# ECON220B Discussion Section 5 Selection on Observables

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#### Roadmap

1. From RCT to Selection on Observables

2. Inverse Probability Weighting

3. Regression Adjustment Estimator

4. Propensity Score and Heavy-Tailed Distributions

#### Potential Outcome Framework

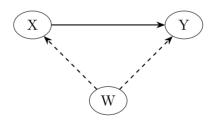
- Treatment:  $x_i = 1\{treated\}$
- Observed outcome:  $y_i = x_i y_i(1) + (1 x_i) y_i(0)$
- RCT assumption:  $x_i \perp (y_i(1), y_i(0))$

RCT: randomization equalizes everything other than the treatment in the treatment and control group. Fine for randomized experiment, what about observational studies?

#### Cancer & Smoking

"Considerable propaganda is now being developed to convince the public that cigarette smoking is dangerous." - Sir. Ronald Fisher (1958)

What was his argument?



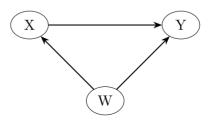
"[...] run the risk of failing to recognize, and therefore failing to prevent, other and more genuine causes"

#### Different Identification Strategies

Starting point is conditional independence, or uncounfoundedness, or selection on observables:

$$x_i \perp (y_i(1), y_i(0))|w_i$$

i.e. all confounders have already been identified and accounted for by the set of covariates. Then two possible ways to estimate  $\tau_{ATE} \equiv E[y_i(1) - y_i(0)]$ 



## Inverse Probability Weighting (1/2)

- We model how the treatment take-up decision  $x_i$  is related with the covariates  $w_i$ , re-weighting each observation by the likelihood of receiving the treatment.
- Overlap condition:  $0 < P(x_i = 1 | w_i = \mathbf{w}) < 1$ , i.e. for a particular characteristic  $w_i = \mathbf{w}$  if we observe some treated unit, then we should be able to observe some untreated unit as well.
- $E\left[\frac{x_iy_i}{e(w_i)}\right] =$

#### Inverse Probability Weighting (2/2)

• The propensity score  $e(w_i) \equiv P(x_i = 1 | w_i = \mathbf{w})$  is a balancing score: after conditioning on the propensity score, the distribution of the treatment is the same for treated and untreated:

$$x_i \perp w_i | e(w_i)$$

•  $e(w_i)$  is all you need to know: sufficient statistic for  $x_i$ .

#### More About Overlap Condition

Note: 
$$0 < P(x_i = 1 | w_i = \mathbf{w}) < 1 \iff f_{w|x_i=1}(\mathbf{w}) > 0, f_{w|x_i=0}(\mathbf{w}) > 0$$
  
Proof

#### Regression Adjustment Estimator

Do we really need overlap condition? No  $\rightarrow$  linearity assumption: interaction effect of covariates and treatment

$$\tau_{ATE} = E[y_i(1) - y_i(0)] = E[y_i(1)] - E[y_i(0)]$$

$$\tau_{ATE} = E[E[y_i(1)|w_i] - E[y_i(0)|w_i]]$$

$$\tau_{ATE} = E[E[y_i(1)|w_i, x_i = 1] - E[y_i(0)|w_i, x_i = 0]]$$

$$\tau_{ATE} = E[E[y_i|w_i, x_i = 1] - E[y_i|w_i, x_i = 0]]$$

$$\tau_{ATE} = E[g_1(w_i) - g_0(w_i)] = E[w_i^T \delta_1 - w_i^T \delta_0] = \mu_w^T (\delta_1 - \delta_0)$$

Consider the inverse probability weighting estimator:

$$\hat{\theta} = \frac{1}{n} \sum_{i=1}^{n} \frac{x_i y_i}{e(w_i)}$$

where  $x_i$  is the binary indicator of treatment status,  $y_i$  is the outcome variable, and  $w_i$  represents the covariates. For simplicity, we assume that the propensity score,  $e(wi) = P[x_i = 1|w_i]$ , is known. In addition, assume  $y_i$  is bounded

#### Heavy-Tailed Distributions

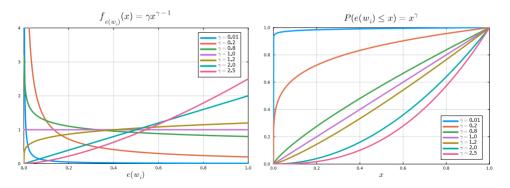
#### Some useful definitions:

- A distribution is heavy-tailed if  $E[tX] = \infty \ \forall t > 0$
- A distribution is light-tailed if it is not heavy-tailed.
- A distribution is light-tailed if  $E[X^k] < \infty \ \forall t > 0$

**Q.1** - Assume strong overlap, show that 
$$E\left[\left|\frac{x_iy_i}{e(w_i)}\right|^2\right] < \infty$$
 then conclude  $\operatorname{Var}(\hat{\theta}) < \infty$ .

**Q.2** - Now assume the propensity score can be arbitrarily close to zero  $P(e(w_i) \leq \delta) = \delta^{\gamma}$ . What will happen if  $\gamma$  is small?

**Q.3** - Take  $\gamma = 2$ . Plot the density function of the propensity score. Is this sufficient to show that  $E\left[\left|\frac{x_iy_i}{e(w_i)}\right|^2\right] < \infty$ ?



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**Q.4** - Discuss what will happen if  $\gamma < 1$ .