

UNIVERSITY OF LONDON

[II(4)E 2000]

B.ENG. AND M.ENG. EXAMINATIONS 2000

For Internal Students of the Imperial College of Science, Technology and Medicine

This paper is also taken for the relevant examination for the Associateship.

PART II : MATHEMATICS 4 (ELECTRICAL ENGINEERING)

Thursday 8th June 2000 2.00 - 4.00 pm

Answer FOUR questions.

[Before starting, please make sure that the paper is complete; there should be 4 pages, with a total of 6 questions. Ask the invigilator for a replacement if your copy is faulty.]

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1. Find the eigenvalues of

$$A = \begin{pmatrix} 3 & -3 & 2 \\ -1 & 5 & -2 \\ -1 & 3 & 0 \end{pmatrix}.$$

Determine the general form of the eigenvector corresponding to the repeated eigenvalue of A .

Obtain three independent eigenvectors and, hence, determine a matrix C such that

$$C^{-1}AC = D$$

where D is a diagonal matrix.

Show that D is the spectral matrix of A , i.e. the diagonal elements of D are the eigenvalues of A .

2. Reduce the quadratic form

$$Q = \mathbf{x}^T A \mathbf{x} = 2x_1^2 + 5x_2^2 + 3x_3^2 + 4x_1x_2$$

to the form

$$Q = \mathbf{y}^T D \mathbf{y} = \lambda_1 y_1^2 + \lambda_2 y_2^2 + \lambda_3 y_3^2$$

where $\mathbf{x} = (x_1, x_2, x_3)^T$, $\mathbf{y} = (y_1, y_2, y_3)^T$, $\mathbf{x} = P \mathbf{y}$ and $D = P^T A P$.

The scalars $\lambda_1, \lambda_2, \lambda_3$ and the matrices P and D are to be determined.

Hence, find y_1, y_2 and y_3 in terms of x_1, x_2 and x_3 .

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3. (i) For any two events A and B which are subsets of a sample space Ω , use a Venn diagram to show that

$$A = (A \cap B) \cup (A \cap \bar{B}) \quad \text{and} \quad B = (A \cap B) \cup (\bar{A} \cap B).$$

Hence show that for events A, B the probability of *one and only one* of them occurring is

$$P(A) + P(B) - 2P(A \cap B).$$

- (ii) If A and B are independent events, prove that \bar{A} and \bar{B} are independent events, and also that \bar{A} and B are independent events.

[Hint: You may use the fact that $\bar{A} \cap \bar{B} = \overline{A \cup B}$.]

- (iii) The probability that generating equipment will still be in use in 10 years time is $1/4$ at power plant 1 and $1/3$ at plant 2. Given that the equipment at the two sites behaves independently, find the probability that in 10 years time

- (a) equipment at both plants will be functioning,
- (b) equipment at at least one plant will be functioning,
- (c) equipment at neither plant will be functioning,
- (d) only the equipment at plant 2 will be functioning,
- (e) equipment at plant 2 will be functioning, given that equipment at plant 1 is no longer functioning.

4. Printed circuit boards (PCBs) leave a production line randomly and independently and each is routinely subject to a quality check. It has been found over a long period of time that a proportion q of PCBs will fail the quality check. Let X be the number of PCBs tested up to and including the first rejection, and let $p = 1 - q$.

- (i) Show that $P(X = j) = p^{j-1}q$, $j = 1, 2, \dots$, i.e., X has a geometric distribution, and find $P(X > j)$.
- (ii) Demonstrate that X has the lack-of-memory property by showing that for j, k positive integers,

$$P(X > j + k | X > j) = P(X > k).$$

Why should it be no surprise that the geometric distribution enjoys the lack-of-memory property?

- (iii) Let Y be the total number of PCBs tested up to and including the second rejection. Determine the joint probability distribution $P(X = j, Y = k)$ and the marginal probability distribution $P(Y = k)$, where $k > j \geq 1$.

5. Let the positive random variable T represent the lifetime of an electrical system. Carefully define the hazard rate $z(t)$ of the system.

