

Taming the bias zoo[☆]Hongqi Liu^a, Cameron Peng^{b,*}, Wei A. Xiong^c, Wei Xiong^d^a Chinese University of Hong Kong, Shenzhen, China^b London School of Economics and Political Science (LSE), Houghton Street, London WC2A 2AE, United Kingdom^c Shenzhen Stock Exchange, China^d Princeton University, United States

ARTICLE INFO

Article history:

Received 5 February 2021

Accepted 17 March 2021

Available online 8 June 2021

JEL classification code:

G11

G41

G50

Keywords:

Bias zoo

Excessive trading

Gambling preference

Perceived information advantage

ABSTRACT

The success of behavioral economics has led to a new challenge: many biases offer observationally similar predictions for a targeted financial anomaly. To tame this bias zoo, we combine subjective survey responses with observational data to propose a new approach, one that is robust to question-specific biases introduced through surveys. We illustrate this approach by administering a nationwide survey of Chinese retail investors to elicit their trading motives. In cross-sectional regressions of respondents' *actual* turnover on survey-based trading motives, perceived information advantage and gambling preference dominate other motives, though they are not the most prevalent biases based on survey responses.

© 2021 Elsevier B.V. All rights reserved.

1. Introduction

Over the last few decades, behavioral economists have used keen insights from psychology to explain many

anomalies in individuals' financial decision making.¹ As a byproduct of these successes, there are now multiple behavioral biases proposed to explain each anomaly, and the set of proposed behavioral explanations often varies from one anomaly to another. This “zoo” of biases is not satisfying: quantitatively, it is unlikely that the biases are equally important, and qualitatively, it is possible that a seemingly relevant bias is just a manifestation of a different yet more fundamental one. For the field of behavioral economics to eventually arrive at a unified conceptual framework (one that uses a small set of biases to explain a wide range of phenomena), it is necessary to consolidate the many biases proposed for each anomaly.²

[☆] We thank David Hirshleifer (the editor), an anonymous referee, Nick Barberis, Hui Chen, Thummim Cho, James Choi, Bing Han, Rawley Heimer, Sam Hartzmark, Xing Huang, Lawrence Jin, Christian Julliard, Terry Odean, Søren Leth-Petersen, Andrei Shleifer, Johannes Stroebe, Martin Weber, and seminar participants at Baruch, the CEPR Household Finance Workshop, Chapman, CUHK Shenzhen, Imperial College, INSEAD, LSE, the NBER Behavioral Finance Meeting, Princeton, the Red Rock Finance Conference, and Tilburg. Cameron Peng acknowledges financial support from the Paul Woolley center at the LSE. Wei A. Xiong acknowledges support from the National Natural Science Foundation of China (Project Number 71703104). An earlier version of this paper was circulated under the title “Resolving the Excessive Trading Puzzle: An Integrated Approach Based on Surveys and Transactions.” This study has received exemption from Princeton IRB and LSE Research Ethics.

* Corresponding author.

E-mail address: c.peng9@lse.ac.uk (C. Peng).

¹ See Barber and Odean (2013), Hirshleifer (2015), and Barberis (2018) for recent literature reviews.

² This effort can be thought of as a response to the “lack-of-discipline critique” about behavioral finance that was common in the 1990s, which said that because people may depart from full rationality in various ways, it is too easy to pick biases for a given anomaly by flipping through the pages of a psychology textbook (Fama, 1998).

This consolidation task is challenging because existing explanations, by design, share similar or identical predictions for a targeted anomaly. Therefore, key moments in observational data may not have sufficient power to differentiate one explanation from the others. Some explanations may offer unique predictions along more subtle dimensions, but testing such predictions often requires particular data that are difficult to collect. Comparing the relative importance of different explanations is even more demanding, as it requires empirical proxies of different explanations in the same data sample.

Choi and Robertson (2020) adopt a survey-based approach to directly compare many factors that may affect investment decisions. Specifically, they administer a survey to elicit individual responses to an exhaustive list of economic mechanisms, ranging from expectations and risk concerns to biases and transactional factors. Survey responses make it possible to rank the relevance of these factors. Despite the appeal, surveys also raise some methodological concerns due to their subjective nature: respondents may not truthfully report their answers and, even when they do, their subjective responses are noisy and may be influenced by the wording and framing of the questions (Bertrand and Mullainathan, 2001). Such question-specific biases may distort the ranking of survey responses and lead to spurious conclusions.

In this paper, we propose a new approach to consolidate the bias zoo. We design and administer a survey to elicit individual responses to an exhaustive list of behavioral biases. However, we depart from a purely survey-based approach by using subjective survey responses to explain respondents' *actual* investment behaviors. This integrated approach enables us to overcome some of the challenges faced by the existing approaches that are based on observational data or survey responses alone.

First, by collecting people's attitudes towards a variety of economic mechanisms, our approach allows for an apples-to-apples comparison of the attitudes' explanatory power for the targeted anomaly. This feature, shared by the purely survey-based approach as in Choi and Robertson (2020), overcomes the difficulty faced by using observational data alone when they are used to differentiate multiple explanations that are observationally equivalent. Second, because the targeted anomaly is measured using field data rather than survey responses, our subject of interest is immune to noise introduced through surveys. This avoids the biases that arise when survey responses are used for both dependent and independent variables, in which case correlated measurement errors on both sides of the regression can significantly bias the coefficients (Bertrand and Mullainathan, 2001). Third, by regressing observational outcomes on survey responses at the individual level, the regression coefficients are not affected by question-specific biases arising from misunderstanding or prejudice. That is, a common bias in survey responses favoring a certain factor (potentially due to the survey's choice of words or framing of questions) does not affect the cross-sectional explanatory power of that factor, even though the common bias may distort the ranking of the factor relative to other factors purely based on survey responses.

