The Dynamics of Price Dispersion, or Edgeworth Variations

Timothy Cason
Department of Economics, Purdue University

Daniel Friedman
Department of Economics, UCSC

and

Florian Wagener CeNDEF, University of Amsterdam

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Abstract:

Hypotheses on the dynamics of dispersed prices are extracted from computer simulations, as well as traditional and recent theory. The hypotheses are tested on existing laboratory data. As predicted in some variations of the Edgeworth hypothesis, the laboratory data exhibit a significant cycle. Relative to the unique stationary distribution, the empirical distribution of posted prices has excess mass in an interval that moves downward over time until it approaches the lower boundary of the stationary distribution. Then the excess mass jumps upward and the downward cycle resumes. The amplitude of the cycle seems fairly constant over the longer experimental sessions. Of the simulations we consider, the one closest to Edgeworth's 1925 account, a hybrid of gradient dynamics and logit dynamics, seems to best reproduce the observed dynamics.

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The Dynamics of Price Dispersion

1. Introduction

The law of one price fails dramatically even in Internet markets where it should have its best shot (e.g., Baye and Morgan, 2001; Pauly, Herring and Song, 2002). This observation renews interest in theoretical debates on price dynamics and price dispersion that go back at least to Bertrand (1883) and Edgeworth (1925).

Bertrand argued for a unified competitive price, noting that at any higher price at least one seller could substantially increase sales volume and profit by slightly decreasing his price. Edgeworth argued that the outcome is more complicated when the sellers have binding capacity constraints at the competitive price. He predicted a price cycle in which sellers reduce price in small increments when there is excess capacity, but jump to much higher prices when the capacity constraints bind. Maskin and Tirole (1988) consider a duopoly with alternating moves (and no capacity constraints) and obtain price cycles as Markov perfect equilibria.

Modern textbook treatments of multi-firm oligopolies (e.g., Tirole, 1988) downplay price dynamics and instead focus on stationary price dispersion, supported as a mixed strategy Nash equilibrium. In particular, the Burdett and Judd (1983) noisy search model has a unique Nash equilibrium price distribution, denoted NSE below. The distribution is truly dispersed—i.e., it has positive density over a non-trivial range of prices—for most parameter values of the buyer search technology, but for extreme parameter values the distribution degenerates to a unified competitive price or to a unified monopoly price.

Hopkins and Seymour (2002) show that the unified monopoly price NSE is dynamically stable but that all relevant dispersed price equilibria are unstable under a wide class of learning dynamics. Their analysis leaves open the question of what behavior we should then observe. The

possibilities include convergence to a unified price, or small or large amplitude cycles in the price distribution, or chaos.

In this paper we further analyze laboratory data originally gathered to investigate NSE. After offering more details on the theory, we describe a set of dynamic computer simulations of the dispersed prices, and derive testable hypotheses. Then we present the laboratory procedures and data. The data feature dispersed prices that tend to cycle. Estimated transition rates among the quartiles of the empirical price distribution confirm an Edgeworth-like cycle. The amplitude of the cycle is not small, and does not appear to decrease over time. Of the simulations we consider, a hybrid of gradient dynamics and logit dynamics seems to best reproduce the observed dynamics.

2. Theory

A textbook account of Edgeworth dynamics (Chamberlin, 1962, includes a classic example) typically begins by describing Bertrand best reply dynamics in a symmetric oligopoly with negligible fixed costs and constant marginal cost. Each period each firm posts a price slightly below the prices posted in the previous period. Eventually prices decline to the firms' marginal cost and profit is zero. The account then recalls Edgeworth's observation that if demand at this low price exceeds the firms' combined capacity, then a firm's best response is to charge the monopoly price for the residual demand. Other firms follow, Bertrand competition resumes, and a new price cycle begins.

We shall refer to the foregoing as the basic Edgeworth cycle (BEC) and, somewhat exaggerating its precision, we shall say that its observable features are a narrow interval of prices each period that moves downward until it hits bottom, then jumps upward, with cycles of

constant amplitude and frequency.

Modern discussions such as Tirole (1988) question whether such a BEC is an equilibrium, and focus on mainly on stationary equilibrium distributions. The relevant equilibrium in the present context is the Burdett and Judd (1983) Noisy Search Equilibrium distribution (NSE), derived in closed form in Cason and Friedman (2003). The presumption seems to be that if cycles were to occur, their amplitude would decrease over time and the distribution would converge to NSE.

Hopkins and Seymour (2002) reverse this presumption. They show that NSE is dynamically unstable under a wide class of learning dynamics. The local linear approximation indicates that a small perturbation of NSE will lead to Edgeworth-like cycles whose amplitude is small at first but increasing. They refer to the cycles as Rock-Paper-Scissors (RPS) to emphasize that an excess of medium prices follows an excess of high prices that follows an excess of low prices that follows an excess of medium prices, in the same way that, in the familiar children's game, Paper beats Rock which beats Scissors which beats Paper. Their analysis does not show whether the cycle amplitude tapers off at some moderate level (i.e., the attractor is a stable limit cycle not far from the NSE distribution), or diverges to a high amplitude BEC, or goes chaotic. The simulations introduced below demonstrate that the answer depends on the specific dynamics and parameters.

The simulations represent buyers via an algorithm that is derived from the Noisy Search assumption that each of a continuum of buyers (total mass normalized to 1) seeks to buy a unit at lowest price, and employs an optimal reservation price search strategy given a search technology that produces random samples of sellers' posted prices. The initial sample contains one price with probability $q \ge 0$ or two prices with probability $1 - q \ge 0$, and the buyer can obtain a fresh such

sample at cost $c \ge 0$. When the distribution of posted prices has cumulative distribution function F, a seller posting price p will then obtain sales volume s(p) = q + 2(1-q)(1-F(p)) if $p \le p^* = c/(1-q)$ and s(p) = 0 if $p > p^*$. The algorithm yields seller profits $\pi = ps(p)$ because production cost is normalized to 0. NSE in this setup is characterized by an equilibrium price dispersion with distribution function $F(p) = 1 + \left(1 - \frac{p^*}{p}\right) \frac{q}{2-2q}$ on the support interval $[qp^*/(2-q), p^*]$.

This distribution comes from an equal profit condition and is nontrivial when c > 0 and 0 < q < 1; see Cason and Friedman (2003) for the derivations.

2.1 Gradient dynamics.

Consider now what happens away from NSE. One possibility is that sellers each period consider nearby prices and (with some inertia) chose those that would have been most profitable last period.

