## Cheatsheet

Leo

## 2020/12/27

## Read files(txt)

header: column name row.names: which column to be the row names : space

```
# bp \leftarrow read.table ("blood_pressure.txt", header = T, row.names = 1, sep='\t')
```

#### Get vector

continuous vector

```
a <- seq(1,20,1)
a
```

```
## [1] 1 2 3 4 5 6 7 8 9 10 11 12 13 14 15 16 17 18 19 20
```

```
b \le seq(1,20,0.5)
## [1] 1.0 1.5 2.0 2.5 3.0 3.5 4.0 4.5 5.0 5.5 6.0 6.5 7.0 7.5 8.0
## [16] 8.5 9.0 9.5 10.0 10.5 11.0 11.5 12.0 12.5 13.0 13.5 14.0 14.5 15.0 15.5
## [31] 16.0 16.5 17.0 17.5 18.0 18.5 19.0 19.5 20.0
replicate
c <- replicate(10,mean(rnorm(10)))</pre>
## [1] -0.56258069 -0.42971760 0.32115199 0.09273652 0.10012562 0.08796602
## [7] -0.45215031 -0.18440508 0.21596631 0.33255610
d \leftarrow rep(1,10)
## [1] 1 1 1 1 1 1 1 1 1 1
date frame
trial <- read.csv("drug_trial.csv")</pre>
summary(trial)
##
   treatment
                          pain
                   Min. :-1.6485
## Length:44
## Class:character 1st Qu.: 0.5069
## Mode :character Median : 0.9995
##
                      Mean : 1.0065
##
                      3rd Qu.: 1.5582
                      Max. : 2.5704
##
nrow(trial)
## [1] 44
ncol(trial)
## [1] 2
trail <- trial[-c(4),] # remove 4th column</pre>
unique(trial$treatment) # a list of unique value
## [1] "placebo" "drugA" "drugB"
```

#### sample(trial,2,replace=FALSE) # sample for columns

```
##
      treatment
                       pain
## 1
        placebo 2.35493558
## 2
        placebo
                1.38270464
## 3
        placebo 0.71267611
## 4
        placebo 0.40831384
## 5
        placebo 1.69831708
## 6
        placebo 1.41274251
## 7
        placebo 2.28024592
## 8
        placebo -0.20054178
## 9
        placebo 1.96714256
## 10
        placebo 2.57041769
## 11
        placebo 2.30697609
## 12
        placebo 0.90166190
## 13
        placebo 0.93660510
## 14
        placebo 1.93873063
## 15
        placebo 0.91111230
## 16
          drugA 0.41230740
## 17
          drugA 1.56798788
## 18
          drugA 1.24273074
## 19
          drugA 1.07840196
## 20
          drugA 1.55094706
## 21
          drugA 1.06230197
## 22
          drugA 2.52645539
## 23
          drugA 1.13096568
## 24
          drugA 1.44379791
## 25
          drugA 1.79646637
## 26
          drugA 0.53843290
## 27
          drugA 1.55498307
## 28
          drugA 0.32647057
## 29
          drugA 1.74414345
## 30
          drugA 1.38454621
## 31
          drugB
                0.69245923
## 32
          drugB 0.19226536
## 33
          drugB -1.64851266
## 34
          drugB 0.07668351
## 35
          drugB -0.27024559
## 36
          drugB -0.32047823
## 37
          drugB 0.91721163
## 38
          drugB
                0.89274028
## 39
          drugB 0.65363774
## 40
          drugB -0.24564002
## 41
          drugB 0.64731693
## 42
          drugB 0.62837703
## 43
          drugB -0.15360754
## 44
          drugB 1.28126479
```

