

**UNIVERSITY OF LONDON**

**[II(4)E 2001]**

**B.ENG. AND M.ENG. EXAMINATIONS 2001**

For Internal Students of the Imperial College of Science, Technology and Medicine

This paper is also taken for the relevant examination for the Associateship.

**PART II : MATHEMATICS 4 (ELECTRICAL ENGINEERING)**

**Thursday 7th June 2001      2.00 - 4.00 pm**

*Answer FOUR questions.*

*[Before starting, please make sure that the paper is complete; there should be 6 pages, with a total of 6 questions. Ask the invigilator for a replacement if your copy is faulty.]*

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1. Find the eigenvalues and normalized eigenvectors of the matrix

$$A = \begin{pmatrix} 2 & 1 & 0 \\ 1 & 2 & 0 \\ 0 & 0 & 4 \end{pmatrix}.$$

Using these, or otherwise, show that the matrix

$$P = \frac{1}{\sqrt{2}} \begin{pmatrix} 1 & 1 & 0 \\ -1 & 1 & 0 \\ 0 & 0 & \sqrt{2} \end{pmatrix}$$

diagonalises  $A$  such that

$$P^{-1}AP = \begin{pmatrix} 1 & 0 & 0 \\ 0 & 3 & 0 \\ 0 & 0 & 4 \end{pmatrix}.$$

2. Show that the quadratic form

$$Q = 4x_1^2 + 4x_1x_2 + x_2^2 + 4x_3^2$$

can be written as

$$Q = \mathbf{x}^T A \mathbf{x},$$

where  $\mathbf{x} = (x_1, x_2, x_3)^T$  and  $A$  is a real symmetric matrix, which is to be found. Hence, show that  $Q$  can be re-expressed in the diagonal form

$$Q = \lambda_1 y_1^2 + \lambda_2 y_2^2 + \lambda_3 y_3^2,$$

where the  $\lambda_i$  are to be determined, by finding a matrix  $P$  that satisfies  $\mathbf{x} = P\mathbf{y}$ , where  $\mathbf{y} = (y_1, y_2, y_3)^T$ . Find  $y_1$ ,  $y_2$  and  $y_3$  in terms of  $x_1$ ,  $x_2$  and  $x_3$  from the matrix  $P$ .

**PLEASE TURN OVER**

3. Use the definition of conditional probability to show that  $P(A|B)P(B) = P(B|A)P(A)$ .

A packet-switching network is, in any day, subject to either standard or heavy use, with probabilities 0.85 and 0.15, respectively. At the end of a day it is in one of three mutually exclusive states: operational, out of order due to a software failure or out of order due to a hardware failure. Given standard use, the probabilities of software and hardware failures are 0.05 and 0.015, respectively. Under heavy use both these probabilities are tripled.

- (i) Find the probability that the network is not operational at the end of a day given
  - (a) standard use,
  - (b) heavy use,during the day.
- (ii) Show that the probability that the network is operational at the end of a day is 0.9155.
- (iii) If the network is not operational at the end of a day, find the probability that it was subjected to heavy use during the day.
- (iv) Given that the network is not operational at the end of the day, find the probability that it experienced software failure during the day.

4. Explain the terms *prior distribution* and *posterior distribution*, in relation to the Bayesian analysis of data, and show how they are related.

Lifetimes,  $T$ , for certain components have a probability density function

$$f(t) = \beta^2 t e^{-\beta t}, \quad 0 \leq t < \infty, \beta > 0,$$

and  $\beta$  has a prior probability density function

$$\Pi(\beta) = \lambda e^{-\lambda \beta}, \quad 0 < \beta < \infty,$$

where  $\lambda$  is a known constant.

