## Package 'pivmet'

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Type Package

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**Description** More about what it does (maybe more than one line)
Use four spaces when indenting paragraphs within the Description.

License What license is it under?

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## **R** topics documented:

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## **Description**

Perform Maxima Units Search (MUS) algorithm on a large and sparse matrix in order to find a set of pivotal units through a sequential search in the given matrix.

## Usage

```
MUS(C, clusters, prec_par = 5)
```

#### **Arguments**

С	NxN matrix with a non-negligible number of zeros. For instance, a similarity matrix estimated from a NxD data matrix whose rows are statistical units, or a co-association matrix resulting from clustering ensembles.
clusters	A vector of integers from 1:k (with k <= 4) indicating a partition of the $N$ units resulting from clustering.
prec_par	Optional argument. The maximum number of alternative pivots for each group.

#### **Details**

Consider H distinct partitions of a set of N d-dimensional statistical units into k groups determined by some clustering technique. A  $N \times N$  co-association matrix C with generic element  $c_{i,j} = n_{i,j}/H$  can be constructed, where  $n_{i,j}$  is the number of times the i-th and the j-th unit are assigned to the same cluster with respect to the clustering ensemble. Units which are very distant from each other are likely to have zero co-occurrences; as a consequence, C is a square symmetric matrix expected to contain a non-negligible number of zeros. The main task of the MUS algorithm is to detect submatrices of small rank from the co-association matrix and extract those units—pivots—such that the  $k \times k$  submatrix of C, determined by only the pivotal rows and columns indexes, is identical or nearly identical. Practically, the resulting units have the desirable property to be representative of the group they belong to.

## Value

pivots The k pivotal units

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#### References

Egidi, L., Pappad\'a, R., Pauli, F., Torelli, N. (2018). Maxima Units Search(MUS) algorithm: methodology and applications. In: Perna, C., Pratesi, M., Ruiz-Gazen A. (eds.) Studies in Theoretical and Applied Statistics, Springer Proceedings in Mathematics & Statistics 227, pp. 71—81.

```
# Data generated from a mixture of three bivariate Gaussian distributions
N <- 620
centers <- 3
n1 <- 20
n2 <- 100
n3 <- 500
x <- matrix(NA, N,2)
truegroup <- c(rep(1,n1), rep(2, n2), rep(3, n3))
for (i in 1:n1){
x[i,]=rmvnorm(1, c(1,5), sigma=diag(2))
for (i in 1:n2){
x[n1+i,]=rmvnorm(1, c(4,0), sigma=diag(2))
for (i in 1:n3){
x[n1+n2+i,]=rmvnorm(1, c(6,6), sigma=diag(2))
# Build a similarity matrix from clustering ensembles
H <- 1000
a <- matrix(NA, H, N)
for (h in 1:H){
   a[h,] \leftarrow kmeans(x,centers)$cluster
sim_matr <- matrix(1, N,N)</pre>
for (i in 1:(N-1)){
  for (j in (i+1):N){
     sim_matr[i,j] \leftarrow sum(a[,i]==a[,j])/H
     sim_matr[j,i] <- sim_matr[i,j]</pre>
}
# Obtain a clustering solution via kmeans with multiple random seeds
cl <- KMeans(x, centers)$cluster</pre>
# Find three pivots
mus_alg <- MUS(C = sim_matr, clusters = cl)</pre>
```

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#### **Description**

Perform classical k-means clustering on a data matrix using pivots as initial centers.

#### Usage

```
piv_KMeans(x, centers, alg.type = c("KMeans", "hclust"),
    piv.criterion = c("MUS", "maxsumint", "maxsumnoint", "maxsumdiff"),
    H = 1000, ...)
```

#### **Arguments**

x A  $N \times D$  data matrix, or an object that can be coerced to such a matrix (such as

a numeric vector or a dataframe with all numeric columns).

centers The number of clusters in the solution.

alg. type The clustering algorithm for the initial partition of the N units into the desired

number of clusters. Possible choices are "KMeans" and "hclust".

piv.criterion The pivotal criterion used for identifying one pivot for each group. Possi-

ble choices are: "MUS", "maxsumint", "maxsumnoint", "maxsumdiff". If centers <= 4, the default method is "MUS"; otherwise, the default method is

"maxsumdiff" (see the details and the vignette).