We simulate these gradient dynamics by choosing a price grid $0 \le p_1 \le ... \le p_i \le ... \le p_n$, initial proportions of sellers $f_1, ..., f_i$, ..., f_n posting these prices, and a positive adjustment rate parameter α . The simulation computes profit π_i as above at each price p_i , and then adjusts the proportion of sellers as follows. For each price p_i ,

(1) if $\pi_{i-1} > \pi_i$ then f_i decreases and f_{i-1} increases by the amount $(\pi_{i-1} - \pi_i) \alpha f_i$; and if $\pi_{i-1} < \pi_i$ then f_i increases and f_{i-1} decreases by the amount $(\pi_i - \pi_{i-1}) \alpha f_{i-1}$.

The index i here is to be understood mod n, so that p_1 and p_n are adjacent. The idea is that the probability mass describing sellers' choices moves between nearest neighbors, up the profit

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¹ That is, the highest relevant price p^* is deemed nearby to the lowest relevant price $qp^*/(2-q)$. Thus we avoid imposing an arbitrary treatment of endpoints, e.g., as absorbing or reflecting. More importantly, thus we allow the possibility of rapid transitions from low to high prices, as hypothesized in the Edgeworth (or RPS) story.

gradient. The flux is proportional to the slope of the gradient and to the mass already there.

These dynamics do not satisfy the hypotheses of Hopkins and Seymour's theorem, so we have no theoretical prediction regarding their convergence behavior. Simulations therefore are especially helpful.

Figure 1A shows a typical simulation. From an initial uniform distribution, the density evolves via heavily damped RPS oscillations to a neighborhood of the NSE distribution. But it seems to converge to a low amplitude limit cycle rather than to the NSE, and the Figure shows the initial part of such a cycle. Behavior is qualitatively the same for other buyer parameters c > 0 and 0 < q < 1. Figure 1B shows that as the price grid gets finer, the time average distribution converges very closely to the NSE distribution.

Figure 1A [dyn_supp2.pdf] about here

Figure 1B [conv 2.pdf] about here

2.2. Replicator dynamics

Replicator dynamics (Taylor and Jonker, 1978; Weibull, 1995) are perhaps the most widely known adjustment processes. They do satisfy the assumptions of Hopkins and Seymour's theorem so we do not expect convergence to NSE. Our simulation uses the same price grid and profit calculations as before, notes the current average profit $\pi_A = \Sigma_i f_i \pi_i$, and computes the next period proportions f_i^* from

(2)
$$f_i' = f_i + ((\pi_i - \pi_A)/\pi_A)\alpha f_i$$
.

That is, probability mass moves to or from any price point at a rate proportional to its payoff (relative to the population average) and to the mass already there.

Figure 2A shows a typical replicator simulation. From an initial uniform distribution, the density quickly diverges via increasing RPS oscillations to an extreme basic Edgeworth cycle (BEC). Figure 2B shows that the time-averaged distribution does not converge to NSE, but rather to something closer to the uniform distribution (but with moderate skew in the same direction as NSE.)

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2.3. Logit dynamics

In the case of replicator dynamics, prices are coupled only through the average profit. For our logit dynamics, coupling between two price levels is more explicit. Using the same price grid and profit calculations as before, compute the "attraction weights" β_i each period from

$$\beta_i = \alpha \frac{e^{\lambda \pi_i}}{\sum_j e^{\lambda \pi_j}}, \quad 1 \le i \le n.$$

Note that β_i is close to its maximum value α if the profit is uniquely maximized at price p_i and β_i is close to its minimum value 0 if the profit at p_i is much lower than the maximum. Note also that larger values of the parameter λ create larger differences in β_i for given profit differentials. However, it turned out that our final results hardly depended on λ . To reduce the number of free parameters, we set it equal to unity.

The dynamics are now specified as follows: for every pair of prices p_i and p_j , each period

(3) a portion $\beta_i f_j$ of the sellers at price p_j moves to p_i , and a portion $\beta_j f_i$ of the sellers at price p_i moves to p_j .

Simulations show that the dynamics depend more strongly on the adjustment rate α than in the two previous cases: if α is small, logit dynamics converge to a stationary distribution but one which is different from the NSE. For larger values of α , the stationary distribution becomes dynamically unstable and cycling behavior is observed. The period of the cycles is much shorter than in replicator dynamics.

Figures 3A,B about here.

[dyn supp3.pd, conv 3.pdf]

2.4. Hybrid dynamics

A close reading of Edgeworth³ shows that he allowed transitions to high prices from medium low as well as from very low prices. To investigate the implications, we constructed a 'hybrid' type of dynamics by adding logit-like terms $\sum_{j\neq i} \beta_i f_j - \sum_{j\neq i} \beta_j f_i$ to the gradient dynamics (1), thus allowing flux between all possible points of our price grid. For small α, this Edgeworth variation seems to converge to something resembling NSE, but again it produces fairly large amplitude cycles for larger values of α, as can be seen in Figure 4A. As the price grid gets finer, the time averages again converge to a distribution more moderately skewed than the NSE.

Figures 4A, B about here

2.5. Testable Hypotheses

² Given our choice of price grid etc, the critical value of α is about 0.15.

³ For at every stage in the fall of the price, and before it has reached its limiting value [...], it is competent to each monopolist to deliberate whether it will pay him better to lower the price against his rival [...] or rather to raise it to a higher [...] level [...]. Edgeworth (1925)

Let Q_{it} , i=1...4, be the observed fraction of sellers posting prices in quartile i, and let $M = ((m_{ij}))$ be the 4×4 matrix of observed proportions of sellers moving from quartile i one period to quartile j the next period. We will examine the following hypotheses.

H0: The transition matrix is uniform: all off-diagonal entries equal, and diagonal entries equal. That is, up to sampling error, $m_{ij} = m_{kl}$ for all $i \neq j$ and all $k \neq l$; and similarly $m_{ii} = m_{jj}$ for all i,j = 1, ..., 4. This hypothesis is consistent with NSE equilibrium and with rational expectations, as well as with no structure.

H1. Bertrand/RPS/Edgeworth cycles: The subdiagonal entries (those representing transitions to next lower quartile) are significantly larger than other off-diagonal entries. That is, up to sampling error, $m_{i,i-1} > m_{ij}$ for all $j \neq i$, i-1 for i = 2, 3, 4. This is the clearest implication of the cycle stories.

H1'. Full cycles: In addition to H1, it is also true that transitions from the lowest quartile are most often to the highest quartile. That is, $m_{1,4} > m_{1,j}$ for j = 2,3. This is the simplest way of completing an Edgeworth-like cycle.

H2. Anticipatory dynamics: The sub-subdiagonal entries are larger than the next lower entries. That is, $m_{4,2} > m_{4,1}$. Sellers who anticipate that other sellers respond as in H1 will best respond with larger price reductions.