#### sample\_n(trial,2,replace=FALSE) # sample for rows

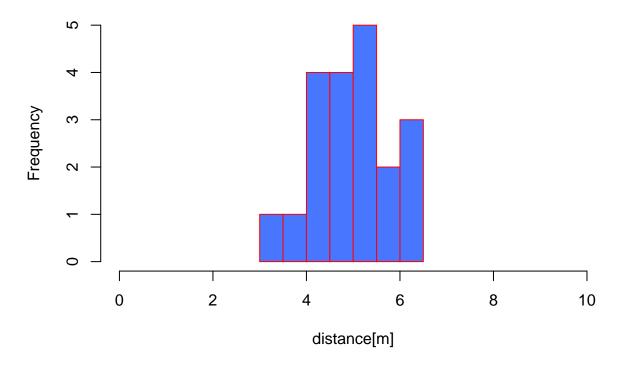
```
## treatment pain
## 1 drugA 1.7441434
## 2 drugA 0.3264706
```

```
# Create data frame
data1 \leftarrow data.frame(x1=c(1,3,5,7),x2=c(2,4,6,8),x3=c(11,12,13,14),x4=c(15,16,17,18))
##
    x1 x2 x3 x4
## 1 1 2 11 15
## 2 3 4 12 16
## 3 5 6 13 17
## 4 7 8 14 18
# add a row
row \leftarrow c(1,1,1,1)
data2 <- rbind(data1[1:2,], row, data1[3:nrow(data1),]) # row bind</pre>
data2
     x1 x2 x3 x4
##
## 1 1 2 11 15
## 2 3 4 12 16
      1 1 1 1
## 31 5 6 13 17
## 4
     7 8 14 18
# add a column
y < -1:4
data2 <- cbind(data1[,1:2],y,data1[,3:ncol(data1)]) # column bind</pre>
data2
##
   x1 x2 y x3 x4
## 1 1 2 1 11 15
## 2 3 4 2 12 16
## 3 5 6 3 13 17
## 4 7 8 4 14 18
# column names
names(data2)
## [1] "x1" "x2" "y" "x3" "x4"
# row names
row.names(data2)
## [1] "1" "2" "3" "4"
```

## Some useful tools to plot the data

looking at the distribution, use histogram

# distance(along the street)[m]



if two vectors, combine them into a data frame with index.

```
response <- c(a,n)
index <- c(rep("esc",20),rep("casc",20))
dataset <- data.frame(response,index)
library(knitr)</pre>
```

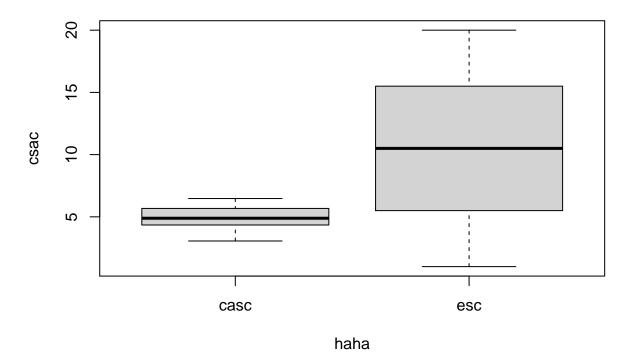
## Warning: package 'knitr' was built under R version 4.0.4

```
kable(dataset[1:5,])
```

response	index
1	esc
2	$\operatorname{esc}$
3	$\operatorname{esc}$
4	$\operatorname{esc}$
5	$\operatorname{esc}$

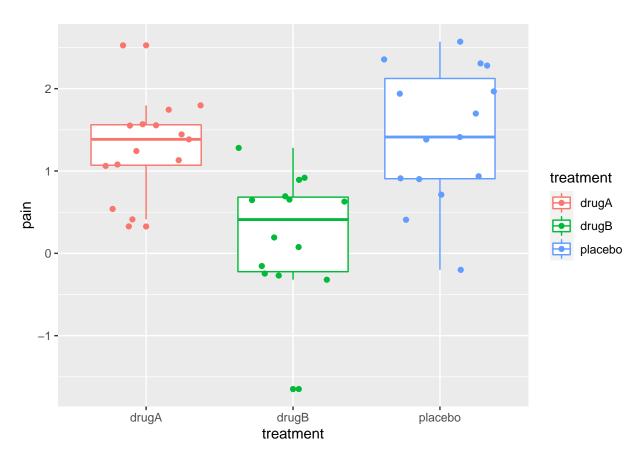
```
boxplot( response ~ index, dataset,xlab = "haha", main = "csd",ylab = "csac")
```