H If MUS is selected, this is the number of distinct k-means partitions used for

building a  $N \times N$  co-association matrix.

... Optional arguments to be passed to MUS or KMeans.

#### **Details**

The function implements a modified version of K-means which aims at improving the clustering solution starting from a careful seeding. In particular, it performs a pivot-based initialization step using pivotal methods to find the initial centers for the clustering procedure. The starting point consists of multiple runs of the classical K-means (which uses random seeds) with a fixed number of clusters in order to build the co-association matrix of data units.

### Value

#### A list with components

cluster A vector of integers indicating the cluster to which each point is allocated.

centers A matrix of cluster centres (centroids).

totss The total sum of squares.

withinss The within-cluster sum of squares for each cluster.

tot.withinss The within-cluster sum of squares summed across clusters.

betwennss The between-cluster sum of squared distances.

size The number of points in each cluster. iter The number of (outer) iterations.

ifault integer: indicator of a possible algorithm problem – for experts.

pivots The pivotal units identified by the selected pivotal criterion.

#### Author(s)

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#### References

Egidi, L., Pappadà, R., Pauli, F., Torelli, N. (2018). K-means seeding via MUS algorithm. Conference Paper, Book of Short Papers, SIS2018, ISBN: 9788891910233.

```
# Data generated from a mixture of three bivariate Gaussian distributions
N <- 620
k <- 3
n1 <- 20
n2 <- 100
n3 <- 500
x <- matrix(NA, N,2)
truegroup <- c( rep(1,n1), rep(2, n2), rep(3, n3))</pre>
for (i in 1:n1){
x[i,]=rmvnorm(1, c(1,5), sigma=diag(2))
for (i in 1:n2){
x[n1+i,]=rmvnorm(1, c(4,0), sigma=diag(2))
for (i in 1:n3){
x[n1+n2+i,]=rmvnorm(1, c(6,6), sigma=diag(2))
# Apply piv_KMeans with MUS as pivotal criterion
res <- piv_KMeans(x, k)</pre>
# Apply piv_KMeans with maxsumdiff as pivotal criterion
res2 <- piv_KMeans(x, k, piv.criterion ="maxsumdiff")</pre>
# Plot the data and the clustering solution
par(mfrow=c(1,2), pty="s")
colors_cluster <- c("grey", "darkolivegreen3", "coral")</pre>
colors_centers <- c("black", "darkgreen", "firebrick")</pre>
plot(x, col = colors_cluster[truegroup],
   bg= colors_cluster[truegroup], pch=21, xlab="x[,1]",
   ylab="x[,2]", cex.lab=1.5,
   main="True data", cex.main=1.5)
plot(x, col = colors_cluster[res$cluster],
   bg=colors_cluster[res$cluster], pch=21, xlab="x[,1]",
   ylab="x[,2]", cex.lab=1.5,
   main="piv_KMeans", cex.main=1.5)
points(x[res$pivots[1],1], x[res$pivots[1],2],
   pch=24, col=colors_centers[1],bg=colors_centers[1],
   cex=1.5)
points(x[res$pivots[2],1], x[res$pivots[2],2],
   pch=24, col=colors_centers[2], bg=colors_centers[2],
   cex=1.5)
points(x[res$pivots[3],1], x[res$pivots[3],2],
   pch=24, col=colors_centers[3], bg=colors_centers[3],
   cex=1.5)
points(res$centers, col = colors_centers[1:k],
```

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pch = 8, cex = 2)

piv\_MCMC

JAGS Sampling for Gaussian Mixture Models and Clustering via Co-Association Matrix.

#### **Description**

Perform MCMC JAGS sampling for Gaussian mixture models, post-process the chains and apply a clustering technique to the MCMC sample. Pivotal units for each group are selected among four alternative criteria.

