Problem 1

Show that the innovations v= Yt-E[Xt|Y1:t-1] in the local level model are mutally independent.

Proof

Let's first find the joint distributions of the observations Yi.

$$\begin{split} \cdot & p(\gamma_{2,1}\gamma_{1}) &= p(\gamma_{2}|\gamma_{1}) \cdot p(\gamma_{1}) \\ \cdot & p(\gamma_{3,1}\gamma_{2,1}\gamma_{1}) &= p(\gamma_{3}|\gamma_{112}) \cdot p(\gamma_{1,1}\gamma_{2}) \\ &= p(\gamma_{3}|\gamma_{12}) \cdot p(\gamma_{2}|\gamma_{1}) \cdot p(\gamma_{1}) \\ \cdot & p(\gamma_{4,1}\gamma_{3,1}\gamma_{2,1}\gamma_{1}) &= p(\gamma_{4}|\gamma_{1:3}) \cdot p(\gamma_{1,1}\gamma_{2,1}\gamma_{3}) \\ &= p(\gamma_{4}|\gamma_{1:3}) \cdot p(\gamma_{3}|\gamma_{1:2}) \cdot p(\gamma_{2}|\gamma_{1}) \cdot p(\gamma_{1}) \end{split}$$

Bosed on the derivations above, we say that the joint density function of the first nobservations is: $P_{Y_1,Y_2,...,Y_n}(y_1,y_2,...,y_n) = P_{Y_1}(y_1) \cdot \prod_{k=2}^{n} P_{Y_1,k+1}(y_k|y_1,y_2,...,y_k)$ (this can be proved easily by induction)

We recall that

$$v_{t} = y_{t} - y_{t} = y_{t} - E[x_{t}|y_{i+1}] = or linear function of of y_{i,...,y_{t-1}}$$
 (2)

Now, our goal is to transform (1) into the joint obensity function of 01,02,..., on. To do this, we will apply Jacobian for multivariate probability density function. It says:

$$f_{V_{1},V_{2},...,V_{n}}(v_{1},v_{2},...,v_{n}) = \frac{f_{Y_{1},Y_{2},...,Y_{n}}(y_{1},y_{2},...,y_{n})}{\left|\frac{\partial(v_{1},v_{2},...,v_{n})}{\partial(y_{1},y_{2},...,y_{n})}\right|}$$
where
$$\frac{\partial(v_{1},v_{2},...,v_{n})}{\partial(y_{1},y_{2},...,v_{n})} = \det \begin{pmatrix} \frac{2y_{1}}{2v_{1}} & ... & \frac{2y_{1}}{2v_{n}} \\ \vdots & & & \\ \frac{2y_{n}}{2v_{1}} & ... & \frac{2y_{n}}{2v_{n}} \end{pmatrix}$$

$$\vdots$$

$$\frac{y_{n}}{2v_{n}} = \frac{1y_{n}}{2v_{n}}$$

Because of (2), we conclude that the matrix is the identity and therefore

$$\left| \frac{\partial (v_1, v_2, ..., v_n)}{\partial (y_1, y_2, ..., y_n)} \right| = 1.$$

Expanding (1), we get:

$$P_{Y_{1},Y_{2},...,Y_{n}}(y_{1},y_{\epsilon,...,Y_{n}}) = P_{Y_{1}}(y_{1}) \cdot P(y_{2}|Y_{1}) \cdot P(y_{3}|Y_{1},y_{2}) \cdot P(Y_{4}|Y_{1:3}) \cdot ... \cdot P(Y_{n}|Y_{1:n-1})$$
(4)

