

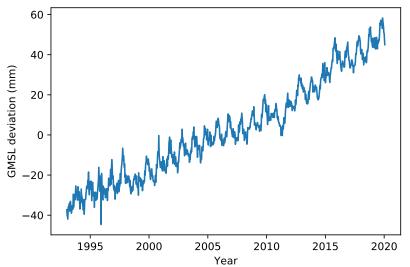
## **Time Series and Sequence Learning**

Lecture 4 - Classical Decomposition and Kalman Filter

Johan Alenlöv, Linköping University 2020-09-09

#### Summary

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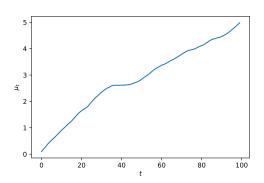
For today we will talk about

- 1. The classical decomposition
- 2. Modelling the trend
- 3. Kalman Filter
- 4. Calculating the likelihood and estimating parameters
- 5. Forecasting and missing data

Our goal for this lecture is to model a **time-series**  $y_1, y_2, \ldots y_n$  by decomposing it into unobserved stochastic processes. The three parts we will decompose the model into is the **trend**, **seasonal**, and **noise**.

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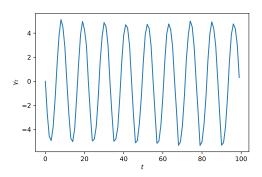
 $\underline{\mu}_t$  The **trend** component of the time-series.



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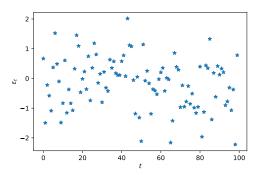
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- $\gamma_t$  The seasonal component of the time-series.
- $\varepsilon_t$  The **noise** component of the time-series.

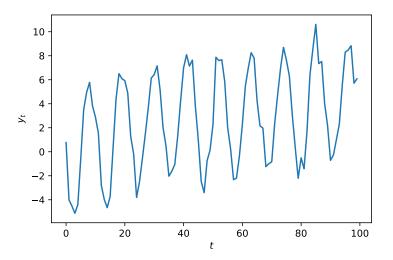


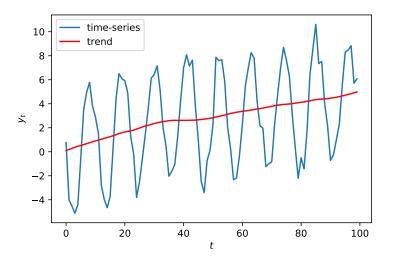
## Time Series Modelling (cont.)

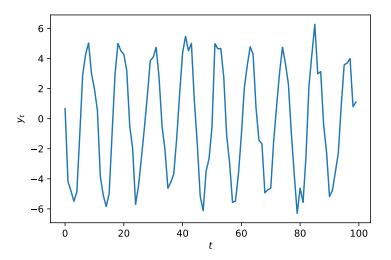
There are two main ways of combining these parts.

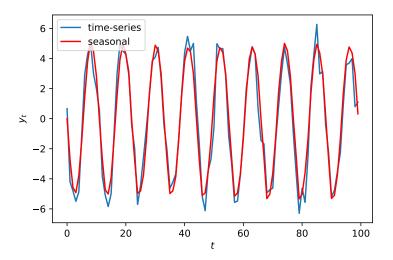
1. The most common way is an additive model where,

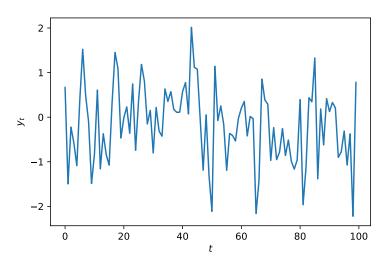
$$y_t = \underbrace{\mu_t}_{\text{trend}} + \underbrace{\gamma_t}_{\text{seasonal}} + \underbrace{\varepsilon_t}_{\text{noise}}.$$

