

# Probability Distribution Workshop

Los Angeles County  
ISAB

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## 1 Introduction

This workshop is a review of common probability distributions. It has two principal goals.

1. Serve as a review for people that have not studied this material for a while.
2. Provide exercise for people to familiarize themselves with their Python and R distributions.

The exercises themselves are detailed in *Jupyter notebooks* for Python and *R markdown* for R.

### 1.1 Availability

The source code for this document along with the Python and R exercises are available from the ISAB GitHub repository.

<https://github.com/lacounty-isab/workshops/tree/master/distributions>

### 1.2 Who are We?

ISAB (Information Systems Advisory Body) is a subcommittee of the CCJCC (Countywide Criminal Justice Coordination Committee) of Los Angeles County, California. ISAB is a multi-jurisdictional organization serving the justice communities within the county. The ISAB Data Science Committee (IDSC) addresses data science issues faced by the community. This article addresses skill development. It is not intended as an endorsement of any one product or technology.

More details on ISAB and CCJCC can be found on their websites.

ISAB <http://ccjcc.lacounty.gov/Subcommittees-Task-Forces/Information-Systems-Advisory-Board-ISAB>

CCJCC <http://ccjcc.lacounty.gov/>

## Random Variables

Since this article is not intended to be a foundational exposition on probability theory, we'll jump right in with a review of random variables. Recall that an event space  $\Omega$  is a set of possible outcomes for an experiment. In the case of flipping a single coin, we would have

$$\Omega = \{\text{heads}, \text{tails}\}$$

A random variable  $X$  is an assignment from the event space to the real numbers.

$$X : \Omega \rightarrow \mathbb{R}$$

An example for the case of a coin flip might be

$$\begin{aligned} X(\text{tails}) &= 0 \\ X(\text{heads}) &= 1 \end{aligned}$$

A less trivial example would be flipping a coin 100 times where the outcome is the number of heads. There are many random variables you could define on this event space.

- the number of heads,
- the number of tails,
- 0 if even number of heads, 1 otherwise,
- greatest number consecutive heads,
- number of heads minus number of tails.

There are many more examples. Usually the mapping is straight forward. In fact, many times the mapping is so straight forward that we forget that the event and the random variable are not the same thing!

The goal of the random variable is to convert the experiment outcomes from an abstract "set of things" to a set of numbers. This isn't hard when the set of things naturally maps to a set

of numbers in a useful way. When it's not so natural, we sometimes have to go back to the definition the right mapping. Ultimately the goal is to get a set of number with which we can analyze the experiment quantitatively.

## 2 Distributions

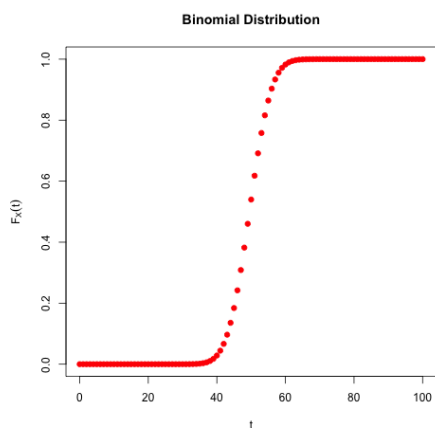
Once we have a random variable defined for an experiment that maps outcomes to real numbers, the next question is to ask how the values of the random variable are likely to be distributed along the real axis after performing the experiment many times. A central concept to answering this question is the *distribution*.

$$F_X(t) = P(X \leq t)$$

If we think back to the example in the previous section of flipping a coin 100 times where the random variable  $X$  represented the number of heads, then  $F_X(t)$  represents the probability that the number of heads is less than or equal  $t$ . Some obvious values are

$$\begin{aligned} F_X(-1) &= 0 \\ F_X(100) &= 1 \end{aligned}$$

The first equality follows from the fact we can't have a negative count of the number of heads. The second equality follows from the fact that we can't encounter more than 100 heads if we only flip the coin 100 times. A graph for all values of  $t$  is shown below.