- (i) Write down the likelihood function (as a function of  $\beta$ ) of a sample  $t_1, \dots, t_n$  of lifetimes.
- (ii) Show that the posterior probability density function of  $\beta$  has the form

$$f(\beta|t_1, \dots, t_n) \propto \beta^{2n} \exp \left\{ -\beta \left( \lambda + \sum_{i=1}^n t_i \right) \right\}, \quad 0 < \beta < \infty,$$

and show that the constant of proportionality,  $C$ , say, has the form

$$C = \frac{(\lambda + \sum_{i=1}^n t_i)^{2n+1}}{(2n)!}.$$

- (ii) Find the posterior mean corresponding to the posterior probability density function  $f(\beta|t_1, \dots, t_n)$ .

*You may use the fact that for any positive integer  $n$ ,  $\int_0^\infty x^n e^{-x} dx = n!$*

**PLEASE TURN OVER**

5. Let the positive random variable  $T$  (in hours) represent the lifetime of an electrical component. Carefully define the hazard rate  $z(t)$  of the component and derive an expression for it in terms of the density and distribution functions of  $T$ .

- (i) Show that for any hazard rate function  $z(t)$ , the reliability  $R(t)$  is given by

$$R(t) = \exp \left[ - \int_0^t z(x) dx \right] .$$

- (ii) If  $z(t)$  has the Weibull form  $z(t) = \alpha\beta t^{\alpha-1}$ , where  $\alpha, \beta$  are positive constants:

- (a) Find the reliability,  $R(t)$ .  
 (b) Show that the probability density function  $f(t)$  of the Weibull random variable  $T$  takes the form

$$f(t) = \alpha\beta t^{\alpha-1} e^{-\beta t^\alpha}, \quad t \geq 0.$$

- (c) Given that for  $x > 0$ ,

$$\int_0^\infty u^{x-1} e^{-u} du = \Gamma(x),$$

where  $\Gamma(x)$  is the Gamma function, show that the mean,  $E\{T\}$ , of  $T$  is given by

$$E\{T\} = \beta^{-\frac{1}{\alpha}} \Gamma\left(1 + \frac{1}{\alpha}\right).$$

- (d) For which value of the Weibull distribution parameter  $\alpha$  is the exponential distribution obtained? Given that  $\Gamma(n+1) = n!$  for non-negative integers  $n$ , show that the mean obtained in (c) is consistent with that of an exponential random variable.

6. Given two random variables  $X$  and  $Y$  with joint probability density function  $f_{X,Y}(x, y)$ , the minimum mean square error estimate of the unobserved value  $Y = y$  in terms of the observed value  $X = x$  is given by  $E\{Y | X = x\}$ .

- (i) Carefully explain the meaning of this statement.

The random variables  $X$  and  $Y$  which represent the amplitudes of two signals, have joint probability density function

$$f_{X,Y}(x, y) = \begin{cases} C(x + 2y), & 0 \leq x \leq 2; 0 \leq y \leq 1, \\ 0, & \text{otherwise,} \end{cases}$$

where  $C$  is a constant.

- (ii) Find the value of  $C$  required for such a joint probability density function.
- (iii) Find  $f_{Y|X=x}(y|x)$ , the conditional probability density function of  $Y$ , given  $X = x$ , and show that it integrates to unity.
- (iv) Hence derive the minimum mean square error estimate of  $Y = y$  in terms of  $X = x$ .

**END OF PAPER**

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$$A = \begin{pmatrix} 2 & 1 & 0 \\ 1 & 2 & 0 \\ 0 & 0 & 4 \end{pmatrix} \quad \lambda_3 = 4 \text{ \& } \lambda_2 \text{ satisfy} \\ (\lambda - 2)^2 - 1 = 0 \quad \lambda_2 = 3, \lambda_1 = 1$$

$$\lambda_3 = 4; \quad \begin{matrix} 2a = b \\ 2b = a \end{matrix} \Rightarrow a = b \text{ \& } c \text{ arb.} \quad \therefore \underline{a}_3 = \begin{pmatrix} 0 \\ 0 \\ 1 \end{pmatrix}$$

$$\lambda_2 = 3 \quad \begin{matrix} a = b \\ c = 0 \end{matrix}$$

$$\underline{a}_2 = \frac{1}{\sqrt{2}} \begin{pmatrix} 1 \\ 1 \\ 0 \end{pmatrix}$$