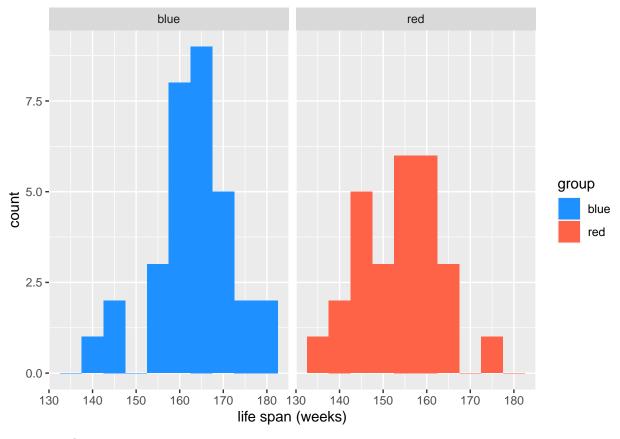


if data frame, ggplot2

```
g <- ggplot(trial,aes(x = treatment,y = pain, colour = treatment))
g <- g + geom_boxplot()
g <- g + geom_jitter()
g</pre>
```



```
hamsters <- read.csv("hamsters.csv")
p <- ggplot(hamsters, aes(x=lifespan,fill=group))
p <- p + facet_grid(cols=vars(group))
p <- p + geom_histogram(binwidth = 5)
p <- p + xlab("life span (weeks)")
p <- p + scale_fill_manual(values=c("dodgerblue", "tomato1"))
p</pre>
```



more on ggplot2

## # multiple curves

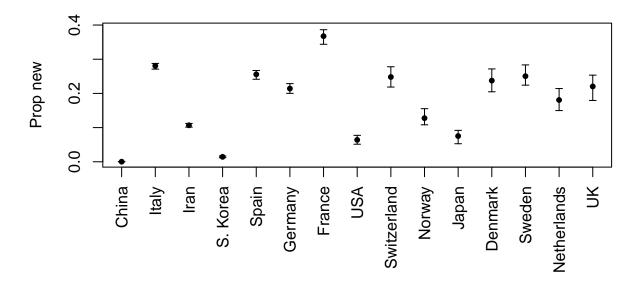
Confidence interval plot

```
covid <- read.csv("coronavirus.csv")</pre>
```

```
obs_values<-vector()
lower_cis<-vector()</pre>
upper_cis<-vector()</pre>
for (a in 1:nrow(covid)){
  country<-covid[a,1]</pre>
  obs_total<-covid[a,2]
  obs_new<-covid[a,3]
  obs<-obs_new/obs_total
  old = obs_total - obs_new
  obs_sample<-c(rep("new", obs_new), rep("old", old))
  bootstrap_new<-vector()</pre>
  for (b in 1:100){
    bootstrap_sample<-sample(obs_sample, length(obs_sample), replace = T)</pre>
    bootstrap_new<-c(bootstrap_new, length(subset(bootstrap_sample, bootstrap_sample == "new")))
  }
  lower_ci<-quantile(bootstrap_new, 0.025)</pre>
  upper_ci<-quantile(bootstrap_new, 0.975)
  obs_new<-obs_new/obs_total
```

```
lower_ci<-lower_ci/obs_total
upper_ci<-upper_ci/obs_total
obs_values<-c(obs_values, obs_new)
lower_cis<-c(lower_cis, lower_ci)
upper_cis<-c(upper_cis, upper_ci)
}</pre>
```

```
ymax<-ceiling(max(upper_cis)*100)/100
par(mar=c(11,4,4,2))
plot(obs_values, xaxt = "n", ylim = c(0,ymax), xlab = "", pch = ".", ylab = "Prop new")
axis(side = 1, at = seq(1,nrow(covid),1), labels = covid$Country, las = 2)
for (a in 1:length(lower_cis)){
   lines(x = c(a,a), y = c(lower_cis[a], upper_cis[a]))
   lines(x = c(a-0.1,a+0.1), y = c(lower_cis[a], lower_cis[a]))
   lines(x = c(a-0.1,a+0.1), y = c(upper_cis[a], upper_cis[a]))
}
points(x = seq(1,nrow(covid),1), y = obs_values, pch = 20)</pre>
```