H3. Divergence from NSE: H1 holds more strongly in later periods than in early periods.

H4. Convergence to NSE: H1 holds more strongly in earlier periods and H0 holds in later periods.

3. The Laboratory Experiment

The laboratory experiment used an electronic posted offer market institution with human subjects

as sellers and (in the data analyzed below) automated buyers. Subjects were recruited from undergraduate classes in Economics and Biology at UCSC and Purdue, and were instructed orally and in writing. The written instructions appear in Cason and Friedman (1999) and are posted at http://www.mgmt.purdue.edu/faculty/cason/papers/NSEinst.pdf. At the end of the session the subjects received total profits, converted from lab dollars into U.S. dollars at a fixed exchange rate and averaging about \$20 for sessions that lasted about 90-100 minutes.

Each period each of six sellers posts a single price. Profit $\pi = ps(p)$ is computed using the formula given earlier. Sales volume is explained to subjects in terms of the equivalent expression $s(p) = q + 2(1-q)(r_{i-1})/(n-1)$, where r_i is the rank of a given seller's price p_i among the $n \le 6$ sellers currently posting prices at or below the reservation price p^* . For example, with c = 0.60 and q = 1/3, the reservation price is $p^* = \$0.90$; we tell the sellers that each of hundreds of small automated buyers immediately accepts the lowest price it sees if it is below \$0.91 and otherwise keeps searching, with the result that sales volume will be $0.33 + 1.33(r_{i-1})/5$ if all 6 sellers post prices below \$0.91. The current values of c and d, and corresponding reservation prices, are shown on the blackboard and on subjects' screens. At the end of the period each seller see a sorted list of all posted prices, with corresponding sales volumes and profits. Seller identities are not shown.

We report here results from four sessions, each divided into several runs of 20-30 consecutive periods in which all treatments are held constant. The treatments include the four combinations of search cost, c=0.20 or 0.60, and sample size, q = 1/3 or 2/3. Two sessions with experienced subjects, one with c = 0.20 and one with c = 0.60, were conducted at UCSC; each included two runs with q = 1/3 and two runs with q = 2/3. Two other sessions, one at Purdue University with c = 0.20 and the other at UCSC with c = 0.60, used inexperienced subjects and

contained one run each with q = 1/3 and q = 2/3. All four sessions had initial runs (not analyzed here) with q = 1 and q = 0; these values theoretically produce unified price. See Cason and Friedman (2003) for a more complete description of the experiment, which included additional sessions with human buyers and with noisier buyer algorithms.

4. Empirical results

We begin by presenting the overall price distributions and the trends in some sample periods.

Different readers may form differing impressions from these qualitative summaries, so we then proceed to estimate transition matrices and to test our hypotheses.

4.1. Price distributions and trends

The equilibrium distribution of posted prices (NSE) shifts up with increases in the probability that buyers observe only one price (q) and with increases buyer search costs (c), as indicated by the lighter histogram bars in Figure 5. With 20-cent search costs, NSE prices range between 6 and 30 cents for q=1/3 and range between 30 and 60 cents for q=2/3. With 60-cent search costs, NSE prices range between 18 and 90 cents for q=1/3 and range between 90 and 180 cents for q=2/3. The observed distributions, indicated by darker bars, shift as predicted and generally are in the predicted range. But they tend to be less skewed than the NSE distributions. Cason and Friedman (2003) offer a detailed analysis of these regularities.

Our present focus is price dynamics. Figures 6 through 8 present the dynamics in a sample of 3 of the 12 runs studied here. The circles represent posted prices, and the short dashed line indicates the mean transaction price for each period. Solid vertical lines separate trading periods, and the horizontal dashed lines indicate the NSE price range.

Figure 6 presents a q=2/3 run with 20-cent search costs. Prices decline gradually in the first part of the run before they jump up toward the reservation price around period 80. Price dispersion decreases in the middle of this run before increasing again—reflecting many price reductions—in later periods. Figure 7 presents a q=1/3 run with 60-cent search costs that begins with one of the best basic Edgeworth cycles (BEC) we have seen in the data. By contrast, Figure 8 (a q=2/3 run with 60-cent search costs) shows prices with higher cross-sectional variance; the distribution shifts down in early periods but seems to stablize in later periods.

4.2. Transition matrices

The analysis begins with the empirical distribution of posted prices F_k , where k indexes the

experimental treatment (i.e., the search cost c and probability q of receiving only one seller's price).⁴ First, define the quartile cutoff points r_{1k} , r_{2k} and r_{3k} , where $0.25 = F_k(r_{1k})$, $0.50 = F_k(r_{2k})$, and $0.75 = F_k(r_{3k})$. Then define the four quartiles as the price observation between the cutoff points. Those between r_{1k} and r_{2k} , for example, are in quartile 2 (denoted Q_2).

The movement of prices from one period to the next can be summarized by a 4×4 transition matrix M_k , where the ij^{th} entry gives the empirical proportion of sellers whose prices move from quartile i in period t to quartile j in period t+1. Large entries on the diagonal indicate a stable price distribution and very little churning within the distribution.

Table 1 presents the estimated transition matrices for sessions with experienced subjects, and Table 2 does the same for inexperienced subjects. Of course, in all the calculations the first transition from the previous treatment k is excluded. The diagonal entries indicate that prices remain in the same quartile for two consecutive periods roughly 40 to 75 percent of the time. The subdiagonal entries indicate that transitions to the next lower quartile typically are more frequent than transitions to any other quartile. The only exceptions are the Q_3 quartile in the c=20 and q=2/3 treatment in Table 1 and the Q_4 quartile in the c=20 and q=1/3 treatment in Table 2. In many cases these transitions are twice as common as any other transition.

Tables 3 and 4 subdivide the experienced subject data into first and second runs. Both relevant sessions featured a first run of 30 periods and a second run of 20 periods for both q treatments, so this is a natural way to divide the data. The tables suggest several regularities. First, particularly for the c=20, q=2/3 and c=60, q=1/3 treatments, the diagonal entries tend to be higher in the early periods. Of the 16 pairwise diagonal comparisons between Tables 3 and 4 the values are lower in early periods for only four cases. Second, the subdiagonal entries tend to be

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⁴ The NSE distribution could be used instead of the empirical distribution, but it turns out to be impractical because its lowest quartile contains very few observations. The reason, as documented in the previous subsection, is that the empirical price distributions do not exhibit the strong skewness of the NSE distributions.

larger for the late data, except for the c=60, q=1/3 treatment. Indeed, these entries are at least twice as great as the next highest off diagonal entry in 7 out of the 12 possible rows in the late period data, compared to only 4 out of the 12 rows in the early period data. Finally, we note a possible regularity not considered in our list of hypotheses. The transition $Q_4^t \rightarrow Q_1^{t+1}$ from very high to very low prices is small for all treatments in the early periods. But this transition is the most frequent for search cost=60 in the late periods. But this could be an artifact of the small sample size for Q_4^t in the late data, since the number of seller price changes in these transition cells is only 8 or 5 in the q=1/3 and q=2/3 cases.

