

# ESE 520 Probability and Stochastic Processes

## Lecture 13

”Convergence of random variables. The law of large numbers.”

Given a sequence of random variables

$$X_n : (\Omega, \mathcal{F}, P) \rightarrow (\mathbb{R}, \mathcal{B}(\mathbb{R})), n = 1, 2, \dots,$$

it is natural to study the convergence of  $\{X_n\}, n = 1, 2, \dots$

There are several types of convergence we can introduce when talking about a sequence of random variables: *convergence in probability*, *convergence with probability one* (almost everywhere - a.e.), *convergence in distribution*, *convergence in mean or mean square* etc.

**Definition 1.** a) A sequence of random variables  $\{X_n\}, n = 1, 2, \dots$  converges to a random variable  $X$  *in probability* if for any  $\epsilon > 0$

$$\lim_{n \rightarrow \infty} P(w : |X_n(w) - X(w)| > \epsilon) = 0. \quad (1)$$

b) We say that a sequence of random variables  $\{X_n\}, n = 1, 2, \dots$  converges to  $X$  *almost everywhere* (a.e.) if there exist a set  $N \in \mathcal{F}$  with  $P(N) = 0$  so that for all  $w \notin N$ :

$$\lim_{n \rightarrow \infty} X_n(w) = X(w). \quad (2)$$

The condition (2) is then equivalent to

$$P(w : \lim_{n \rightarrow \infty} X_n(w) = X(w)) = 1.$$

Convergence almost everywhere is also called convergence with probability one.

Our particular interest here is the following: consider a sequence of i.i.d. (independent and identically distributed) random variables  $\{X_n\}, n = 1, 2, \dots$  with  $E(X_n) = \mu$  and  $Var(X_n) = \sigma^2$  for all  $n = 1, 2, \dots$

Define

$$\bar{X}_n := \frac{X_1 + X_2 + \dots + X_n}{n},$$

the moving average.

Question: What can be said about the *asymptotic behavior* of  $\bar{X}_n$ ? In other words, what is

$$\lim_{n \rightarrow \infty} \bar{X}_n?$$

(if it exists).

First, we notice that

$$E[\bar{X}_n] = E\left[\frac{1}{n} \sum_{i=1}^n X_i\right] = \frac{1}{n} \sum_{i=1}^n E(X_i) = \mu$$

and, using independence of random variables, we also get

$$Var[\bar{X}_n] = Var\left[\frac{1}{n} \sum_{i=1}^n X_i\right] = \frac{1}{n^2} \sum_{i=1}^n Var(X_i) = \frac{1}{n^2} n \sigma^2 = \frac{\sigma^2}{n}.$$

We see that

$$Var(\bar{X}_n) \rightarrow 0 \text{ as } n \rightarrow \infty$$

so that it is naturally to expect that

$$\bar{X}_n \rightarrow \mu \text{ as } n \rightarrow \infty.$$

Indeed, we will show that it holds:

$$\bar{X}_n \rightarrow \mu \text{ as } n \rightarrow \infty \text{ in probability} \quad (3)$$

(*weak law of large numbers: WLLN*)

and

$$\bar{X}_n \rightarrow \mu \text{ as } n \rightarrow \infty \text{ a.e.} \quad (4)$$

(*strong law of large numbers: SLLN*).

Statements (3) and (4) play an important role in probability theory and its applications. In particular, law of large numbers provide the basis of so-called Monte Carlo simulations and Monte Carlo computations of probabilities and expected values.

**Example 1.** (probability  $P(A)$  of an event  $A \subset \Omega$  is the limit of its relative frequency). Let  $A \in \mathcal{F}$  and define the random variable

$$X(w) = \begin{cases} 1, & w \in A \\ 0, & w \notin A. \end{cases}$$

Clearly,

$$\mu = E(X) = 1 \times P(A) + 0 \times P(A^c) = P(A).$$

Now, let  $X_n, n = 1, 2, \dots$  be the outcomes of independent repetitions of an experiment with  $A$  as "success" ("Bernoulli trial"). Then all  $X_n$  are independent random variables having the same distribution as  $X$ .

By the construction, we can interpret

$$\bar{X}_n = \frac{X_1 + X_2 + \dots + X_n}{n} = \frac{\text{number of successes in } n \text{ experiments}}{\text{number of all experiments}}$$

as the *relative frequency* of event  $A$ .