## Randomisation

```
# forearms$newGender1 <- sample(forearms$Gender,nrow(forearms), replace=FALSE)
```

## **Concatenate Strings**

```
paste0("a","b")

## [1] "ab"

c="Leo"
paste0("I am ",c)

## [1] "I am Leo"
```

#### Some functions related to normal distribution

```
rnorm(10) # number, mean(0), sd(1)

## [1] -0.26767013 -3.26417857 -0.01590474 -1.28237580  2.57880839 -0.48256762

## [7]  0.40790757 -0.14279013  0.23027685  1.46574027

pnorm(1.96) # cumulative distribution function

## [1]  0.9750021

qnorm(0.975) # the inverse of the second

## [1]  1.959964
```

## Some functions related to uniform distribution

```
runif(26,0,100) # number,min,max

## [1] 15.858808 82.700428 22.641563 90.757996 51.139930 0.726628 35.229499
## [8] 45.692756 25.230992 10.477067 78.201328 97.996398 84.649117 63.664794
## [15] 89.117741 62.809955 17.791786 6.824928 59.174370 73.170481 63.834266
## [22] 53.119184 32.122643 94.791871 9.365563 82.314189
```

#### Some functions related to t distribution

From t value to p value: x=2.093, df=10 - and +: same show the area t>2.093 and t<-2.093 if significant: one-tailed:<0.1, two tailed<0.05

```
dt(1.729,10) #one tailed
```

```
## [1] 0.09232579
```

```
dt(2.093,10) #two tailed
```

```
## [1] 0.05275807
```

## Hypothesis testing

#### t-test

Assumptions:

- 1. Independent random sampling (For independent random sampling, we assume that this assumption has been met from looking at the experimental design.)
- 2. Normal distribution of data in both groups
- 3. Sample sizes are small(n < 100)

one-tailed vs two-tailed

```
# one-tailed (alternative=greater, less)
t.test(n,mu=6,alternative="greater")
```

```
##
## One Sample t-test
##
## data: n
## t = -5.0113, df = 19, p-value = 7.752e-05
## alternative hypothesis: true mean is not equal to 6
## 95 percent confidence interval:
## 4.571692 5.413286
## sample estimates:
## mean of x
## 4.992489
```

one sample vs two sample(mu)

```
t.test(n,mu=6)
##
##
    One Sample t-test
##
## data: n
## t = -5.0113, df = 19, p-value = 7.752e-05
## alternative hypothesis: true mean is not equal to 6
## 95 percent confidence interval:
## 4.571692 5.413286
## sample estimates:
## mean of x
## 4.992489
1 < - seq(1,20)
t.test(1,n)
##
   Welch Two Sample t-test
##
##
## data: l and n
## t = 4.116, df = 19.877, p-value = 0.0005422
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## 2.715249 8.299773
## sample estimates:
## mean of x mean of y
## 10.500000 4.992489
paired vs unpaired (paired)
c \leftarrow rep(1,10)
t.test(c,d,paired=TRUE)
##
## Paired t-test
##
## data: c and d
## t = NaN, df = 9, p-value = NA
\#\# alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## NaN NaN
## sample estimates:
## mean of the differences
##
```

#### Power of t test

statistical power is the likelihood that a study will detect an effect when there is an effect there to be detected. sample size/power, delta: difference in means, type: t-test type, alternative:

```
power.t.test(n=10, delta = 3, sd =10, sig.level = 0.05, type = "one.sample",
             alternative = "one.sided")
##
##
        One-sample t test power calculation
##
##
                 n = 10
             delta = 3
##
##
                sd = 10
##
         sig.level = 0.05
##
             power = 0.2216581
       alternative = one.sided
power.t.test(delta = 26, sd =30, sig.level = 0.05, power=0.8, type = "two.sample",
             alternative = "one.sided")
##
##
        Two-sample t test power calculation
##
##
                 n = 17.18096
             delta = 26
##
##
                sd = 30
##
         sig.level = 0.05
##
             power = 0.8
##
       alternative = one.sided
##
## NOTE: n is number in *each* group
Wilcoxon test
1-sample wilcoxon test
wilcox.test(a,mu=0)
##
## Wilcoxon signed rank exact test
##
## data: a
## V = 1, p-value = 1
\#\# alternative hypothesis: true location is not equal to 0
Wilcoxon signed-rank test
wilcox.test(c,d,paired = TRUE)
## Warning in wilcox.test.default(c, d, paired = TRUE): cannot compute exact p-
## value with zeroes
##
## Wilcoxon signed rank test with continuity correction
```

```
##
## data: c and d
## V = 0, p-value = NA
## alternative hypothesis: true location shift is not equal to 0
```