4.3 Hypothesis tests

Consider now the formal hypotheses H0 to H4. The data clearly reject H0 because the estimated transition matrices are obviously not uniform, even allowing for randomly distributed errors.

The data seem generally consistent with H1, since the subdiagonal entries are usually the largest off-diagonal entries. The null hypothesis implies that the target (subdiagonal) entry would be largest of the three off-diagonal entries in about one-third of the rows, or in about 4 of the 12 cases. A simple binomial test rejects this null hypothesis in favor of H1. For Table 1 as well as for Table 2, we have subdiagonal entries largest in 11 of 12 cases (p-value < 0.01). For Table 3 (early periods of experienced data), 9 of the 12 these transitions to the next lower quartile are highest (p-value < 0.01). And for Table 4 (late periods of experienced data), 8 of the 12 subdiagonal entries are largest (p-value < 0.05).

The results do not support Hypothesis H1', however, since transitions from the lowest to the highest quartile are greater than transitions to other quartiles less than half the time. In three of the four relevant cases shown in Table 1, the transition to the next higher quartile $(m_{1,2})$ is

greater than the transition to the highest quartile $(m_{1,4})$. And in the four cases shown in Table 2, the $m_{1,4}$ transition is greater than the $m_{1,2}$ and $m_{1,3}$ transitions only for the c=60 treatment. Pooling these eight cases together for a binomial test fails to reject the null hypothesis that transitions out of the lowest quartile are equally likely to fall into the three other quartiles (p-value = 0.53).

The data provide some support for H2, since the 2-below entries are usually greater than the 3-below entries in the transition matrices. Here quartile Q_4^t is relevant. The $Q_4^t \rightarrow Q_1^{t+1}$ transition $(m_{4,1})$ is smaller than the $Q_4^t \rightarrow Q_2^{t+1}$ transition $(m_{4,2})$ in 4 out of the 4 matrices of Table 2 (inexperienced data), and in 3 out of the 4 matrices of Table 1 (experienced data). Nevertheless, these 2-below and 3-below transitions are still rather infrequent, occurring typically about 10 to 15 percent of time.

Hypotheses H3 and H4 require a comparison between Table 3 (early periods) and Table 4 (later periods); inexperienced subject data do not have second runs in a given q condition and so are not useful for this comparison. As already noted in the previous subsection, the transition frequency to the next lower quartile (H1) tends to be greater in the later data than in the early data. To evaluate H3 formally, we differenced the 12 relevant transitions ($Q_i^t \rightarrow Q_{i-1}^{t+1}$) between Tables 4 and 3. Under the null hypothesis that the transitions to next lower quartile are equally strong in the early and later data, these 12 differences would be centered on zero. One is exactly zero, 7 are greater in the later data (+.13, +.04, +.15, +.10, +.28, +.22, and +.06), and 4 are greater in the early data (-.01, -.04, -.01 and -.12). A Wilcoxon signed rank test indicates that the central tendency of these 12 differences is significantly different from zero at the marginal 10-percent level (Wilcoxon test statistic=19.5). We conclude that the evidence favors H3 (divergence from NSE) over H4 (convergence), but the statistical significance is rather weak.

4.4 Fitting the simulations

Simulations may provide evidence on the crucial question of convergence. The strategy is to fit the observed transition matrices to simulated transition matrices obtained from each of the four dynamics described earlier. We standardize on a 24-point uniform price grid over the support of the NSE, and a simulation length 2000 time iterations. (Gradient dynamics was the exception; we needed a longer simulation and used 50,000 iterations.) A preliminary simulation produces the time-average distribution from which we determine which of the 24 price points belong to each quartile. The main simulation then yields average transition probabilities between quartiles over varying numbers ω of time iterations that might correspond to a single period of the experiment. The other parameter to fit is the adjustment speed α .

We computed the 4×4 quartile transition matrix for α values 0.05, 0.10, 0.15, 0.20, and 0.25 (higher values produced numerical instabilities in some cases) and a wide range of ω for each of the four dynamics. In each case we took the differences d_{ij} between the simulated matrix entries and the corresponding laboratory data matrix entries for experienced subjects in Table 1. The best fit is defined as that which produces the smallest Hilbert-Schmidt norm, the square root of the sum of the squares of the matrix elements d_{ij} .

The results are shown in Figure 9 and in Table 5 for c=20, q=1/3. Replicator dynamics give the worst fit; gradient dynamics give fair fits, but at very large ω ; logit and hybrid dynamics fit almost equally well. The overall best fit is given by hybrid dynamics for α =0.25: the corresponding dynamics are cycling. All best-fit matrices have large subdiagonal (and upper right hand corner) entries, consistent with H1 (and even H1').

The large ω for gradient dynamics may arise from its nearest-neighbors nature: information on higher profits needs far more time to travel through the grid. This translates into a

large number of simulation iterations corresponding to single period of laboratory time.

Similar results hold for the other transition matrices in Table 1. In all cases, hybrid dynamics provide the best fit. For the best fitting parameter values, the dynamics display cycles.

5. Discussion

Are dispersed prices a transient phenomenon, destined to disappear as buyers and sellers learn to obey the law of one price? All our evidence indicates that the answer is no. Experienced sellers in our laboratory experiment change prices more frequently in later periods than in earlier periods, and in all treatments the average distributions of posted prices resembles the dispersed equilibrium (NSE) of the Noisy Search model of Burdett and Judd (1983).

Do prices settle down into stable distribution such as NSE? Again, the answer appears to be negative. Consistent with the theoretical predictions of Hopkins and Seymour (2002), the NSE seems to be unstable and we observe a price cycle. Observed price transitions from higher prices tend to be to slightly lower prices, consistent with hypothesis H1.

