

THE ANSWERS

Notations:

- (a) B - Bookwork
- (b) E - New example
- (c) A - New application

1. a) This question checks your understanding of basic concepts of probability. It is very similar to many problems we have seen in the lectures/classes. Most students answered i) and ii) correctly. Most students could not answer iii) correctly! This is an A level question with a simple application of L'hospital rule to compute the limit. d) was answered correctly by about half of the students and results from a direct application of Chebyshev's inequality.

The CDF is given by $F_X(x) = \int_{-\infty}^x f_X(x)dx$ which leads to

$$F_X(x) = \begin{cases} 0, & x \leq 0, \\ 2x - x^2, & 0 < x < 1, \\ 1, & x \geq 1. \end{cases}$$

[2 - A]

- ii) $E(X) = \int_{-\infty}^{\infty} xf_X(x)dx$. We get $E(X) = \int_0^1 x2(1-x)dx = 1/3$.

[2 - A]

$\text{Var}(X) = E(X^2) - E(X)^2$. $E(X^2) = 1/6$. So $\text{Var}(X) = 1/6 - 1/9 = 1/18$.

[2 - A]

- iii) We write $m_X(t) = E(e^{tX}) = \int_{-\infty}^{\infty} e^{tx} f_X(x)dx$. By integration by part,

$$m_X(t) = \begin{cases} \frac{2e^t}{t^2} - \frac{2}{t^2} - \frac{2}{t}, & t \neq 0, \\ 1, & t = 0. \end{cases}$$

[1 - A]

We can compute $E(X) = m'_X(0)$ and $E(X^2) = m''_X(0)$.

[1 - A]

We get $m'_X(t) = \frac{2e^t t - 4e^t + 4 + 2t}{t^3}$. Applying L'hospital rule 3 times, we get $E(X) = 1/3$.

[1 - A]

Similarly $m''_X(t) = \frac{2e^t t^2 - 8te^t + 12e^t - 12 - 4t}{t^4}$. Applying L'hospital rule 4 times, we get $E(X) = 1/6$ such that $\text{Var}(X) = 1/6 - 1/9 = 1/18$.

[1 - A]

- iv) By Chebyshev's inequality $P\left(\left|X - \frac{1}{3}\right| \geq \frac{1}{4}\right) \leq \frac{1}{1/16} E[(X - 1/3)^2] = 16\text{Var}(X) = 8/9$.

[2 - A]

The exact value can be computed as follows

$$\begin{aligned} P\left(\left|X - \frac{1}{3}\right| \geq \frac{1}{4}\right) &= 1 - P\left(\left|X - \frac{1}{3}\right| \leq \frac{1}{4}\right) \\ &= 1 - P\left(-\frac{1}{4} \leq X - \frac{1}{3} \leq \frac{1}{4}\right) \\ &= 1 - P\left(\frac{1}{12} \leq X \leq \frac{7}{12}\right) \\ &= 1 - F_X\left(\frac{7}{12}\right) + F_X\left(\frac{1}{12}\right) \\ &= \frac{1}{3} \end{aligned}$$

[2 - A]

- b) **This question checks your understanding of basic concepts of statistics. It is very similar to many problems we have seen in the lectures/classes. Part iv) requires you to make the connection with CLT.**

By the method of moments, we aim at finding the estimator by equating the sample mean with corresponding population mean, i.e. $E(X) = \bar{X}$.

[2 - A]

We can compute

$$E(X) = \int_0^\theta \frac{2x^2}{\theta^2} dx = \frac{2}{3}\theta.$$

We choose $\tilde{\theta}$ as $\frac{2}{3}\tilde{\theta} = \bar{X}$, i.e. $\tilde{\theta} = \frac{3}{2}\bar{X}$.

