

ESE 520 Probability and Stochastic Processes

Lecture 6

"Random vectors and their probability distributions. Multidimensional cdf's and pdf's".

Now, we can extend what we said about a single random variable to a random vector.

◦ Formally, under a *random vector* we understand a map

$$X = (X_1, X_2, \dots, X_n) : \Omega \rightarrow \mathbb{R}^n,$$

where each $X_i : \Omega \rightarrow \mathbb{R}, i = 1, 2, \dots, n$ is random variable.

◦ An equivalent definition would be to say that X is a *measurable map* from (Ω, \mathcal{F}) to $(\mathbb{R}^n, \mathcal{B}(\mathbb{R}^n))$, that is for every Borel set $B \in \mathcal{B}(\mathbb{R}^n)$:

$$X^{-1}(B) \in \mathcal{F}.$$

◦ The probability measure

$$P_X(B) := P(w \in \Omega : X(w) \in B)$$

is then called the *probability distribution of X* .

However, similarly as in the 1-dimensional case ($n = 1$), it is then enough to know $P_X(B)$ only for a "specific collection" of sets $B \in \mathcal{B}(\mathbb{R}^n)$ leading to the introduction of the function

$$F_X(x_1, x_2, \dots, x_n) := P(w : X_1(w) \leq x_1, \dots, X_n(w) \leq x_n), \quad (1)$$

where $(x_1, x_2, \dots, x_n) \in \mathbb{R}^n$.

In other words, $F(x_1, x_2, \dots, x_n)$ is obtained if we restrict the distribution $P_X(\bullet)$ onto the collection of sets $\mathcal{A} = \{(-\infty, x_1] \times \dots \times (-\infty, x_n], x_i \in \mathbb{R}, i = 1, 2, \dots, n\}$ which generates the Borel σ -algebra $\mathcal{B}(\mathbb{R}^n)$.

The function in (1) is then called the *joint cumulative distribution function* (cdf) of the vector X and it completely describes the distribution of X . Again, that fact follows then from the Caratheodory Extension Theorem.

In summary, when working with a random vector X , we can obtain everything what we possibly can be interested in when working with the cdf $F_X(x_1, x_2, \dots, x_n)$ of X .

Given the cdf $F_X(x_1, x_2, \dots, x_n)$, we always can find the 1-dimensional (or, more generally, of any dimension less than n) cdfs $F_{X_i}(x_i), i = 1, 2, \dots, n$ (also called *marginal cdfs*) as

$$F_{X_i}(x_i) = P(X_i \leq x_i) = P(X_1 \in \mathbb{R}, \dots, X_i \leq x_i, \dots, X_n \in \mathbb{R}) = F_X(\infty, \dots, x_i, \dots, \infty).$$

In other words, the marginal cdfs can be always obtained from the joint cdf of the vector $X = (X_1, \dots, X_n)$.

When studying a random vector $X = (X_1, \dots, X_n)$, a particular important case is the case when random variables X_1, X_2, \dots, X_n are independent.

Definition 1. Random variables X_1 and X_2 are said to be *independent* if they assume their values independently. That is, if for all Borel sets B_1 and B_2 , it holds

$$P(X_1 \in B_1, X_2 \in B_2) = P(X_1 \in B_1)P(X_2 \in B_2). \quad (2)$$

Clearly, for $B_1 = (-\infty, x_1]$ and $B_2 = (-\infty, x_2]$, the relation becomes

$$F_X(x_1, x_2) = F_{X_1}(x_1)F_{X_2}(x_2). \quad (3)$$

The relations (2) and (3) can be then extended for any number $n = 2, 3, \dots$

Also, one can easily prove then the following statement:
random variables X_1, X_2, \dots, X_n are independent if and only if

$$F_X(x_1, x_2, \dots, x_n) = F_{X_1}(x_1) \dots F_{X_n}(x_n). \quad (4)$$

As in the case of $n = 1$, there is one particularly important class of vectors.

Definition 2. A random vector $X = (X_1, X_2, \dots, X_n)$ is said to be *jointly continuous* if there is a function $f : \mathbb{R}^n \rightarrow \mathbb{R}$ such that

i) $f(x) \geq 0$ for all $x \in \mathbb{R}^n$;

ii) $\int_{\mathbb{R}^n} f(x) dx = 1$;

and

$$P(w : X(w) \in B) = \int_B f(x) dx \quad (5)$$

for all Borel sets $B \in \mathcal{B}(\mathbb{R}^n)$.

