

## Some of my white board writing in week 10 ①

1.) Regarding the multinomial distribution:

I explained the formula

$$P(X_1=x_1, X_2=x_2, \dots, X_k=x_k) = \frac{n!}{x_1! x_2! \dots x_k!} w_1^{x_1} w_2^{x_2} \dots w_k^{x_k},$$

$0 < w_i < 1$ ,  $\sum_{i=1}^k w_i = 1$  for calculating the probability of a particular outcome  $\begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ x_k \end{pmatrix}$  with  $x_1 + x_2 + \dots + x_k = n$ ,  $n$  being the number of independent trials.

I also discussed two examples:

i) If a die is tossed 6 times, what is the probability that each number (1, 2, 3, 4, 5, 6) turns up once.

Applying the above formula with  $k=6$ ,  $x_i=1$ ,  $i=1, 2, \dots, 6$ , and  $w_i = \frac{1}{6}$ ,  $i=1, 2, \dots, 6$  we get  $6! \left(\frac{1}{6}\right)^6 = \frac{5}{324}$

ii) Out of 7 tosses, what is the probability that each number (1, 2, 3, 4, 5, 6) turns up at least once.

Answer:  $6 \cdot \frac{7!}{2! (1!)^6} \left(\frac{1}{6}\right)^7 = \frac{35}{648}$

2.) I discussed in detail the proof of Theorem 7.2. (p. 60) but I see that all details are presented in the lecture note so I will abstain from reproducing them again here.

3.) I discussed a simple method to derive the density of the  $r$ -th order statistic as stated in Theorem 7.3, p. 61

(2)

$Y = \{ \text{number of realisations } X_1, X_2, \dots, X_n \text{ that happen to be } \leq x \}$   
Then  $Y \sim \text{Bin}(n, F_X(x))$ .

Now we first derive  $F_{X_{(n)}}(x)$  (the cdf of  $X_{(n)}$ ) and then differentiate it to find the density. The main observation we make is that

$$F_{X_{(r)}}(x) = P(X_{(r)} \leq x) = P(Y \geq r)$$

Hence we can state that

$$F_{X(r)}(x) = \sum_{k=r}^n \binom{n}{k} (F_X(x))^k (1 - F_X(x))^{n-k}$$

Now to get the density we need to differentiate each of the summands in  $\sum_{k=r}^n$  by applying the  $(uv)' = u'v + v'u$  formula each time.

We get:

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$$f_{X(r)}(x) = \binom{n}{r} r f_X(x) F_X^{r-1}(x) (1-F_X(x))^{n-r} - \binom{n}{r} (n-r) F_X^r(x) (1-F_X(x))^{n-r-1} f_X(x)$$

$$+ \binom{n}{r+1} (r+1) f_X(x) F_X^r(x) (1-F_X(x))^{n-r-1} - \binom{n}{r+1} (n-r) F_X^{r+1}(x) (1-F_X(x))^{n-r-2} f_X(x)$$

$$+ \binom{n}{r+2} (r+2) f_X(x) F_X^{r+1}(x) (1-F_X(x))^{n-r-2} - \binom{n}{r+2} (n-r-1) F_X^{r+2}(x) (1-F_X(x))^{n-r-3} f_X(x)$$

$$+ \binom{n}{r+3} (r+3) f_X(x) F_X^{r+2}(x) (1-F_X(x))^{n-r-3} - \binom{n}{r+3} (n-r-2) F_X^{r+3}(x) (1-F_X(x))^{n-r-4} f_X(x)$$

$$+ \binom{n}{n} (n-n) f_X(x) F_X^{n-1}(x) (1-F_X(x))^{n-n} = 0$$

Huge cancellation happens and, because of the equality  $\binom{n}{r} (n-r) = \binom{n}{r+1} (r+1)$  each of the summands after the first one disappears. Hence

$$f_{X(r)}(x) = \frac{n!}{(r-1)!(n-r)!} f_X(x) F_X(x)^{r-1} (1-F_X(x))^{n-r} \text{ holds}$$

4) I also discussed the idea of the proof of Theorem 7.4 on p. 62. Again, we first get the cdf and then find the mixed partial derivative

$$\frac{\partial^2}{\partial u \partial v} F_{X_{(i)}, X_{(j)}}(u, v) \text{ to calculate the density } f_{X_{(i)}, X_{(j)}}(u, v).$$

With the discrete variables  $U$  and  $V$  as introduced on p. 62 we see that

$$(U, V, n - U - V) \sim \text{Multinomial}(n; \frac{F(u)}{X}, \frac{F(v) - F(u)}{X}, 1 - \frac{F(v)}{X})$$