Do we then get a basic Edgeworth cycle (BEC), where sellers' prices move downward together until they hit bottom, then jump together to the top? Our direct evidence supports neither the closeness each period nor the jump to the top. Simulations of replicator dynamics exhibit a strong BEC, but they fit the data poorly.⁵

The data instead seem more consistent with a gentler variant of Edgeworth dynamics. Relative to the stationary distribution NSE, each period there is an excess of prices (but not usually a majority of prices) in one of the quartiles, and transitions are more frequent to the next

⁵ Many other evolutionary and learning dynamics not investigated here also satisfy the assumptions of Hopkins and Seymour's main theorem. The presumption now would be that when such dynamics produce a BEC they would fit the data poorly. As noted in the Appendix, it is possible to find learning dynamics that fit the data rather well.

lower quartile than to other quartiles. In the best fitting simulations—logit dynamics and gradient dynamics and especially a hybrid of the two—the cycle maintains a constant amplitude, neither very small (which would be hard to distinguish from convergence to NSE) nor very large (as in the BEC). Direct tests of the data (for hypotheses H3 and H4) indicate divergence from NSE, not convergence. Both direct examination of the data (H1' tests) and the success of the logit and hybrid rules indicate that sellers sometimes move to higher prices before they have dropped to the bottom of the price distribution.

The laboratory and simulation results may help organize examinations of field data. For example, Lach (2002) analyzes individual seller price data on coffee, flour, chicken, and refrigerators. The reported transition matrices on close examination show a now familiar pattern: the largest off diagonal entries are almost always on the subdiagonal, suggesting some variant of Edgeworth dynamics.

We close with a caveat and call for further work. Four laboratory sessions at two sites is a good start, but strong claims for robustness will require more data from different labs and from the field. Four dynamical specifications, each with two fitted parameters again is a reasonable start, but other investigators may want to consider other variations and other criteria for goodness of fit. Our hope is simply that the questions and methods introduced here will turn out to be fruitful in dealing with the dynamics of price dispersion.

Appendix: Simulation Details

We collect some details here for investigators interested in replicating and extending our simulation results. Comparable details regarding the laboratory data can be found in Cason and Friedman (1999, 2003).

The simulations were coded in MATLAB.⁶ Each entry $m_{i,j}$ in the 4×4 transition matrix M_t in a given simulation is simply the proportion of sellers charging prices in quartile i in period t who charge a price in quartile j in the next period. The length ω average is the time average of the products M_{t+1} ·...· $M_{t+\omega}$. Quartile bins might at times carry very little probability mass. In order to prevent them from influencing unduly the averages in that situation, the term M_{t+1} ·...· $M_{t+\omega}$ in the average is weighted with the mass that resides on average in the quartile bin over the period from t+1 up to $t+\omega$.

The above description of the transition matrix is insufficient in the case of replicator dynamics, since in that case the migration of sellers between quartiles is not modeled explicitly, but only in reference to the average profit. In order to be able to compute transition matrices we supplement the model by an independence condition: we assume that all price-changing sellers gather together in one place, forget what prices they came from, and are distributed anew over the prices according to the replicator rule.

The concluding remarks call for investigating new variations on the dynamics. We now sketch preliminary results from a specification inspired by the remarks of an anonymous referee. The referee conjectured that we might be able to better replicate the behavior of six human sellers if we used heterogeneous learning models (such as stochastic fictitious play) with individual parameters and "forgetting." Accordingly, we created heterogeneity among six simulated sellers by drawing individual values of the logit parameter beta, and we explored

⁶ Code is available on request from the third author, by e-mail to: f.o.o.wagener@uva.nl.

memory length using a parameter similar to that in Cheung and Friedman (1997). The best fit parameters involved immediate forgetting (as in the Cournot learning model) and moderate heterogeneity. The best fit had Hilbert-Schmidt error about 10% lower than the best fit reported in the paper. Perhaps still better fits are possible with other sorts of heterogeneous dynamics!

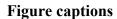
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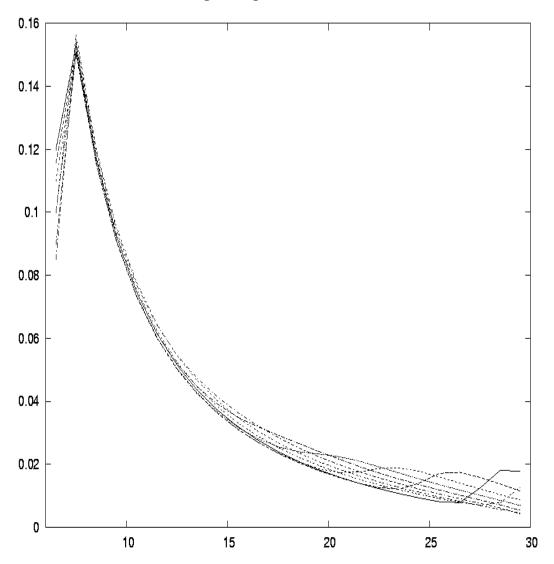


Figure 1A:

Simulated gradient dynamics. Buyer parameters are c=0.20 and q=1/3, the rate constant is alpha = 0.20 and the 24 point price grid is evenly spaced over the theoretical support interval [0.06, 0.30]. The density of posted prices is graphed at 0, 50 and 100 time periods following approximate convergence the NSE distribution. Note that a small hump (excess of prices) enters at the high end of the distribution in period 0. It flattens out and moves to lower prices by period 100.