## Usage

```
piv_MCMC(y, k, priors, nMC, piv.criterion = c("MUS", "maxsumint",
    "maxsumnoint", "maxsumdiff"), clustering = c("diana", "hclust"))
```

## **Arguments**

y N-dimensional data vector/matrix.k Number of mixture components.

nMC Number of MCMC iterations for the JAGS function execution.

piv.criterion The pivotal criterion used for identifying one pivot for each group. Possi-

ble choices are: "MUS", "maxsumint", "maxsumnoint", "maxsumdiff". If  $k \le 4$ , the default method is "MUS"; otherwise, the default method is "maxsumdiff"

(see the details and the vignette).

clustering The clustering technique adopted for partitioning the N observations into k groups.

Possible choices: "diana" (default), "hclust".

#### **Details**

The function fits univariate and bivariate Bayesian Gaussian mixture models of the form (here for univariate only):

$$(Y_i|Z_i=j) \sim f(y;\mu_i,\phi_i),$$

where the  $Z_i$ ,  $i=1,\ldots,N$ , are i.i.d. random variables,  $j=1,\ldots,k$ ,  $\phi_j$  is the variance,  $Z_i\in \{1,\ldots,k\}$ , and

$$P(Z_i = j) = \pi_j.$$

The likelihood of the model is then

$$L(y; \mu, \pi, \phi) = \prod_{i=1}^{n} \sum_{j=1}^{k} \pi_j \mathcal{N}(\mu_j, \phi),$$

with  $\mu=(\mu_1,\ldots,\mu_k)$  component-specific parameters and  $\pi=(\pi_1,\ldots,\pi_k)$  mixture weights. Let  $\nu$  denote a permutation of  $1,\ldots,k$ , and let  $\nu(\mu)=(\mu_{\nu(1)},\ldots,\mu_{\nu(k)}),\,\nu(\pi)=(\pi_{\nu(1)},\ldots,\pi_{\nu(k)})$  be the corresponding permutations of  $\mu$  and  $\pi$ . Denote by V the set of all the permutations of the indexes  $1,\ldots,k$ , the likelihood above is invariant under any permutation  $\nu\in V$ , that is

$$L(y; \mu, \pi, \phi) = L(y; \nu(\mu), \nu(\pi), \phi).$$

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As a consequence, the model is unidentified with respect to an arbitrary permutation of the labels. When Bayesian inference for the model is performed, if the prior distribution  $p_0(\mu, \pi, \phi)$  is invariant under a permutation of the indices, then so is the posterior. That is, if  $p_0(\mu, \pi, \phi) = p_0(\nu(\mu), \nu(\pi), \phi)$ , then

$$p(\mu, \pi, \phi|y) \propto p_0(\mu, \pi, \phi) L(y; \mu, \pi, \phi)$$

is multimodal with (at least) k! modes.

Priors are chosen as weakly informative. For univariate mixtures, the specification is the same as the function BMMmodel of the bayesmix package:

$$\mu_{j} \sim \mathcal{N}(0, 1/B0inv)$$
 
$$\phi_{j} \sim \text{invGamma}(nu0Half, nu0S0Half)$$
 
$$\pi \sim \text{Dirichlet}(1, \dots, 1)$$
 
$$S0 \sim \text{Gamma}(g0Half, g0G0Half),$$

with default values: B0inv = 0.1, nu0Half = 10, S0 = 2, nu0S0Half = nu0Half\*S0, g0Half = 5e-17, g0G0Half = 5e-33, in accordance with the default specification priors=list(kind = "independence", par (see bayesmix for further details and choices).

