Polygenic Risk Score (PRS) Introduction 501 LD and Concluding Remarks

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At the end of this lecture, some basic understanding of our **limited understanding of LD**

- ▶ No heterogeneity between ex.data and my.data.
- ▶ Only SNPs in perfect LD, $r^2 = 1$, and
- ▶ But vary the number tagging SNPs
- No allelic heterogeneity, and
- No multiple causual SNPs within a locus.

First

Live Final Quiz 8

Adding two perfectly tagging SNPs to each of the $10\ \text{causal SNP},$ the AUC of PRS.gw will

A: decrease

B: increase

C: ~same

D: identical

Recall the baseline model without any heterogeneity

10 out 5000 indep. SNPs with varying 'moderate-large' effects are truly associated with Y (all $\beta=0.3$ but MAF vary).

$$Y_i = \sum_{j=1}^{10} \beta_j G_{ij} + e$$
, where $\beta_j = 0.3$

 $\mathsf{MAF} \sim \ \mathsf{Unif}(0.05, 0.5), \ e \sim \mathit{N}(0, 1).$

```
# external data
ex.nsnp.true=10; ex.beta.true=0.3
ex.nsnp=5000; ex.nsample=1000; ex.sigma=1; ex.seed=101
ex.sumstat=generate.ex.sumstat(ex.seed,ex.nsample,ex.nsnp,ex.nsnp.true,ex.beta.true,ex.sigma)
# my data
my.nsnp.true=10; my.beta.true=0.3; my.maf=ex.sumstat[,"MAF"]
my.nsnp=5000; my.nsample=1000; my.sigma=1; my.seed=102
my.data=generate.my.data(my.seed,my.nsnp)my.nsnp.true,my.beta.true,my.sigma,my.maf)
```

Total and SNP h^2 of the external model

```
## [1] 0.243
```

[1] 0.023 0.009 0.032 0.031 0.019 0.021 0.029 0.022 0.030 0.028

The index of the six 'genome-wide' significant ones

[1] 1 3 4 6 7 10

Total and SNP h^2 of my model

```
## [1] 0.243
```

[1] 0.023 0.009 0.032 0.031 0.019 0.021 0.029 0.022 0.030 0.028

Recall the different $PRS_i = \sum_{j=1}^{J} \hat{\beta}_j G_{ij}$

Using the GW threshold on the external data

$$\textit{my.PRS}_{\textit{GW}} = \sum_{j=1}^{6} \hat{eta}^{\textit{external}}_{j} imes \textit{G}^{\textit{my.data}}_{ij}$$

Using $\alpha=0.01$ (and also add $\alpha=0.1$) on the external data

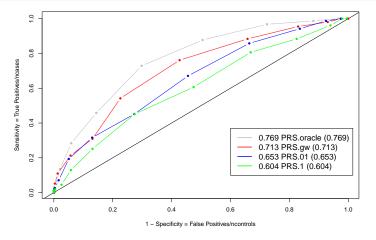
$$my.PRS_{.01}$$
 (or $my.PRS_{.1}$) = $\sum_{i=1}^{66 \text{ (or } 492)} \hat{\beta}_j^{\text{external}} \times G_{ij}^{my.data}$

The oracle one (benchmarking the upper bound)

$$my.PRS_{oracle} = \sum_{i=1}^{10} 0.3 \times G_{ij}^{my.data}$$

The baseline model ROC and AUC

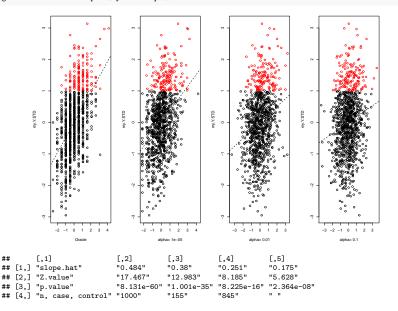
```
# generate the ROC plots
my.PRS.output=generate.my.PRS.output(ex.sumstat,my.data,alpha.level,1.threshold)
generate.ROC.plot(my.PRS.output)
```



```
## alpha J TP FP
## [1,] 1e-05 6 6 0
## [2,] 1e-02 66 10 56
## [3,] 1e-01 492 10 482
```