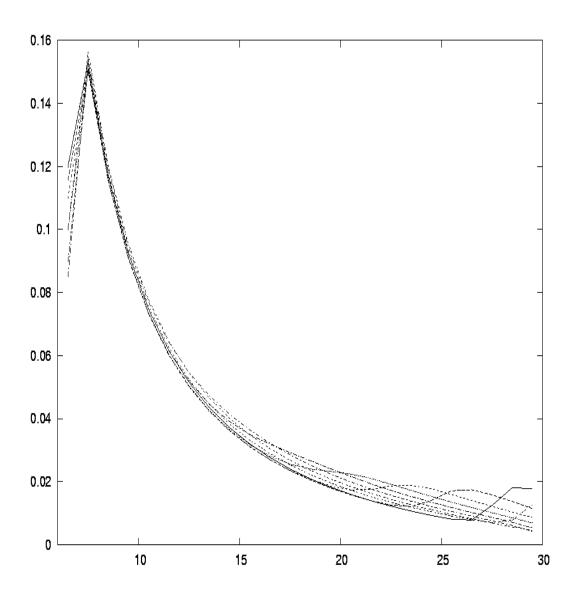
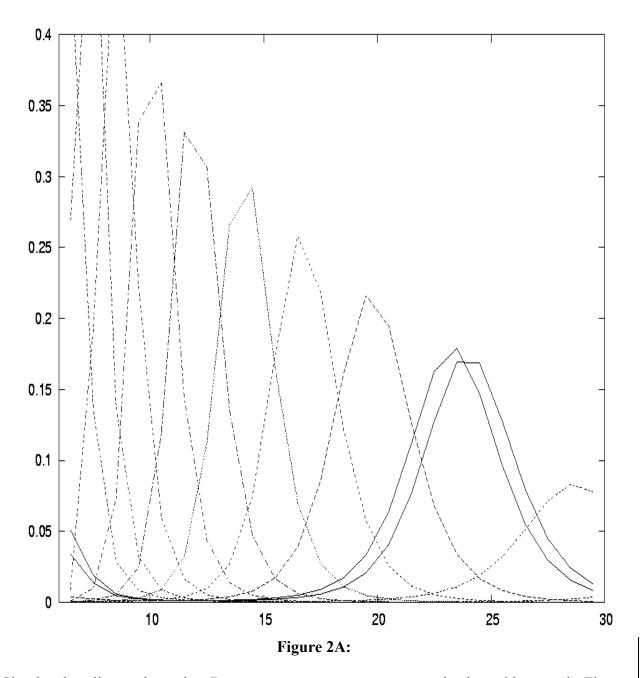
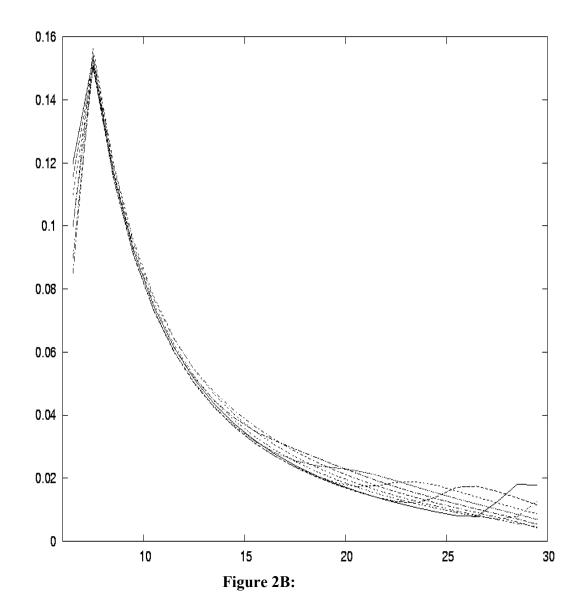


Figure 1B:

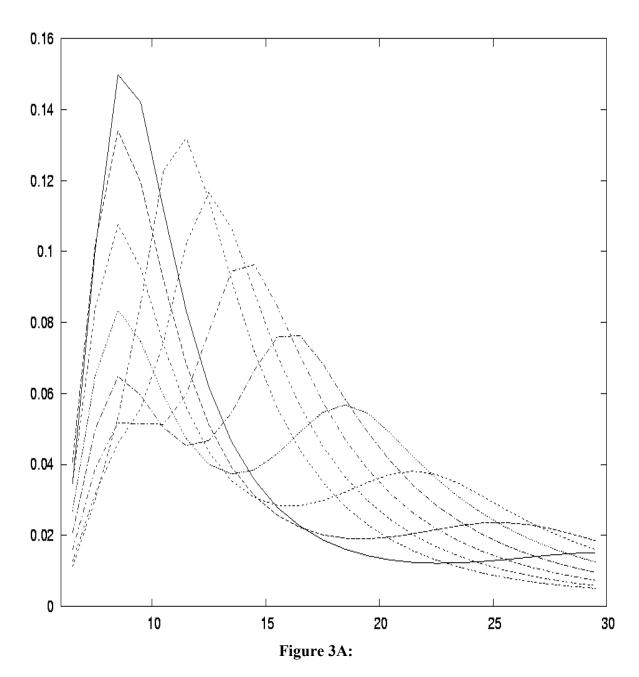
Time averages of gradient dynamics. Buyer and rate parameters are the same as in Figure 1A. The solid lines graph densities of posted prices averaged over 50,000 time periods for evenly spaced price grids of 10, 20, 40 and 80 points. The dashed line graphs the NSE distribution, which assumes a price continuum.



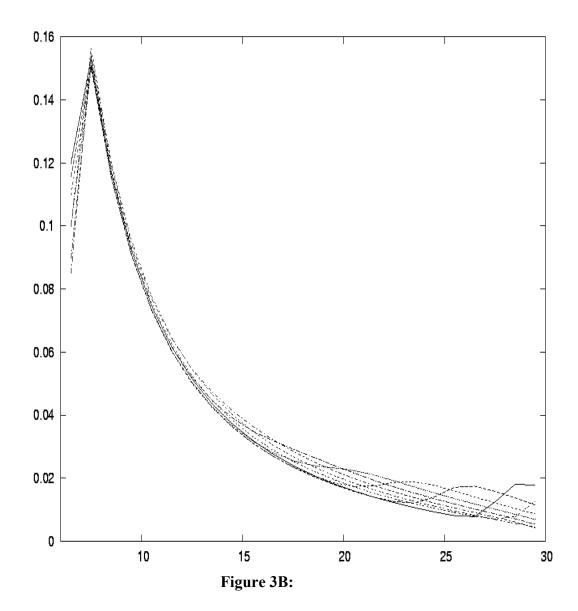
Simulated replicator dynamics. Buyer parameters, rate constant and price grid are as in Figure 1A. The density of posted prices is graphed following approximate convergence to a BEC. At t=0, a peak enters at high prices. It becomes higher and sharper as it moves towards lower prices at t=40, 80, 120. At t=160, the peak arrives at the bottom of the price range. After that, high prices have become profitable and the cycle repeats.



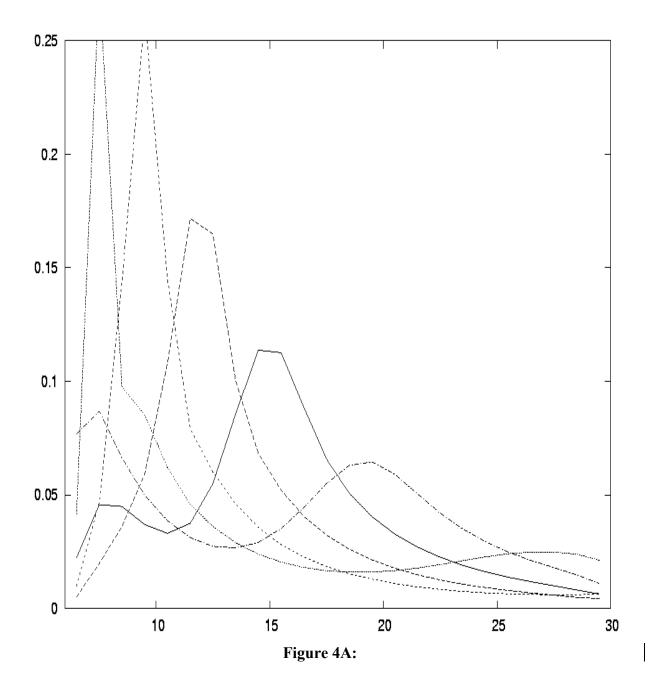
Time averages of replicator dynamics. Parameters are as in Figure 1B. As the price grid becomes finer, the discrete distributions (solid lines) converge to a continuous distribution that is much less skewed (more uniform than) the NSE distribution (dashed line).