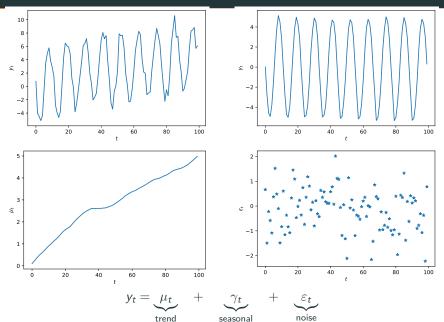












## Time Series Modelling (cont.)

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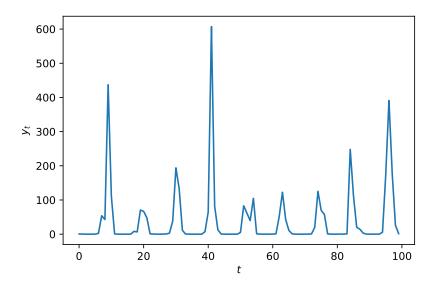
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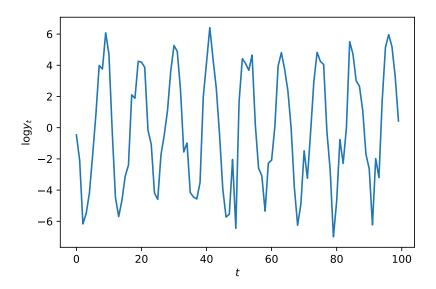
2. The other way is a multiplicative model where,

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$$y_t = \underbrace{\mu_t}_{\text{trend}} \quad \times \quad \underbrace{\gamma_t}_{\text{seasonal}} \quad \times \quad \underbrace{\varepsilon_t}_{\text{noise}}.$$

As we could see, it is often possible to look at the **logarithm** of the multiplicative model and get an additive model

$$\log y_t = \underbrace{\log \mu_t}_{\text{trend}} + \underbrace{\log \gamma_t}_{\text{seasonal}} + \underbrace{\log \varepsilon_t}_{\text{noise}}.$$

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- Focus on modelling of the **trend**  $\mu_t$  today.
- Recall the AR model of order p from previous lecture,

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Will start with a simple random walk.

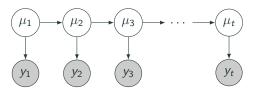
Simple Random Walk: The stochastic process  $\mu_1, \mu_2, \ldots$  is a simple random walk if  $\mu_{t+1} = \mu_t + \eta_t$  where  $\eta_t \stackrel{\text{iid}}{\sim} \mathcal{N}(0, \sigma_\eta^2)$ 

### **Basic State-Space Model**

The Local Level Model Two stochastic processes  $y_1, y_2, y_3,...$  and  $\mu_1, \mu_2, \mu_3,...$ 

$$\underline{y_t} = \mu_t + \varepsilon_t, \qquad \qquad \varepsilon_t \stackrel{\text{iid}}{\sim} \mathcal{N}(0, \sigma_{\varepsilon}^2) \\
\underline{\mu_{t+1}} = \mu_t + \eta_t, \qquad \qquad \underline{\eta_t \stackrel{\text{iid}}{\sim} \mathcal{N}(0, \sigma_{\eta}^2)} \\
\underline{\mu_1} \sim \mathcal{N}(a_1, P_1)$$

This is a **state-space model (SSM)**, where  $y_t$  is the **observed** time-series and  $\mu_t$  is the **unobserved** process.

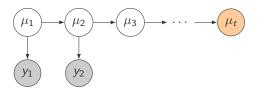


## Filtering, Smoothing, and Predicting

Given a time-series  $y_{1:n} = (y_1, y_2, \dots, y_n)$  we wish to calculate the distribution of  $\mu_t$  conditioned on the observed time-series  $y_{1:n}$ .

This problem changes depending on the relationship of n and t.

n < t: This is known as the **forecasting** problem.



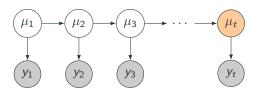
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## Filtering, Smoothing, and Predicting

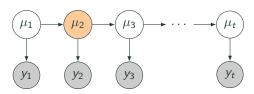
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n < t: This is known as the **forecasting** problem.

n = t: This is known as the **filtering** problem.

n > t: This is known as the **smoothing** problem.



### Calculating the Filter Distribution

We will now focus on finding the **filter distribution** and **one step ahead predictor**.