A system consists of two components connected in parallel so that $T = \max \{T_1, T_2\}$ where T_1 and T_2 are independent lifetimes of component 1 and component 2, respectively, both having exponential failure time distributions, i.e.,

$$f_{T_1}(t) = \lambda_1 e^{-\lambda_1 t}, \quad t \geq 0, \quad \text{and} \quad f_{T_2}(t) = \lambda_2 e^{-\lambda_2 t}, \quad t \geq 0, \quad \text{with} \quad \lambda_1, \lambda_2 > 0.$$

- (i) Find the cumulative distribution function, $F_T(t)$, of T .
(ii) Show that the probability density function $f_T(t)$ of T is given by

$$f_T(t) = \lambda_1 e^{-\lambda_1 t} + \lambda_2 e^{-\lambda_2 t} - (\lambda_1 + \lambda_2) e^{-(\lambda_1 + \lambda_2)t}, \quad t \geq 0,$$

and demonstrate that it integrates to unity.

- (iii) When $\lambda_1 = \lambda_2 = \lambda > 0$, show that the hazard rate is given by

$$z(t) = \frac{2\lambda}{\frac{1}{1-e^{-\lambda t}} + 1},$$

and comment on its form as $t \rightarrow \infty$.

6. Given two random variables X and Y with joint probability density function $f_{X,Y}(x, y)$, the minimum mean square error estimate of the unobserved value $Y = y$ in terms of the observed value $X = x$ is given by $E\{Y | X = x\}$.

- (i) Carefully explain the meaning of the above statement.

The random variables X and Y which represent the amplitudes of two signals have joint probability density function

$$f_{X,Y}(x, y) = \begin{cases} x^{-1}, & 0 \leq y \leq x \leq 1, \\ 0, & \text{otherwise.} \end{cases}$$

- (ii) Show that the double integral of $f_{X,Y}(x, y)$ over the specified ranges of x and y takes the value unity, as required for such a joint probability density function.
(iii) Find $f_{Y|X=x}(y|x)$, the conditional probability density function of Y given $X = x$. What type of continuous distribution is this ?
(iv) Hence derive the minimum mean square error estimate of $Y = y$ in terms of $X = x$.

END OF PAPER

MATHS 4
2000

MATHEMATICS FOR ENGINEERING STUDENTS
EXAMINATION QUESTION / SOLUTION
SESSION : 1999/2000

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PAPER

II 4 EE

QUESTION

SOLUTION

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$$\begin{vmatrix} 3-\lambda & -3 & 2 \\ -1 & 5-\lambda & -2 \\ -1 & 3 & -\lambda \end{vmatrix} \begin{matrix} R_1 = R_1 + R_2 \\ R_3 = R_3 - R_2 \end{matrix} \rightarrow \begin{vmatrix} 2-\lambda & 2-\lambda & 0 \\ -1 & 5-\lambda & -2 \\ 0 & \lambda-2 & 2-\lambda \end{vmatrix}$$

$$= (2-\lambda)^2 \begin{vmatrix} 1 & 1 & 0 \\ -1 & 5-\lambda-2 & -2 \\ 0 & -1 & 1 \end{vmatrix} = (2-\lambda)^2 (5-\lambda-2+1) = (2-\lambda)^2 (4-\lambda)$$

\therefore E-values : 2, 2, 4.

$$\lambda = 4 \quad \begin{pmatrix} -1 & -3 & 2 \\ -1 & 1 & -2 \\ -1 & 3 & -4 \end{pmatrix} \begin{pmatrix} x \\ y \\ z \end{pmatrix} = 0 \rightarrow \begin{pmatrix} -1 & -3 & 2 \\ 0 & 4 & -4 \\ 0 & 6 & -6 \end{pmatrix} \begin{pmatrix} x \\ y \\ z \end{pmatrix} = 0$$

$$\therefore e_4 = \begin{pmatrix} -1 \\ 1 \\ 1 \end{pmatrix}$$

$$\lambda = 2 \quad \begin{pmatrix} 1 & -3 & 2 \\ -1 & 3 & -2 \\ -1 & 3 & -2 \end{pmatrix} \begin{pmatrix} x \\ y \\ z \end{pmatrix} = 0 \rightarrow x - 3y + 2z = 0$$

$$\therefore e_2 = \begin{pmatrix} 3\alpha - 2\beta \\ \alpha \\ \beta \end{pmatrix} = \alpha \begin{pmatrix} 3 \\ 1 \\ 0 \end{pmatrix} + \beta \begin{pmatrix} -2 \\ 0 \\ 1 \end{pmatrix}$$

$$\therefore e_4 = \begin{pmatrix} -1 \\ 1 \\ 1 \end{pmatrix}, e_{21} = \begin{pmatrix} 3 \\ 1 \\ 0 \end{pmatrix}, e_{22} = \begin{pmatrix} -2 \\ 0 \\ 1 \end{pmatrix}$$

$$\text{and } C = \begin{pmatrix} -1 & 3 & -2 \\ 1 & 1 & 0 \\ 1 & 0 & 1 \end{pmatrix}$$