##

generate.association.plot(my.PRS.output)



```
ex.sumstat[1:13,]
```

```
##
               MAF MAF.hat beta
                                   beta.hat
                                                         Z.value
                                                                      p.value
                                                    se
   [1.] 0.21748927
                    0.2215 0.3
                                 0.29257288 0.06536792 4.4757871 8.489445e-06
##
                    0.0610
                            0.3
##
   [2,] 0.06972117
                                 0.33145758 0.10935747 3.0309551 2.500692e-03
   [3.] 0.36935781
                    0.3780
                            0.3
                                 0.23908858 0.05323916 4.4908404 7.922031e-06
##
##
   [4.] 0.34596068
                    0.3480
                            0.3
                                 0.38889542 0.05565755 6.9872894 5.116550e-12
   [5,] 0.16243508
                    0.1695
                            0.3 0.30892955 0.07052329 4.3805323 1.308960e-05
##
##
   [6.] 0.18502467
                    0.1995
                            0.3 0.37606430 0.06503910 5.7821265 9.859505e-09
   [7.] 0.31318998
                    0.3375
                            0.3
                                 0.33166110 0.05410586 6.1298559 1.264930e-09
   [8,] 0.20006021
                    0.2020
                            0.3
                                 0.28159164 0.06670313 4.2215657 2.647447e-05
   [9.] 0.32990538
                    0.3360
                            0.3
                                 0.23025579 0.05661344 4.0671574 5.134017e-05
## [10.] 0.29562285
                    0.2905
                            0.3
                                 0.28906539 0.05841261 4.9486810 8.766086e-07
## [11.] 0.44590808
                    0.4445
                            0.0 0.09584075 0.05424572 1.7667892 7.756916e-02
## [12,] 0.36809363
                    0.3745
                            0.0 -0.02245388 0.05302784 -0.4234356 6.720687e-01
## [13.] 0.37938767 0.3750 0.0 -0.06366768 0.05424574 -1.1736899 2.407993e-01
```

my.data\$my.sumstat[1:13,]

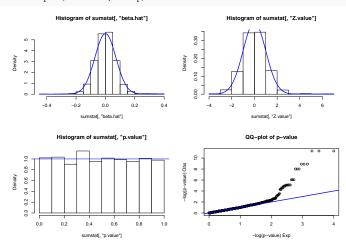
```
##
               MAE MAE hat beta
                                  beta.hat
                                                         Z.value
                                                                      p.value
                                                   se
   [1.] 0.21748927 0.2270 0.3 0.17164714 0.06066349 2.8294965 4.755586e-03
   [2,] 0.06972117 0.0755
                            0.3 0.33447059 0.09604226 3.4825358 5.182232e-04
##
   [3.] 0.36935781
                    0.3555
                            0.3 0.32234988 0.05230483 6.1629085 1.034940e-09
   [4.] 0.34596068
                    0.3545
                            0.3 0.25019642 0.05335395 4.6893703 3.121234e-06
##
   [5.] 0.16243508
                    0.1530
                            0.3
                                 0.32262395 0.06958963 4.6360925 4.021553e-06
##
   [6,] 0.18502467
                    0.1800
                            0.3
                                 0.28017270 0.06679596 4.1944557 2.978552e-05
##
   [7.] 0.31318998
                    0.3045
                            0.3
                                 0.36190034 0.05517189 6.5595060 8.652840e-11
   [8.] 0.20006021
                    0.1780
                            0.3
                                 0.35342514 0.06681630 5.2895046 1.507249e-07
   [9,] 0.32990538
                    0.3300
                            0.3
                                 0.31052039 0.05197947 5.9739043 3.218822e-09
## [10.] 0.29562285
                    0.2960
                            0.3
                                 0.33840898 0.05429097 6.2332464 6.731015e-10
## [11.] 0.44590808
                    0.4465
                            0.0 0.04515026 0.05043254 0.8952605 3.708637e-01
## [12,] 0.36809363
                    0.3580
                            0.0 -0.02127391 0.05509213 -0.3861515 6.994668e-01
## [13.] 0.37938767
                    0.3870 0.0 -0.02908571 0.05264218 -0.5525171 5.807178e-01
```