Strong LLN says that the relative frequency of an event  $A$  in independent repetitions of the experiment converges to the probability of  $A$  as  $n \rightarrow \infty$  a.e.

Let us prove the weak LLN and the strong LLN.

**Proposition.** (*Chebyshev's inequality*). Assume that  $EX^2 < \infty$ . Then, for any  $\epsilon > 0$ :

$$P(|X| \geq \epsilon) \leq \frac{EX^2}{\epsilon^2}.$$

*Proof.* (for continuous case; in a discrete case - similar).

$$EX^2 = \int_{-\infty}^{\infty} x^2 f_X(x) dx \geq \int_{\{|x| \geq \epsilon\}} x^2 f_X(x) dx \geq$$

$$\int_{\{|x| \geq \epsilon\}} \epsilon^2 f_X(x) dx = \epsilon^2 \int_{\{|x| \geq \epsilon\}} f_X(x) dx = \epsilon^2 P(|X| \geq \epsilon)$$

implying the statement.

Using Chebyshev's inequality, one has for any  $\epsilon > 0$ :

$$P(|\bar{X}_n - \mu| > \epsilon) \leq \frac{1}{\epsilon^2} E(\bar{X}_n - \mu)^2 =$$

$$\frac{1}{\epsilon^2} \text{Var}(\bar{X}_n) = \frac{1}{\epsilon^2} \frac{\sigma^2}{n} \rightarrow 0 \text{ as } n \rightarrow \infty$$

so that the weak LLN (3) is proven.

The proof of strong LLN (4) is more difficult and relies on the use of some inequalities of independent interest.

**Lemma** (Toeplitz): For  $a_n \geq 0$ , define  $b_n := \sum_{i=1}^n a_i$  and assume that  $\lim_{n \rightarrow \infty} b_n = +\infty$  ( $b_n$  diverges). Then, if  $x_n \rightarrow x$  as  $n \rightarrow \infty$ , it also holds that

$$\frac{1}{b_n} \sum_{i=1}^n a_i x_i \rightarrow x \text{ as } n \rightarrow \infty.$$

*Proof.* Let  $\epsilon > 0$ . We can choose an integer  $N_0 \in \mathbb{N}$  such that  $|x_n - x| \leq \frac{\epsilon}{2}$  for all  $n \geq N_0$ . We also can choose  $N_1 \geq N_0$  so that

$$\frac{1}{b_{N_1}} \sum_{i=1}^{N_0} a_i |x_i - x| \leq \frac{\epsilon}{2}.$$

As the result, for all  $n \geq N_1$ , it holds:

$$\begin{aligned} \left| \frac{1}{b_n} \sum_{i=1}^n a_i x_i - x \right| &= \left| \frac{1}{b_n} \sum_{i=1}^n a_i (x_i - x) \right| \leq \\ &\frac{1}{b_n} \sum_{i=1}^{N_0} a_i |x_i - x| + \frac{1}{b_n} \sum_{i=N_0+1}^n a_i |x_i - x| \leq \frac{\epsilon}{2} + \frac{\epsilon}{2} = \epsilon \end{aligned}$$

since  $b_{N_1} \leq b_n$  and  $\frac{1}{b_n} \sum_{i=N_0+1}^n a_i \leq 1$ .  $\square$

**Corollary.** If  $x_n \rightarrow x$  as  $n \rightarrow \infty$ , then  $\frac{1}{n} \sum_{i=1}^n x_i \rightarrow x$  as  $n \rightarrow \infty$ .  
(Choose  $a_n = 1$ ,  $b_n = n$  and apply the Lemma above).

**Lemma** (Kronecker). Assume  $b_n > 0$  so that  $b_n \uparrow \infty$  as  $n \rightarrow \infty$  and  $\sum_{i=1}^{\infty} x_i < \infty$ . Then

$$\frac{1}{b_n} \sum_{i=1}^n b_i x_i \rightarrow 0 \text{ as } n \rightarrow \infty.$$

*Proof.* Let

$$S_n := \sum_{i=1}^n x_i \text{ with } S_0 = 0, b_0 = 0.$$

Summation by parts yields:

$$\begin{aligned}\sum_{i=1}^n b_i x_i &= \sum_{i=1}^n b_i (S_i - S_{i-1}) = b_n S_n - S_0 b_0 - \sum_{i=1}^n (b_i - b_{i-1}) S_{i-1} = \\ &= b_n S_n - \sum_{i=1}^n (b_i - b_{i-1}) S_{i-1}.\end{aligned}$$

Dividing both sides by  $b_n$ , we obtain

$$\frac{1}{b_n} \sum_{i=1}^n b_i x_i = S_n - \frac{1}{b_n} \sum_{i=1}^n a_i S_{i-1} \quad (5)$$

where  $a_i := b_i - b_{i-1}$  so that  $\sum_{i=1}^n a_i = b_n \uparrow \infty$  as  $n \rightarrow \infty$ .

