

Question 4

(b).

From the result of (a). we know that for optimization problem

$$\min_{\beta} E[q(Y - \beta)^+ + (1 - q)(Y - \beta)^-] \quad (1)$$

The optimal solution ( $\beta^*$ ) is the  $q$ -th quantile  $y_q$  of the random variable  $Y$  as any point that satisfies the equation  $F(y_q) = q$ , where  $F(y)$  is the CDF of random variable  $Y$ . According to the definition of CDF,  $F(y) = \text{Prob}(Y \leq y)$ .

For question (b), we are looking for the optimal solution of optimization problem

$$\min_{\{\beta(x): R^d \rightarrow R\}} E[q(Y - \beta(X))^+ + (1 - q)(Y - \beta)^-] \quad (2)$$

As  $\beta(x)$  is a function that map  $x$  from  $R^d$  to  $R$ , and  $X \in R^d$ , expression (1) and (2) are equivalent. While the  $q$ -th quantile  $y_q$  of the random variable  $Y$  is the optimal solution for (1), given condition  $X = x$ , the conditional  $q$ -quantile  $\beta^*(x)$  of random variable  $Y$  is the optimal solution for expression (2).

Therefore,  $q$ -quantile  $\beta^*(x)$  assess how much  $y$  will change for distribution  $y$  as  $x$  change by 1 unit at quantile point  $q$  for a given set of other covariates.

(c).

The goal is to find  $w$  that  $\min_{w \in R^{d+1}} \|y - Mw\|_1 + (2q - 1)1^T(y - Mw)$

For  $\|y - Mw\|_1$ , if  $y > Mw$ ,  $\|y - Mw\|_1 = Y - Mw$   
if  $y < Mw$ ,  $\|y - Mw\|_1 = Mw - Y$

Therefore,

$$\begin{aligned} & \min_{w \in R^{d+1}} \|y - Mw\|_1 + (2q - 1)1^T(y - Mw) \\ &= \min \sum_{i \in y_i > M_i w_i}^N y_i - M_i w_i + (2q - 1)(y_i - M_i w_i) + \sum_{i \in y_i < M_i w_i}^N M_i w_i - y_i + (2q - 1)(y_i - M_i w_i) \\ &= \min \sum_{i \in y_i > M_i w_i}^N y_i - M_i w_i + (2q)(y_i - M_i w_i) - (1)(y_i - M_i w_i) + \sum_{i \in y_i < M_i w_i}^N -(y_i - M_i w_i) + (2q)(y_i - M_i w_i) - (1)(y_i - M_i w_i) \\ &= \min \sum_{i \in y_i > M_i w_i}^N (2q)(y_i - M_i w_i) + \sum_{i \in y_i < M_i w_i}^N (2q - 2)(y_i - M_i w_i) \end{aligned}$$

According to what is defined in the question,

$\beta(x)$  is restricted to be of the form  $\beta(x) = [X_i^T 1] * w_i$ , and  $M_i$  is  $[X_i^T 1]$

We could rewrite the expression

$$\begin{aligned}
& \min \sum_{i \in y_i > M_i w_i}^N (2q)(y_i - M_i w_i) + \sum_{i \in y_i < M_i w_i}^N (2q - 2)(y_i - M_i w_i) \\
&= \min \sum_{i \in y_i > M_i w_i}^N (2q)(y_i - \beta(x_i)) + \sum_{i \in y_i < M_i w_i}^N (2q - 2)(y_i - \beta(x_i)) \\
&= \min \sum_{i \in y_i > M_i w_i}^N q(y_i - \beta(x_i)) + \sum_{i \in y_i < M_i w_i}^N (q - 1)(y_i - \beta(x_i)) \\
&= \min \sum_{i \in y_i > M_i w_i}^N q(y_i - \beta(x_i)) + \sum_{i \in y_i < M_i w_i}^N (1 - q)(-(y_i - \beta(x_i)))
\end{aligned}$$

Define  $(x)^+ := \max\{x, 0\}$ , when  $x < 0$ ,  $(x)^+ = 0$ , and when  $x > 0$ ,  $(x)^+ = x$

So

$$\begin{aligned}
& \sum_{i \in y_i > M_i w_i}^N q(y_i - \beta(x_i)) \\
&= q \sum_{i \in y_i < M_i w_i}^N 0 + \sum_{i \in y_i > M_i w_i}^N (y_i - \beta(x_i)) \\
&= q \sum_i^N (y_i - \beta(x_i))^+
\end{aligned}$$

Define  $(x)^- := \max\{-x, 0\}$ , when  $x < 0$ ,  $(x)^- = -x$ , and when  $x > 0$ ,  $(x)^- = 0$

So

$$\begin{aligned}
& \sum_{i \in y_i < M_i w_i}^N (1 - q)(-(y_i - \beta(x_i))) \\
&= (1 - q) \sum_{i \in y_i < M_i w_i}^N -(y_i - \beta(x_i)) + \sum_{i \in y_i > M_i w_i}^N 0 \\
&= (1 - q) \sum_i^N (y_i - \beta(x_i))^-
\end{aligned}$$

Therefore

$$\begin{aligned}
& \min \sum_{i \in y_i > M_i w_i}^N q(y_i - \beta(x_i)) + \sum_{i \in y_i < M_i w_i}^N (1-q)(-(y_i - \beta(x_i))) \\
&= \min q \sum_i^N (y_i - \beta(x_i))^+ + (1-q) \sum_i^N (y_i - \beta(x_i))^- \\
&= \min \sum_i^N [q(y_i - \beta(x_i))^+ + (1-q)(y_i - \beta(x_i))^-] \\
&= \min E[q(Y - \beta(X))^+ + (1-q)(Y - \beta(X))^-]
\end{aligned}$$

Therefore, we prove that  $\min E[q(Y - \beta(X))^+ + (1-q)(Y - \beta(X))^-]$

$$\approx \min_{w \in R^{d+1}} \lVert y - Mw \rVert_1 + (2q-1)1^T(y - Mw)$$