Consider LD now

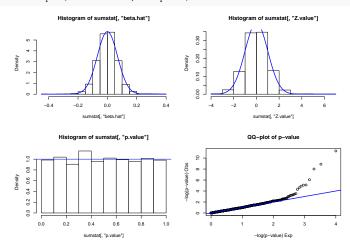
```
# First save the baseline noLD model results for later comparison
ex.nsnp.noLD=ex.nsnp
ex.sumstat.noLD=ex.sumstat
my.data.noLD=my.data
my.nsnp.noLD=my.nsnp
my.PRS.output.noLD=my.PRS.output
# Second, specify the ntag.T for each of the nsnp.true, start with 2 for each
# For now, no null SNPs are tagged, the extreme version of the assumption
# that `truly associated SNPs are more likely to be tagged than null SNPs'
# external data
ex.nsnp.true=10; ex.beta.true=0.3
# specify the ntag. T for each of the nsnp. true
ex.ntag.T=rep(2,ex.nsnp.true)
\#ex.ntag.T=c(1,2,3,4,5,6,7,8,9,10)
\#ex.ntag.T=c(10,9,8,7,6,5,4,3,2,1)
\#ex.ntag.T=c(2,2,1,4,3,10,2,3,7,5)
ex.nsnp=(5000+sum(ex.ntag.T)); ex.nsample=1000; ex.sigma=1; ex.seed=101
# use a new LD-aware data and summary stat function
ex.sumstat=generate.ex.sumstat.LD(ex.seed,ex.nsample,ex.nsnp,ex.nsnp.true,ex.be
```

##		MAF	MAF.hat	beta	beta.hat	se	Z.value	p.value
##	[1,]	0.21748927	0.2215	0.3	0.29257288	0.06536792	4.4757871	8.489445e-06
##	[2,]	0.06972117	0.0610	0.3	0.33145758	0.10935747	3.0309551	2.500692e-03
##	[3,]	0.36935781	0.3780	0.3	0.23908858	0.05323916	4.4908404	7.922031e-06
##	[4,]	0.34596068	0.3480	0.3	0.38889542	0.05565755	6.9872894	5.116550e-12
##	[5,]	0.16243508	0.1695	0.3	0.30892955	0.07052329	4.3805323	1.308960e-05
##	[6,]	0.18502467	0.1995	0.3	0.37606430	0.06503910	5.7821265	9.859505e-09
##	[7,]	0.31318998	0.3375	0.3	0.33166110	0.05410586	6.1298559	1.264930e-09
##	[8,]	0.20006021	0.2020	0.3	0.28159164	0.06670313	4.2215657	2.647447e-05
##	[9,]	0.32990538	0.3360	0.3	0.23025579	0.05661344	4.0671574	5.134017e-05
##	[10,]	0.29562285	0.2905	0.3	0.28906539	0.05841261	4.9486810	8.766086e-07
##	[11,]	0.44590808	0.4445	0.0	0.09584075	0.05424572	1.7667892	7.756916e-02
##	[12,]	0.36809363	0.3745	0.0	-0.02245388	0.05302784	-0.4234356	6.720687e-01
##	[13,]	0.37662575	0.3720	0.0	0.03703145	0.05527881	0.6699033	5.030744e-01
##	[14,]	0.37629499	0.3815	0.0	-0.07125913	0.05474273	-1.3017095	1.933161e-01
##	[15,]	0.21748927	0.2215	0.0	0.29257288	0.06536792	4.4757871	8.489445e-06
		0.21748927	0.2215	0.0	0.29257288	0.06536792	4.4757871	8.489445e-06
##	[17,]	0.06972117	0.0610	0.0	0.33145758	0.10935747	3.0309551	2.500692e-03
##	[18,]	0.06972117	0.0610	0.0	0.33145758	0.10935747	3.0309551	2.500692e-03
##	[19,]	0.36935781	0.3780	0.0	0.23908858			7.922031e-06
		0.36935781	0.3780	0.0	0.23908858			7.922031e-06
		0.34596068	0.3480	0.0	0.38889542			5.116550e-12
##		0.34596068	0.3480	0.0	0.38889542			5.116550e-12
##		0.16243508	0.1695	0.0	0.30892955			1.308960e-05
##		0.16243508	0.1695	0.0	0.30892955			1.308960e-05
##		0.18502467	0.1995	0.0	0.37606430			9.859505e-09
##		0.18502467	0.1995	0.0	0.37606430			9.859505e-09
##		0.31318998	0.3375	0.0	0.33166110			1.264930e-09
##		0.31318998	0.3375	0.0	0.33166110			1.264930e-09
##		0.20006021	0.2020	0.0	0.28159164			2.647447e-05
##		0.20006021	0.2020	0.0	0.28159164			2.647447e-05
##		0.32990538	0.3360	0.0		0.05661344		5.134017e-05
		0.32990538	0.3360	0.0		0.05661344		5.134017e-05
##	133.1	0.29562285	0.2905	0.0	0.28906539	0.05841261	4.9486810	8.766086e-07