Then we observe that

$$F_{X_{(i)}, X_{(j)}}(u, v) = P(U \geq i \cap U + V \geq j) =$$

$$= \sum_{k=i}^{j-1} \sum_{m=j-k}^{n-k} P(U=k, V=m) + P(U \geq j)$$

Since the second summand does not involve  $v$ , its mixed partial derivative w.r.  $u$  and  $v$  will be zero and hence

$$f_{X_{(i)}, X_{(j)}}(u, v) = \frac{\partial^2}{\partial u \partial v} \sum_{k=i}^{j-1} \sum_{m=j-k}^{n-k} \frac{n!}{k! m! (n-k-m)!} \left( \frac{F(u)}{X} \right)^k \left( \frac{F(v) - F(u)}{X} \right)^m \left( 1 - \frac{F(v)}{X} \right)^{n-k-m}$$

Again, a huge cancellation happens when we calculate the partial derivatives by using the product rule for differentiation and we end up with

$$f_{X_{(i)}, X_{(j)}}(u, v) = \frac{n!}{(i-1)!(j-1-i)!(n-j)!} f_X(u) f_X(v) \left( \frac{F(u)}{X} \right)^{i-1} \left( \frac{F(v) - F(u)}{X} \right)^{j-1-i} \left( 1 - \frac{F(v)}{X} \right)^{n-j}$$

for  $n \geq j > i \geq 1$ ,  $-\infty < u < v < \infty$  (and = 0 else)



(5) I also discussed in detail the example stating that (4)  
 for the range  $R = X_{(n)} - X_{(1)}$  for order statistic  
 $X_{(1)} < X_{(2)} < \dots < X_{(n)}$  from the uniform  $(0,1)$   
 distribution, it holds  $f_R(u) = \begin{cases} n(n-1)u^{n-2}(1-u), & 0 < u < 1 \\ 0 & \text{else} \end{cases}$

Proof:

To this end, first we note that by using the formula  
 from Theorem 7.4 we have (with  $i=1$  and  $j=n$ ):

$$f_{X_{(1)}, X_{(n)}}(x, y) = n(n-1)(F_X(y) - F_X(x))^{n-2} f_X(x) f_X(y)$$

We introduce the variable of interest

$U = X_{(n)} - X_{(1)}$  and one auxiliary variable

$V = X_{(n)}$  so that we could apply

the density transformation formula

$$f_{\begin{pmatrix} U \\ V \end{pmatrix}}(u, v) = f_{\begin{pmatrix} X_{(1)} \\ X_{(n)} \end{pmatrix}}(x(u, v), y(u, v)) \left| \begin{vmatrix} \frac{\partial x}{\partial u} & \frac{\partial x}{\partial v} \\ \frac{\partial y}{\partial u} & \frac{\partial y}{\partial v} \end{vmatrix} \right|$$

Note:  $\begin{cases} X_{(1)} = V - U =: x \\ X_{(n)} = V =: y \end{cases}$  and  $\left| \begin{vmatrix} \frac{\partial x}{\partial u} & \frac{\partial x}{\partial v} \\ \frac{\partial y}{\partial u} & \frac{\partial y}{\partial v} \end{vmatrix} \right| = \left| \begin{vmatrix} -1 & 1 \\ 0 & 1 \end{vmatrix} \right| = 1$

Hence  $f_{\begin{pmatrix} U \\ V \end{pmatrix}}(u, v) = n(n-1)(F_X(v) - F_X(v-u))^{n-2} f_X(v-u) f_X(v) \times 1$

To get  $f_U(u)$  (which we are interested in) we need to  
 integrate out the unwanted variable  $V$  from  
 the joint density  $f_{\begin{pmatrix} U \\ V \end{pmatrix}}(u, v)$ . We need to be care-  
 ful with the integration range when doing this:

(5)

Since  $0 < X_{(1)} < X_{(n)} < 1$  we get

$$0 < V - U < V < 1$$

$$0 < U < V < 1$$
 This means that for

a fixed  $u$ ,  $V$  ranges in the interval  $(u, 1)$ .

Therefore:

$$\begin{aligned} f_R(u) &= \int_u^1 f(u, v) dv = \int_u^1 n(n-1) (1-u)^{n-2} dv \\ &= \begin{cases} n(n-1)u^{n-2}(1-u) & \text{if } 0 < u < 1 \\ 0 & \text{else} \end{cases} \end{aligned}$$

I also advised you to repeat this exercise by using

$X_{(1)} = V$  as an auxiliary variable. The intermediate calculations will be slightly different but at the end after integrating out  $V$  again (BUT THIS TIME in the range  $(0, 1-u)$  (!)) you will get the same final result for the density  $f_R(u)$ .

6.) I also discussed one more problem in class

(7/d) from tutorial sheet 4) but because it is completely solved in the solutions to tutorial set 4, I abstain from reproducing the derivation here.