

A power and prediction analysis for knockoffs with Lasso statistics

By Weinstein, Barber, Candes

Mengqi Lin

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Overview

1. Recall the powerful theorem we learned in class.
2. Introduce the knockoff procedure which is aimed for dealing with the Lasso statistics in practice.
3. Show you how the result in 1 can be applied in 2.

Recall the beautiful result we learned in class

Under our "old friend" setting : Consider linear model

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{z}$$

where $\mathbf{X} \in \mathbb{R}^{n \times p}$, $X_{ij} \sim_{i.i.d.} \mathcal{N}(0, 1/n)$, $\mathbf{z} \in \mathbb{R}^n$, $z_i \sim_{i.i.d.} \mathcal{N}(0, \sigma^2)$,
 $\boldsymbol{\beta} \in \mathbb{R}^p$, $\beta_j \sim_{i.i.d.} \Pi$, where $\mathbb{P}(\Pi \neq 0) = \epsilon$, $\mathbb{P}(\Pi = 0) = 1 - \epsilon$
 $n/p \rightarrow \delta$, $n \rightarrow \infty$.

So the expected number of nonzero (nonnull) is ϵp .

The lasso estimator :

$$\hat{\beta}(\lambda) = \underset{\mathbf{b} \in \mathbb{R}^p}{\operatorname{argmin}} \frac{1}{2} \|\mathbf{y} - \mathbf{X}\mathbf{b}\|^2 + \lambda \|\mathbf{b}\|_1$$

Let $V(\lambda) = |\{j : \hat{\beta}_j(\lambda) \neq 0, \beta_j = 0\}|$, $T(\lambda) = |\{j : \hat{\beta}_j(\lambda) \neq 0, \beta_j \neq 0\}|$,
 $R(\lambda) = |\{j : \hat{\beta}_j(\lambda) \neq 0\}| = V(\lambda) + T(\lambda)$, $k = |\{j : \beta_j \neq 0\}| = \epsilon p$

Lemma (Lec11, 12, AMP)

$$\frac{V(\lambda)}{p} \xrightarrow{\mathbb{P}} 2(1 - \epsilon)\Phi(-\alpha),$$

$$\frac{T(\lambda)}{p} \xrightarrow{\mathbb{P}} \mathbb{P}(|\Pi + \tau W| > \alpha\tau, \Pi \neq 0) = \epsilon \mathbb{P}(|\Pi^* + \tau W| > \alpha\tau),$$

where $W \sim \mathcal{N}(0, 1)$ is indep of Π , and $\tau, \alpha > \max\{\alpha_0, 0\}$ is the unique solution to

$$\begin{aligned} \tau^2 &= \sigma^2 + \frac{1}{\delta} \mathbb{E}(\eta_{\alpha\tau}(\Pi + \tau W) - \Pi)^2 \\ \lambda &= \left(1 - \frac{1}{\delta} \mathbb{P}(|\Pi + \tau W| > \alpha\tau)\right) \alpha\tau. \end{aligned} \tag{1}$$

It follows immediately from Lemma 1 that for a fixed $\lambda > 0$, the limits of FDP and TPP are

$$\text{FDP}(\lambda) = \frac{V(\lambda)}{1 \vee R(\lambda)} \xrightarrow{\mathbb{P}} \frac{2(1 - \epsilon)\Phi(-\alpha)}{2(1 - \epsilon)\Phi(-\alpha) + \epsilon\mathbb{P}(|\Pi^* + \tau W| > \alpha\tau)} \quad (2)$$

and

$$\text{TPP}(\lambda) = \frac{T(\lambda)}{1 \vee k} \xrightarrow{\mathbb{P}} \mathbb{P}(|\Pi^* + \tau W| > \alpha\tau) \quad (3)$$

philosophy question

We don't know ϵ and Π^* in practice !

In practice

How do scientists decide the threshold λ for the Lasso under FDR-TPP criteria?