generate.sumstat.plot(ex.sumstat,ex.nsnp)



generate.sumstat.plot(ex.sumstat.noLD,ex.nsnp.noLD)



The true h^2 of our LD model should stay the same, because $\beta = 0$ for all the tagging SNPs.

GWAS-based estimates may be a different story!

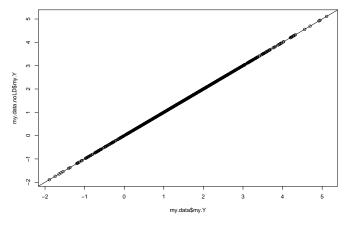
```
Total and SNP h² of the ex.model
# the trait h2
V.G=sum(ex.sumstat[1:ex.nsnp.true,"beta"]^2*(2*ex.sumstat[1:ex.nsnp.true,"MAF"]*(1-ex.sumstat[1:ex.nsnp.true,"MAF"]*(1-ex.sumstat[1:ex.nsnp.true,"defa"]^2*(2*ex.sumstat[1:ex.nsnp.true,"MAF"]*(1-ex.sumstat[1:ex.nsnp.true,"defa"]^2*(2*ex.sumstat[1:ex.nsnp.true,"MAF"]*(1-ex.sumstat[1:ex.nsnp.true,"defa"]^2*(2*ex.sumstat[1:ex.nsnp.true,"MAF"]*(1-ex.sumstat[1:ex.nsnp.true,"defa"]^2*(2*ex.sumstat[1:ex.nsnp.true,"defa"]^2*(2*ex.sumstat[1:ex.nsnp.true,"defa"]^2*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(1-ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(ex.ntag.T)+1)*(2*ex.sumstat[1:ex.nsnp-sum(
```

Move to my.data now

```
my.ntag.T=ex.ntag.T # SAME tagging; no heterogeneity between my. and ex.
my.nsnp.true=10; my.beta.true=0.3; my.maf=ex.sumstat[,"MAF"]
my.nsnp=(5000+sum(my.ntag.T)); my.nsample=1000; my.sigma=1; my.seed=102
my.data=generate.my.data.LD(my.seed,my.nsample,my.nsnp,my.nsnp.true,my.beta.true,my.sigma,my.maf,my.ntag.*
```

Checking our data: my.Y = my.Y.noLD (given the same seed etc.)

plot(my.data\$my.Y,my.data.noLD\$my.Y); abline(0,1)



Tagging SNPs have large GWAS Z and small p-values, but their true $\beta=0$. Thus, they do not affect how we generate Y based on the true model.