$$\lambda_1 = 1 \quad \begin{matrix} b = -a \\ c = 0 \end{matrix}$$

$$\underline{a}_1 = \frac{1}{\sqrt{2}} \begin{pmatrix} 1 \\ -1 \\ 0 \end{pmatrix}$$

$$\therefore P = \frac{1}{\sqrt{2}} \begin{pmatrix} 1 & 1 & 0 \\ -1 & 1 & 0 \\ 0 & 0 & \sqrt{2} \end{pmatrix} = (\underline{a}_1, \underline{a}_2, \underline{a}_3) \quad \begin{matrix} A \text{ is symmetric} \\ \underline{a}_i^T \underline{a}_j = \delta_{ij} \end{matrix}$$

Either

$$\therefore P^T P = \{ \underline{a}_i^T \underline{a}_j \} = \{ \delta_{ij} \} = I \Rightarrow P^T = P^{-1}$$

$$\text{Now } A \underline{a}_i = \lambda_i \underline{a}_i \Rightarrow AP = P\Lambda \Rightarrow P^{-1}AP = \Lambda$$

$$\therefore P^T A P = \Lambda$$

$$\Lambda = \begin{pmatrix} \lambda_1 & 0 & 0 \\ 0 & \lambda_2 & 0 \\ 0 & 0 & \lambda_3 \end{pmatrix}$$

OR (directly)

$$AP = \frac{1}{\sqrt{2}} \begin{pmatrix} 2 & 1 & 0 \\ 1 & 2 & 0 \\ 0 & 0 & 4 \end{pmatrix} \begin{pmatrix} 1 & 1 & 0 \\ -1 & 1 & 0 \\ 0 & 0 & \sqrt{2} \end{pmatrix} = \frac{1}{\sqrt{2}} \begin{pmatrix} 1 & 3 & 0 \\ -1 & 3 & 0 \\ 0 & 0 & 4\sqrt{2} \end{pmatrix}$$

$$P^T = \frac{1}{\sqrt{2}} \begin{pmatrix} 1 & -1 & 0 \\ 1 & 1 & 0 \\ 0 & 0 & \sqrt{2} \end{pmatrix}$$

$$\therefore P^T A P = \frac{1}{2} \begin{pmatrix} 1 & -1 & 0 \\ 1 & 1 & 0 \\ 0 & 0 & \sqrt{2} \end{pmatrix} \begin{pmatrix} 1 & 3 & 0 \\ -1 & 3 & 0 \\ 0 & 0 & 4\sqrt{2} \end{pmatrix}$$

$$= \frac{1}{2} \begin{pmatrix} 2 & 0 & 0 \\ 0 & 6 & 0 \\ 0 & 0 & 8 \end{pmatrix} = \begin{pmatrix} 1 & 0 & 0 \\ 0 & 3 & 0 \\ 0 & 0 & 4 \end{pmatrix}$$

MATHEMATICS FOR ENGINEERING STUDENTS

EXAMINATION QUESTION / SOLUTION

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PAPER

4

QUESTION

SOLUTION

2

$$Q = 4x_1^2 + 4x_1x_2 + x_2^2 + 4x_3^2 = \underline{x}^T \begin{pmatrix} 4 & 2 & 0 \\ 2 & 1 & 0 \\ 0 & 0 & 4 \end{pmatrix} \underline{x}$$

$$\text{Evs of } A: \lambda = 4 \text{ \& } (\lambda - 4)(\lambda - 1) - 4 = 0 \quad \text{so } \lambda = 0, 5.$$

$$\lambda_1 = 0 \quad \lambda_2 = 4 \quad \lambda_3 = 5$$

$$\underline{a}_1 = \frac{1}{\sqrt{5}} \begin{pmatrix} 1 \\ -2 \\ 0 \end{pmatrix} \quad \underline{a}_2 = \begin{pmatrix} 0 \\ 0 \\ 1 \end{pmatrix} \quad \underline{a}_3 = \frac{1}{\sqrt{5}} \begin{pmatrix} 2 \\ 1 \\ 0 \end{pmatrix} \quad \underline{a}_i^T \underline{a}_j = \delta_{ij} \text{ as } A^T = A$$