We illustrate this integrated approach with an attempt to resolve the so-called "excessive trading puzzle." Initially documented by Odean (1999) and Barber and Odean (2000) for U.S. retail investors and later found to be prevalent across many markets, the puzzle is characterized by three robust facts about retail investor behavior: (1) retail investors perform poorly relative to the market index before fees, (2) transaction costs make their performance even worse, and (3) those who trade more often perform worse. The literature has proposed several behavioral explanations, including overconfidence, realization utility, gambling preference, sensation seeking, social interaction, and low financial literacy, on top of standard arguments such as portfolio rebalancing and liquidity needs (see Table 1 for a complete list of the explanations along with the references). This myriad of behavioral explanations represents a classic example of a bias zoo, and it remains unclear which biases matter the most for excessive trading.

We ran a nationwide survey among Chinese retail investors, with respondents randomized across regions and brokers. As of the end of 2018, the Chinese equity market stood as the second largest in the world. As highlighted by Allen et al. (2020), one of the most striking features of the Chinese stock market is the coexistence of low returns and high trading volume, with more than 80% of total trading volume coming from retail investors. To understand this phenomenon, our survey asked a series of questions related to financial literacy and return expectations, and, most importantly, included an exhaustive list of behavioral biases and motives as potential explanations of excessive trading. Implicit in our empirical design is the assumption that trading motives persist over time and explain cross-sectional differences in trading behavior. The survey took place in September 2018 and received responses from more than 10,000 investors. We then merged these responses with account-level transaction data from the Shenzhen Stock Exchange.

We conduct three sets of exercises that elucidate our approach. In the first set of exercises, we show that, by and large, there is robust statistical consistency between what people *say* and what they *do*: our survey responses are largely in line with the trading patterns they are designed to capture. For the four trading motives that can be directly matched with our observational data, we find: (1) survey-based measures of gambling preference explain the tendency to buy lottery-like stocks, (2) survey-based measures of extrapolation explain the tendency to buy stocks with positive recent returns, (3) survey-based measures of risk aversion explain the holding of stocks with greater volatility, and (4) survey-based return expectations explain changes in stock holdings. By demonstrating this consistency for a wide range of questions, we provide external support for prior studies that test finance theories based on surveys alone.

In our second set of exercises, we illustrate how our approach can be used to tame the bias zoo. In the first step, we run a series of cross-sectional regressions of turnover on each trading motive *alone*. These regressions confirm that many of the previous explanations for excessive trading also hold true in our sample. In the second step, we

Table 1
Summary of theories on trading volume.

Theory	Forms of Representation	Papers
Overconfidence	1. overplacement 2. miscalibration of uncertainty	Daniel et al. (1998, 2001), Odean (1998), Benos (1998), Glaser and Weber (2007), Dorn and Huberman (2005), Graham et al. (2009), Ben-David et al. (2013)
Extrapolation	1. upward trend to continue 2. downward trend to continue	Barberis et al. (2018), Jin and Sui (2021), Da et al. (2021), Liao et al. (2021)
Neglect of trading costs	1. underestimation of transaction fees 2. knowledge about the bid-ask spread 3. salience of transaction fees	Barber and Odean (2000), Barber et al. (2009), Bordalo et al. (2012)
Gambling preferences	1. blockbuster 2. lotteries	Barber and Odean (2000), Shefrin and Statman (2000); Barberis and Huang (2008), Kumar (2009), Barber et al. (2008), Bordalo et al. (2012)
Realization utility	1. utility from realizing gains 2. disutility from realizing losses	Barberis and Xiong (2009, 2012), Ingersoll and Jin (2013), Frydman et al. (2014), Liao et al. (2021)
Sensation seeking	1. novelty seeking 2. volatility seeking	Grinblatt and Keloharju (2009), Dorn and Sengmueller (2009), Gao and Lin (2015)
Information	1. perceived information advantage 2. dismissiveness (of others' information)	Kyle and Wang (1997), Odean (1998), Gervais and Odean (2001), Scheinkman and Xiong (2003), Eyster et al. (2019)
Social and advisor influences	1. advisor influence 2. social influence	Shiller (1989), Kelly and Grada (2000), Hong et al. (2004), Hong et al. (2008), Pool et al. (2015), Han et al. (2020)
Financial or investment literacy	1. compounding 2. inflation 3. diversification 4. asset risk 5. definition of stocks 6. definition of bonds 7. the PE ratio 8. definition of mutual funds	Van Rooij et al. (2011), Lusardi and Mitchell (2007, 2011, 2014), Grinblatt et al. (2011)

include *all* survey-based trading motives as regressors to compare their explanatory power in a horse race.

Our findings from the second set of exercises are three-fold. First, two trading motives stand out in the horse race as the dominant drivers of excessive trading: gambling preference and perceived information advantage. Both motives have sizable explanatory power: while the standard deviation of the monthly turnover rate in our sample is 123%, gambling preference can explain up to 21% and perceived information advantage can explain up to 24%. They contribute to annualized transaction fees of 0.6% and 0.7%, respectively, implying substantial investment consequences. Despite strong cross-sectional explanatory power, these two motives are only supported by 37% and 18% of the respondents and rank much lower than several other trading motives in the survey based on supporting rates. Therefore, for survey responses, there is an important difference between their cross-sectional explanatory power and their simple ranking based on supporting rates, as the latter may be biased due to question-specific biases in the survey.

Second, for several trading motives, coefficients turn from large and significant in the baseline to small and insignificant in the horse race. For instance, we have constructed two measures of sensation seeking, one for novelty seeking and the other for volatility seeking. While

both measures exhibit positive and significant explanatory power in univariate regressions, their explanatory power is significantly reduced in the horse race. In comparison, the explanatory power of both gambling preference and perceived information advantage is robust across various specifications. Therefore, having an apples-to-apples comparison among a large set of behavioral biases, which would be virtually impossible to conduct based on observational data alone, allows us to narrow down to the few behavioral biases that are the most important.

