

Answer the following questions either prove the assertions or give counter examples. If you can't prove them, try to give reasons for your answers. The deadline is 11:59 pm Dec 8. Please latex your solution.

1. (15 pts) For a positive function f , consider two quantities

$$A := \int dy \left[\int f(x, y)^p dx \right]^{1/p}$$

$$B := \left[\int dx \left(\int f(x, y) dy \right)^p \right]^{1/p}$$

For $1 \leq p < \infty$. Assume all quantities are integrable and finite. Do we know that $A \geq B$ or $A \leq B$ for all functions f ? Prove your assertion. Since this is a known result, you cannot just cite a theorem. Try to prove it directly by using the fact that

$$\|h\|_p = \sup_{\|g\|_q \leq 1} \int hg, \quad \frac{1}{p} + \frac{1}{q} = 1.$$

2. (20 points) Suppose $f_n \rightarrow f$ a.e. for all $x \in X = (0, 1)$ and $\sup_n \|f_n\|_{L^2(X)} \leq M$ for some M fixed. Do we know that $\lim_{n \rightarrow \infty} \|f_n - f\|_{L^2(X)} = 0$? If, in addition, $\lim_{n \rightarrow \infty} \|f_n\|_{L^2(X)} = \|f\|_{L^2(X)} < \infty$. Do we know that $\lim_{n \rightarrow \infty} \|f_n - f\|_{L^2(X)} = 0$?
3. (15 pts) Suppose μ is a probability measure (on a set, say, \mathbb{R}^n) and $f \geq 0$ with $\int f d\mu = 1$. Assuming that the following integral is finite, prove that

$$S(f) := \int f \log f d\mu \geq 0.$$

Prove that, for any bounded real function h ,

$$\int h f d\mu - S(f) = \int [h - (\log f)] f d\mu \leq \log \left[\int e^h d\mu \right]$$

Notice that you will need to use the condition $\int f d\mu = 1$.

4. (20 pts) Let X_1, X_2, \dots, X_n be identically independent random variables with $\mathbb{E}X_j = 0$, $\mathbb{E}X_j^2 = 1$ and $\mathbb{E}|X_j|^3 \leq M < \infty$. This problem gives a proof of CLT with an error bound. Let ϕ be a real function such that the first three derivatives are bounded, i.e., $\sum_{j=0}^3 \|\phi^{(j)}\|_\infty \leq M < \infty$. Let Y_j be i.i.d. normal distribution with mean zero and variance one. Rewrite

$$\begin{aligned} & \mathbb{E}\phi\left(\frac{1}{\sqrt{N}} \sum_j X_j\right) - \mathbb{E}\phi\left(\frac{1}{\sqrt{N}} \sum_j Y_j\right) \\ &= \mathbb{E}\phi\left(\frac{1}{\sqrt{N}} [X_1 + X_2 + \dots + X_N]\right) - \mathbb{E}\phi\left(\frac{1}{\sqrt{N}} [Y_1 + X_2 + \dots + X_N]\right) \\ &+ \mathbb{E}\phi\left(\frac{1}{\sqrt{N}} [Y_1 + X_2 + X_3 + \dots + X_N]\right) - \mathbb{E}\phi\left(\frac{1}{\sqrt{N}} [Y_1 + Y_2 + X_3 + \dots + X_N]\right) \\ &+ \dots + \mathbb{E}\phi\left(\frac{1}{\sqrt{N}} [Y_1 + \dots + Y_{N-1} + X_N]\right) - \mathbb{E}\phi\left(\frac{1}{\sqrt{N}} [Y_1 + Y_2 + \dots + Y_N]\right). \end{aligned}$$

Prove by using Taylor theorem that

$$\mathbb{E}\phi\left(\frac{1}{\sqrt{N}}[X_1 + X_2 + \dots + X_N]\right) - \mathbb{E}\phi\left(\frac{1}{\sqrt{N}}[Y_1 + X_2 + \dots + X_N]\right) \leq C_M N^{-3/2}$$

Similar bound clearly holds for the differences in the telescoping sum. From here, prove that for any ϕ satisfying $\sum_{j=0}^3 \|\phi^{(j)}\|_\infty \leq M < \infty$,

$$\mathbb{E}\phi\left(\frac{1}{\sqrt{N}} \sum_j X_j\right) - \mathbb{E}\phi(\xi) \leq C_M N^{-1/2},$$

where ξ is a normal distribution with mean zero and variance one.

5. (30 points) Let X_1, X_2, \dots, X_n be identically independent random variables with $\mathbb{E}X_j = 0$ and $\mathbb{E}X_j^2 = \sigma^2$. Let $S_n = X_1 + \dots + X_n$. The weak law of large numbers states that Then for any $\varepsilon > 0$

$$\mathbb{P}\left(\left|\frac{S_n}{n}\right| > \delta\right) \leq \frac{\sigma^2}{n\delta^2} \quad (+)$$

Suppose that, instead of $\mathbb{E}X_j^2 = \sigma^2$, we only know that $(\mathbb{E}|X_j|^p)^{1/p} = M < \infty$ for some $1 < p < 2$. As in the class, let $\hat{X}_j = X_j \mathbf{1}(|X_j| \leq c)$, $\hat{Y}_j = X_j \mathbf{1}(|X_j| > c)$, $a_c = \mathbb{E}\hat{X}_j$ and $b_c = \mathbb{E}\hat{Y}_j$. Clearly, $X_j = \hat{X}_j + \hat{Y}_j$. Then we have

$$\begin{aligned} \mathbb{E}\left|\sum_j (\hat{X}_j + \hat{Y}_j)\right| &\leq \mathbb{E}\left|\sum_j (\hat{X}_j - a_c)\right| + \mathbb{E}\left|\sum_j \hat{Y}_j - b_c\right| \\ &\leq \left[\mathbb{E}\left(\sum_j \hat{X}_j - a_c\right)^2\right]^{1/2} + 2n\mathbb{E}|\hat{Y}_j| \\ &= \sqrt{n}\left[\mathbb{E}(\hat{X}_1 - a_c)^2\right]^{1/2} + 2n\mathbb{E}|\hat{Y}_1| \end{aligned}$$

Prove that

$$\begin{aligned} \mathbb{E}|\hat{Y}_1| &\leq c^{1-p} M^p \\ \mathbb{E}(\hat{X}_1 - a_c)^2 &\leq 2\mathbb{E}\hat{X}_1^2 + 2a_c^2 \leq 4c^{2-p} M^p \\ P\left(\left|\sum_j (\hat{X}_j + \hat{Y}_j)\right| \geq n\delta\right) &\leq 4\delta^{-1} \inf_{c>0} [c^{1-p/2} M^{p/2} n^{-1/2} + c^{1-p} M^p] \end{aligned}$$

To carry out the inf, prove by calculus that, for any $\alpha, \beta > 0$, there is a constant K such that

$$\inf_{x>0} Ax^\alpha + Bx^{-\beta} = KA^{\frac{1}{1+\gamma}} B^{\frac{\gamma}{1+\gamma}}, \quad \gamma = \frac{\alpha}{\beta}$$

(Hint: Using $y = x^\alpha$ will make calculation much simpler.) Now take $\alpha = 1 - p/2, \beta = p - 1$, you can carry out the inf to finally get a bound on $\mathbb{P}\left(\left|\frac{S_n}{n}\right| > \delta\right)$. Besides unimportant constants, what is your final answer?