

# ESE 520 Probability and Stochastic Processes

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## EXAM 2-Sample

Name (Please print):

Solutions

Instructions:

1. You must show all work to completely justify your answers in order to receive any credit.
2. You can use One one-sided sheet of paper with your own formulas.

1. a) Assume  $X$  has a normal distribution with mean  $\mu$  and variance  $\sigma^2$ . Find the moment-generating function of  $X$ .

$$\begin{aligned}
 & \bullet X \sim N(\mu, \sigma^2) \Rightarrow X = \sigma \cdot Y + \mu \text{ with } Y \sim N(0, 1) \\
 & \bullet \psi_Y(t) = E e^{ty} = \int_{\mathbb{R}} e^{ty} \frac{1}{\sqrt{2\pi}} e^{-y^2/2} dy = e^{t^2/2} \underbrace{\int_{\mathbb{R}} \frac{1}{\sqrt{2\pi}} e^{-\frac{(y-t)^2}{2}} dy}_{=1} \\
 & = e^{t^2/2}, \quad t \in \mathbb{R} \\
 & \bullet \psi_X(t) = E(e^{t(\mu + \sigma Y)}) = e^{\mu t} \cdot E e^{t\sigma Y} = e^{\mu t} e^{\frac{\sigma^2 t^2}{2}} \\
 & = e^{\mu t + \frac{\sigma^2 t^2}{2}}
 \end{aligned}$$

b) Use part a) to find  $E(X)$ ;

$$\begin{aligned}
 \psi'_X(t) &= (\mu + t\sigma^2) \cdot e^{\mu t + \frac{\sigma^2 t^2}{2}} \\
 \& \quad E(X) &= \psi'_X(0) = \mu.
 \end{aligned}$$

2. Let  $X$  and  $Y$  be two random variables having jointly continuous distribution with the joint pdf given as

$$f(x, y) = \begin{cases} 8xy, & 0 < y < x < 1 \\ 0, & \text{otherwise} \end{cases}$$

Find the best linear mean-square estimate of  $Y$  by  $X$ .

$$\bullet E(Y|X) = \mu_Y + \frac{\text{cov}(X, Y)}{\text{Var}(X)} \cdot (X - \mu_X),$$

$$\bullet f_X(x) = \int_0^x 8xy \, dy = 4x y^2 \Big|_0^x = 4x^3;$$

$$\bullet \mu_X = E(X) = \int_0^1 4x^4 \, dx = \frac{4}{5} x^5 \Big|_0^1 = \frac{4}{5};$$

$$\bullet f_Y(y) = \int_y^1 8xy \, dx = 4x^2 y \Big|_y^1 = 4(y - y^2);$$

$$\bullet \mu_Y = E(Y) = \int_0^1 4(y^2 - y^4) \, dy = 4 \left( \frac{y^3}{3} - \frac{y^5}{5} \right) \Big|_0^1 = \frac{8}{15};$$

$$\bullet E(X^2) = \int_0^1 4x^5 \, dx = \frac{2}{3};$$

$$\bullet \text{Var}(X) = E(X^2) - (E(X))^2 = \frac{2}{3} - \left(\frac{4}{5}\right)^2 = \frac{2}{75};$$

$$\begin{aligned} \bullet E(X \cdot Y) &= \int_0^1 \left( \int_0^x 8x^2 y^2 \, dy \right) dx = 8 \int_0^1 x^2 \frac{x^3}{3} \Big|_0^x dx \\ &= \frac{8}{3} \int_0^1 x^5 \, dx = \frac{8}{3 \cdot 6} = \frac{4}{9}. \end{aligned}$$

$$\bullet \text{cov}(X, Y) = E(XY) - E(X) \cdot E(Y) = \frac{4}{9} - \frac{4}{5} \cdot \frac{8}{15} = \frac{4 \cdot 75 - 9 \cdot 32}{9 \cdot 75}$$

$$\bullet \frac{\text{cov}(X, Y)}{\text{Var}(X)} = \frac{4 \cdot 75 - 9 \cdot 32}{9 \cdot 75} \cdot \frac{75}{2} = \frac{2 \cdot 75 - 9 \cdot 16}{9} = \frac{2}{3}$$

$$\bullet E(Y|X) = \frac{8}{15} + \frac{2}{3} \left( X - \frac{4}{5} \right)$$

3. Suppose  $X$  and  $Y$  have joint density  $e^{-x}$  for  $0 < y < x$ . Find  $E(X|Y)$  and  $\text{Var}(X|Y)$ .

$$\bullet f_Y(y) = \int_y^{\infty} e^{-x} dx = e^{-y}, \quad y > 0$$

$$\bullet E(X|Y=y) = \int_y^{\infty} x \cdot \frac{e^{-x}}{e^{-y}} dx = e^y \left( -x \cdot \frac{e^{-x}}{y} + \int_y^{\infty} e^{-x} dx \right) \\ = e^y \left( y \cdot e^{-y} - e^{-x} \right) \Big|_y^{\infty} = y + 1$$

$$\bullet E(X^2|Y=y) = \int_y^{\infty} x^2 \cdot e^y \cdot e^{-x} dx = e^y \left( -x^2 \cdot \frac{e^{-x}}{y} + \right. \\ \left. + 2 \int_y^{\infty} x \cdot e^{-x} dx \right) = e^y \left( y^2 \cdot e^{-y} + 2(y e^{-y} + e^{-y}) \right) \\ \text{calculated before} = y^2 + 2(y + 1)$$

$$\bullet \text{Var}(X|Y) = E(X^2|Y) - (E(X|Y))^2 = \\ = y^2 + 2(y + 1) - (y + 1)^2 = 1$$