$$\text{Write } P = \{ \underline{a}_1, \underline{a}_2, \underline{a}_3 \}$$

$$\text{then } P^T P = \{ \underline{a}_i^T \underline{a}_j \} = \{ \delta_{ij} \} = I$$

$$\therefore P^T = P^{-1}$$

$$\text{Writing } \underline{x} = P \underline{y} \Rightarrow \underline{y} = P^T \underline{x}$$

and

$$Q = \underline{x}^T A \underline{x} = (P \underline{y})^T A (P \underline{y}) = \underline{y}^T (P^T A P) \underline{y}$$

$$\text{Moreover } A \underline{x}_i = \lambda_i \underline{x}_i \Rightarrow A P = P \Lambda \quad \Lambda = \begin{pmatrix} 0 & 0 & 0 \\ 0 & 4 & 0 \\ 0 & 0 & 5 \end{pmatrix}$$

$$\text{so } P^T A P = \Lambda$$

$$\therefore Q = \underline{y}^T \Lambda \underline{y}$$

$$= 4y_2^2 + 5y_3^2.$$

$$\text{and } \underline{y} = P^T \underline{x} \text{ where } P = \frac{1}{\sqrt{5}} \begin{pmatrix} 1 & 0 & 2 \\ -2 & 0 & 1 \\ 0 & \sqrt{5} & 0 \end{pmatrix}$$

$$\therefore y_1 = \frac{1}{\sqrt{5}} (x_1 - 2x_2)$$

$$y_2 = x_3$$

$$y_3 = \frac{1}{\sqrt{5}} (2x_1 + x_2)$$

$$P^T = \frac{1}{\sqrt{5}} \begin{pmatrix} 1 & -2 & 0 \\ 0 & 0 & \sqrt{5} \\ 2 & 1 & 0 \end{pmatrix}$$

Setter : J. D. GIDDON

Checker: HERBERT

Setter's signature: J. D. Giddon

Checker's signature: Herbert



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EE II (4)

QUESTION

SOLUTION

Stats 3

By definition  $P(A|B) = \frac{P(A \cap B)}{P(B)}$ , but  $P(A \cap B) = P(B \cap A)$

and  $P(B|A) = \frac{P(B \cap A)}{P(A)}$  so that  $P(A|B)P(B) = P(B|A)P(A)$ .

Let

$S$  = standard use ;  $O$  = operational ;  $R$  = software fail ;  $H$  = hardware fail

$$P(R|S) = 0.05 \quad P(H|S) = 0.015$$

$$P(R|\bar{S}) = 0.15 \quad P(H|\bar{S}) = 0.045$$

$$P(S) = 0.85 \quad P(\bar{S}) = 0.15$$

$$(a) P(\bar{O}|S) = P(R|S) + P(H|S) = 0.05 + 0.015 = \underline{0.065}.$$

$$P(\bar{O}|\bar{S}) = P(R|\bar{S}) + P(H|\bar{S}) = 0.15 + 0.045 = \underline{0.195}$$

$$(b) P(O) = P(O|S)P(S) + P(O|\bar{S})P(\bar{S})$$

$$= (1 - P(\bar{O}|S))P(S) + (1 - P(\bar{O}|\bar{S}))P(\bar{S})$$

$$= (.935 \times .85) + (.805 \times .15) = \underline{.9155}$$

$$(c) P(\bar{S}|\bar{O}) = \frac{P(\bar{O}|\bar{S})P(\bar{S})}{1 - P(O)} = \frac{P(\bar{O}|\bar{S})P(\bar{S})}{.0845} = \frac{.195 \times .15}{.0845}$$