Wilcoxon Rank Sum test

```
wilcox.test(a,n)
```

```
##
## Wilcoxon rank sum exact test
##
## data: a and n
## W = 20, p-value = 0.09524
## alternative hypothesis: true location shift is not equal to 0
```

#### **ANOVA**

Formulate null and alternative hypothesis and decide which type of ANOVA to use

#### **Assumptions:**

- 1. Independent random sampling
- 2. Normality of residuals (residuals: distances from group mean)
- 3. Equality of Variances

**Check assumptions** Independent random sampling: we assume that this assumption has been met from looking at the experimental design.

In order to test the other two assumptions, we need to make an ANOVA model.

```
model <- aov(pain~treatment,data=trial)</pre>
```

Normality of residuals

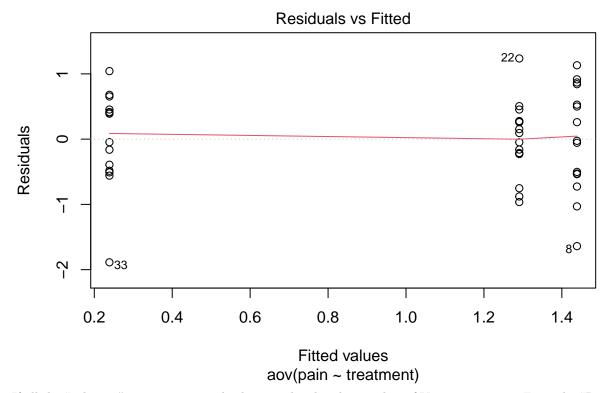
```
shapiro.test(resid(model))
```

```
##
## Shapiro-Wilk normality test
##
## data: resid(model)
## W = 0.97317, p-value = 0.3895
```

The null hypothesis of the shapiro-wilk test is that the distribution is normal. The p-value is >/< 0.05, so we can(not) reject the null hypothesis. We conclude that the residuals is (not) normally distributed.

Equality of Variances

```
plot(model,1)
```



If all the "columns" are approximately the same height, the equality of Variances is met. From the "Residual vs Fitted" plot, we conclude that.

#### Perform ANOVA

We can now perform an ANOVA.

## summary(model)

```
## Df Sum Sq Mean Sq F value Pr(>F)
## treatment    2 12.27 6.133 11.9 8.4e-05 ***
## Residuals 41 21.13 0.515
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

conclusion of ANOVA:

**Perform TukeyHSD test(a type of post hoc test)** We can now perform TukeyHSD test to find out what groups are exactly different.

## TukeyHSD(model)

```
## Tukey multiple comparisons of means
## 95% family-wise confidence level
```

conclusion of TukeyHSD test:

## About masking

(from formative) This is a type of masking to avoid any kind of human bias in the experiment or the data analysis. If you have ideas about what you would expect the lifespan of both populations to be, you may be tempted to throw out data points that don't fit in your theory as "outliers", for instance.

## Correlation and linear regression

import sample data

