) = E(X^2) - (E(X))^2 = \frac{2}{3} - \left(\frac{4}{5}\right)^2 = \frac{2}{75}$;
- $E(XY) = \int_0^1 \left(\int_0^x 8xy^2 dy \right) dx = 8 \int_0^1 x^2 \frac{x^3}{3} |_0^x dx$
 $= \frac{8}{3} \int_0^1 x^5 dx = \frac{8}{3 \cdot 6} = \frac{4}{9}$,
- $\text{cov}(X, Y) = E(XY) - E(X) \cdot E(Y) = \frac{4}{9} - \frac{4}{5} \cdot \frac{8}{15} = \frac{4.75 - 9.32}{9.75}$
- $\frac{\text{cov}(X, Y)}{\text{Var}(X)} = \frac{4.75 - 9.32}{9.75} \cdot \frac{75}{2} = \frac{2.75 - 9.16}{9} = \frac{2}{3}$
- $E(\bar{Y}|X) = \frac{8}{15} + \frac{2}{3} \left(X - \frac{4}{5}\right)$

3. Suppose X and Y have joint density e^{-x} for $0 < y < x$. Find $E(X|Y)$ and $Var(X|Y)$.

$$\begin{aligned}
 \cdot f_Y(y) &= \int_y^{\infty} e^{-x} dx = e^{-y}, \quad y > 0 \\
 \cdot E(X|Y=y) &= \int_y^{\infty} x \cdot \frac{e^{-x}}{e^{-y}} dx = e^y / (-x \cdot e^{-y}) \Big|_y^{\infty} + \int_y^{\infty} e^{-x} dx \\
 &= e^y \left(y \cdot e^{-y} - e^{-y} \Big|_y^{\infty} \right) = y+1 \\
 \cdot E(X^2|Y=y) &= \int_y^{\infty} x^2 \cdot e^y \cdot e^{-x} dx = e^y \left(-x^2 \cdot e^{-x} \Big|_y^{\infty} + \right. \\
 &\quad \left. + 2 \int_y^{\infty} x \cdot e^{-x} dx \right) = e^y \left(y^2 e^{-y} + 2(y e^{-y} + e^{-y}) \right) \\
 &\quad \text{calculated before} \\
 &= y^2 + 2(y+1) \\
 \cdot Var(X|Y) &= E(X^2|Y) - (E(X|Y))^2 = \\
 &= y^2 + 2(y+1) - (y+1)^2 = 1
 \end{aligned}$$

4. Suppose you play roulette and bet \$1 on black each time. Your net winnings on the i th play, X_i , are 1 with probability $1/3$ and -1 with probability $2/3$. What is the probability your net winnings are ≥ 0 after 100 plays?

Use the Central Limit Theorem and the enclosed normal table to find the approximate probability.

- CLT: $\frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \sim N(0,1)$, $\bar{X}_n := \frac{X_1 + \dots + X_n}{n}$
 with $E(X_i) = 1 \cdot \frac{1}{3} - 1 \cdot \frac{2}{3} = -\frac{1}{3} = \mu$ &
 $\text{Var}(X_i) = E(X_i^2) - (E(X_i))^2 = 1 - \frac{1}{9} = \frac{8}{9} = \sigma^2$
- $P\left(\sum_{i=0}^{100} X_i \geq 0\right) = P\left(\bar{X}_{100} \geq 0\right) = P\left(\bar{X}_{100} - \mu \geq \frac{1}{3}\right)$
 $= P\left(\frac{\bar{X}_{100} - \mu}{\sigma/\sqrt{100}} \geq \frac{\frac{1}{3}}{\sqrt{8/100}}\right) = P(Z \geq \frac{10}{\sqrt{8}})$
 $\approx P(Z \geq 3.5) \approx 0$

5. Suppose U, V, W are independent and have mean μ and variance σ^2 . Find $\rho(U + V, V + W)$.

$$\begin{aligned} \cdot \rho(U+V, V+W) &= \frac{\text{cov}(U+V, V+W)}{\sqrt{\text{Var}(U+V) \cdot \text{Var}(V+W)}} \\ \cdot \text{cov}(U+V, V+W) &= E(U+V)(V+W) - E(U+V) \cdot E(V+W) \\ &= E(U \cdot V) + EV^2 + E(U \cdot W) + EW^2 - \left[\underbrace{E(U) + E(V)}_{\mu} \right] \left[\underbrace{(EV) + EW}_{\mu} \right] \\ &= \text{independence} \\ &= (EU)(EV) + EV^2 + (EU)(EW) + (EV)(EW) - 4\mu^2 \\ &= \mu^2 + \sigma^2 + \mu^2 + \mu^2 + \mu^2 - 4\mu^2 = \sigma^2 \\ \cdot \text{Var}(U+V) &= \text{Var}(U+W) = 2\sigma^2 \\ \Rightarrow \rho &= \frac{\sigma^2}{\sqrt{2\sigma^2 \cdot 2\sigma^2}} = \frac{1}{2}. \end{aligned}$$

6. Suppose X and Y are independent exponential random variables with the same parameter $\lambda > 0$. Find the density of $\min\{X, Y\}$.

$$Z = \min\{X, Y\} \quad f(x) = \lambda \cdot e^{-\lambda x}, x \geq 0$$

- $F_Z(z) = 1 - P(Z \geq z) =$
- $= 1 - P(\min\{X, Y\} \geq z) = 1 - P(X \geq z) \cdot P(Y \geq z) =$
- $= 1 - (1 - F_X(z))^2 = 1 - (1 - e^{-\lambda z})^2$
- $f_Z(z) = F'_Z(z) = 2(1 - e^{-\lambda z}) \cdot \lambda e^{-\lambda z}, z \geq 0$

7. Let X_1, X_2 be independent random variables each of which has a normal distribution $\mathcal{N}(\mu_1, \sigma_1^2)$ and $\mathcal{N}(\mu_2, \sigma_2^2)$, respectively. Define the random variable $Y := X_1 + X_2$. Using characteristic functions, show that Y also has a normal distribution $\mathcal{N}(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$. Hint: The characteristic function of a normal rv with mean μ and variance σ^2 has the form $\phi(t) = e^{i\mu t - \frac{1}{2}\sigma^2 t^2}$.

$$\begin{aligned} \cdot \quad \varphi_Y(t) &= \varphi_{X_1}(t) \cdot \varphi_{X_2}(t) = e^{i(\mu_1 + \mu_2)t - \frac{1}{2}(\sigma_1^2 + \sigma_2^2)t^2} \\ \cdot \quad \Rightarrow Y &\sim \mathcal{N}(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2) \end{aligned}$$