For bivariate mixtures, the prior specification is the following:

$$\mu_j \sim \mathcal{N}_2(\boldsymbol{m}\boldsymbol{u}_0, S2)$$
  
 $1/\Sigma \sim \text{Wishart}(S3, 3)$   
 $\pi \sim \text{Dirichlet}(1, \dots, 1),$ 

where S2 and S3 are diagonal matrices with diagonal elements (the variances) equal to 1e+05. The user may specify other values for the hyperparameters  $\mu_0$ , S2, S3 via priors argument in such a way:

priors = list(mu0 = c(1,1), S2 = matrix(c(0.002,0,0,0.1),2,2, byrow=TRUE),S3 = matrix(c(0.1,0,0,0.1),0.1)

The function performs JAGS sampling using the bayesmix package for univariate Gaussian mixtures, and the runjags package for bivariate Gaussian mixtures. After MCMC sampling, this function clusters the units in k groups, calls the  $piv\_sel()$  function and yields the pivots obtained from one among four different methods (the user may specify one among them via  $piv\_criterion$  argument): "maxsumint", "maxsumnoint", "maxsumdiff" and "MUS" (available only if k < 5) (see the vignette for thorough details).

## Value

The function gives the MCMC output, the clustering solutions and the pivotal indexes. Here is a complete list of outputs.

Freq	k x 2 matrix where: the first column reports the number of units allocated to each group as given by JAGS program; the second column reports the same number of units as given by the chains' post-processing.
true.iter	The number of MCMC iterations for which the number of JAGS groups exactly coincides with the prespecified number of groups k.
Z	N x k x true.iter array with values: 1, if the $i$ -th unit belongs to the $j$ -th group at the $h$ -th iteration; 0, otherwise.
ris	MCMC output matrix as provided by JAGS.

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groupPost	true.iter $\boldsymbol{x}$ N matrix with values from 1: $\boldsymbol{k}$ indicating the post-processed group allocation vector.
mu_switch	If y is a vector, a true.iter $x$ k matrix with the post-processed MCMC chains for the mean parameters; if y is a matrix, a true.iter $x$ 2 $x$ k array with the post-processed MCMC chains for the mean parameters.
mu_raw	If y is a vector, a nMC $\times$ k matrix with the raw MCMC chains for the mean parameters as given by JAGS; if y is a matrix, a nMC $\times$ 2 $\times$ k array with the raw MCMC chains for the mean parameters as given by JAGS.
С	Co-association matrix constructed from the MCMC sample.
grr	Group vector allocation as provided by "diana" or "hclust".
pivots	The pivotal units identified by the selected pivotal criterion.
piv.criterion	Gives the pivotal criterion used for identifying the pivots.

#### Author(s)

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## References

Egidi, L., Pappadà, R., Pauli, F. and Torelli, N. (2018). Relabelling in Bayesian Mixture Models by Pivotal Units. Statistics and Computing, 28(4), 957-969.

```
# Bivariate simulation
N <- 200
k <- 4
nMC <- 1000
M1 <-c(-.5,8)
M2 < -c(25.5,.1)
M3 < -c(49.5,8)
M4 < -c(63.0,.1)
Mu <- matrix(rbind(M1,M2,M3,M4),c(4,2))</pre>
stdev \leftarrow cbind(rep(1,k), rep(200,k))
Sigma.p1 \leftarrow matrix(c(stdev[1,1],0,0,stdev[1,1]), nrow=2, ncol=2)
Sigma.p2 \leftarrow matrix(c(stdev[1,2],0,0,stdev[1,2]), \ nrow=2, \ ncol=2)
W \leftarrow c(0.2, 0.8)
sim <- piv_sim(N,k,Mu, stdev, Sigma.p1,Sigma.p2,W)</pre>
res <- piv_MCMC(y = sim$y, k = k, nMC = nMC)
#changing priors
res2 <- piv_MCMC(y = sim$y,
                  priors = list (
                  mu0=c(1,1),
                  S2 = matrix(c(0.002,0,0, 0.1),2,2, byrow=TRUE),
                  S3 = matrix(c(0.1,0,0,0.1), 2,2, byrow = TRUE)),
                  k = k, nMC = nMC)
# Fishery data (bayesmix package)
data(fish)
```

piv\_plot

piv\_plot

Plotting outputs from pivotal relabelling

## **Description**

Plot and visualize MCMC outputs, posterior relabelled chains and estimates and diagnostics.