```
## Selecting by val
```

```
## # A tibble: 6 x 4
## # Groups:
               measure_name [3]
##
     measure_name location_name
                                  year
                                               val
##
     <chr>>
                   <chr>>
                                  <int>
                                             <dbl>
## 1 Deaths
                   Nigeria
                                   2008
                                          280604.
## 2 Deaths
                   Nigeria
                                   2009
                                          276715.
## 3 Incidence
                   Nigeria
                                   2011 63597364.
## 4 Incidence
                   Nigeria
                                   2010 64045849.
                   Nigeria
## 5 Prevalence
                                   2005 50823690.
## 6 Prevalence
                   Nigeria
                                   2006 51296796.
```

Assumptions (from ADS2 week 18 lecture, slide 18):

- 1. The residuals are normally distributed.
- 2. The errors are independent (independent random sampling)
- 3. The relationships are linear.

We can see there is linear relationship (not "U" shaped etc..) between year and prevalence in all age groups. Also visual inspection shows that there is no obvious outliers. We think it is appropriate to do linear regression.

Although the linear regression lecture mentions the assumption for linear regression is that the residuals need to be normally distributed. However, due to the small sample size, it is quite likely that the residuals are not normally distributed. We think r.squared is enough to assess whether the linear model fits our data points.

r.squared describes the degree of interpretation of input variables to output variables. Formula: In linear regression, the larger the R-square (the closer to 1), the better the linear regression is.

We want to set the threshold of r.squared to 0.7 (from ADS2 week 18 lecture, slide 21). If r.squared is larger than 0.7, we think the result of linear regression can reflect our data points.

### Perform linear regression

```
# fit <- lm(val~year,ayvs1[[cnt]])
# slope[j,i] <- as.numeric(fit$coefficients[2])
# slope[j,i+1] <- summary(fit)$r.squared</pre>
```

#### Calculate the correlation coefficient

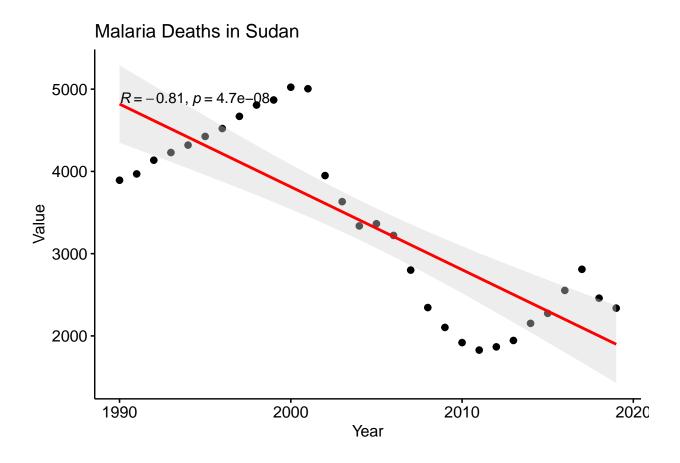
```
# cor()
```

#### Add straight line to the plot

```
# abline(fit,col="red") fit=lm results
```

#### Add straight line and confidence interval using ggplot2

```
## 'geom_smooth()' using formula 'y ~ x'
```



## **Bootstrapping**

Case resampling (without absolute)

```
# all <- c()
# for (i in 1:10000){
# ab <- sample(a$points,4,replace=TRUE)
# ac <- sample(a$points,5,replace=TRUE)
# all <- c(all,median(ab)-median(ac))
# }
# hist(all)
# exact <- median(Female$points)-median(Male$points)
# mean(all>=exact)
```

#### Confidence interval

```
# bootstrapping for 100 times
#
   for (b in 1:100){
      bootstrap_sample <- sample(obs_sample, length(obs_sample), replace = T)</pre>
#
#
      bootstrap_new <- c(bootstrap_new,
#
                           length(subset(bootstrap_sample,
#
                                          bootstrap_sample == "obese")))
#
    }
#
   # get 95% confidence interval
# lower_ci <- quantile(bootstrap_new, 0.025)</pre>
#
  upper_ci <- quantile(bootstrap_new, 0.975)</pre>
#
  lower_cis <- c(lower_cis, lower_ci/1000*100)</pre>
#
    upper_cis <- c(upper_cis, upper_ci/1000*100)</pre>
# }
# big$upper <- upper_cis</pre>
# big$lower <- lower_cis</pre>
```

## Some data cleaning function