(Multiple causal SNPs in LD is a different story!)

##		MAG	MAF.hat	hota	beta.hat	se	Z.value	p.value
##	F1 1	0.21748927	0.2270	0.3		0.06066349		4.755586e-03
##		0.06972117	0.0755	0.3	0.33447059			5.182232e-04
##		0.36935781	0.3555	0.3		0.05230483		1.034940e-09
##	/-	0.34596068	0.3545	0.3	0.25019642			3.121234e-06
##	- /-	0.16243508	0.1530	0.3		0.06958963		4.021553e-06
##		0.18502467	0.1800	0.3		0.06679596		2.978552e-05
##	/ -	0.31318998	0.3045	0.3		0.05517189		8.652840e-11
##	- /-	0.20006021	0.1780	0.3		0.06681630		1.507249e-07
##		0.32990538	0.3300	0.3	0.31052039			3.218822e-09
##	- /-	0.29562285	0.2960	0.3	0.33840898			6.731015e-10
##		0.44590808	0.4465	0.0		0.05043254		3.708637e-01
##		0.36809363	0.3580	0.0	-0.02127391	0.05509213		6.994668e-01
##	[13,]	0.37662575	0.3820	0.0	-0.02435158	0.05358273	-0.4544670	6.495915e-01
##	[14,]	0.37629499	0.3780	0.0	0.01024406	0.05300796	0.1932551	8.467985e-01
##	[15,]	0.21748927	0.3780	0.0	0.17164714	0.06066349	2.8294965	4.755586e-03
##	[16,]	0.21748927	0.3780	0.0	0.17164714	0.06066349	2.8294965	4.755586e-03
##	[17,]	0.06972117	0.3780	0.0	0.33447059	0.09604226	3.4825358	5.182232e-04
##	[18,]	0.06972117	0.3780	0.0	0.33447059	0.09604226	3.4825358	5.182232e-04
##	[19,]	0.36935781	0.3780	0.0	0.32234988	0.05230483	6.1629085	1.034940e-09
##	[20,]	0.36935781	0.3780	0.0	0.32234988	0.05230483	6.1629085	1.034940e-09
##	[21,]	0.34596068	0.3780	0.0	0.25019642	0.05335395	4.6893703	3.121234e-06
##	[22,]	0.34596068	0.3780	0.0	0.25019642	0.05335395	4.6893703	3.121234e-06
##	[23,]	0.16243508	0.3780	0.0	0.32262395	0.06958963	4.6360925	4.021553e-06
##	[24,]	0.16243508	0.3780	0.0	0.32262395	0.06958963	4.6360925	4.021553e-06
##	[25,]	0.18502467	0.3780	0.0	0.28017270	0.06679596	4.1944557	2.978552e-05
##		0.18502467	0.3780	0.0	0.28017270	0.06679596	4.1944557	2.978552e-05
##	[27,]	0.31318998	0.3780	0.0	0.36190034	0.05517189	6.5595060	8.652840e-11
##	[28,]	0.31318998	0.3780	0.0	0.36190034	0.05517189	6.5595060	8.652840e-11
##	[29,]	0.20006021	0.3780	0.0	0.35342514	0.06681630	5.2895046	1.507249e-07
		0.20006021	0.3780	0.0		0.06681630		1.507249e-07
		0.32990538	0.3780	0.0	0.31052039			3.218822e-09
		0.32990538	0.3780	0.0		0.05197947		3.218822e-09
##	<pre>[33.]</pre>	0.29562285	0.3780	0.0	0.33840898	0.05429097	6.2332464	6.731015e-10