The function  $F_X$  is often called the *cumulative distribution* since it represents the accumulation of probability as  $t$  covers more of the real axis. In the figure above for a binomial cumulative

distribution for 100 trials of a fair coin, we can see that probability of observing less than 40 heads is close to zero. The probability of observing less than 60 is close to one. That tells us that the most likely scenarios are between 40 and 60.

## 2.1 Discrete Densities

*Discrete distributions* describe experiments where the random variable takes on discrete values. These are often associated with counts. Their notation will often be written as  $P(X \leq n)$  to emphasize the discrete character of  $n$ . (Note that  $n$  doesn't have to be an integer.)

A *discrete density*  $f_X$  corresponding to a discrete distribution  $F_X$  assigns a probability to each discrete value of the random variable. If the random variable only takes integer values, then the density and the distribution are related in the following ways.

$$\begin{aligned} f(n) &= F(n) - F(n-1) \\ F(n) &= \sum_{i \leq n} f(i) \\ \sum_{i=-\infty}^{\infty} f(i) &= 1 \end{aligned}$$

The  $X$  subscripts were dropped for brevity; but in general one writes  $F_X$  or  $f_X$  to associate them with the random variable they represent.

Many useful quantities can be expressed in terms of a density. First, there is the *mean*.

$$E(X) = \sum_{i=-\infty}^{\infty} i \cdot f_X(i) \quad (1)$$

And the variance.

$$\text{Var}(X) = \sum_{i=-\infty}^{\infty} (i - E(X))^2 \cdot f_X(i) \quad (2)$$

A popular alternative expression for the variance can be obtained by expanding the square.

$$\text{Var}(X) = \sum_{i=-\infty}^{\infty} (i - E(X))^2 \cdot f_X(i)$$

$$\begin{aligned}
&= \sum_{i=-\infty}^{\infty} (i^2 - 2iE(X) + E(X)^2) \cdot f_X(i) \\
&= \sum_{i=-\infty}^{\infty} i^2 \cdot f_X(i) - 2E(X) \sum_{i=-\infty}^{\infty} i \cdot f_X(i) - E(X)^2 \sum_{i=-\infty}^{\infty} f_X(i) \\
&= E(X^2) - 2E(X)E(X) - E(X)^2 \\
&= E(X^2) - E(X)^2
\end{aligned} \tag{3}$$

This is often expressed in words as “the mean squared minus the square of the mean.” In the special case where the mean is zero, the variance is equal the mean squared.

The mean and variance are special cases of *moments*. The  $n^{\text{th}}$  moment is defined as

$$\text{Var}(X) = \sum_{i=-\infty}^{\infty} (i - E(X))^n \cdot f_X(i)$$

We’ve already seen that the first moment is the mean. The second moment is the variance. Higher moments are not discussed as often, but they still come in handy. The third moment is the *skew*. It describes the degree to which a distribution is lopsided with respect to its mean. The fourth moment is the *kurtosis*. It describes the extent to which a distribution avoids its mean.

An important tool used in working with distributions is the *moment generating function*. It’s defined as

$$m_X(t) = E(e^{it}) = \sum_{i=-\infty}^{\infty} e^{it} f_X(i) \tag{4}$$

At first glance it looks like way more trouble than it could possibly be worth. But moment generating functions turn out to be important both practically and theoretically.

The theoretical importance derives from analytic function theory which describes a class of functions that are completely determined in some neighborhood by the derivatives of all orders at a single point. Moment generating functions are used to show how, in most cases, a distribution is uniquely determined by the value of all its moments. This result is encountered in many proofs to show how a particular distribution is, in fact, equal to a known distribution.

The practical use is that moment generating functions often simplify the symbolic calculation of the mean and variance for many distributions. The results we need for the mean and variance are the following.

Given  $m_X(t)$  as described in (4), the first and second moments of  $X$  are, respectively,

$$E(X) = m'_X(0) \quad (5)$$

$$E(X^2) = m''_X(0) \quad (6)$$

It's not immediately obvious how this makes things any easier. The examples below will bear it out.