Third, in both the baseline regressions and the horse race, we report a number of “null” results. Contrary to popular accounts, low financial literacy, social interaction, and neglect of trading costs do not appear to contribute to more trading in our setting. Perhaps the most consistent, yet surprising set of results concerns neglect of trading costs. Although we have constructed three different measures, none of them explain turnover as predicted. Furthermore, in a randomized experiment, we give half of the respondents a “nudge” by having them read a message with pictures illustrating how excessive trading hurts their investment performance due to transaction costs. The treatment group, however, does not exhibit any difference in turnover after the “nudge,” leading to a further questioning of the role of neglect of trading costs in driving excessive trading.

In the third and last set of exercises, we compare transaction-based and survey-based measures of trading motives by constructing two measures of gambling preference. While the transaction-based measure shows greater explanatory power for turnover, it is also correlated with several other trading motives. We therefore conclude by discussing the pros and cons of these two approaches. On the one hand, carefully designed survey questions can directly target a specific trading motive without being confounded by other trading motives. However, survey responses are subject to measurement noise at the individual level and are thus less powerful. On the other hand, although transaction-based measures are less subject to measurement noise, they may simultaneously capture multiple trading motives and are less reliable in isolating a single economic mechanism.

Our paper contributes to the growing literature that uses surveys to construct economic variables that are otherwise difficult to measure. For example, Dorn and Huberman (2005), Glaser and Weber (2007), and Dorn and Sengmueller (2009) have previously combined survey data with observational data to study the excessive trading puzzle. They each focus on one or two behavioral biases or trading motives: risk aversion and perceived financial knowledge in Dorn and Huberman (2005), two forms of overconfidence (overplacement and miscalibration) in Glaser and Weber (2007), and sensation seeking in Dorn and Sengmueller (2009). Our study expands the idea of combining survey responses with observational data by running a horse race among an exhaustive list of trading motives. In the absence of such a horse race, significant effects associated with one motive may simply reflect other motives, as in the case of sensation seeking in our analysis. Furthermore, by showing consistency between survey-based trading motives and observed trading behaviors, we provide external validation to the survey responses in our sample.

Another strand of the literature (e.g., Greenwood and Shleifer, 2014; Barberis et al., 2018; Giglio et al., 2021a, 2021b) uses survey-based expectations to analyze people's belief dynamics. Similar to our paper, Giglio et al. (2021a, 2021b) combine survey expectations with mutual fund holdings data to validate the consistency between survey expectations and actual investments. In other related studies, Chincio et al. (2021) use surveys to uncover subjective perceptions of consumption risk in investors' portfolio choice decisions, while Epper et al. (2020) use experiments to measure the time discount rate and examine its relation to wealth accumulation over time. These studies again tend to focus on a single variable or bias. In this regard, our paper is most closely related to Choi and Robertson (2020), who also use survey responses to compare a large number of potentially relevant factors for investment decisions. Employing a different framework by combining survey responses with observational data, our study not only provides external validation to survey responses but also overcomes question-specific biases that might distort a simple ranking of survey responses.

The rest of the paper is organized as follows. In Section 2, we explain the survey design and report some stylized facts about Chinese investors from the survey. In Section 3, we validate survey responses using actual

trading data, compare survey-based trading motives in a horse race, and discuss the implications of these results. In Section 4, we compare survey-based and transaction-based measures. We conclude in Section 5. We also report detailed information about the survey and additional analysis in an Internet Appendix.

2. The survey

In this section, we first discuss the survey design to further elaborate, from an econometric point of view, the advantages of our approach and the concerns that may arise in our framework. We then explain the procedure for survey distribution and data collection. Finally, we summarize some basic facts from the survey.

2.1. Survey design

We designed the survey to test and differentiate a large set of trading motives developed by the literature. Table 1 provides a summary of all the trading motives we consider. A trading motive may take several forms. For instance, overconfidence comes in at least three forms. The first is overplacement, which means that people have overly rosy views of their abilities relative to others. The second is miscalibration of uncertainty, which means people are too confident in the accuracy of their beliefs. The third is perceived information advantage, which means that people believe they have superior information over others. The survey included at least one question for each form of overconfidence, as detailed in Sections 1 and 2 of the Internet Appendix.

In our research design, we first survey a pool of investors about their trading motives and then compare the different motives' explanatory power for an actual trading behavior. Specifically, we consider a standard linear model to relate investor turnover y to a list of trading motives x_1, \dots, x_K :

$$y^i = \beta_0 + \beta_1 \tilde{x}_1^i + \dots + \beta_K \tilde{x}_K^i + \varepsilon_i, \quad (1)$$

where i indexes individuals. Surveys allow us to collect noisy measures of the trading motives $\{\tilde{x}_k^i\}$, where $\tilde{x}_k^i = x_k^i + u_k^i$ with u_k^i representing the measurement error of variable x_k induced through the survey. In our approach, we rank the trading motives not by the values of their noisy measures $\{\tilde{x}_k^i\}$ but instead by their cross-sectional explanatory power $\{\beta_k\}$ for the observed turnover. Many respondents may agree with a particular trading motive, but we can confirm its relevance *only if* we also observe that these respondents trade more than other respondents.

A first advantage of our approach is that by directly observing dependent variable y from transaction data, we can avoid spurious coefficients due to mismeasurement in dependent variables. To see why, suppose that instead we use the survey-based measure \tilde{y}^i , where $\tilde{y}^i = y^i + \delta^i$ and δ^i reflects the survey-induced measurement error in y^i . In our analysis, this corresponds to using self-reported turnover rather than actual turnover. When δ^i is white noise and uncorrelated with x_k^i , the OLS coefficients will not be biased. However, as discussed by Bertrand and Mullainathan (2001), if measurement error δ^i is correlated

with x_k , which is highly likely, OLS coefficients can be severely biased. For example, it may be more difficult for overconfident investors to recall bad trading experiences in the past, resulting in a negative bias (δ^i) in the self-reported turnover rate. If we use overconfidence as an explanatory variable x_k , coefficient β_k can be substantially biased downward due to the negative correlation between x_k and δ .