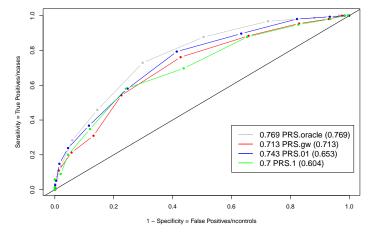
Move to PRS now

We can use the same PRS functions, as PRS construction only depend on ex.sumstat that, which were already generated and checked.

```
# Even the ex.nsnp is slightly bigger now, use the same alpha due to LD
alpha.level=c(0.00001,0.01,0.1)
l.threshold=1
```

the LD my.model, ntag.T=2

```
# generate the ROC plots
my.PRS.output=generate.my.PRS.output(ex.sumstat,my.data,alpha.level,l.threshold)
generate.ROC.plot(my.PRS.output)
```

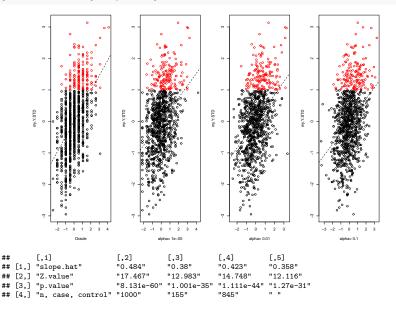


```
## alpha J TP FP
## [1,] 1e-05 18 6 12
## [2,] 1e-02 86 10 76
## [3,] 1e-01 512 10 502
```

the LD my.model, ntag.T=2

##

generate.association.plot(my.PRS.output)



Did we make a mistake?! Recall

$$PRS_{i}^{my.data} = \sum_{i=1}^{J} \hat{eta}_{j}^{external} G_{ij}^{my.data}$$

When $r^2 = 1$ and ntag.T = 2 for all the causal SNPS:

$$\begin{split} PRS_i^{my.LD} &= \sum_{j=1}^{J_{LD}} \hat{\beta}_j^{\text{ex.LD}} G_{ij}^{my.LD} \\ &= 3 \times \sum_{j=1;TP}^{J_{noLD}} \hat{\beta}_j^{\text{ex.noLD}} G_{ij}^{my.noLD} + \sum_{j=1;FP}^{J_{noLD}} \hat{\beta}_j^{\text{ex.noLD}} G_{ij}^{my.noLD} \end{split}$$

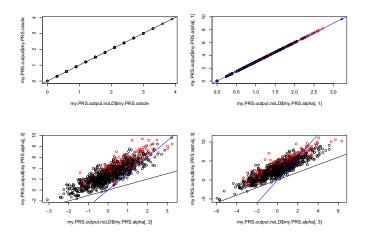
When FP=0

$$PRS_{:}^{my.LD} = 3 \times PRS_{:}^{my.noLD}$$

PRS.oracle should stay the same

$$PRS_{i,oracle}^{my.LD} = \sum_{i=1}^{J=10} \beta_j \cdot G_{ij}^{my.LD} = \sum_{i=1}^{J=10} \beta_j \cdot G_{ij}^{my.noLD} = PRS_{i,oracle}^{my.noLD}, \text{ where } \beta_j = 0.3$$

Check the NON-STD PRS data



In this simplest setting ($r^2=1$ and only causal SNPs are tagged), NOT adjusting LD '=' leveraging LD to improve performance!

Could consider different types of ntag.T, e.g.

ex.nsnp=(5000+sum(ex.ntag.T)); ex.nsample=1000; ex.sigma=1; ex.seed=101

my.ntag.T=ex.ntag.T # SAME tagging; no heterogeneity between my. and ex. my.nsnp.true=10; my.beta.true=0.3; my.maf=ex.sumstat[,"MAF"] mv.nsnp=(5000+sum(mv.ntag.T)): mv.nsamble=1000; mv.sigma=1: mv.seed=102