4. Suppose you play roulette and bet \$1 on black each time. Your net winnings on the  $i$ th play,  $X_i$ , are 1 with probability  $1/3$  and -1 with probability  $2/3$ . What is the probability your net winnings are  $\geq 0$  after 100 plays?

Use the Central Limit Theorem and the enclosed normal table to find the approximate probability.

$$\bullet \text{ CLT: } \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \sim \mathcal{N}(0,1) \text{ apprx.}, \quad \bar{X}_n := \frac{X_1 + \dots + X_n}{n}$$

$$\text{with } E(X_i) = 1 \cdot \frac{1}{3} - 1 \cdot \frac{2}{3} = -\frac{1}{3} = \mu$$

$$\text{Var}(X_i) = E(X_i^2) - (E(X_i))^2 = 1 - \frac{1}{9} = \frac{8}{9} = \sigma^2$$

$$\bullet \quad P\left(\sum_{i=1}^{100} X_i \geq 0\right) = P\left(\bar{X}_{100} \geq 0\right) = P\left(\bar{X}_{100} - \mu \geq \frac{1}{3}\right)$$

$$= P\left(\frac{\bar{X}_{100} - \mu}{\sigma/\sqrt{100}} \geq \frac{1/3}{\sqrt{8/90}}\right) = P(Z \geq \frac{10}{\sqrt{8}})$$

$$\approx P(Z \geq 3.5) \approx 0$$

5. Suppose  $U, V, W$  are independent and have mean  $\mu$  and variance  $\sigma^2$ . Find  $\rho(U + V, V + W)$ .

$$\rho(U+V, V+W) = \frac{\text{cov}(U+V, V+W)}{\sqrt{\text{Var}(U+V) \cdot \text{Var}(V+W)}}$$

$$\begin{aligned} \text{cov}(U+V, V+W) &= E(U+V)(V+W) - E(U+V) \cdot E(V+W) \\ &= E(U \cdot V) + E(V^2) + E(U \cdot W) + E(V \cdot W) - \underbrace{[E(U) + E(V)]}_{\mu} \underbrace{[E(V) + E(W)]}_{\mu} \end{aligned}$$

= independence

$$= (E U)(E V) + E V^2 + (E U)(E W) + (E V)(E W) - 4\mu^2$$

$$= \mu^2 + \sigma^2 + \mu^2 + \mu^2 + \mu^2 - 4\mu^2 = \sigma^2$$

$$\text{Var}(U+V) = \text{Var}(U+W) = 2\sigma^2$$

$$\Rightarrow \rho = \frac{\sigma^2}{\sqrt{2\sigma^2 \cdot 2\sigma^2}} = \frac{1}{2}$$

6. Suppose  $X$  and  $Y$  are independent exponential random variables with the same parameter  $\lambda > 0$ . Find the density of  $\min\{X, Y\}$ .

$$Z = \min\{X, Y\}$$

$$f(x) = \lambda \cdot e^{-\lambda x}, x > 0$$

$$\bullet F_Z(z) = 1 - P(Z \geq z) =$$

$$= 1 - P(\min\{X, Y\} \geq z) = 1 - P(X \geq z) \cdot P(Y \geq z) =$$

$$= 1 - (1 - F_X(z))^2 = 1 - (1 - e^{-\lambda z})^2$$

$$\bullet f_Z(z) = F_Z'(z) = 2(1 - e^{-\lambda z}) \cdot \lambda e^{-\lambda z}, z > 0$$

7. Let  $X_1, X_2$  be independent random variables each of which has a normal distribution  $\mathcal{N}(\mu_1, \sigma_1^2)$  and  $\mathcal{N}(\mu_2, \sigma_2^2)$ , respectively. Define the random variable  $Y := X_1 + X_2$ . Using characteristic functions, show that  $Y$  also has a normal distribution  $\mathcal{N}(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$ . Hint: The characteristic function of a normal rv with mean  $\mu$  and variance  $\sigma^2$  has the form  $\phi(t) = e^{i\mu t - \frac{1}{2}\sigma^2 t^2}$ .

$$\cdot \quad \varphi_Y(t) = \varphi_{X_1}(t) \cdot \varphi_{X_2}(t) = e^{i(\mu_1 + \mu_2)t - \frac{1}{2}(\sigma_1^2 + \sigma_2^2)t^2}$$

$$\cdot \quad \Rightarrow Y \sim \mathcal{N}(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$$