A second advantage of our approach is that question-specific biases in the measurement of x_k will not bias the OLS coefficients. Suppose that $u_k^i = u_k + \eta_k^i$, where $u_k \neq 0$ and η_k^i is pure white noise. For instance, if a trading motive is poorly phrased and subsequently invites misperception or prejudice, there is a question-specific bias, u_k , against that motive among all survey respondents. This bias reduces the mean of the survey responses, \bar{x}_k^i , and thus may distort the ranking of the motive relative to other motives in the purely survey-based approach used by Choi and Robertson (2020).³ In contrast, in the cross-sectional regression (1), the question-specific bias u_k will not bias the OLS estimate of β_k , as the bias will be absorbed by the intercept. Therefore, when a question is poorly worded and generates, on average, less-favorable responses from the respondents, it will not bias the OLS estimation as long as the downward bias is common to all respondents and does not interact with individual characteristics that we do not control for.

Although our approach is not subject to question-specific measurement bias, other measurement issues may still arise. We now discuss their implications and our solutions.

White noise. We start with η_k^i , the white noise component in the measurement error. This component will produce an attenuation bias in the estimate of the regression coefficients, β_k . The magnitude of this bias may differ across motives, depending on the variance-covariance matrix of the K explanatory variables and the variance of each type of white noise. For instance, a larger variance of white noise contributes a greater attenuation bias, which leads to the common concern that insignificant results from horse races may simply reflect a lot of white noise. To the extent that white noise in measurement errors makes it more difficult to detect significant factors, any significant factor from our analysis would be even more important in practice.

Wording, scaling, and mental effort. Measurement errors could also arise due to the wording of questions and scaling of answer options. For example, people give rather different answers to the following two questions: “Do you think the United States should *forbid* public speeches against democracy?” and “Do you think that the United

States should *allow* public speeches against democracy?” Similarly, when the scaling of answer options changes, subjects may report their answers differently, as they might be anchored by the choice of options. Lack of mental effort typically makes these issues worse, as subjects may not read the questions in detail and may be more likely to choose answers that appear first or last in the list of options. As discussed above, when wording or scaling induces a question-specific bias, it will be absorbed by the intercept and will not bias the OLS coefficients. When the bias is individual-specific and more prevalent in certain demographic groups, then individual characteristics should be properly accounted for. In our main regressions, we control for an exhaustive list of demographic variables.

To mitigate biases induced by wording, we adopted a jargon-free protocol. We phrased the questions as accurately as possible when describing the underlying concept, while ensuring that they remained comprehensible to the average respondent. To confirm that respondents could immediately understand each question, we ran a series of pilot tests among the general population on a Chinese version of Mechanical Turk and solicited feedback on the survey design. The overwhelming majority of respondents found the questions easy to understand. This also ensured that subjects typically did not find it mentally burdensome to complete the survey.

To deal with biases induced by scaling, we designed all questions to be multiple choice with a standardized menu of answer options. There are two types of qualitative questions. The first type, “agreement,” asked respondents whether they agree or disagree with a statement that describes a particular motive driving trading decisions. Answer options included “strongly agree,” “agree,” “neutral,” “disagree,” “strongly disagree,” “do not know,” and “decline to answer.” The second type, “frequency,” asked respondents how often they consider a particular motive when they trade. Answer options included “always,” “often,” “sometimes,” “rarely,” “never,” “do not know,” and “decline to answer.” We also sought quantitative answers for certain trading motives (e.g., estimates of transaction fees to measure neglect of trading costs). In such cases, we provided several options, each covering a specific value range. The standardization of answer options ensured that any bias resulting from the design of answer options would be small and consistent across all the questions.

Attitudes. Survey questions typically elicit people’s attitudes toward a certain description or statement. However, a clear attitude may not always exist. If forced to give an answer, people may randomly pick one, causing further noise in measurement. To deal with this “no-attitude” issue, we include two answer options, “do not know” and “decline to answer,” so that respondents do not feel compelled to give an answer when they do not have a clear one in mind.

Social desirability. Another concern, particularly relevant to eliciting biases and mistakes, is that respondents may want to look good in front of others and avoid giving answers that may sound stupid or wrong. This concern arises naturally in interview-based surveys, in which respondents di-

³ More precisely, Choi and Robertson (2020) asked respondents in their survey to rank competing mechanisms specifically for a given decision variable, the y variable in Eq. (1). Therefore, an alternative way to interpret their survey responses is that the responses may already capture the respondents’ own estimates of the beta coefficients, β_k . In our view, this interpretation may further complicate the task assigned to the respondents and result in other sources of bias due to the more complex inference process.

rectly interact with the interviewer. Since our survey was conducted online, our respondents had less of a need to appear socially desirable. Moreover, we carefully phrased the questions to be objective and avoided making any inference about a certain behavior being right or wrong. For instance, to elicit a measure of overconfidence, instead of asking “How overconfident do you think you are?” we asked respondents to only self-assess their investment performance. Later, we would compare it to their actual performance to get our measures of overconfidence.