```
gather
```

## 2

1

1

0

```
data <- read.csv("really_tiny_dataset.csv",header = F)</pre>
data[c(1,2),c(1,2,784,785)]
   V1 V2 V784 V785
##
## 1 5 0
           0
## 2 0 0
              0
names(data) \leftarrow c("num", seq(1,784))
# c(-num): gather according column (which column is x), y is the header
data <- gather(data,key="order",value="intensity",c(-num))</pre>
head(data)
     num order intensity
##
## 1 5
           1
                       0
## 2
     0
             1
                       0
                       0
## 3
      4
             1
                       0
     1
           1
## 4
## 5 9
           1
                       0
## 6
     2
# aline
aggregate (aggregate according to column)
data_aggr <- aggregate(intensity ~ num+order,data=data,mean)</pre>
head(data_aggr)
     num order intensity
##
## 1
       0
```

```
## 3 2 1 0
## 4 3 1 0
## 5 4 1 0
## 6 5 1
```

## Categorical data

## chi-square

input data

```
survey <- matrix(c(84,82,34,82,57,11),nrow = 3)
row.names(survey) <- c("Good", "Fair", "Bad")
colnames(survey) <- c("A","B")
survey_df <- as.data.frame(as.table(survey))
names(survey_df) <- c("Rating", "School", "Freq")</pre>
```

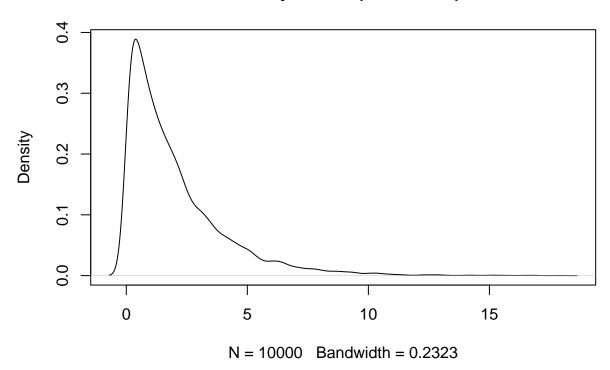
Assumptions for chi-square test:

- 1. Discrete, categorical data.
- 2. No expected cell frequencies are less than 1.
- 3. No more than 20% are less than 5.

Generate chi square distribution

```
chi.df2 <- rchisq(10000, df=2)
plot(density(chi.df2))</pre>
```

# density.default(x = chi.df2)



Goodness of fit (actual vs expected)

```
# Given probability (expected)
chisq.test(c(84,82,34), p=c(0.45,0.43,0.12))

##
## Chi-squared test for given probabilities
##
## data: c(84, 82, 34)
## X-squared = 4.7527, df = 2, p-value = 0.09289