To find C^{-1} :

$$\begin{array}{ccc|ccc} -1 & 3 & -2 & 1 & 0 & 0 & R_1 = -R_1 & +1 & -3 & +2 & -1 & 0 & 0 \\ 1 & 1 & 0 & 0 & 1 & 0 & R_2 = R_2 + R_1 & 0 & 4 & -2 & 1 & 1 & 0 \\ 1 & 0 & 1 & 0 & 0 & 1 & R_3 = R_3 + R_1 & 0 & 3 & -1 & 1 & 0 & 1 \end{array}$$

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$$\begin{aligned} R_1 &= R_1 + R_3 \\ R_3 &= R_3 - \frac{3}{4}R_2 \end{aligned} \Rightarrow \begin{array}{ccc|ccc} 1 & 0 & 1 & 0 & 0 & 1 \\ 0 & 1 & -\frac{1}{2} & \frac{1}{4} & \frac{1}{4} & 0 \\ 0 & 0 & \frac{1}{2} & \frac{1}{4} & -\frac{3}{4} & 1 \end{array} \quad \begin{aligned} R_1 &= R_1 - 2R_3 \\ R_2 &= R_2 + R_3 \\ R_3 &= 2R_3 \end{aligned}$$

$$\Rightarrow \begin{array}{ccc|ccc} 1 & 0 & 0 & -\frac{1}{2} & \frac{3}{2} & -1 \\ 0 & 1 & 0 & \frac{1}{2} & -\frac{1}{2} & 1 \\ 0 & 0 & 1 & \frac{1}{2} & -\frac{3}{2} & 2 \end{array} \quad \therefore C^{-1} = \frac{1}{2} \begin{pmatrix} -1 & 3 & -2 \\ 1 & -1 & 2 \\ 1 & -3 & 4 \end{pmatrix}$$

$$\therefore C^{-1}AC = \frac{1}{2} \begin{pmatrix} -1 & 3 & -2 \\ 1 & -1 & 2 \\ 1 & -3 & 4 \end{pmatrix} \begin{pmatrix} 3 & -3 & 2 \\ -1 & 5 & -2 \\ -1 & 3 & 0 \end{pmatrix} \begin{pmatrix} -1 & 3 & -2 \\ 1 & 1 & 0 \\ 1 & 0 & 1 \end{pmatrix}$$

$$= \frac{1}{2} \begin{pmatrix} -1 & 3 & -2 \\ 1 & -1 & 2 \\ 1 & -3 & 4 \end{pmatrix} \begin{pmatrix} -4 & 6 & -4 \\ 4 & 2 & 0 \\ 4 & 0 & 2 \end{pmatrix}$$

$$= \begin{pmatrix} -1 & 3 & -2 \\ 1 & -1 & 2 \\ 1 & -3 & 4 \end{pmatrix} \begin{pmatrix} -2 & 3 & -2 \\ 2 & 1 & 0 \\ 2 & 0 & 1 \end{pmatrix}$$

$$= \begin{pmatrix} 4 & 0 & 0 \\ 0 & 2 & 0 \\ 0 & 0 & 2 \end{pmatrix}$$

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$$Q = (x_1, x_2, x_3) \begin{pmatrix} 2 & 2 & 0 \\ 2 & 5 & 0 \\ 0 & 0 & 3 \end{pmatrix} \begin{pmatrix} x_1 \\ x_2 \\ x_3 \end{pmatrix}$$

$$\text{Eigenvalues of } A : \begin{vmatrix} 2-\lambda & 2 & 0 \\ 2 & 5-\lambda & 0 \\ 0 & 0 & 3-\lambda \end{vmatrix} = 0$$