Since the Jocobian = 1 and Pv+(v+) = Py+1yn+1 (y+1yn+1), (4) becomes:

$$\begin{array}{l} \displaystyle \widehat{P}_{V_{1},V_{2_{1},...,V_{n}}}\left(\ \boldsymbol{\sigma}_{v_{1}}\boldsymbol{\sigma}_{\boldsymbol{\varepsilon}_{1},...,v_{n}} \ \boldsymbol{\sigma}_{n} \ \right) \ = \ \displaystyle \widehat{P}\left(\boldsymbol{\sigma}_{1}\right) \cdot \ \boldsymbol{\varphi}\left(\boldsymbol{\sigma}_{\boldsymbol{\varepsilon}}\right) \cdot ... \ \cdot \ \boldsymbol{P}\left(\boldsymbol{\mathcal{V}}_{n}\right) \\ = \ \displaystyle \prod_{t=1}^{n} \ \boldsymbol{\varphi}\left(\boldsymbol{\sigma}_{t}\right) \end{array}$$

Since the joint density factors, we conclude that vi, vi..., on one motivally independent.

MTH9893_HW5_Group13 3/23/17, 9:39 PM

MTH9893 Time Series Analysis HW5

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Problem 1

From the lecture note, $v_1, v_2, \dots v_t$ are joint normal with expection equal zero.

So to prove v_t are mutually independent, it just need to prove v_s , v_t are uncorrelated for any \$s

The prove using "Tower Property":

If
$$H \subset G$$
, then $E[E[Y|H]|G] = E[Y|H] = E[E[Y|G]|H]$
$$Cov(v_t, v_s) = E[v_t v_s]$$

$$= E[E[v_t v_s | Y_{1:s}]]$$

$$= E[E[v_t | Y_{1:s}]v_s]$$

$$= E[E[Y_t - E[X_t | Y_{1:t-1}]|Y_{1:s}]v_s]$$

$$= E[(E[Y_t | Y_{1:s}] - E[E[X_t | Y_{1:t-1}]|Y_{1:s}])v_s]$$

$$= E[(E[X_t + \eta_t | Y_{1:s}] - E[X_t | Y_{1:s}])v_s]$$

$$= E[(E[X_t | Y_{1:s}] - E[X_t | Y_{1:s}])v_s]$$

$$= 0$$

MTH9893_HW5_Group14

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- Wenli Dong

```
In [73]: #import nessesary package
  import math
  import numpy as np
  import pandas as pd
  import statsmodels.api as sm
  from pykalman import KalmanFilter
  from sklearn.metrics import mean_squared_error
  from matplotlib import pyplot as plt
```

```
In [59]: #parsing and reading in data
    # data parsing

SPX = pd.read_excel("HW5.xlsx", skiprows=1, parse_cols='A:B')
APPL= pd.read_excel("HW5.xlsx", skiprows=1, parse_cols='D:E')
data = SPX.merge(APPL,how = 'inner',on = 'Date')

# Change price to daily return
APPL = ((data.PX_LAST_y- data.PX_LAST_y.shift())/ data.PX_LAST_y)[1:]
SPX = ((data.PX_LAST_x- data.PX_LAST_x.shift())/ data.PX_LAST_x)[1:]
```