That is the distribution of  $\underline{\mu_t} \mid \underline{y_{1:t}}$  and  $\underline{\mu_t} \mid \underline{y_{1:t-1}}$  in the local-level model with Gaussian noise.

Let 
$$\hat{\mu}_{i|j} = \mathbb{E}[\underbrace{\mu_{i}^{j}}_{\mathbf{z}_{i}}]\underbrace{y_{1:j}}_{\mathbf{z}_{i}}$$
 and  $P_{i|j} = \operatorname{Var}[\mu_{i} | y_{1:j}]$ .

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Will proceed as follows:

- 1. Assume that  $\mu_t | y_{1:t-1} \sim \mathcal{N}(\hat{\mu}_{t|t-1}, P_{t|t-1})$ .
- 2. Show that  $\mu_t \mid y_{1:t} \sim \mathcal{N}(\hat{\mu}_{t|t}, P_{t|t})$
- 3. Show that  $\mu_{t+1} | y_{1:t} \sim \mathcal{N}(\hat{\mu}_{t+1|t}, P_{t+1|t})$

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And find the expressions for the mean and variance of the Gaussian distributions.

#### The Kalman Filter

#### Kalman Filter for local-level model

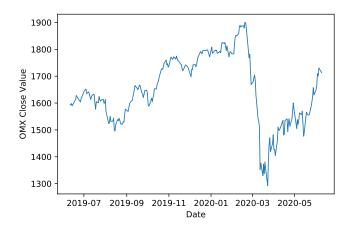
For each iteration  $t = 1, 2, 3, \dots$  repeat the following steps:

- Measurement updates
  - 1. Forecasting error:  $v_t = y_t \hat{\mu}_{t|t-1}$
  - 2. Forecasting variance:  $\underline{F_t} = P_{t|t-1} + \sigma_{\varepsilon}^2$
  - 3. Kalman gain:  $K_t = P_{t \mid t-1}/F_t$
  - 4. Filter mean:  $\hat{\mu}_{t \mid t} = \hat{\mu}_{t \mid t-1} + K_t v_t$
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- Time updates
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Initialized using  $\hat{\mu}_{1|0} = a_1$  and  $P_{1|0} = P_1$ .

## ex) The Kalman Filter

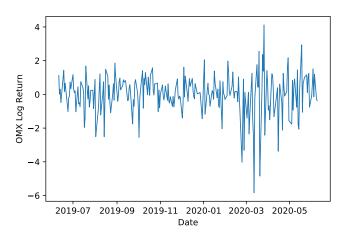
We look at the time-series  $S_1, \ldots, S_{248}$  which are the **OMXS30** closing values over a year.



## ex) The Kalman Filter

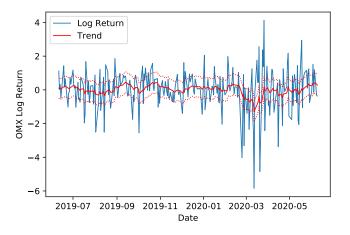
Typically we look at the log-returns

$$y_t = 100 \cdot \log \left( \frac{S_t}{S_{t-1}} \right).$$



## ex) The Kalman Filter

We run the Kalman filter using  $\hat{\mu}_1=0, P_1=0.1, \sigma_{\varepsilon}^2=1, \text{ and } \sigma_{\eta}^2=0.01$ 



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More commonly we use the log-likelihood,

$$\ell(\boldsymbol{\theta}) = \log L(\boldsymbol{\theta}) = \log p_{\boldsymbol{\theta}}(y_1) + \sum_{t=2}^{n} \log p_{\boldsymbol{\theta}}(y_t \mid y_{1:t-1}).$$

#### Parameter Estimation Using Kalman Filter

From the derivation of the Kalman filter we have that:

$$y_t | y_{1:t-1} \sim \mathcal{N}(\hat{\mu}_{t|t-1}(\boldsymbol{\theta}), \mathcal{F}_t(\boldsymbol{\theta})),$$

given that we get that the components of the log-likelihood is calculated by

$$\log p_{\theta}(y_t | y_{1:t-1}) = -\frac{1}{2} \left( \log 2\pi + \log F_t(\theta) + \frac{(y_t - \hat{\mu}_{t | t-1}(\theta))^2}{F_t(\theta)} \right).$$