$$= \underline{.3462} \text{ to 4 dp}$$

$$(d) P(R|\bar{O}) = \frac{P(R \cap \bar{O})}{P(\bar{O})} = \frac{P(R)}{P(\bar{O})} = \frac{.065}{.0845} = \underline{.7692} \text{ to 4 dp}$$

$$\text{Since } P(R) = P(R|S)P(S) + P(R|\bar{S})P(\bar{S})$$

$$= (.05 \times .85) + (.15 \times .15)$$

$$= .065$$

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Checker's signature : DJH

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If  $\beta$  is an unknown parameter, the prior distribution of  $\beta$  represents our prior knowledge on  $\beta$  before we collect the data. The posterior represents our knowledge of  $\beta$  after we've collected the data. The prior is modified by the likelihood of the data, given  $\beta$ , to give the posterior distribution

$$f(\beta | \text{data}) \propto f(\text{data} | \beta) \pi(\beta)$$

$$f(t) = \beta^2 t e^{-\beta t} \quad \text{and} \quad \pi(\beta) = \lambda e^{-\lambda \beta} \quad \text{so}$$

$$(i) \quad L = \prod_{i=1}^n f(t_i) = \beta^{2n} \left( \prod_{i=1}^n t_i \right) e^{-\beta \sum_{i=1}^n t_i}$$

$$(ii) \quad \pi(\beta | \underline{t}) \propto \beta^{2n} e^{-\beta \sum_{i=1}^n t_i} \lambda e^{-\lambda \beta} \propto \beta^{2n} e^{-\beta(\lambda + \sum t_i)}$$

$0 < \beta < \infty.$

$$\text{But} \quad \int_0^{\infty} x^n e^{-x} dx = n! \quad \text{so}$$

$$c \int_0^{\infty} \beta^{2n} e^{-\beta(\lambda + \sum t_i)} d\beta = 1 \Rightarrow \frac{c}{\lambda + \sum t_i} \int_0^{\infty} \frac{x^{2n}}{(\lambda + \sum t_i)^{2n}} e^{-x} dx = 1$$

$$\Rightarrow \frac{c(2n)!}{(\lambda + \sum t_i)^{2n+1}} = 1 \Rightarrow c = (\lambda + \sum t_i)^{2n+1} / (2n)!.$$

$$(iii) \quad E\{\beta | \underline{t}\} = c \int_0^{\infty} \beta^{2n+1} e^{-\beta(\lambda + \sum t_i)} d\beta$$

$$= \frac{c}{(\lambda + \sum t_i)^{2n+2}} \int_0^{\infty} x^{2n+1} e^{-x} dx$$

$$= \frac{c}{(\lambda + \sum t_i)^{2n+2}} (2n+1)! = \frac{2n+1}{\lambda + \sum t_i}.$$

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PAPER  
EE II (4)

QUESTION

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SOLUTION  
Stats 5

- a) Suppose a component has survived for  $t$  time units and we want the probability it will not survive for an additional time  $\Delta t$ . Then, with  $T$  the time to failure,

$$P(T \in [t, t + \Delta t] | T > t) = P(T \in [t, t + \Delta t]) / P(T > t) \\ = f(t) \Delta t / [1 - F(t)]$$

The hazard rate  $z(t) = f(t) / [1 - F(t)]$  represents the conditional probability intensity that a  $t$ -unit-old system will fail.

$$\int_0^t z(x) dx = \int_0^t \frac{f(x)}{1 - F(x)} dx = - \int_0^t \frac{u'(x)}{u(x)} dx \text{ where } u(x) = 1 - F(x) \\ = - \ln [1 - F(t)]$$

$$\text{So } R(t) = 1 - F(t) = \exp \left[ - \int_0^t z(x) dx \right].$$

b) (i)  $R(t) = \exp \left[ - \int_0^t \alpha \beta x^{\alpha-1} dx \right] = \exp \left[ - \beta x^{\alpha} \Big|_0^t \right] = e^{-\beta t^{\alpha}}$