By assumption, the sequence  $\{S_n\}$  converges, say to  $S$  and, by Toeplitz's Lemma, the sequence  $\frac{1}{b_n} \sum_{i=1}^n a_i S_{i-1}$  converges to the same limit  $S$  so that the right side in (5) goes to zero as  $n \rightarrow \infty$ .  $\square$

**Corollary.** Let  $\sum_{n=1}^{\infty} \frac{y_n}{n} < \infty$ . Then

$$\frac{1}{n} \sum_{i=1}^n y_i \rightarrow 0 \text{ as } n \rightarrow \infty.$$

(Choose  $b_n = n, x_n = \frac{y_n}{n}$ )

Now, since

$$\frac{1}{n} \sum_{i=1}^n X_i - \mu = \frac{1}{n} \sum_{i=1}^n (X_i - \mu),$$

the strong LLN will be proven if we are able to show (by Kronecker's Lemma) that the series

$$\sum_{n=1}^{\infty} \left( \frac{X_n - \mu}{n} \right)$$

converges a.e.

Set

$$S_k := \sum_{n=1}^k \left( \frac{X_n - \mu}{n} \right).$$

Then we have to show that the sequence  $\{S_k\}$  converges a.e.

Let  $N := \{w : S_k(w) \text{ does not converge}\}$ .

For all  $w \notin N$ ,  $S_k(w)$  is a Cauchy sequence. That is, for any  $\epsilon > 0$  there is  $\tilde{N} \in \mathbb{N}$  so that for all  $m \geq \tilde{N}$  it holds  $|S_m(w) - S_{\tilde{N}}(w)| \leq \epsilon$ .

Using the negation argument, we can write that:

For all  $w \in N$  there exists a number  $\epsilon_0 > 0$  so that for all  $\tilde{N} \in \mathbb{N}$  there exist  $m \geq \tilde{N}$  so that  $|S_m(w) - S_{\tilde{N}}(w)| > \epsilon_0$ .

Without loss of generality, we can always choose  $\epsilon_0 = \frac{1}{r}$  for some  $r \in \mathbb{N}$ .

Using the above, we can write  $N$  as

$$N = \cup_{r=1}^{\infty} \cap_{\tilde{N}=1}^{\infty} \cup_{m=\tilde{N}}^{\infty} \{w : |S_m(w) - S_{\tilde{N}}(w)| > \frac{1}{r}\}. \quad (6)$$

We note that the representation (6) shows that  $N \in \mathcal{F}$  so that  $P(N)$  exists.

We can write

$$N = \cup_{r=1}^{\infty} N_r.$$

Since

$$P(N) = P(\cup_{r=1}^{\infty} N_r) \leq \sum_{r=1}^{\infty} P(N_r),$$

it is enough to show that

$$P(N_r) = 0 \text{ for all } r \in \mathbb{N}.$$

Fix  $r$ . Then

$$N_r \subset \cup_{m=\tilde{N}}^{\infty} \{w : |S_m(w) - S_{\tilde{N}}(w)| > \frac{1}{r}\}$$

for all  $\tilde{N} = 1, 2, \dots$

Therefore,

$$P(N_r) \leq P(\cup_{m=\tilde{N}}^{\infty} \{w : |S_m(w) - S_{\tilde{N}}(w)| > \frac{1}{r}\}) =$$

$$P(\sup_{m \geq \tilde{N}} \{w : |S_m(w) - S_{\tilde{N}}(w)| > \frac{1}{r}\})$$

for all  $\tilde{N} = 1, 2, \dots$  so that

$$P(N_r) \leq \lim_{\tilde{N} \rightarrow \infty} P(\sup_{m \geq \tilde{N}} \{w : |S_m(w) - S_{\tilde{N}}(w)| > \frac{1}{r}\}). \quad (7)$$

To finish the proof of the strong LLN, we have to show that the right side of (7) is equal to zero. For that, we need one more inequality.