Others. We discuss three final considerations in our survey design. First, survey responses are subjective: they capture how people consciously perceive themselves to be making investment decisions.⁴ A common criticism of subjective surveys in economic analysis is the “as if” critique: respondents may not consciously perceive a factor to be important, but they still behave as if it were (Friedman, 1953). However, subjective perceptions are still useful for several reasons: they shed light on the true decision-making process, they help differentiate competing theories, and they have predictive power for implications of debiasing mechanisms on individuals’ future behaviors (Choi and Robertson, 2020). It is also inherently interesting to know about people’s subjective reasoning. We add that subjective perceptions are also relevant for nudge interventions: if a nudge is targeting a bias that people are not even aware of, it is unlikely that the intervention would successfully produce the desired outcome (DellaVigna and Linos, 2020).

Second, at a general level, there is a significant trade-off between “being rigorous” and “being intuitive” in the design of survey questions. To be fully rigorous in investigating trading motives, the corresponding survey questions needed to comprehensively capture all their aspects. For instance, to fully grasp realization utility requires calibrating a utility function that captures not only the different attitudes between gains and losses but also the shape of the utility function in different regions. Such a design would make the survey exceedingly long and would unavoidably include academic jargon, which, as discussed above, would immediately raise issues related to wording and mental efforts. The psychology literature also documents an attribute substitution bias, whereby participants do not respond to complicated questions but rather answer a related question that is easier to respond (Kahneman and Frederick, 2002). In light of these concerns, we used the “being intuitive” design to make the phrasing as intuitive as possible to laypeople.

Third, post survey, we designed our empirical strategy with the aforementioned measurement issues in mind. First, we validated survey responses with actual trading behavior and found strong consistency between survey responses and transaction data. This provides further validation of our survey design. Second, we encoded all survey-based trading motives into dummy variables. This standardization minimizes the variation of measurement errors

across all survey-based trading motives and facilitates an apples-to-apples comparison.

The final survey consisted of four main parts. The first part contained eight questions measuring financial literacy. These questions included the classic “big three” questions as well as several other widely used questions to measure financial literacy (Lusardi and Mitchell, 2007, 2011). At the end of this section, we also asked respondents to assess how many questions they answered correctly. This allowed us to construct a measure for overconfidence based on financial literacy. The second part represented the core of the survey, in which we asked respondents to answer a series of questions related to various trading motives. We postpone a more detailed discussion of this part to Section 2.3. The third part asked about basic demographic characteristics, including name, gender, date of birth, province, city, education, income, net worth, phone number, brokerage firm, and broker branch. While many of these variables served as control variables in subsequent analysis, they also provided crucial identifying information that enabled us to locate each correspondent in the transaction database. Finally, for a randomly selected group of respondents (the treatment group), we also included a fourth “nudge” section. We explain the “nudge” and discuss the results in more detail in Section 3.8.

2.2. Data

We administered the survey through the Investor Education Center of the Shenzhen Stock Exchange (SZSE). As part of its regular operations, the Investor Education Center annually surveys domestic retail investors to assess their financial literacy and trading motives. In 2018, we began to collaborate with the center to redesign the survey with the aforementioned research question in mind. Our target sample size was 10,000 investors, a size that would provide sufficient statistical power and was feasible to implement. To ensure that the survey sample was nationally representative, we randomized across branch offices of China’s ten largest brokers. Specifically, we selected 500 branch offices across 29 provinces (and regions) and required each branch office to collect at least 20 valid responses. The number of branch offices allocated to each province (region) was proportional to the total trading volume from that province (region) in 2017.

The survey took place in September 2018, and respondents were given two weeks to complete it.⁵ A valid response had to be completed within 30 min of the start time. Respondents could open the survey using their personal computers or their smartphones.⁶ We collected an

⁴ In the language of Adam Smith, respondents are effectively asked to act as the “impartial spectators” to evaluate the reasons and drivers behind their own decisions (Gramp, 1948).

⁵ The SZSE Center first distributed the survey link to each broker’s headquarter. The headquarter then redistributed it to the preselected branches, after which local client managers sent the survey to their clients (investors), likely via phone calls or WeChat messages. Once an investor had completed the survey, the manager would record the investor’s name, phone number, and branch name. This information was then sent back to us for verification purposes.

⁶ To boost the response rate, we included the logos of both the SZSE and the Shenzhen Finance Institute on the front page of the survey. We also explicitly included a confidentiality agreement to make respondents feel more secure about their answers. Finally, we used monetary rewards

initial sample of 12,856 respondents. We report the distribution of respondents across brokers and provinces in Table A1 of the Internet Appendix. By design, respondents are evenly distributed across the ten brokers, with only slight variation. In terms of geographic variation, areas that are more financially developed (e.g., Guangdong, Zhejiang, Jiangsu, and Shanghai) are more strongly represented in our sample.

Table 2 reports a more detailed summary of the sample's demographic characteristics. Overall, the sample is balanced in gender and highly educated: more than half of the respondents have a college or higher degree. Respondents are primarily middle-aged: about half of the sample are between ages 30 and 50. They were also quite wealthy: the median annual income is around 200,000 RMB, and the median household net worth is around 500,000 RMB. Overall, our sample represents a relatively well-educated, wealthy set of retail investors, and this means that any results we find should not be simply interpreted as an average effect. Instead, to the extent that rich and sophisticated investors are less affected by behavioral biases in their investment decision making, our results may serve as a lower bound.

Finally, we eliminate respondents who shirked by examining the total amount of time spent on the survey. We show, in Fig. A2 of the Internet Appendix, that it took a median respondent about eight minutes to complete the survey, and that 95% of respondents finished within 20 min. Respondents who spent less than three minutes on the survey experienced a sharp drop in their financial literacy score, suggesting that they may have shirked during the survey. In the subsequent analysis, we dropped these observations, reducing our sample size to 11,268.

2.3. Survey results

We now summarize the survey questions and their responses. Our empirical strategy is not to rank trading motives by their supportive rates but to compare their explanatory power for actual turnover in a cross-sectional regression. Nevertheless, it is useful to have an overall picture of the survey responses.