only one tagging SNP

mu data

```
# ex.ntag.T=rep(1,ex.nsnp.true)

# heterogeneity in the number of tagging SNPs
# ex.ntag.T=c(1,2,3,4,5,6,7,8,9,10)
# ex.ntag.T=c(10,9,8,7,6,5,4,3,2,1)
# ex.ntag.T=c(2,2,1,4,3,10,2,3,7,5)

# my.ntag.T=ex.ntag.T # SAME tagging; no heterogeneity between my. and ex.

Results are not drastically different, as expected, e.g.
# external data
ex.nsnp.true=10; ex.beta.true=0.3
# specify the ntag.T for each of the nsnp.true
ex.ntag.T=c(2,2,1,4,3,10,2,3,7,5)
```

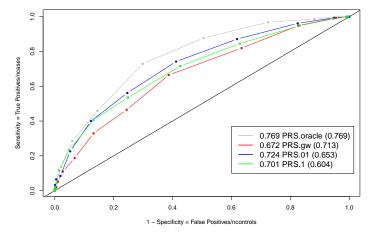
ex.sumstat=generate.ex.sumstat.LD(ex.seed,ex.nsample,ex.nsnp,ex.nsnp.true,ex.beta.true,ex.sigma,ex.ntag.T

mv.data=generate.mv.data.LD(mv.seed.mv.nsample.mv.nsnp.mv.nsnp.true.mv.beta.true.mv.sigma.mv.maf.mv.ntag.

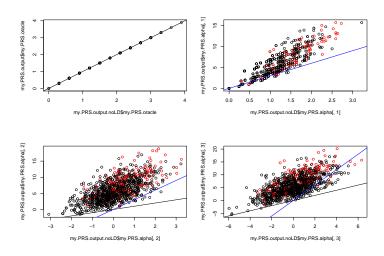
##		MAF	MAF.hat	hota	beta.hat	se	Z.value	p.value
##	F1 1	0.21748927	0.2270	0.3	0.17164714			4.755586e-03
##		0.06972117	0.0755	0.3	0.33447059			5.182232e-04
##		0.36935781	0.3555	0.3	0.32234988			1.034940e-09
##	/-	0.34596068	0.3545	0.3	0.25019642			3.121234e-06
##	- /-	0.16243508	0.1530	0.3	0.32262395			4.021553e-06
##		0.18502467	0.1800	0.3	0.28017270			2.978552e-05
##	/ -	0.31318998	0.3045	0.3	0.36190034			8.652840e-11
##	- /-	0.20006021	0.1780	0.3	0.35342514			1.507249e-07
##		0.32990538	0.3300	0.3	0.31052039			3.218822e-09
##	- /-	0.29562285	0.2960	0.3	0.33840898			6.731015e-10
##		0.44590808	0.4465	0.0	0.04515026			3.708637e-01
		0.36809363	0.3580					6.994668e-01
	- /-	0.37662575	0.3820					6.495915e-01
		0.37629499	0.3780	0.0	0.01024406			8.467985e-01
		0.21748927	0.3780	0.0	0.17164714	0.06066349	2.8294965	4.755586e-03
##	[16,]	0.21748927	0.3780	0.0	0.17164714	0.06066349	2.8294965	4.755586e-03
##	[17,]	0.06972117	0.3780	0.0	0.33447059	0.09604226	3.4825358	5.182232e-04
##	[18,]	0.06972117	0.3780	0.0	0.33447059	0.09604226	3.4825358	5.182232e-04
##	[19,]	0.36935781	0.3780	0.0	0.32234988	0.05230483	6.1629085	1.034940e-09
##	[20,]	0.34596068	0.3780	0.0	0.25019642	0.05335395	4.6893703	3.121234e-06
##	[21,]	0.34596068	0.3780	0.0	0.25019642	0.05335395	4.6893703	3.121234e-06
##	[22,]	0.34596068	0.3780	0.0	0.25019642	0.05335395	4.6893703	3.121234e-06
##	[23,]	0.34596068	0.3780	0.0	0.25019642	0.05335395	4.6893703	3.121234e-06
##	[24,]	0.16243508	0.3780	0.0	0.32262395	0.06958963	4.6360925	4.021553e-06
##	[25,]	0.16243508	0.3780	0.0	0.32262395	0.06958963	4.6360925	4.021553e-06
##	[26,]	0.16243508	0.3780	0.0	0.32262395	0.06958963	4.6360925	4.021553e-06
##	[27,]	0.18502467	0.3780	0.0	0.28017270	0.06679596	4.1944557	2.978552e-05
##	[28,]	0.18502467	0.3780	0.0	0.28017270	0.06679596	4.1944557	2.978552e-05
##	[29,]	0.18502467	0.3780	0.0	0.28017270	0.06679596	4.1944557	2.978552e-05
		0.18502467	0.3780	0.0	0.28017270			2.978552e-05
		0.18502467	0.3780	0.0	0.28017270			2.978552e-05
		0.18502467	0.3780	0.0	0.28017270			2.978552e-05
##	Γ33.1	0.18502467	0.3780	0.0	0.28017270	0.06679596	4.1944557	2.978552e-05

