► Limit Theorems

Some Particular results about S_n and \overline{X}_n

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The following are important cases when we can find exactly the Distribution of S_n and/or \overline{X}_n :

▶ If $X_k \sim \mathcal{N}(\mu, \sigma^2)$, k = 1, ..., n, are Independent, then

$$S_n = X_1 + ... + X_n \sim \mathcal{N}(n \cdot \mu, n \cdot \sigma^2)$$

and

$$\overline{X}_n = \frac{X_1 + X_2 + ... + X_n}{n} \sim \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right).$$

▶ If $X_k \sim Pois(\lambda)$, k = 1, ..., n, are Independent, then

$$S_n = X_1 + ... + X_n \sim Pois(n \cdot \lambda).$$

▶ If $X_k \sim Bernoulli(p)$, k = 1, ..., n are Independent, then $S_n = X_1 + ... + X_n \sim Binom(n, p)$;

If $X_k \sim Binom(m,p)$, k=1,...,n, are Independent, then $S_n = X_1 + ... + X_n \sim Binom(n \cdot m,p).$

The Weak LLN

Now, what can be said about S_n and \overline{X}_n in the general case? LLN and CLT help us in this matter, they describe the *asymptotic* properties of these guys:

The Weak Law of Large Numbers, WLLN:

If $X_1, X_2, ..., X_n$ are IID, with finite $\mathbb{E}(X_1)$ and Variance $Var(X_1)$, then

$$\overline{X}_n = \frac{X_1 + X_2 + ... + X_n}{n} \stackrel{\mathbb{P}}{\to} \mathbb{E}(X_1), \qquad n \to +\infty,$$

i.e., for any $\varepsilon > 0$,

$$\mathbb{P}\left(\left|\frac{X_1+X_2+\ldots+X_n}{n}-\mathbb{E}(X_1)\right|\geq\varepsilon\right)\to 0, \qquad n\to+\infty.$$

Note: This means that for any $\varepsilon > 0$, the chances that \overline{X}_n is far from $\mathbb{E}(X_1)$ more than ε , is very small, if n is large.

The Strong LLN

The Strong Law of Large Numbers, SLLN, Kolmogorov If $X_1, X_2, ..., X_n$ are IID, with finite $\mathbb{E}(|X_1|)$, then

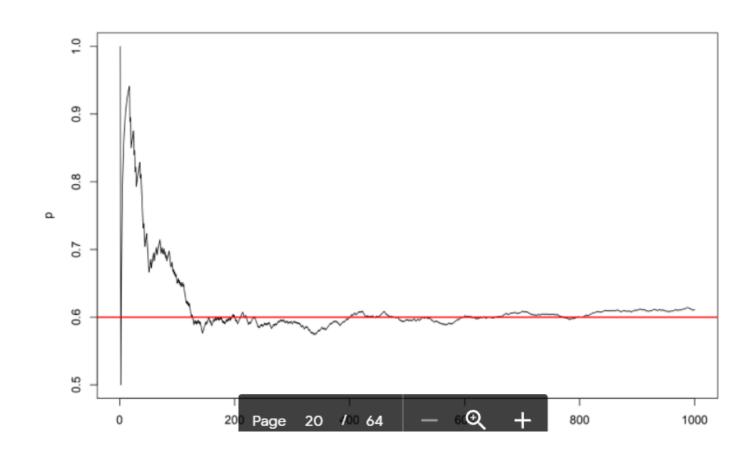
$$\frac{X_1+X_2+\ldots+X_n}{n}\stackrel{a.s.}{\to} \mathbb{E}(X_1), \qquad n\to+\infty$$

that is,

$$\mathbb{P}\left(\lim_{n\to+\infty}\frac{X_1+X_2+\ldots+X_n}{n}=\mathbb{E}(X_1)\right)=1.$$

Visualization of the LLN

```
set.seed(111); n <- 1000; expect <- 0.6
X <- rbinom(n, 1, expect)
S <- cumsum(X); p <- S/(1:n)
plot(p, type = "l")
abline(expect,0, col = "red", lwd = 2)</pre>
```



Supplements, LLN

Sometimes we are required to calculate limits of the form:

$$\lim_{n\to+\infty}\frac{g(X_1)+g(X_2)+\ldots+g(X_n)}{n}$$

in the *Probability* or *a.s.* sense, for some nice function g.Clearly, under the condition that $\mathbb{E}(g(X_1))$ and $Var(g(X_1))$ are finite, or $\mathbb{E}(|g(X_1)|) < +\infty$, we will have

$$\frac{g(X_1)+g(X_2)+...+g(X_n)}{n} \stackrel{\mathbb{P},a.s.}{\longrightarrow} \mathbb{E}(g(X_1)), \qquad n \to +\infty.$$

Say, for example,

$$\frac{X_1^2 + X_2^2 + \ldots + X_n^2}{n} \xrightarrow{\mathbb{P}, a.s.} \mathbb{E}(X_1^2), \qquad n \to +\infty.$$

CLT

The LLN says that the values of \overline{X}_n are concentrated around $\mathbb{E}(X_1)$. But it is not giving us an idea about how the values are distributed around that mean value. CLT helps us in this regard.