Financial literacy. Table 3 reports the summary statistics for the eight questions on financial literacy. In addition to the classic “big three” questions on interest rates, inflation, and diversification (Lusardi and Mitchell, 2014), we included five other questions to capture additional dimensions of financial (or investment) literacy.⁷ Panel A shows that, out of all eight questions, seven have a correct-answer rate above 75%. The only exception is the question about the relation between interest rates and bond prices. Panel B shows that more than 80% of the respondents correctly an-

swered at least six questions. In fact, one-third were correct on all eight questions. Panel B also shows the distribution of self-assessed scores, which is similar to that of the actual scores. Overall, investors in our sample display a high level of financial literacy.⁸

Overconfidence. Overconfidence is an important concept in behavioral finance and has been adopted by various models to explain a wide range of anomalies in financial markets, including excessive trading, use of leverage, price momentum and reversals, and asset bubbles.⁹ The literature also suggests that overconfidence may be present in several closely related, albeit distinct, forms: overplacement of ability, miscalibration of uncertainty, and overprecision of information. We designed questions to capture each of these forms.

Overplacement of one's own ability is perhaps the most direct form of overconfidence. We constructed two measures of this form, one by the difference between self-assessed and actual performance and the other by the difference between self-assessed and actual literacy scores.¹⁰ In Table 4, Panel A reports their summary statistics. In constructing overplacement of performance, self-assessed performance is measured by the self-reported rank of investment performance among all investors in 2017, while actual performance is measured by the actual rank in the population. At this point, we have not yet merged survey responses with transaction data, so Panel A only reports the distribution of self-assessed performance. The patterns suggest that respondents are optimistic about their performance: almost two-thirds believe that their performance is better than average, while only a quarter believe that it is below average. Panel A also reports the second measure, overplacement of literacy. Overall, respondents do *not* overestimate their financial literacy, which is perhaps not surprising, given the sample's overall high level of financial literacy.

Overconfidence may also show up as miscalibration of uncertainty, as suggested by Alpert and Raiffa (1982).¹¹ We measure miscalibration of uncertainty by the difference between two estimates: one for upside returns and the other for downside returns. This is based on two questions asking respondents to estimate how much the stock market will go up or down next year with 10% probability. The difference between the two estimates gives an 80% confidence interval. The rational benchmark (based on historical market volatility) suggests that this difference should be

as incentives. Specifically, among those who completed the survey, 20 would be randomly selected to receive a gift card worth 500 RMB (around 80 USD), and 1,000 would receive a gift card worth 50 RMB (around 8 USD).

⁷ These questions are related to the concept of risks and volatility (Question 4), the definitions of shareholders, the price-to-earnings ratio, and mutual funds (Question 5, 7, and 8), and the relation between interest rates and bond prices (Question 6).

⁸ Lusardi and Mitchell (2014) show that the fraction of respondents who correctly answer all “big three” questions ranges from 3% (Russia) to 57% (Germany). In contrast, 70.4% of investors in our survey correctly answer all “big three” questions. One possible reason is that their surveys typically draw respondents from the general population, whereas ours draws from investors already participating in the stock market.

⁹ See, for example, Kyle and Wang (1997), Daniel et al. (1998, 2001), Odean (1998), Gervais and Odean (2001), Scheinkman and Xiong (2003), Glaser and Weber (2007), and Barber et al. (2020).

¹⁰ Dorn and Huberman (2005) and Barber et al. (2020) use a similar measure for perceived financial knowledge.

¹¹ Ben-David et al. (2013) show that 80% confidence intervals provided by firm executives for the subsequent year's stock market return only cover 36% of the realizations, and they use the surveyed confidence interval to measure the executives' overconfidence.

Table 2

Summary statistics for the investor population, survey respondents, and investors in the main sample.

This table shows the summary statistics for the investor population, all respondents, and the main sample. The population's characteristics are obtained from the centralized database at the Shenzhen Stock Exchange. Survey respondents are the 12,856 investors who have completed the survey. The main sample includes 4,671 survey respondents that: (1) can be identified in the Shenzhen Stock Exchange centralized database, and (2) held at least one SZSE stock during the two-year window before the survey. Gender, education, and age are from either the SZSE centralized database or the survey answers. Investment age and trading characteristics in 2017 are calculated from trading records and thus missing for the survey respondents who cannot be identified in the SZSE centralized database. Income and net worth are obtained from the survey and thus missing for the investor population. See Table A2 in the Internet Appendix for more details about variable definitions.

Gender	Population	All Respondents	Main Sample	Income (RMB)	Population	All Respondents	Main Sample
Male	71.70%	54.00%	54.40%	< 20K	NA	3.80%	2.08%
Female	28.30%	46.00%	45.60%	20 K to 100K	NA	17.20%	16.42%
				100 K to 200K	NA	29.50%	30.08%
				200 K to 500K	NA	29.50%	30.16%
				500 K to 1M	NA	12.60%	13.74%
				1 M to 2M	NA	4.20%	4.47%
				2 M to 10M	NA	2.10%	2.46%
				10 M and above	NA	1.20%	0.58%
				Net worth (RMB)			
				< 20K	NA	4.80%	2.57%
				20 K to 100K	NA	12.30%	9.59%
				100 K to 500K	NA	27.50%	25.50%
				500 K to 1M	NA	22.30%	23.91%
				1 M to 2M	NA	21.90%	25.16%
				2 M to 10M	NA	6.50%	8.05%
				10 M and above	NA	4.80%	5.22%
				Trading characteristics in 2017			
				Maximum value of investment (in thousand RMB)	639	NA	1,250
				Annual turnover rate	9.4	NA	8.3
				Annual raw return rate	–3.90%	NA	–1.20%
Education							
Middle school or blow	7.30%	8.60%	8.10%				
High school	24.70%	15.60%	18.35%				
Professional school	26.00%	21.90%	24.83%				
College	23.60%	44.90%	40.95%				
Graduate school and above	3.40%	9.20%	7.77%				
Others	14.80%	0.00%	0.00%				
Age							
< 30	21.30%	27.80%	26.10%				
30 to 40	27.40%	29.10%	27.40%				
40 to 50	24.50%	19.90%	22.40%				
50 to 60	15.10%	14.80%	16.00%				
> 60	11.70%	8.50%	8.10%				
Investment age (in years)							
< 2	10.00%	NA	21.20%				
2 to 6	29.80%	NA	26.20%				
6 to 10	18.00%	NA	17.40%				
> 10	42.20%	NA	35.10%				

Table 3

Survey responses on questions on financial literacy.