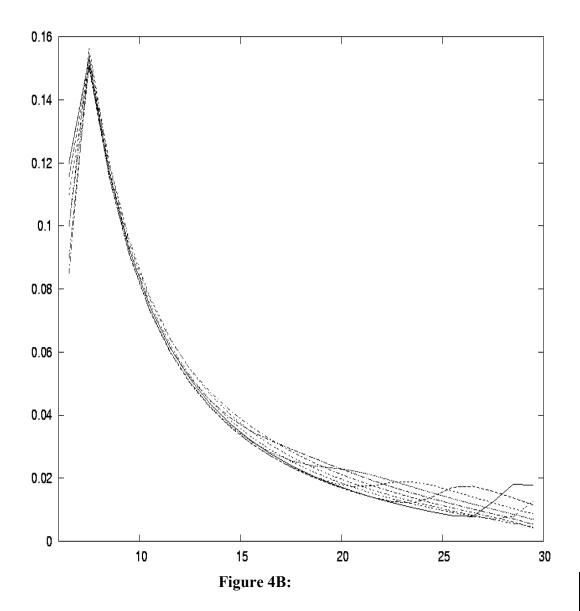
Simulated logit dynamics. Buyer parameters, rate constant and price grid are as in Figure 1A. An Edgeworth cycle can be seen as excess probability mass moves from medium prices at t=0, to lower prices at t=2 and t=4, to high prices at t=6 and to medium high prices at t=8.



Time averages of logit dynamics. Parameters are as in Figure 1B. As the price grid becomes finer, the discrete distributions (solid lines) converge rapidly to a smooth distribution that is somewhat less skewed (more uniform than) the NSE distribution (dashed line).



Simulated hybrid dynamics. Buyer parameters, rate constant and price grid are as in Figure 1A. A larger amplitude Edgeworth cycle can be seen as excess probability mass moves from medium prices at t=0, to lower prices at t=2 and t=4, to high prices at t=6 and to medium high prices at t=8.



Time averages of hybrid dynamics. Parameters are as in Figure 1B. As the price grid becomes finer, the discrete distributions (solid lines) converge rapidly to a smooth distribution that is somewhat less skewed (more uniform than) the NSE distribution (dashed line).

Table 1: Estimated Transition Matrices for All Periods in Experienced Sessions

			search	cost=20	q=2/3						
Observ.		$Q_{\scriptscriptstyle m l}^{\scriptscriptstyle t+1}$	Q_2^{t+1}	Q_3^{t+1}	Q_4^{t+1}	Observ.		$Q_{ m l}^{\scriptscriptstyle t+1}$	Q_2^{t+1}	Q_3^{t+1}	Q_4^{t+1}
82	Q_1^t	0.73	0.11	0.09	0.07	83	Q_1^t	0.75	0.10	0.06	0.10
70	Q_2^t	0.24	0.59	0.13	0.04	63	Q_2^t	0.22	0.62	0.08	0.08
65	Q_3^t	0.08	0.25	0.58	0.09	75	Q_3^t	0.13	0.12	0.59	0.16
71	Q_4^t	0.03	0.11	0.14	0.72	61	Q_4^t	0.02	0.11	0.36	0.51
		search	cost=60	q = 1/3				search	cost=60	q=2/3	
Observ.		$Q_{\scriptscriptstyle m l}^{\scriptscriptstyle t+1}$	Q_2^{t+1}	Q_3^{t+1}	Q_4^{t+1}	Observ.		$Q_{ m l}^{\scriptscriptstyle t+1}$	Q_2^{t+1}	Q_3^{t+1}	Q_4^{t+1}
71	Q_1^t	0.63	0.18	0.07	0.11	74	Q_1^t	0.66	0.16	0.09	0.08
74	Q_2^t	0.24	0.55	0.14	0.07	71	Q_2^t	0.24	0.55	0.07	0.14
74	$Q_3^{\bar{t}}$	0.04	0.24	0.62	0.09	71	$Q_3^{\overline{t}}$	0.03	0.25	0.52	0.20
69	Q_4^t	0.14	0.06	0.16	0.64	72	Q_4^t	0.10	0.11	0.28	0.51

Table 2: Estimated Transition Matrices for All Periods in Inexperienced Sessions

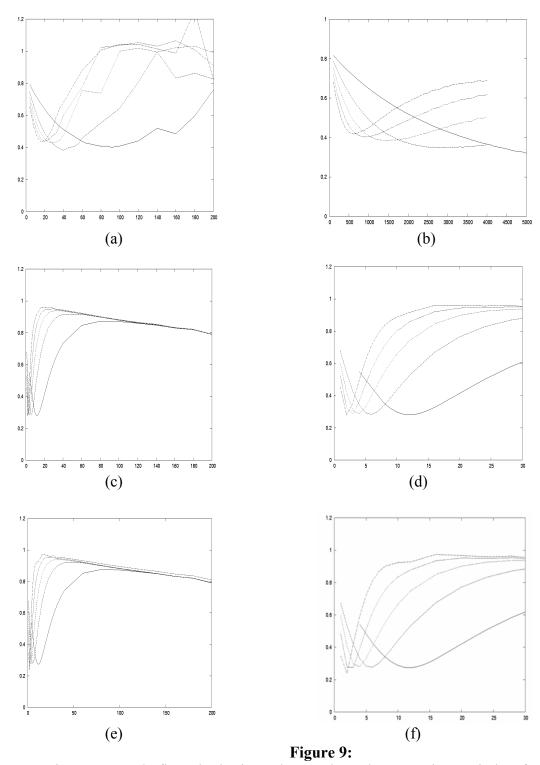
		search	cost=20	q = 1/3				search	cost=20	q=2/3	
Observ.		Q_1^{t+1}	\mathcal{Q}_2^{t+1}	Q_3^{t+1}	Q_4^{t+1}	Observ.		Q_1^{t+1}	Q_2^{t+1}	Q_3^{t+1}	Q_4^{t+1}
44	Q_1^t	0.64	0.07	0.16	0.14	46	Q_1^t	0.61	0.20	0.09	0.11
41	Q_2^t	0.32	0.49	0.15	0.05	47	Q_2^t	0.28	0.40	0.15	0.17
43	Q_3^t	0.05	0.30	0.53	0.12	38	Q_3^t	0.11	0.34	0.45	0.11
46	Q_4^t	0.02	0.20	0.17	0.61	43	Q_4^t	0.02	0.16	0.23	0.58
		search	cost=60	q = 1/3				search	cost=60	q = 2/3	
Observ.		Q_1^{t+1}	\mathcal{Q}_2^{t+1}	Q_3^{t+1}	Q_4^{t+1}	Observ.		Q_1^{t+1}	Q_2^{t+1}	Q_3^{t+1}	Q_4^{t+1}
48	Q_1^t	0.56	0.17	0.08	0.19	46	Q_1^t	0.76	0.07	0.07	0.11
38	Q_2^t	0.34	0.45	0.11	0.11	39	Q_2^t	0.23	0.64	0.10	0.03
42	$Q_3^{\overline{t}}$	0.14	0.24	0.57	0.05	44	$Q_3^{\overline{t}}$	0.11	0.18	0.55	0.16
46	Q_4^t	0.07	0.11	0.24	0.59	45	$O_{\scriptscriptstyle A}^{t}$	0.02	0.04	0.27	0.67