#### Usage

```
piv_plot(y, mcmc, rel_est, type = c("chains", "estimates", "hist"))
```

## Arguments

```
y Data vector or matrix.

rel_est Pivotal estimates as provided by piv_rel.

type Type of plots required. Choose among: "chains", "estimates", "hist".
```

## Author(s)

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```
# Fishery data

data(fish)
y <- fish[,1]
N <- length(y)
k <- 5
nMC <- 5000
res <- piv_MCMC(y = y, k = k, nMC = nMC)
rel <- piv_rel(mcmc=res, nMC = nMC)
piv_plot(y, res, rel, "chains")
piv_plot(y, res, rel, "estimates")
piv_plot(y, res, rel, "hist")</pre>
```

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piv\_rel

Perfroming the pivotal relabelling step and computing the relabelled posterior estimates

## **Description**

This function allows to perform the pivotal relabelling procedure described in Egidi et al. (2018) and to obtain the relabelled posterior estimates.

#### Usage

piv\_rel(mcmc, nMC)

#### **Arguments**

mcmc The output of the MCMC sampling from piv\_MCMC.

nMC The number of total MCMC iterations (given in input to the piv\_MCMC function,

or any function suited for MCMC sampling).

#### **Details**

Prototypical models in which the label switching problem arises are mixture models, where for a sample  $y = (y_1, \dots, y_N)$  we assume

$$(Y_i|Z_i=j) \sim f(y;\mu_i,\phi),$$

where the  $Z_i$ ,  $i=1,\ldots,N$ , are i.i.d. random variables,  $j=1,\ldots,k$ ,  $\phi$  is a parameter which is common to all components,  $Z_i \in \{1,\ldots,k\}$ , and

$$P(Z_i = j) = \pi_i$$
.

This model is unidentified with respect to an arbitrary permutation of the labels 1, ..., k. Relabelling means permuting the labels at each iteration of the Markov chain in such a way that the relabelled chain can be used to draw inferences on component-specific parameters.

We assume here that an MCMC sample is obtained from the posterior distribution for model above—for instance via piv\_MCMC function—with a prior distribution which is labelling invariant. Furthermore, suppose that we can find k units, one for each group, which are (pairwise) separated with (posterior) probability one (that is, the posterior probability of any two of them being in the same group is zero). It is then straightforward to use the k units, called pivots in what follows, to identify the groups and to relabel the chains: for each MCMC iteration  $h=1,\ldots,H$  (H corresponds to the argument nMC) and group  $j=1,\ldots,k$ , set

$$[\mu_j]_h = [\mu_{[Z_{i_j}]_h}]_h;$$

$$[Z_i]_h = j$$
 for  $i : [Z_i]_h = [Z_{i_j}]_h$ .

The applicability of this strategy is limited by the existence of the pivots, which is not guaranteed. The existence of the pivots is a requirement of the method, meaning that its use is restricted to those chains—or those parts of a chain—for which the pivots are present. First, although the model is based on a mixture of k components, each iteration of the chain may imply a different number of non-empty groups. Let then  $[k]_k \le k$  be the number of non-empty groups at iteration k,

$$[k]_h = \#\{j : [Z_i]_h = j \text{ for some } i\},$$

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where #A is the cardinality of the set A. Hence, the relabelling procedure outlined above can be used only for the subset of the chain for which  $[k]_h = k$ ; let it be

$$\mathcal{H}_k = \{h : [k]_h = k\},\,$$

which correspond to the argument true.iter given by  $piv_MCMC$ . This means that the resulting relabelled chain is not a sample (of size H) from the posterior distribution, but a sample (of size  $\#\mathcal{H}_k$ ) from the posterior distribution conditional on there being (exactly) k non-empty groups. Even if k non-empty groups are available, however, there may not be k perfectly separated units. Let us define

$$\mathcal{H}_{k}^{*} = \{ h \in \mathcal{H}_{k} : \exists k, s \text{ s.t. } [Z_{i_{k}}]_{h} = [Z_{i_{s}}]_{h} \}$$

that is, the set of iterations where (at least) two pivots are in the same group. In order for the pivot method to be applicable, we need to exclude iterations  $\mathcal{H}_k^*$ ; that is, we can perform the pivot relabelling on  $\|-\mathcal{H}_k^*$ , corresponding to the argument Final\_It.

#### Value

This function gives the relabelled posterior estimates—both mean and medians—obtained from the Markov chains of the MCMC sampling.

mu\_rel\_mean k-vector (in case of univariate mixture) or k x 2 matrix (in case of bivariate

mixture) of estimated posterior means for the mean parameters.

 $mu\_rel\_median$  k-vector (in case of univariate mixture) or k x 2 matrix (in case of bivariate

mixture) of estimated posterior medians for the mean parameters.

mu\_rel\_complete

Complete relabelled chains

Final\_It The final number of valid MCMC iterations, as explained in Details

## Author(s)

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#### References

Egidi, L., Pappada, R., Pauli, F. and Torelli, N. (2018). Relabelling in Bayesian Mixture Models by Pivotal Units. Statistics and Computing, 28(4), 957-969.

#### **Examples**

#Univariate simulation