the LD my.model, ntag.T=c(2,2,1,4,3,10,2,3,7,5)

```
# generate the ROC plots
my.PRS.output=generate.my.PRS.output(ex.sumstat,my.data,alpha.level,1.threshold)
generate.ROC.plot(my.PRS.output)
```



```
## alpha J TP FP
## [1,] 1e-05 30 6 24
## [2,] 1e-02 105 10 95
## [3,] 1e-01 531 10 521
```



The effect of adding ntag.F for null SNPs?

To do... Educated guess:

- ► AUC stays the same if ntag.T = ntag.F
- ► AUC drops if ntag.T ≠ ntag.F
- ► AUC drops more significantly if ntag.T < ntag.F
- AUC drops more significantly for less stringent alpha

Some models/methods not discussed

- GxG and GxE interactions
- Bayesian methods
- Rare variants
- ► The X chromosome
- ▶ Pitfalls of standardization (how do we define a case in practice?)
- Many more...

Recap of the learning goal: a deeper understanding of

1. PRS foundation: GWAS, h2 and prediction

- the multiple hypothesis testing issue inherent in GWAS
- the (high) variability inherent in the h2 estimates
- ▶ h2 as a function of both genetic effect beta and MAF
- ▶ the 'genetic effect size' of a SNP as a function of beta and MAF
- ▶ a conceptual PRS construction based on the ground truth, PRS.oracle
- ▶ DIY ROC plotting and AUC calculation for a PRS-based prediction

2. PRS basic: PRS calculation and performance evaluation

- ▶ the complexity of constructing a good PRS even under the simplest setting without LD or any heterogeneities; 10 out 5000 independent SNPs are truly associated with the same effect size of 0.3 but varying MAFs.
- the trouble introduced by false positives, due to multiple hypothesis testing and low power.
- 'the more is not always better' statement: PRS based on 6 'genome-wide' significant SNPs vs. 66 SNPs significant at 0.01.
- the various over-fitting or selection biases, and winner's curse in beta estimates for both false positives and true positives.

Learning goal Cont'd, a deeper understanding of

3. PRS basic-plus: some obvious or not so obvious follow-up Qs

- Effects of ex.nsample and ex.beta.true on AUC: easy to answer.
- Answers to these Qs are less obvious: If we decrease ex.beta.true from 0.3 to 0.1 but increase ex.nsnp.true from 10 to 90,

 h^2 and SNP h^2 ? AUC in general? AUC between PRS.gw and PRS.01?

4. PRS heterogeneity and transportability

- First, why reference allele (genome build) matching is so consequential
- Then, population and locus heterogeneity including my.maf ≠ ex.maf my.beta.true ≠ ex.beta.true my.nsnp.true ≠ ex.nsnp.true

5. PRS LD consideration

Some basic understanding of our limited understanding of LD.

End of the (hopefully fun) ride!