### 2.1.1 Binomial Distribution

We've been using the binomial distribution as an example for much of the introduction. Recall that the experiment is that we perform a sequence of  $n$  independent Bernoulli trials, each of which has probability  $p$  of being successful. The outcome is the number of successful trials. Let's consider what the binomial density looks like.

Let  $i$  be the number of successes. Since  $i$  is a count between zero and  $n$ ,  $i$  cannot be less than zero or greater than  $n$ . So

$$p(i) = 0, i \notin 0, 1, 2, \dots, n$$

For  $i \in 0 \dots n$  there are  $i$  successes and  $n - i$  failures. The probability of such a sequence is  $i^p \cdot (n - i)^{(1-p)}$  since each outcome is independent of the one before it. We then have to account for the number of positions the  $i$  occurrences could have occurred among the  $n$  trials. There are  $n$  ways to pick the first one. For each one of these, there are  $n - 1$  ways to pick the next one. This continues until we have picked  $i$  times down to  $n - i + 1$ . This leads to the following expression with  $i$  factors for the  $i$  chosen positions.

$$n \cdot n - 1 \cdots n - i + 1$$

But this product distinguishes between the order of the successes. For example, if the first two are successes, it counts first then second separately from second and then first. The order doesn't matter. So we divide by the number of permutations of  $i$  elements, which is  $i!$ :  $i$  ways to choose the first one,  $i - 1$  ways to choose the second one, down to 1. So the number of ways we can choose  $i$  elements from  $n$  elements without replacement divided by the number of ways we could have ordered  $i$  elements, we get

$$\begin{aligned} \frac{n(n-1) \cdots (n-i+1)}{i!} &= \frac{n(n-1) \cdots (n-i+1)}{i!} \cdot \frac{n-i}{n-i} \cdot \frac{n-i-1}{n-i-1} \cdots \frac{2}{2} \cdot \frac{1}{1} \\ &= \frac{n!}{i!(n-i)!} \\ &= \binom{n}{i} \end{aligned}$$

The notation in the last expression is read “ $n$  choose  $i$ ”.

Whew!

So a particular occurrence of  $i$  successes and  $n - i$  failures occurs with probability  $p^i \cdot (1 - p)^{n-i}$ . Since there are  $\binom{n}{i}$  ways this can happen, the probability this can happen in *some way* (whichever way, we don't care) is

$$f(i) = \binom{n}{i} p^i (1 - p)^{n-i}, i \in 0, \dots, n \quad (7)$$

This is a probability, the sum of all possible values should add to one. Using the binomial formula, we have

$$\begin{aligned} \sum_{i=-\infty}^{\infty} f(i) &= \sum_{i=0}^n f(i) \\ &= \sum_{i=0}^n \binom{n}{i} p^i (1 - p)^{n-i} \\ &= [p + (1 - p)]^n \\ &= 1^n \end{aligned}$$

Great! The probabilities sum to one. Now let's evaluate the mean and variance of this distribution. From (1) we have to evaluate

$$\sum_{i=0}^n i \binom{n}{i} p^i (1 - p)^{n-i}$$

This is not easy to solve in a closed form. It's one of those times where we appeal to the moment generating function.

$$\begin{aligned} m_X(t) &= E(e^{it}) \\ &= \sum_{i=0}^n e^{it} \binom{n}{i} p^i (1 - p)^{n-i} \\ &= \sum_{i=0}^n \binom{n}{i} (pe^{it})^i (1 - p)^{n-i} \\ &= (pe^{it} + 1 - p)^n \\ &= (pe^{it} + q)^n \end{aligned}$$

In the last step,  $1 - p$  is replaced with  $q$ . We just remember that  $p + q = 1$ . Now differentiate the moment generating function twice with respect to  $t$ .

$$\begin{aligned}
 m'_X(t) &= n(pe^t + q)^{(n-1)}pe^t \\
 m''_X(t) &= n(n-1)(pe^t + q)^{(n-2)}pe^tpe^t + n(pe^t + q)^{(n-1)}pe^t \\
 &= npe^t(pe^t + q)^{(n-2)}((n-1)pe^t + pe^t + q) \\
 &= npe^t(pe^t + q)^{(n-2)}(npe^t + q)
 \end{aligned}$$