```
N <- 250
nMC <- 2500
k <- 3
p <- rep(1/k,k)
x <- 3
stdev <- cbind(rep(1,k), rep(200,k))
Mu <- seq(-trunc(k/2)*x,trunc(k/2)*x,length=k)
W <- c(0.2,0.8)
sim <- piv_sim(N,k,Mu,stdev,W=W)
res <- piv_MCMC(y = sim$y, k =k, nMC = nMC)
rel <- piv_rel(mcmc=res, nMC = nMC)</pre>
```

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```
N <- 200
k <- 3
nMC <- 5000
M1 < c(-.5,8)
M2 < c(25.5,.1)
M3 <- c(49.5,8)
Mu \leftarrow matrix(rbind(M1,M2,M3),c(k,2))
stdev <- cbind(rep(1,k), rep(200,k))
Sigma.p1 <- matrix(c(stdev[1,1],0,0,stdev[1,1]),
                    nrow=2, ncol=2)
Sigma.p2 \leftarrow matrix(c(stdev[1,2],0,0,stdev[1,2]),
                    nrow=2, ncol=2)
W \leftarrow c(0.2, 0.8)
sim \leftarrow piv\_sim(N,k,Mu,stdev,Sigma.p1,Sigma.p2,W)
res <- piv_MCMC(y = sim$y, k = k, nMC = nMC)
rel <- piv_rel(mcmc = res, nMC = nMC)</pre>
piv_plot(y=sim$y, mcmc=res, rel_est = rel, type="chains")
piv_plot(y=sim$y, mcmc=res, rel_est = rel,
         type="hist")
```

piv\_sel

Pivotal Selection via Co-Association Matrix

## **Description**

Finding the pivots according to three different methods involving a co-association matrix C.

## Usage

```
piv_sel(C, clusters)
```

#Bivariate simulation

## **Arguments**

C A NxN co-association matrix, i.e. a matrix whose elements are co-occurences

of pair of units in the same cluster among H distinct partitions.

clusters A vector of integers indicating a partition of the N units into, say, k groups.

#### Value

pivots A matrix with k rows and three columns containing the indexes of the pivotal

units for each method.

## Author(s)

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#### References

Egidi, L., Pappadà, R., Pauli, F. and Torelli, N. (2018). Relabelling in Bayesian Mixture Models by Pivotal Units. Statistics and Computing, 28(4), 957-969.