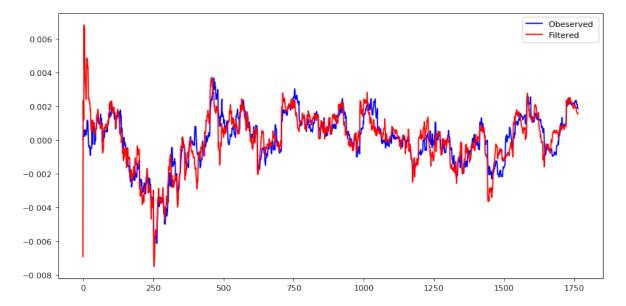
```
In [60]: #get the legth of the return data
length = len(data.index)-1

# add a constant term for the ols regreesion
SPX_sm = sm.add_constant(SPX)
#Xs to store alphas, betas to store betas
Xs = []
betas = []

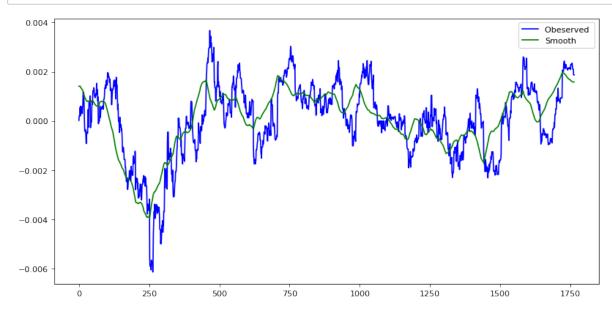
# OLS for window = 63
for i in xrange(63,length +1):
    res = sm.OLS(APPL[i-63:i], SPX_sm[i-63:i]).fit()
    Xs.append(res.params[0])
    betas.append(res.params[1])

# Calculate Ys by definition
Ys = np.array(APPL[62:length+1]) - np.array(betas)*np.array(SPX[62:length+1])
```

```
In [62]: # plot the filtered values with observed value
   plt.figure(figsize=(12, 6), dpi=80)
   plt.plot(Xs,linestyle='-',color='blue')
   plt.plot(filtered_state_means,linestyle='-',color='red')
   plt.legend(['Obeserved ','Filtered'])
   plt.show()
```



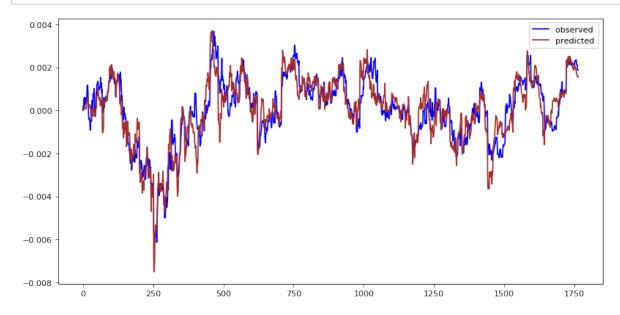
```
In [63]: # plot the smoothing values with observed value
   plt.figure(figsize=(12, 6), dpi=80)
   plt.plot(Xs,linestyle='-',color='blue')
   plt.plot(smoothed_state_means,linestyle='-',color='green')
   plt.legend(['Obeserved ','Smooth'])
   plt.show()
```



In [83]: # Now we try to say something about its predict power
#initalized two list with the length as observations
(predicted_state_means, predicted_state_covariances) = kf.filter(Ys)
predicted_state_means[0] = 0
predicted_state_covariances[0] = 0

use observation at t and mean/variance estimates at t-1 to predict m
ean at time t
for i in range(1,len(Ys)):
 (predicted_state_means[i], predicted_state_covariances[i]) = kf.f
ilter_update(predicted_state_means[i-1], predicted_state_covariances[i])

In [65]: # plot the smoothing values with observed value
 plt.figure(figsize=(12, 6), dpi=80)
 plt.plot(Xs,linestyle='-',color='blue')
 plt.plot(predicted_state_means,linestyle='-',color='brown')
 plt.legend(['observed ','predicted'])
 plt.show()



In [84]: mse = mean_squared_error(Xs,predicted_state_means)
 print "mse over varriance of observed value: ",mse/np.var(Xs)
 print "\nFrom the picture above, we can see that the predicted value d
 eviate a lot from the observed value sometime,\
 but generally speaking, it does show the direction and trend of the ob
 served value."
 print "\nFurthur more, by comparing the mean square error and the vari
 ance of the observed alpha\
 we find that the mse is small, which indicated a good prediction."

mse over varriance of observed value: 0.198049182801

From the picture above, we can see that the predicted value deviate a lot from the observed value sometime, but generally speaking, it do es show the direction and trend of the observed value.

Furthur more, by comparing the mean square error and the variance of the observed alphawe find that the mse is small, which indicated a g ood prediction.