# ggplot(data=dat, aes(x=chi2.value, color=df)) + geom_density()
```

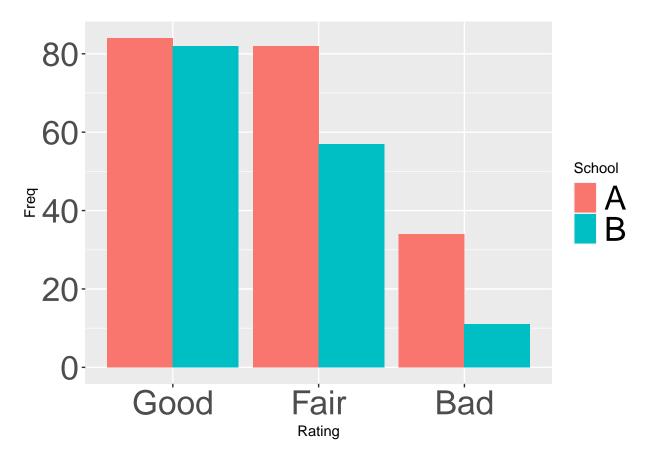
homogeneity (2-way, same or different?)

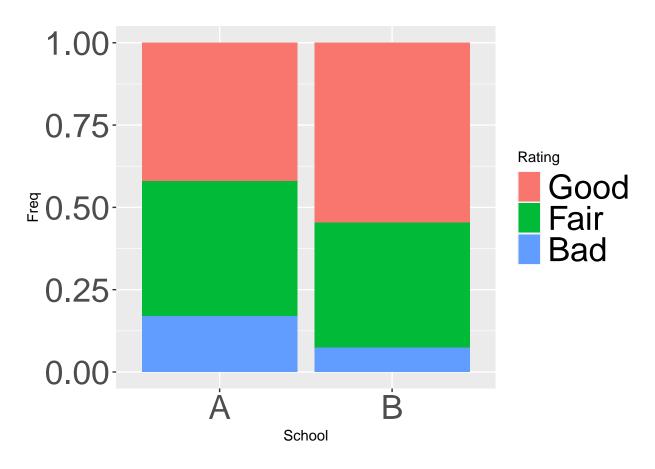
```
chisq.test(survey)
```

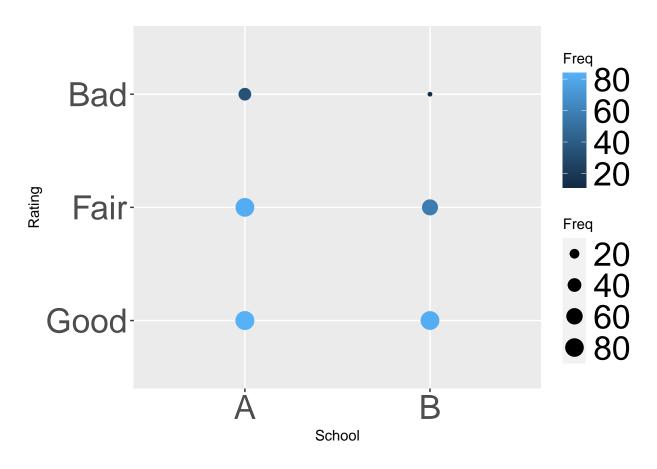
```
##
## Pearson's Chi-squared test
##
## data: survey
## X-squared = 9.3235, df = 2, p-value = 0.00945
independency (2-way) (similar to homogeneity)
```

```
interdependency (3-way, rcl)
```

Data visualization

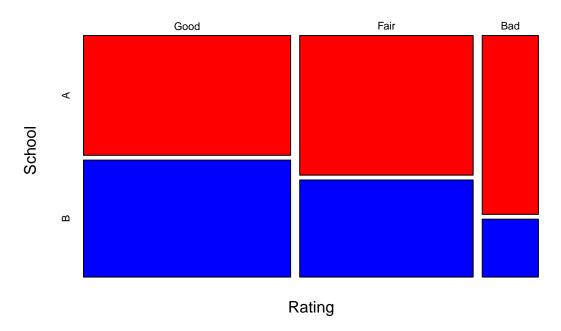






#mosaicplot
mosaicplot(survey, color = c("red","blue"), xlab="Rating", ylab="School")

## survey



Fisher's exact test (for small sample size)

```
survival < -matrix(c(7,3,2,7),nrow = 2)
row.names(survival) <- c("Alive", "Dead")</pre>
colnames(survival) <- c("WT","KO")</pre>
#4.2 Chi-square test with Yates's correction on/off
chisq.test(survival, correct = F)
## Warning in chisq.test(survival, correct = F): Chi-squared approximation may be
## incorrect
##
   Pearson's Chi-squared test
##
##
## data: survival
## X-squared = 4.3372, df = 1, p-value = 0.03729
chisq.test(survival, correct = T)
## Warning in chisq.test(survival, correct = T): Chi-squared approximation may be
## incorrect
```

```
##
## Pearson's Chi-squared test with Yates' continuity correction
##
## data: survival
## X-squared = 2.6324, df = 1, p-value = 0.1047
#.4.3 Fisher's exact test
fisher.test(survival)
##
## Fisher's Exact Test for Count Data
## data: survival
## p-value = 0.06978
## alternative hypothesis: true odds ratio is not equal to 1
## 95 percent confidence interval:
## 0.7520079 113.4668907
## sample estimates:
## odds ratio
## 7.166282
# str_detect & grep
```

group\_by (bet mea value of the same group, similar to aggregate)

```
# pixel_summary <- pixels_gathered %>%
# group_by(x, y, label) %>%
# summarize(mean_value = mean(value)) %>%
# ungroup()
```