To give the general idea of the CLT, let us use the following transform: for a r.v. X, let us denote

$$Z = Standardize(X) = \frac{X - \mathbb{E}(X)}{\sqrt{Var(X)}} = \frac{X - \mathbb{E}(X)}{SD(X)},$$

the Standardization (normalization, scaling) of X. Clearly,

$$\mathbb{E}(Z) = 0$$
 and $Var(Z) = 1$.

Basic Idea of the CLT

The basic idea of the CLT is the following: if we have a sequence of IID r.v. X_n , and we consider their sum S_n or their average \overline{X}_n , then

$$Standardize(S_n) \stackrel{D}{\longrightarrow} \mathcal{N}(0,1)$$

and

$$Standardize(\overline{X}_n) \stackrel{D}{\longrightarrow} \mathcal{N}(0,1).$$

Btw, trivially,

$$Standardize(\overline{X}_n) = Standardize(S_n),$$

and these two versions of CLT are the same.

So for large n, the Distribution of the $Standardize(S_n)$ or $Standardize(\overline{X}_n)$ is approximately Standard Normal. And this independent of the initial Distribution of X_k !

Easy and beautiful, isn't it?

CLT, in the Sums form

Assume X_n be a sequence of IID r.v. with finite expectation $\mu = \mathbb{E}(X_i)$ and variance $\sigma^2 = Var(X_i)$.

We consider sums

$$S_n = X_1 + X_2 + ... + X_n$$
.

We Standardize S_n :

$$Standardize(S_n) = \frac{S_n - \mathbb{E}(S_n)}{\sqrt{Var(S_n)}}.$$

Now we use

$$\mathbb{E}(S_n) = n \cdot \mu, \quad Var(S_n) = n \cdot \sigma^2.$$

Then,

$$Standardize(S_n) = \frac{S_n - \mathbb{E}(S_n)}{\sqrt{Var(S_n)}} = \frac{S_n - n \cdot \mu}{\sqrt{n} \cdot \sigma}.$$

The CLT states:

$$\frac{S_n - n \cdot \mu}{\sqrt{n} \cdot \sigma} \stackrel{D}{\longrightarrow} \mathcal{N}(0,1).$$

CLT, in the Averages form

Again, assume X_n be a sequence of IID r.v. with finite expectation $\mu = \mathbb{E}(X_i)$ and variance $\sigma^2 = Var(X_i)$. We consider the Averages

$$\overline{X}_n = \frac{X_1 + X_2 + \dots + X_n}{n}.$$

We Standardize \overline{X}_n :

$$Standardize(\overline{X}_n) = \frac{\overline{X}_n - \mathbb{E}(\overline{X}_n)}{\sqrt{Var(\overline{X}_n)}}.$$

We use

$$\mathbb{E}(\overline{X}_n) = \mu, \quad Var(\overline{X}_n) = \frac{\sigma^2}{n}.$$

Then,

$$Standardize(\overline{X}_n) = \frac{\overline{X}_n - \mathbb{E}(\overline{X}_n)}{\sqrt{Var(\overline{X}_n)}} = \frac{\overline{X}_n - \mu}{\sigma/\sqrt{n}}.$$

The CLT states:

$$\frac{\overline{X}_n - \mu}{\sigma/\sqrt{n}} \xrightarrow{D} \mathcal{N}(0,1).$$

Two forms of CLT

Of course, these two forms of the CLT are the same: we have

$$Standardize(S_n) = \frac{S_n - \mathbb{E}(S_n)}{\sqrt{Var(S_n)}} = \frac{S_n - n \cdot \mu}{\sqrt{n} \cdot \sigma}$$

and

$$Standardize(\overline{X}_n) = \frac{\overline{X}_n - \mathbb{E}(\overline{X}_n)}{\sqrt{Var(\overline{X}_n)}} = \frac{\overline{X}_n - \mu}{\sigma/\sqrt{n}}.$$

Now,

$$\frac{S_n - n \cdot \mu}{\sqrt{n} \cdot \sigma} = \frac{n \cdot (\frac{S_n}{n} - \mu)}{\sqrt{n} \cdot \sigma} = \frac{\frac{S_n}{n} - \mu}{\frac{\sigma}{\sqrt{n}}} = \frac{\overline{X}_n - \mu}{\sigma/\sqrt{n}},$$

SO

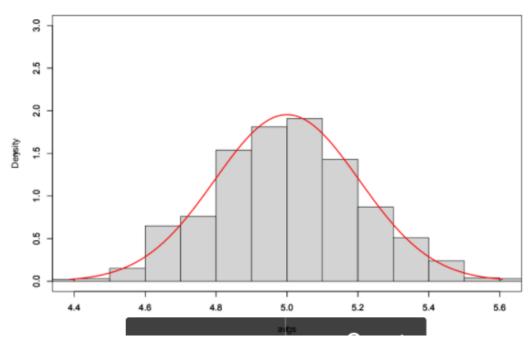
$$Standardize(S_n) = Standardize(\overline{X}_n).$$

Hence, the above two versions of CLT are the same, just one is in terms of S_n , the other one is in terms of \overline{X}_n .