$$\therefore (3-\lambda)((2-\lambda)(5-\lambda)-4) = (3-\lambda)(\lambda^2-7\lambda+6) = 0$$

$$\text{i.e. } (3-\lambda)(\lambda-6)(\lambda-1) = 0 \quad \therefore \text{E-values : } 1, 3, 6.$$

$$\text{Eigenvectors : } \lambda=1 \quad \begin{pmatrix} 1 & 2 & 0 \\ 2 & 4 & 0 \\ 0 & 0 & 2 \end{pmatrix} \begin{pmatrix} x \\ y \\ z \end{pmatrix} = 0 \quad \therefore \underline{e}_1 = \begin{pmatrix} 2 \\ -1 \\ 0 \end{pmatrix}$$

$$\lambda=3 : \begin{pmatrix} -1 & 2 & 0 \\ 2 & 2 & 0 \\ 0 & 0 & 0 \end{pmatrix} \begin{pmatrix} x \\ y \\ z \end{pmatrix} = 0 \quad \therefore \underline{e}_3 = \begin{pmatrix} 0 \\ 0 \\ 1 \end{pmatrix}$$

$$\lambda=6 : \begin{pmatrix} -4 & 2 & 0 \\ 2 & -1 & 0 \\ 0 & 0 & -3 \end{pmatrix} \begin{pmatrix} x \\ y \\ z \end{pmatrix} = 0 \quad \therefore \underline{e}_6 = \begin{pmatrix} 1 \\ 2 \\ 0 \end{pmatrix}$$

$$\therefore \text{Normalised e-vectors are } \begin{pmatrix} \frac{2}{\sqrt{5}} \\ -\frac{1}{\sqrt{5}} \\ 0 \end{pmatrix}, \begin{pmatrix} 0 \\ 0 \\ 1 \end{pmatrix}, \begin{pmatrix} \frac{1}{\sqrt{5}} \\ \frac{2}{\sqrt{5}} \\ 0 \end{pmatrix}$$

$$\therefore P = \begin{pmatrix} \frac{2}{\sqrt{5}} & 0 & \frac{1}{\sqrt{5}} \\ -\frac{1}{\sqrt{5}} & 0 & \frac{2}{\sqrt{5}} \\ 0 & 1 & 0 \end{pmatrix}$$

$$\underline{x} = P\underline{y} \quad \therefore Q = \underline{x}^T A \underline{x} = \underline{y}^T P^T A P \underline{y}$$

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$$P^T A P = \begin{pmatrix} \frac{2}{15} & -\frac{1}{15} & 0 \\ 0 & 0 & 1 \\ \frac{1}{15} & \frac{2}{15} & 0 \end{pmatrix} \begin{pmatrix} 2 & 2 & 0 \\ 2 & 5 & 0 \\ 0 & 0 & 3 \end{pmatrix} \begin{pmatrix} \frac{2}{15} & 0 & \frac{1}{15} \\ -\frac{1}{15} & 0 & \frac{2}{15} \\ 0 & 1 & 0 \end{pmatrix}$$

$$= \begin{pmatrix} \frac{2}{15} & -\frac{1}{15} & 0 \\ 0 & 0 & 3 \\ \frac{6}{15} & \frac{12}{15} & 0 \end{pmatrix} \begin{pmatrix} \frac{2}{15} & 0 & \frac{1}{15} \\ -\frac{1}{15} & 0 & \frac{2}{15} \\ 0 & 1 & 0 \end{pmatrix}$$

$$= \begin{pmatrix} 1 & 0 & 0 \\ 0 & 3 & 0 \\ 0 & 0 & 6 \end{pmatrix} \quad \therefore Q = y_1^2 + 3y_2^2 + 6y_3^2$$

$$\underline{y} = P^{-1} \underline{x} = P^T \underline{x} = \begin{pmatrix} \frac{2}{15} & -\frac{1}{15} & 0 \\ 0 & 0 & 1 \\ \frac{1}{15} & \frac{2}{15} & 0 \end{pmatrix} \begin{pmatrix} x_1 \\ x_2 \\ x_3 \end{pmatrix}$$