[2 - A]

- ii) Expectation

$$E(\tilde{\theta}) = E\left(\frac{3}{2}\bar{X}\right) = \frac{3}{2}E(\bar{X}) = \frac{3}{2}E(X) = \theta.$$

[2 - A]

Variance

$$\text{Var}(\tilde{\theta}) = \text{Var}\left(\frac{3}{2}\bar{X}\right) = \frac{9}{4}\text{Var}(\bar{X}) = \frac{9}{4}\frac{\text{Var}(X)}{n}.$$

We can compute $E(X^2) = \int_0^\theta \frac{2x^3}{\theta^2} dx = \frac{1}{2}\theta^2$ such that $\text{Var}(X) = \frac{1}{2}\theta^2 - \frac{4}{9}\theta^2 = \frac{1}{18}\theta^2$. This leads to $\text{Var}(\tilde{\theta}) = \frac{\theta^2}{8n}$.

[2 - A]

- iii) Since $E(\tilde{\theta}) = \theta$, the estimator is unbiased.

[4 - A]

- iv) For large n , we can make use of the Central Limit Theorem and write

$$\begin{aligned} P(\tilde{\theta} \geq \theta) &= P\left(\frac{3}{2}\bar{X} \geq \theta\right) = P\left(\bar{X} \geq \frac{2}{3}\theta\right) \\ &= P\left(Z \geq \frac{\frac{2}{3}\theta - E(\bar{X})}{\sqrt{\text{Var}(\bar{X})}}\right) = P(Z \geq 0) = 1/2 \end{aligned}$$

where Z is a standard normal random variable.

[4 - A]

2. This question checks your understanding of basic concepts of probability. It is very similar to many problems we have seen in the lectures/classes. One difficulty is in part b) that requires to make a change of variable and use the Jacobian. Part g) is also a subquestion that was not correctly answered by most students. Always pay attention to the domain of the RVs.

We need to compute $E(X)$. We can first compute the marginal of X as

$$f_X(x) = \begin{cases} 2e^{-x} \int_x^\infty e^{-y} dy = 2e^{-2x}, & x > 0, \\ 0, & \text{otherwise.} \end{cases}$$

[2 - A]

This is an EXPO(2). Hence $E(X) = \frac{1}{2}$.

[2 - A]

- b) We can first compute the Jacobian and write

$$\begin{vmatrix} \frac{\partial X}{\partial U} & \frac{\partial X}{\partial V} \\ \frac{\partial Y}{\partial U} & \frac{\partial Y}{\partial V} \end{vmatrix} = \begin{vmatrix} 0 & 1 \\ 1 & 1 \end{vmatrix} = 1$$

[2 - A]

We then write

$$f_{U,V}(u,v) = \begin{cases} 2e^{-(u+2v)}, & u > 0, v > 0, \\ 0, & \text{otherwise.} \end{cases}$$

[2 - A]

- c) The marginal are obtained by integration of the joint pdf as follows

$$f_U(u) = \begin{cases} e^{-u}, & u > 0, \\ 0, & \text{otherwise.} \end{cases}$$

$$f_V(v) = \begin{cases} 2e^{-2v}, & v > 0, \\ 0, & \text{otherwise.} \end{cases}$$

U is EXPO(1) and V is EXPO(2).

[2 - A]

- d) Since $f_{U,V}(u,v) = f_U(u)f_V(v)$, U and V are two independent random variables.

[2 - A]

- e) The conditional pdf $f_{U|V}(u|v)$ is given as

$$f_{U|V}(u|v) = f_U(u) = \begin{cases} e^{-u}, & u > 0, \\ 0, & \text{otherwise.} \end{cases}$$

[2 - A]

- f) $E(U|V) = E(U) = 1$.

[2 - A]

- g) $1 = E(U|V) = E(Y - X|X) = E(Y|X) - E(X|X) = E(Y|X) - X$. Hence $E(Y|X) = 1 + X$.

[2 - A]

- h) $E(Y) = E_X E(Y|X) = E(1 + X) = 1 + E(X) = 1 + \frac{1}{2} = \frac{3}{2}$.

[2 - A]