## **Examples**

```
# Iris data
data(iris)
x<- iris[,1:4]
N <- length(iris[,1])</pre>
H <- 1000
a <- matrix(NA, H, N)
# Perform H k-means partitions
for (h in 1:H){
a[h,] <- kmeans(x, centers = 3)$cluster</pre>
}
# Build the co-association matrix
C <- matrix(1, N,N)</pre>
for (i in 1:(N-1)){
for (j in (i+1):N){
   C[i,j] \leftarrow sum(a[,i]==a[,j])/H
   C[j,i] \leftarrow C[i,j]
 }}
km <- kmeans(x, centers =3)</pre>
# Find the pivots according to the three possible pivoyal criterion
ris <- piv_sel(C, clusters = km$cluster)</pre>
plot(iris[,1], iris[,2], xlab ="Sepal.Length", ylab= "Sepal.Width",
col = km$cluster)
 \ensuremath{\text{\#}}\xspace Add the pivots according to maxsumdiff criterion
points( x[ris$pivot[,3], c( "Sepal.Length", "Sepal.Width" )], col = 1:3,
cex = 2, pch = 8)
```

piv\_sim

Generate Data from a Gaussian Nested Mixture

## **Description**

Simulate N observations from a nested Gaussian mixture model with k pre-specified components under uniform group probabilities 1/k, where each group is in turn drawn from a further level consisting of two subgroups.

piv\_sim

#### Usage

```
piv_sim(N, k, Mu, stdev, Sigma.p1 = matrix(c(1, 0, 0, 1), 2, 2, byrow = TRUE), Sigma.p2 = matrix(c(100, 0, 0, 100), 2, 2, byrow = TRUE), W = c(0.5, 0.5))
```

#### **Arguments**

N The desired sample size.

k The desired number of mixture components.

Mu The input mean vector/matrix.

stdev A k x2 matrix of input standard deviations, one for each group (from 1 to k) and for each subgroup (from 1 to 2). For univariate mixtures only.

Sigma.p1 The covariance matrix for the first subgroup. For bivariate mixtures only.

Sigma.p2 The covariance matrix for the second subgroup. For bivariate mixtures only.

W The vector for the mixture weights of the two subgroups,

#### **Details**

The functions allows to simulate values from a double (nested) univariate Gaussian mixture:

$$(Y_i|Z_i=j) \sim \sum_{s=1}^{2} p_{js} \mathcal{N}(\mu_j, \sigma_{js}^2),$$

or from a bivariate nested Gaussian mixture:

$$(Y_i|Z_i=j) \sim \sum_{s=1}^2 p_{js} \mathcal{N}_2(\boldsymbol{\mu}_j, \Sigma_s),$$

where  $\sigma_{js}^2$  is the variance for the group j and the subgroup s (stdev is the argument for specifying the kx2 standard deviations for univariate mixtures);  $\Sigma_s$  is the covariance matrix for the subgroup s, s = 1, 2, where the two matrices are specified by Sigma.pl and Sigma.p2 respectively;  $\mu_j$  and  $\mu_j, j = 1, \ldots, k$  are the mean input vector and matrix respectively, specified by the argument Mu; W is a vector of dimension 2 for the subgroups weights.

#### Value

y The N simulated observations. true.group A vector of integers from 1:k indicating the values of the latent variables  $Z_i$ . subgroups A 2 x N matrix with values 1 or 2 indicating the subgroup to which each obser-

vation is drawn from.

#### **Examples**

# Bivariate mixture simulation with three components

N <- 2000 k <- 3 M1 <- c(-45,8) M2 <- c(45,.1) piv\_sim 15

```
M3 <- c(100,8)
Mu <- matrix(rbind(M1,M2,M3),c(k,2))
stdev <- cbind(rep(1,k), rep(200,k))
Sigma.p1 <- matrix(c(stdev[1,1],0,0,stdev[1,1]),
nrow=2, ncol=2)
Sigma.p2 <- matrix(c(stdev[1,2],0,0,stdev[1,2]),
nrow=2, ncol=2)
W <- c(0.2,0.8)
sim <- piv_sim(N, k, Mu, Sigma.p1 = Sigma.p1,
Sigma.p2 = Sigma.p2, W)
plot(sim$y, xlab="y[,1]", ylab="y[,2]")</pre>
```

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