$$\therefore y_1 = \frac{1}{15} (2x_1 - x_2)$$

$$y_2 = x_3$$

$$y_3 = \frac{1}{15} (x_1 + 2x_2).$$

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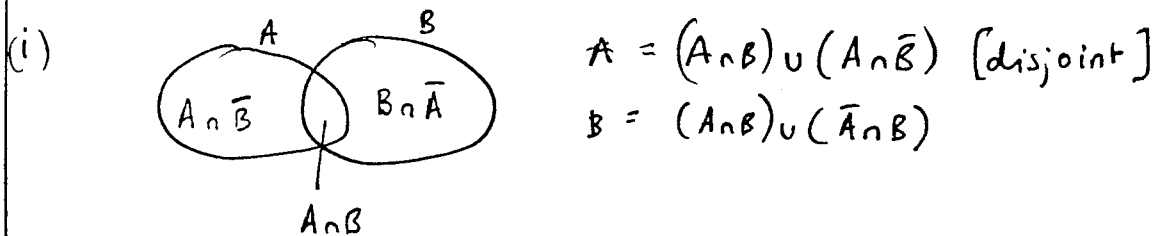
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Hence $P(A \cap \bar{B}) = P(A) - P(A \cap B)$; $P(B \cap \bar{A}) = P(B) - P(A \cap B)$

4 $P(\text{one and only one}) = P((A \cap \bar{B}) \cup (B \cap \bar{A})) = P(A \cap \bar{B}) + P(B \cap \bar{A})$
 $= P(A) + P(B) - 2P(A \cap B).$

(ii) $P(\bar{A} \cap \bar{B}) = P(\overline{A \cup B}) = 1 - P(A \cup B) = 1 - P(A) - P(B) + P(A \cap B)$
 $= 1 - P(A)P(B) + P(A)P(B)$ [indep]
 $= [1 - P(A)][1 - P(B)] = P(\bar{A})P(\bar{B}).$ \bar{A}, \bar{B} ind.

$P(\bar{A} \cap B) = P(A \cup B) - P(A)$ [$A, \bar{A} \cap B$ disjoint]
 $= P(A) + P(B) - P(A)P(B) - P(A) = P(B)[1 - P(A)] = P(B)P(\bar{A})$
 \bar{A}, B ind.

(iii) Let $A =$ equip at plant 1 still functions in 10 yrs time
 $B =$ 2

$P(A) = 1/4$; $P(B) = 1/3$; $P(\bar{A}) = 3/4$; $P(\bar{B}) = 2/3$

(a) $P(A \cap B) = P(A)P(B) = 1/4 \cdot 1/3 = 1/12$

(b) $P(A \cup B) = P(A) + P(B) - P(A)P(B) = 1/4 + 1/3 - 1/12 = 1/2$

(c) $P(\bar{A} \cap \bar{B}) = P(\bar{A})P(\bar{B})$ [by part (ii)] $= 3/4 \cdot 2/3 = 1/2$

(d) $P(\bar{A} \cap B) = P(B)P(\bar{A})$ [by part (iii)] $= 1/3 \cdot 3/4 = 1/4$

(e) $P(B|\bar{A}) = P(\bar{A} \cap B) / P(\bar{A}) = P(\bar{A})P(B) / P(\bar{A}) = P(B) = 1/3$

or, since \bar{A} and B are indep, $P(B|\bar{A}) = P(B) = 1/3.$

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4

$$(i) P(X=j) = P(\text{first } (j-1) \text{ PCBs ok, } j\text{th is faulty}) \\ = (1-q)^{j-1} q = p^{j-1} q, j=1,2,\dots$$

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$$P(X>j) = P(\text{first } j \text{ are all o.k.}) = p^j \quad \text{or} \\ p^j q + p^{j+1} q + \dots = p^j q (1 + p + \dots) \\ = p^j q / (1-p) = p^j.$$

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$$(ii) P(X>j+k | X>j) = P(X>j+k \cap X>j) / P(X>j) \\ = P(X>j+k) / P(X>j) \quad j, k +ve \text{ int} \\ = p^{j+k} / p^j \quad \text{from (i)} \\ = p^k = P(X>k).$$

2

We know (class) that the exponential distⁿ has lack-of-memory, and geometric is discrete analogue of exponential.