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The log-likelihood for the local-level model is given by

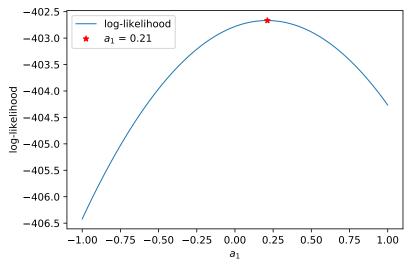
$$\ell(\boldsymbol{\theta}) = -\frac{n}{2}\log 2\pi - \frac{1}{2}\sum_{t=1}^{n} \left(\log F_t(\boldsymbol{\theta}) + \frac{(y_t - \hat{\mu}_{t \mid t-1}(\boldsymbol{\theta}))^2}{F_t(\boldsymbol{\theta})}\right)$$

## ex) Estimating Initial Parameters

We now look at estimating  $\underline{a_1}$  while keeping the other parameters fixed in the OMXS30 log returns.

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- Direct calculations give

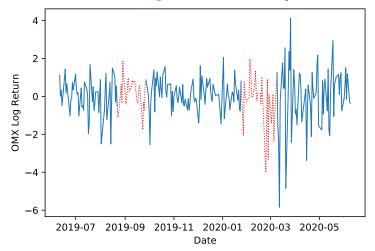
$$\begin{split} \hat{\mu}_{t+j\,|\,t} &= \mathbb{E}[\mu_{t+j}\,|\,y_{1:t}] = \mathbb{E}[\mu_t + \sum_{i=0}^{j-1} \eta_{t+i}\,|\,y_{1:t}] = \hat{\mu}_{t\,|\,t}, \\ P_{t+j\,|\,t} &= \mathrm{Var}[\mu_{t+j}\,|\,y_{1:t}] = \mathrm{Var}[\mu_t + \sum_{i=0}^{j-1} \eta_{t+i}\,|\,y_{1:t}] \\ &= P_{t\,|\,t} + \sum_{i=0}^{j-1} \sigma_{\eta}^2 = P_{t\,|\,t} + j \cdot \sigma_{\eta}^2. \end{split}$$

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- Assume that we are **missing** a set of observations  $y_{\tau:\tau^*}$ .
- Running the <u>Kalman filter</u> for these <u>missing observations</u> will be like <u>forecasting</u>.

For  $t = \tau, \ldots, \tau^*$ :

$$\begin{split} \hat{\mu}_{t\,|\,t} &= \hat{\mu}_{t\,|\,\tau-1} = \hat{\mu}_{\tau\,|\,\tau-1}, \\ \hat{\mu}_{t+1\,|\,t} &= \hat{\mu}_{t+1\,|\,\tau-1} = \hat{\mu}_{\tau\,|\,\tau-1}, \\ P_{t\,|\,t} &= P_{t\,|\,\tau-1} = P_{\tau\,|\,\tau-1} + (t-\tau)\sigma_{\eta}^{2}, \\ P_{t+1\,|\,t} &= P_{t+1\,|\,\tau-1} = P_{\tau\,|\,\tau-1} + (t-\tau+1)\sigma_{\eta}^{2}. \end{split}$$

## The Kalman Filter Again

#### Kalman Filter for local-level model

For each iteration  $t = 1, 2, 3, \dots$  repeat the following steps:

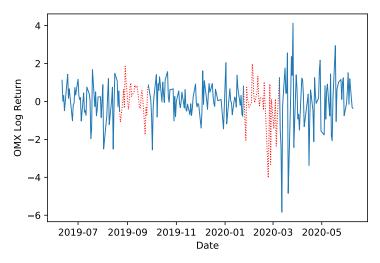
If  $y_t$  is missing, set  $\hat{\mu}_{t\mid t} = \underline{\hat{\mu}_{t\mid t-1}}$ ,  $P_{t\mid t} = \underline{P_{t\mid t-1}}$  and skip to the prediction updates

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Initialized using  $\hat{\mu}_{1\,|\,\underline{0}} = a_1$  and  $P_{1\,|\,0} = P_1$ .

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We again look at the OMXS30 log-returns. We remove some of the observations from the time-series and again run the Kalman filter on this data.



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