1 Methods

1.1 Model

Our model assumes that there are T time periods t = 1, 2, ..., T. A company has a stock return y_t at t time period which is observable. At time period t, there's also a hidden variable $h_t \in \{1, 2\}$ representing whether that company is in a good state or a bad state. The model assumes that y_t is generated from h_t by

$$y_t | h_t \sim N(\mu_{h_*}, \sigma_{h_*}) \tag{1}$$

That means, the return is drawn from two normal distributions, a bad distribution for bad state and a good distribution for good state. Presumably, $\mu_2 > \mu_1$, which means that good state tends to get higher expected return.

The h_t satisfies Markov property such that

$$Pr(h_t|h_1^{t-1}) = Pr(h_t|h_{t-1})$$

That means, h_t is conditional independent with h_1, h_2, \dots, h_{t-2} given h_{t-1} . Therefore the transition probability can be simply characterized as

$$\begin{array}{c|c|c|c} Pr(h_t|h_{t-1}) & h_t = 1 & h_t = 2 \\ \hline h_{t-1} = 1 & q_1 & 1 - q_1 \\ h_{t-1} = 2 & q_2 & 1 - q_2 \\ \end{array}$$

In summary, our model assumes that a company's stock return is a Hidden Markov Model (HMM) of T time periods, where observable data \boldsymbol{y} (stock return) is drawn from normal distributios with mean and standard deviation $\boldsymbol{\mu}, \boldsymbol{\sigma}$ depending on the hidden state \boldsymbol{h} . The hidden state transition probability are \boldsymbol{q} . Overall, our model has data \boldsymbol{y} , hidden variables \boldsymbol{h} and parameters $\boldsymbol{\mu}, \boldsymbol{\sigma}, \boldsymbol{q}$. In later sections, we also use $\boldsymbol{\theta}$ to denote all parameters $(\boldsymbol{\mu}, \boldsymbol{\sigma}, \boldsymbol{q})$.

1.2 Priors and Full Conditionals for MCMC

To infer the parameters and predict future stock return, we run MCMC with the following priors and full conditionals.

For $\mu, \sigma | y, h, q$, we use the Normal-Gamma as the prior (for $\mu, \phi = 1/\sigma^2$) like what we did in ordinary Gaussian (normal) models:

$$1/\sigma^2 \sim G(\nu_0/2, SS_0/2)$$
 (2)

$$\mu|\phi \sim N(\mu_0, 1/(\kappa_0 \phi)) \tag{3}$$

The full conditionals of μ , σ are also Normal-Gamma. The only difference is that we need to use h to split y into two sets $Y_{h=1}, Y_{h=2}$. Then we use $Y_{h=k}$ to update μ_k, σ_k :

$$\mu_k | \phi_k, Y_{h=k} \sim N\left(\mu_{n,k}, \frac{1}{\kappa_{n,k}\phi_k}\right)$$
 (4)

$$\phi_k | Y_{h=k} \sim G(\nu_{n,k}/2, SS_{n,k}/2)$$
 (5)

where

$$\kappa_{n,k} = \kappa_0 + n_k = \kappa_0 + |Y_{h=k}| \tag{6}$$

$$\mu_{n,k} = \frac{\phi_k n_k \overline{Y_{h=k}} + \phi_k \kappa_0 \mu_0}{\phi_k \kappa_{n,k}} \tag{7}$$

$$\nu_n = \nu_0 + n_k \tag{8}$$

$$SS_n = SS_0 + SS_{h=k} + \frac{n_k \kappa_0}{\kappa_{n_k}} (\overline{Y_{h=k}} - \mu_0)^2$$
 (9)

All those priors and full conditionals are similar to what's covered in lecture 6. For q, we use Beta as prior

$$q \sim Beta(a_0, b_0) \tag{10}$$

The full conditionals are also Beta distributions. Suppose n_{ij}^t is the number of transitions in the form $h_{t-1} = i, h_t = j$, we have

$$q_1|\mathbf{h} \sim Beta(a_0 + n_{11}^t, b_0 + n_{12}^t)$$
 (11)

$$q_2|\mathbf{h} \sim Beta(a_0 + n_{21}^t, b_0 + n_{22}^t)$$
 (12)

The full conditionals of h can be derived in two ways:

1. **Direct Gibbs (DG)**: In the first approach, we calculate full conditional for each h_t (thus assuming that all other h_{-t} are fixed):

$$Pr(h_t|\boldsymbol{h_{-t}},\boldsymbol{q},\boldsymbol{y}) \propto Pr(h_t|h_{t-1},\boldsymbol{q})Pr(h_{t+1}|h_t,\boldsymbol{q})Pr(y_t|h_t,\mu_{h_t},\sigma_{h_t})$$
 (13)

where
$$Pr(h_i|h_{i-1}, \mathbf{q}) = q_{h_{i-1}h_i}$$
 if we set $q_{11} = q_1, q_{12} = 1 - q_1, q_{21} = q_2$ and $q_{22} = 1 - q_2$.

2. Forward Backward (FB) Recursion: In the second approach, we calculate full conditionals of the whole h using Forward Backward Recursion algorithms (actually, we will only use forward step here) that are well known in previous HMM research. The full conditionals are (recall that $\theta = (q, \mu, \sigma)$ represents all parameters)

$$Pr(\boldsymbol{h}|\boldsymbol{y},\boldsymbol{\theta}) = Pr(h_T|\boldsymbol{y}_1^T,\boldsymbol{\theta}) \prod_{t=n-1}^{1} Pr(h_t|h_{t+1}^T,\boldsymbol{y},\boldsymbol{\theta})$$

$$= Pr(h_T|\boldsymbol{y}_1^T,\boldsymbol{\theta}) \prod_{t=n-1}^{1} Pr(h_t|h_{t+1},\boldsymbol{y}_1^{t+1},\boldsymbol{\theta})$$
(14)

where $Pr(h_t|h_{t+1}, y_1^{t+1}, \boldsymbol{\theta}) \propto Pr(h_t, h_{t+1}|y_1^{t+1}, \boldsymbol{\theta})$. Note that $Pr(h_T|y_t^T, \boldsymbol{\theta}), Pr(h_t, h_{t+1}|y_1^{t+1}, \boldsymbol{\theta})$ can be calculated efficiently by Forward Backward Recursion:

$$Pr(h_t, h_{t+1}|y_1^{t+1}, \boldsymbol{\theta}) \propto Pr(h_t|y_1^t, \boldsymbol{\theta}) Pr(h_{t+1}|h_t, \boldsymbol{q}) Pr(y_{t+1}|h_{t+1}, \boldsymbol{\theta})$$
 (15)

$$Pr(h_t|y_1^t, \theta) = \sum_{h_{t-1}} Pr(h_t, h_{t-1}|y_1^t, \theta)$$
(16)