3

$$(iii) P(X=j, Y=k) = p^{j-1} q P(\text{next } k-j-1 \text{ o.k., } k\text{th is faulty}) \\ = p^{j-1} q p^{k-j-1} q \\ = p^{k-2} q^2 \quad k > j \geq 1.$$

3

$$P(Y=k) = \sum_{j=1}^{k-1} P(X=j, Y=k) = \sum_{j=1}^{k-1} p^{k-2} q^2 \\ = (k-1) p^{k-2} q^2.$$

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A. Waleh

S. Schwaiblmair

The hazard rate is the conditional probability intensity that a t -unit-old system will fail. It is defined as $z(t) = f(t) / 1 - F(t) = f(t) / R(t)$ where $f(t)$ is p.d.f., $F(t)$ is c.d.f. and $R(t)$ is reliability.

$$\begin{aligned} \text{(i)} \quad F_T(t) &= P(T \leq t) = P(\max\{T_1, T_2\} \leq t) \\ &= P(T_1 \leq t) P(T_2 \leq t) = F_{T_1}(t) F_{T_2}(t) \\ &= (1 - e^{-\lambda_1 t})(1 - e^{-\lambda_2 t}), \quad t \geq 0. \end{aligned}$$

$$\begin{aligned} \text{(ii)} \quad f_T(t) &= \frac{d}{dt} F_T(t) = \frac{d}{dt} (1 - e^{-\lambda_1 t} - e^{-\lambda_2 t} + e^{-(\lambda_1 + \lambda_2)t}) \\ &= \lambda_1 e^{-\lambda_1 t} + \lambda_2 e^{-\lambda_2 t} - (\lambda_1 + \lambda_2) e^{-(\lambda_1 + \lambda_2)t}, \quad t \geq 0. \end{aligned}$$

All three components are p.d.f.s (of exponentials with parameters λ_1 , λ_2 and $\lambda_1 + \lambda_2$). Hence when integrated from 0 to ∞ we get $1 + 1 - 1 = 1$.

(iii) Put $\lambda_1 = \lambda_2 = \lambda$. Then

$$\begin{aligned} z(t) &= \frac{f(t)}{1 - F(t)} = \frac{2\lambda e^{-\lambda t} - 2\lambda e^{-2\lambda t}}{2e^{-\lambda t} - e^{-2\lambda t}} = \frac{2\lambda e^{-\lambda t} (1 - e^{-\lambda t})}{e^{-\lambda t} (2 - e^{-\lambda t})} \\ &= \frac{2\lambda (1 - e^{-\lambda t})}{(2 - e^{-\lambda t})} = \frac{2\lambda (1 - e^{-\lambda t})}{1 + (1 - e^{-\lambda t})} \\ &= \frac{2\lambda}{\frac{1}{1 - e^{-\lambda t}} + 1}. \end{aligned}$$

As $t \rightarrow \infty$, $z(t) \rightarrow \lambda$, which is the same hazard rate as obtained with a single component with $\text{Exp}(\lambda)$ lifetime.

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(i) The minimum mean square error estimate of $Y=y$ in terms of $X=x$ is got by finding that function $c(x)$ which minimizes $E\{(Y-c(x))^2\} = \iint (y-c(x))^2 f_{X,Y}(x,y) dx dy$.

It turns out that $c(x) = E\{Y|X=x\}$ is the minimizing function.

Given $f_{X,Y}(x,y) = \begin{cases} x^{-1} & 0 \leq y \leq x \leq 1 \\ 0 & \text{otherwise.} \end{cases}$

(ii) $\int_0^1 dx \int_0^x \frac{1}{x} dy = \int_0^1 \frac{1}{x} \left\{ y \Big|_0^x \right\} dx = \int_0^1 dx = x \Big|_0^1 = 1.$

(iii) $f_{Y|X=x}(y|x) \equiv f_{X,Y}(x,y) / f_X(x).$

But $f_X(x) = \int_0^x f_{X,Y}(x,y) dy = \int_0^x \frac{1}{x} dy = 1, 0 \leq x \leq 1.$

So $f_{Y|X=x}(y|x) = \frac{1}{x}, 0 \leq y \leq x.$

Hence, conditional on $X=x, 0 \leq x \leq 1, Y$ is uniform on $[0, x].$

(iv) $E\{Y|X=x\} = \int_0^x y f_{Y|X=x}(y|x) dy$
 $= \int_0^x \frac{y}{x} dy = \frac{1}{x} \left\{ \frac{y^2}{2} \Big|_0^x \right\} = \frac{x}{2}.$