CIT Vicually

```
n <- 600 # Sample Size
m <- 1000 # no of Samples
rate <- 0.2
x <- rexp(n*m, rate = rate)
theo.mean <- 1/rate #theoretical mean
theo.sd <- 1/rate #theoretical SD
m <- matrix(x, ncol = m); d <- data.frame(m)
avgs <- sapply(d, mean)
a = theo.mean-3*theo.sd/sqrt(n); b = theo.mean+3*theo.sd/sqrt(n)
hist(avgs, freq = F, xlim = c(a, b), ylim=c(0,3))
par(new = T)
t <- seq(a,b, 0.01)
y <- dnorm(t, mean = theo.mean, sd = theo.sd/sqrt(n))
plot(t,y, type = "l", col="red", lwd = 2, xlim = c(a,b), ylim=c(0,3))</pre>
```

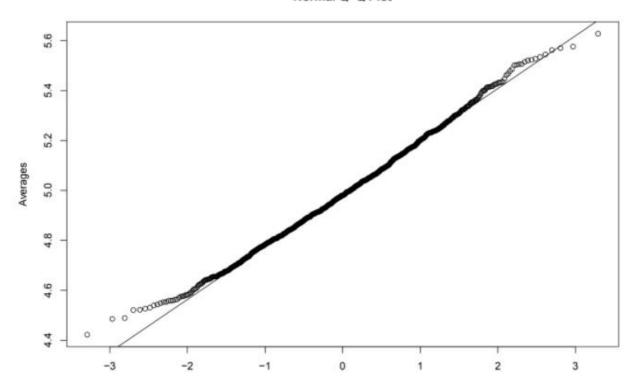
Histogram of avgs



CLT, Visually, v2

```
n <- 600 # Sample Size
m <- 1000 # no of Samples
rate <- 0.2
x <- rexp(n*m, rate = rate)
m <- matrix(x, ncol = m); d <- data.frame(m)
avgs <- sapply(d, mean)
qqnorm(avgs, ylab = "Averages"); qqline(avgs)</pre>
```

Normal Q-Q Plot



CLT, Roughly

In a non-rigorous way, we can write, for large n (here \approx means approximately distributed as):

$$\frac{S_n - n \cdot \mu}{\sqrt{n} \cdot \sigma} \approx \mathcal{N}(0, 1)$$
 and $\frac{\overline{X}_n - \mu}{\sigma/\sqrt{n}} \approx \mathcal{N}(0, 1)$.

or

$$S_n pprox \mathcal{N}(n\mu, n\sigma^2)$$
 and $\overline{X}_n pprox \mathcal{N}\left(\mu, rac{\sigma^2}{n}
ight)$.

CLT, Roughly -2

Let us summarize:

If X_k -s are independent, have the Mean $\mathbb{E}(X_k) = \mu$ and $Var(X_K) = \sigma^2$, and **are Normally Distributed**, i.e., $X_k \sim \mathcal{N}(\mu, \sigma^2)$, then

$$S_n \sim \mathcal{N}(n\mu, n\sigma^2)$$
 and $\overline{X}_n \sim \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right)$;

so we know the **exact Distributions** of S_n and \overline{X}_n .

If X_k -s are independent, have the Mean $\mathbb{E}(X_k) = \mu$ and $Var(X_K) = \sigma^2$, and **from any Distribution** (but the same Distribution), then

$$S_n \approx \mathcal{N}(n\mu, n\sigma^2)$$
 and $\overline{X}_n \approx \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right)$;

and we know the **asymptotic Distributions** (approximate Distributions for large n) of S_n and \overline{X}_n .

CLT, Berry-Esseen Inequality

Now, quickly about the convergence rate of CLT:

Theorem(18+, Berry-Esseen): Assume X_k are IID r.v.s with finite $\mathbb{E}(X_1) = \mu$, $Var(X_1) = \sigma^2$ and $\mathbb{E}(|X_1|^3)$. Then, for any $n \in \mathbb{N}$,

$$\sup_{x\in\mathbb{R}}|\mathbb{P}(Z_n\leq x)-\Phi(x)|\leq \frac{\mathbb{E}(|X_1-\mu|^3)}{\sigma^3\cdot\sqrt{n}},$$

where

$$Z_n = Standardize(S_n) = Standardize(\overline{X}_n),$$

and $\Phi(x)$ is the CDF of $\mathcal{N}(0,1)$.