This table shows the summary statistics of investors' responses to questions on financial literacy. In Panel A, we show the correct rate by question. In Panel B, we compare their actual and self-assessed performances, where actual performance is measured by the total number of questions answered correctly, and self-assessed performance is measured by the total number of questions one reports to have answered correctly.

Panel A: Correct Rate by Question		
Question		Correct Rate
1. Suppose you had \$100 in a savings account and the interest rate was 2% per year. After 5 years, how much do you think you would have in the account if you left the money to grow?		88.4%
2. Imagine that the interest rate on your savings account was 1% per year and inflation was 2% per year. After 1 year, how much will you be able to buy with the money in this account?		91.5%
3. Do you agree with the following statement? Buying an individual stock is usually less risky than buying a stock mutual fund.		86.2%
4. Normally, which asset displays the highest fluctuation over time?		95.2%
5. Which of the following statements is correct? If somebody buys a stock of firm B in the stock market....		76.3%
6. Normally, when the market interest rate falls, the price of an existing bond will		54.7%
7. What is the P/E ratio?		75.8%
8. Which of the following statements about mutual funds is correct?		90.3%
Panel B: Distribution of Financial Literacy Scores		
Score	Actual	Self-assessed
0	0.4%	0.6%
1	0.7%	0.7%
2	1.7%	1.8%
3	2.3%	4.6%
4	5.1%	6.9%
5	8.9%	13.0%
6	17.9%	16.2%
7	30.1%	17.7%
8	33.0%	32.7%
N/A	0.0%	5.8%

76%, but Panel A of Table 4 shows that most respondents report a much narrower range.

Overconfidence may also show up as overprecision about one's own information. We will describe this measure later, when we discuss information-related questions.

Extrapolation. The behavioral finance literature has also emphasized the tendency of investors to extrapolate past returns as a key driver of stock return predictability and excessive trading.¹² In Table 4, Panel B reports the summary statistics for two questions concerning whether investors form expectations about future returns based on past returns. These two questions elicit investors' extrapolative beliefs in two scenarios. In the first scenario, a stock's price keeps rising, and in the second scenario, a stock's price keeps falling. Respondents are asked whether they believe the stock's price will rise or fall in the future. In both scenarios, more respondents believe in price continuation than reversal, suggesting that Chinese investors, on average, exhibit extrapolative beliefs.

Neglect of trading costs. Barber and Odean (2000) and Barber et al. (2009) show that trading causes retail investors in the United States and Taiwan to underperform relative to the overall market, with more than 60% of underperformance directly due to commissions and transaction taxes. These findings suggest that investors who trade a lot may be neglecting the various fees and taxes associated with trading. As financial regulators across the world use Tobin taxes to curb speculative trading, investors' ne-

glect of trading costs could undermine the effectiveness of such financial policies.

Neglect of trading costs stems from at least two possible sources. The first is underestimation: investors systematically believe the fee is lower than it actually is, possibly due to a lack of financial knowledge. The second is a lack of attention: even when investors have full knowledge about trading costs, it could matter little to their trading because the amount associated with each transaction is small and the concept of trading costs may not come to mind at the time of trading.¹³

To capture these two forms of neglect of trading costs, we constructed three different measures. Panel C of Table 4 reports the summary statistics. First, we directly asked investors to estimate the total transaction costs associated with a round-trip buy and sell at 10,000 RMB. The results show that respondents significantly underestimated trading costs: while, on average, such a round-trip transaction should incur a fee of 15 to 26 RMB, almost 70% of the respondents reported an estimate below the lower bound. The second question asked how often an investor considers transaction costs when trading stocks. Similarly, more than half of the respondents said that they never or rarely do so. The third question targeted the implicit cost of the bid-ask spread by asking whether the respondent agrees that bid-ask spread is a form of trading cost. Around 60% of respondents agreed, while 23% disagreed. Overall, there

¹² See, for example, Barberis et al. (1998), Hong and Stein (1999), Barberis et al. (2015, 2018), Liao et al. (2021) and Jin and Sui (2021).

¹³ Several papers show that manipulating the salience of a stock's purchase price affects the level of the disposition effect (e.g., Frydman and Rangel, 2014; Birru, 2015; Frydman and Wang, 2020). Other papers find that manipulating the salience of taxes affects consumer responsiveness to taxes (e.g., Chetty et al., 2009; Taubinsky and Rees-Jones, 2018).

Table 4

Summary statistics for responses to questions, part I

This table tabulates the distribution of investors' answers to questions related to overconfidence (Q10, Q11, Q13, Q14), extrapolation (Q26, Q27), and neglect of trading costs (Q15, Q16, Q17).