Table 3: Estimated Transition Matrices for **Early** Periods in Experienced Sessions

search cost= $20q=1/3$								search	cost=20	q=2/3	
Observ.		Q_1^{t+1}	Q_2^{t+1}	Q_3^{t+1}	Q_4^{t+1}	Observ.		Q_{l}^{t+1}	Q_2^{t+1}	Q_3^{t+1}	Q_4^{t+1}
41	Q_1^t	0.73	0.15	0.07	0.05	44	Q_1^t	0.75	0.07	0.07	0.11
42	Q_2^t	0.19	0.57	0.21	0.02	33	Q_2^t	0.15	0.67	0.12	0.06
44	Q_3^t	0.07	0.25	0.55	0.14	54	Q_3^t	0.11	0.09	0.67	0.13
47	Q_4^t	0.04	0.09	0.13	0.74	43	Q_4^t	0.02	0.09	0.28	0.60
										2 /2	
		search	cost=60	q = 1/3				search	cost=60	q=2/3	
Observ.		search Q_1^{t+1}	$\frac{\cos t=60}{Q_2^{t+1}}$	$\frac{q=1/3}{Q_3^{t+1}}$	Q_4^{t+1}	Observ.		search Q_1^{t+1}	$\frac{\text{cost=60}}{Q_2^{t+1}}$	$\frac{q=2/3}{Q_3^{t+1}}$	Q_4^{t+1}
Observ. 38	Q_1^t				Q_4^{t+1} 0.11	Observ.	Q_1^t				$\frac{\mathcal{Q}_4^{t+1}}{0.07}$
-	$egin{array}{c} Q_1^t \ Q_2^t \end{array}$	Q_{l}^{t+1}	Q_2^{t+1}	Q_3^{t+1}			$egin{array}{c} Q_1^t \ Q_2^t \end{array}$	Q_1^{t+1}	Q_2^{t+1}	Q_3^{t+1}	
38		Q_1^{t+1} 0.74	$Q_2^{t+1} = 0.08$	Q_3^{t+1} 0.08	0.11	28		Q_1^{t+1} 0.61	Q_2^{t+1} 0.21	Q_3^{t+1} 0.11	0.07

Table 4: Estimated Transition Matrices for <u>Late</u> Periods in Experienced Sessions

		search	cost=20	q=1/3		search cost= $20 q=2/3$
Observ.		Q_1^{t+1}	Q_2^{t+1}	Q_3^{t+1}	Q_4^{t+1}	Observ. Q_1^{t+1} Q_2^{t+1} Q_3^{t+1} Q_4^{t+1}
41	Q_1^t	0.73	0.07	0.10	0.10	39 Q_1^t 0.74 0.13 0.05 0.08
28	Q_2^t	0.32	0.61	0.00	0.07	30 Q_2^t 0.30 0.57 0.03 0.10
21	$Q_3^{\overline{t}}$	0.10	0.24	0.67	0.00	$Q_3^{\bar{t}} = 0.19 = 0.19 = 0.38 = 0.24$
24	Q_4^t	0.00	0.17	0.17	0.67	18 Q_4^i 0.00 0.17 0.56 0.28
		search	cost=60	q=1/3		search cost= $60 q=2/3$
Observ.		Q_1^{t+1}	\mathcal{Q}_2^{t+1}	Q_3^{t+1}	Q_4^{t+1}	Observ. Q_1^{t+1} Q_2^{t+1} Q_3^{t+1} Q_4^{t+1}
33	Q_1^t	0.52	0.30	0.06	0.12	46 Q_1^t 0.70 0.13 0.09 0.09
35	Q_2^t	0.22	0.54	0.11	0.11	26 Q_2^t 0.38 0.46 0.04 0.12
21	$Q_3^{\overline{t}}$	0.10	0.24	0.57	0.10	$Q_3^{t} = 0.00 0.29 0.54 0.17$
25	Q_4^t	0.32	0.08	0.08	0.52	18 Q_4^i 0.28 0.11 0.28 0.33



Errors in c=20, q=1/3 fits. The horizontal axes show the averaging period ω for the simulated transition matrices, and the vertical axes show the (Hilbert-Schmidt) distance between the simulated matrix and the first transition matrix in Table 1. One curve is shown for each value of the parameter α : 0.05: solid line, 0.10: long dashes, 0.15: short dashes, 0.20: dots, 0.25: dash-dotted. Replicator dynamics are shown in (a), gradient in (b), logit in (c) and (d), hybrid in (e) and (f).

Table 5: Best Simulation Fits for c=20, q=1/3

	Replicator		Gradient		Lo	git	Hybrid		
α	period	error	period	error	period	error	period	error	
0.05	92	0.399	7600	0.262	12	0.280	12	0.274	
0.10	40	0.383	3000	0.348	6	0.283	6	0.276	
0.15	28	0.428	1400	0.384	4	0.285	4	0.280	
0.20	20	0.438	900	0.403	3	0.286	3	0.276	
0.25	16	0.437	600	0.418	2	0.279	2	0.240	

(a) Best fitting transition matrix at given parameters and quality of fit.

	Q_1^{t+1}	Q_2^{t+1}	Q_3^{t+1}	Q_4^{t+1}
Q_1^t	0.59	0.10	0.12	0.19
\mathcal{Q}_2^t	0.21	0.58	0.09	0.12
Q_3^t	0.08	0.21	0.63	0.08
Q_4^t	0.04	0.09	0.20	0.67

(b) Overall best fitting simulated transition matrix (hybrid, α =0.25).