(ii)  $z(t) = f(t) / R(t)$  so  $f(t) = z(t) R(t) = \alpha \beta t^{\alpha-1} e^{-\beta t^{\alpha}}$

(iii)  $E\{T\} = \int_0^{\infty} t \alpha \beta t^{\alpha-1} e^{-\beta t^{\alpha}} dt = \int_0^{\infty} \alpha \beta t^{\alpha} e^{-\beta t^{\alpha}} dt$ . Put  $u = \beta t^{\alpha}$  so  $t = \left(\frac{u}{\beta}\right)^{1/\alpha}$ ,  $du = \alpha \beta \left(\frac{u}{\beta}\right)^{1/\alpha - 1} dt$  and  $\therefore dt = \frac{1}{\alpha \beta} \left(\frac{u}{\beta}\right)^{\frac{1}{\alpha} - 1} du$  and

$$E\{T\} = \int_0^{\infty} \alpha \beta \left(\frac{u}{\beta}\right)^{1/\alpha} e^{-u} \frac{1}{\alpha \beta} \left(\frac{u}{\beta}\right)^{\frac{1}{\alpha} - 1} du = \beta^{-1/\alpha} \int_0^{\infty} u^{1/\alpha} e^{-u} du = \beta^{-1/\alpha} \Gamma(1 + 1/\alpha).$$

(iv) The pdf in (ii) matches the exponential for  $\alpha = 1$ .

Then  $E\{T\} = \beta^{-1} \Gamma(2) = \beta^{-1} 1! = 1/\beta$

as for the exponential distribution.

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MATHEMATICS FOR ENGINEERING STUDENTS

EXAMINATION QUESTION / SOLUTION

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EE II(4)

QUESTION

SOLUTION 6  
Stats

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- (i) The minimum mean square error estimate of  $Y=y$  in terms of  $X=x$  is got by finding that function  $c(x)$  s.t.  
 $E\{(Y-c(x))^2\} = \iint (y-c(x))^2 f_{X,Y}(x,y) dx dy = \min.$   
 It turns out that  $c(x) = E\{Y|X=x\}$  is the minimizing fn.

$$f_{X,Y}(x,y) = \begin{cases} C(x+2y), & 0 \leq x \leq 2; 0 \leq y \leq 1 \\ 0 & \text{otherwise.} \end{cases}$$

$$\begin{aligned} \text{(ii)} \quad \int_{x=0}^2 \int_{y=0}^1 C(x+2y) dx dy &= C \int_{x=0}^2 \left\{ xy + y^2 \right\} \Big|_0^1 dx = C \int_{x=0}^2 (1+x) dx \\ &= C \int_0^2 \left( \frac{x^2}{2} + x \right) dx = 4C \Rightarrow \underline{C = 1/4}. \end{aligned}$$

$$\text{(iii)} \quad f_{Y|X=x}(y|x) = \frac{f_{X,Y}(x,y)}{f_X(x)} = \frac{x+2y}{x+1} \quad 0 \leq x \leq 2; 0 \leq y \leq 1$$

$$\text{Since } f_X(x) = \frac{1}{4} \int_0^1 (x+2y) dy = xy + y^2 \Big|_0^1 / 4 = \frac{x+1}{4}, \quad 0 \leq x \leq 2.$$

And

$$\int_0^1 f_{Y|X=x}(y|x) dy = \int_0^1 \frac{x+2y}{x+1} dy = \frac{1}{x+1} [xy + y^2]_0^1 = \frac{x+1}{x+1} = 1.$$

$$\begin{aligned} \text{(iii)} \quad E\{Y|X=x\} &= \int_0^1 y \left( \frac{x+2y}{x+1} \right) dy = \frac{1}{x+1} \int_0^1 (yx + 2y^2) dy \\ &= \frac{1}{x+1} \left[ \frac{xy^2}{2} + \frac{2}{3} y^3 \right]_0^1 = \frac{1}{x+1} \left\{ \frac{x}{2} + \frac{2}{3} \right\}. \end{aligned}$$

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