In particular, (5) yields that

$$F_X(x_1, \dots, x_n) = \int_{-\infty}^{x_1} \dots \int_{-\infty}^{x_n} f(y_1, \dots, y_n) dy_n \dots dy_1. \quad (6)$$

The function $f(x_1, \dots, x_n)$ is then called *the joint pdf* of the vector X .

It follows that if the vector $X = (X_1, \dots, X_n)$ is jointly continuous, then each X_i is a continuous random variable.

Indeed: For a fixed $i = 1, 2, \dots, n$, we can write that

$$F_{X_i}(x_i) = F_X(\infty, \dots, x_i, \dots, \infty) = \int_{-\infty}^{\infty} \dots \int_{-\infty}^{x_i} \dots \int_{-\infty}^{\infty} f(y_1, \dots, y_n) dy_n \dots dy_1$$

and after differentiating the last relation

$$f_{X_i}(x_i) = F'_{X_i}(x_i) = \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f(y_1, \dots, x_i, \dots, y_n) dy_n \dots dy_{i+1} dy_{i-1} \dots dy_1. \quad (7)$$

It shows that $f_{X_i}(x_i)$ exists and can be found from the joint pdf of the vector by integrating the joint density in all other variables except x_i as (7) indicates.

In case of a jointly continuous vector the relation (4) becomes then:
random variables X_1, X_2, \dots, X_n are independent if and only if

$$f_X(x_1, x_2, \dots, x_n) = f_{X_1}(x_1) \dots f_{X_n}(x_n). \quad (8)$$

Finally, we provide an important example of a jointly continuous vector.

Definition 3. A random vector $X = (X_1, \dots, X_n)$ is said to have a (*non-degenerate*) *jointly continuous Gaussian distribution* if there exist a vector $\mu \in \mathbb{R}^n$ and a *positive definite* matrix R of size $n \times n$ such that the joint pdf of X is given by

$$f_X(x) = \frac{1}{\sqrt{(2\pi)^n \det R}} \exp\left\{-\frac{1}{2}(x - \mu)^T R^{-1}(x - \mu)\right\}. \quad (9)$$

where $x = (x_1, \dots, x_n) \in \mathbb{R}^n$.

Some comments to notations here:

- vectors $x \in \mathbb{R}^n$ are understood as column-vectors not row-vectors;
- R^{-1} stands for the inverse matrix of R which exists because of R being positive definite;
- R being positive definite means that $x^T R x > 0$ for all $x \in \mathbb{R}^n$ and $x \neq 0$;
- x^T stands for the transpose of x ;

A simple but important example of a jointly Gaussian vector is a vector obtained as follows:

Let $X_i, i = 1, 2, \dots, n$ be independent random variables each of which has $\mathcal{N}(\mu_i, \sigma_i^2)$ distribution.

Then the vector $X = (X_1, \dots, X_n)$ has the joint pdf

$$\begin{aligned} f_X(x) &= f_{X_1}(x_1) \dots f_{X_n}(x_n) = \\ &= \frac{1}{\sqrt{2\pi}\sigma_1} e^{-1/2(x_1 - \mu_1) \frac{1}{\sigma_1^2}(x_1 - \mu_1)} \dots \frac{1}{\sqrt{2\pi}\sigma_n} e^{-1/2(x_n - \mu_n) \frac{1}{\sigma_n^2}(x_n - \mu_n)} = \\ &= \frac{1}{\sqrt{(2\pi)^n \det R}} \exp\left\{-\frac{1}{2}(x - \mu)^T R^{-1}(x - \mu)\right\}, \end{aligned}$$

where $\mu = (\mu_1, \dots, \mu_n)$ and R is a diagonal matrix with the diagonal elements σ_i^2 .

Thus, a random vector $X = (X_1, \dots, X_n)$ whose components X_i are independent Gaussian random variables is a jointly continuous Gaussian vector.

It turns out that the converse is also true: if a jointly continuous random vector $X = (X_1, \dots, X_n)$ has the pdf of the form (9), where R is a diagonal matrix as described above, then the random variables X_1, \dots, X_n are independent Gaussian random variables.