Knockoff procedure (Candes, Barber, 2015)

Idea: For each λ , propose an estimator for $\text{FDP}(\lambda)$: $\widehat{\text{FDP}}$ (possibly overestimate FDP), and greedily choose λ so that the procedure stops when $\widehat{\text{FDP}} < q$. But to estimate FDP, we need to estimate $V(\lambda)$, which is unknown, and this is where the knockoff variable plays the role.

In this section, we work on fixed p, n . And augment the matrix \mathbf{X} with $\tilde{\mathbf{X}} \in \mathbb{R}^{n \times r}$

$$\mathbb{X} := [\mathbf{X} \tilde{\mathbf{X}}] \in \mathbb{R}^{n \times (p+r)}$$

where the whole $\mathbb{X}_{ij} \sim_{i.i.d.} \mathcal{N}(0, 1/n)$, and the corresponding Lasso estimator

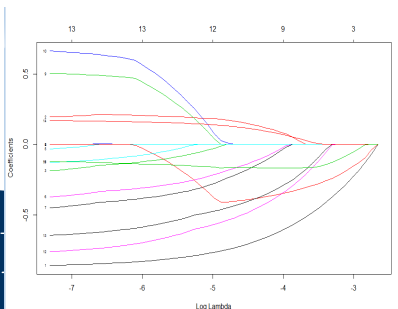
$$\hat{\boldsymbol{\beta}}(\lambda) = \underset{\mathbf{b} \in \mathbb{R}^{p+r}}{\operatorname{argmin}} \frac{1}{2} \|\mathbf{y} - \mathbb{X}\mathbf{b}\|^2 + \lambda \|\mathbf{b}\|_1. \quad (4)$$

Denote sets:

$$\mathcal{H} = \{1, \dots, p\}, \quad \mathcal{H}_0 = \{j \in \mathcal{H} : \beta_j = 0\}, \quad \mathcal{K}_0 = \{p+1, \dots, p+r\}$$

To effectively decide λ , we work on the sufficient statistics for Lasso

$$\tau_j = \sup\{\lambda : \hat{\beta}_j(\lambda) \neq 0\}, \quad j = 1, \dots, p+r. \quad (5)$$



Let

$$v_0(\lambda) = |\{j \in \mathcal{H}_0 : T_j \geq \lambda\}|, \quad v_1(\lambda) = |\{j \in \mathcal{K}_0 : T_j \geq \lambda\}|, \quad R(\lambda) = |\{j \in \mathcal{H} : T_j \geq \lambda\}|$$

$$\widehat{\text{FDP}}(\lambda) = \frac{(1 + v_1(\lambda)) \cdot \frac{|\mathcal{H}| \hat{\pi}_0}{1 + |\mathcal{K}_0|}}{R(\lambda)}$$

And the procedure finds the greedy λ s.t. $\widehat{\text{FDP}}(\lambda) \leq q$, i.e

$$\hat{\lambda} = \inf \{ \lambda \in \Lambda : \widehat{\text{FDP}}(\lambda) \leq q \}$$

and reject $j \in \mathcal{H}$, if $T_j \geq \hat{\lambda}$.

Optimality of knockoff procedure

Let's take $r = \rho p$ for a constant $\rho > 0$, we have that $n/(p + r) \rightarrow \delta' := \delta/(1 + \rho)$, $\epsilon' := \epsilon/(1 + \rho)$ as $n, p \rightarrow \infty$. Notice that our linear model has become:

$$\mathbb{E}\mathbf{y} = \mathbb{X} \begin{bmatrix} \beta \\ \mathbf{0} \end{bmatrix}$$

As the similar analysis in section 1, we can derive the oracle FDP versus TPP for each λ in this augment setting with a slight modification. With that being said, when knowing ϵ , Π^* , we already know that there's an oracle "procedure".

$$\text{FDP}_{\text{aug}}(\lambda) \longrightarrow \frac{2(1 - \epsilon)\Phi(-\alpha')}{2(1 - \epsilon)\Phi(-\alpha') + \epsilon\mathbb{P}(|\Pi^* + \tau'W| > \alpha'\tau')}$$

$$\text{TPP}_{\text{aug}}(\lambda) \longrightarrow \mathbb{P}(|\Pi^* + \tau'W| > \alpha'\tau')$$

simulation figure