Panel A: Overconfidence											
1. What fraction of retail investors do you think earned higher returns than you in 2017?	< 10% 11.8%	10–20% 13.8%	20–30% 15.8%	30–40% 13.5%	40–50% 12.4%	50–60% 10.4%	60–70% 5.8%	70–80% 3.8%	80–90% 2.2%	> 90% 3.4%	N/A 7.2%
2. Actual score–Self-assessed score	< –4 0.8%	–4 1.8%	–3 5.4%	–2 11.4%	–1 19.7%	0 35.1%	1 17.7%	2 5.6%	3 1.7%	4 0.6%	> 4 0.4%
3. Upside return–Downside return	0% 32.7%	5% 14.9%	10% 9.2%	15% 6.9%	20% 5.2%	25% 5.2%	30% 4.3%	35% 3.4%	40% 3.1%	45% 2.5%	> 50% 12.7%
Panel B: Extrapolation											
1. After a stock’s price keeps rising for a while, I usually believe that the price will rise even further in the future.	Strongly Agree			Agree		Neutral		Disagree		Strongly Disagree	N/A
	4.8%			26.9%		39.3%		22.8%		1.3%	5.0%
2. After a stock’s price keeps falling for a while, I usually believe that the price will fall even further in the future.	Strongly Agree			Agree		Neutral		Disagree		Strongly Disagree	N/A
	4.4%			29.1%		41.9%		18.2%		1.3%	5.3%
Panel C: Neglect of Trading Costs											
1. Estimating the cost of a round-trip buy and sell at the value of 10,000 RMB	0–5 17.3%		5–10 27.7%	10–15 23.6%	15–20 12.8%	20–25 8.4%	25–30 3.7%	30–35 2.1%	>35 5.5%		
2. How often do you consider transaction costs when you trade?	Never 14.6%		Rarely 37.7%	Sometimes 27.0%	Often 13.8%	Always 4.6%	N/A 2.5%				
3. The bid-ask spread is one form of transaction cost (The bid-ask spread is the difference between the lowest ask price and the highest bid price).	Agree		Disagree	Don’t Understand	Don’t Know	N/A					
	59.8%		23.1%	8.5%	7.2%	1.4%					

Table 5

Summary statistics for responses to questions, part II

This table tabulates the distribution of investors' answers to questions related to gambling preference (Q18, Q19), realization utility (Q20, Q21), and sensation seeking (Q22, Q23).

Panel A: Gambling Preference						
<i>Blockbusters</i>						
1. When I trade stocks, I aim to select those stocks whose price would rise sharply in a short period of time so that I can make a lot of money quickly.	Strongly Agree 10.4%	Agree 25.4%	Neutral 33.9%	Disagree 23.0%	Strongly Disagree 4.6%	N/A 2.7%
<i>Lotteries</i>						
2. When I trade stocks, I often think of them as lotteries: I am willing to accept small losses in exchange for the possibility of a big upside.	Strongly Agree 5.5%	Agree 24.9%	Neutral 27.2%	Disagree 32.5%	Strongly Disagree 7.3%	N/A 2.7%
Panel B: Realization Utility						
<i>Winners</i>						
1. Normally, if the price of a stock in your portfolio rose substantially since you bought it, which of these two actions would make you feel happier: holding on to the stock or selling that stock?	Sell 37.2%	Same 23.7%	Hold 25.3%	No Feeling 9.2%	N/A 4.5%	
<i>Losers</i>						
2. Normally, if the price of a stock in your portfolio dropped substantially since you bought it, which of these two actions would make you feel more painful: holding on to the stock or selling that stock?	Sell 22.9%	Same 28.0%	Hold 32.1%	No Feeling 12.2%	N/A 4.8%	
Panel C: Sensation Seeking						
<i>Novelty</i>						
1. I feel excited about getting to know new stocks and new firms.	Strongly Agree 5.9%	Agree 20.3%	Neutral 43.9%	Disagree 21.0%	Strongly Disagree 3.2%	N/A 5.7%
<i>Volatility</i>						
2. I feel excited about the stock market moving up and down.	Strongly Agree 5.4%	Agree 23.4%	Neutral 36.7%	Disagree 26.2%	Strongly Disagree 4.3%	N/A 4.1%

is strong evidence that retail investors in China underestimate or neglect trading costs.

If neglect of trading costs is due to (a lack of) attention, then presenting transaction costs in a more salient manner or reminding investors of these costs more frequently may lead them to trade less. To test this hypothesis, we gave half of the respondents a “nudge”: we increased the salience of trading costs by presenting them in annualized terms and by reminding investors about the negative impact of excessive trading to overall returns. We discuss these results in [Section 3.8](#).

Gambling preference. [Barberis and Huang \(2008\)](#) show that the cumulative prospect theory of [Tversky and Kahneman \(1992\)](#) can generate a preference for gambling stocks, meaning stocks with positively skewed returns. [Bordalo et al. \(2012\)](#) suggest that salience may cause investors to exaggerate the probability of salient payoffs, also leading to a preference for gambling stocks.¹⁴ [Barber and Odean \(2000\)](#) argue that if gambling stocks change over time due to fluctuations of volatility and tail distribution, gambling preference may also contribute to excessive trading by leading some investors to chase gambling stocks and thus trade with other investors.

In [Table 5](#), Panel A shows the responses on the two questions about gambling preference. The first question was whether the respondent aims to select a few blockbuster stocks in order to get rich quickly. The second ques-

tion was whether the respondent consciously perceives trading stocks as buying lotteries in that the respondent is willing to exchange small losses for the small probability of a big gain. About one-third of the respondents agreed or strongly agreed with each statement. In what follows, we differentiate these two questions by labeling the first one as representing “blockbusters” and the second one as representing “lotteries.”

In phrasing these two questions, we had the following design in mind: the “blockbusters” question focuses on the salient upside and deliberately tones down the fact that blockbusters are rare. Therefore, investors who agree with this statement are drawn to the large upside without necessarily assessing its small probability. In the language of prospect theory, these investors tend to over-weight small probabilities. In contrast, the “lotteries” question contains a direct description of lotteries by explicitly stating that large payoffs rarely happen. Therefore, the two questions not only help identify the gamblers among the respondents, but also help differentiate their assessments of the tail probabilities. As we will show, the “blockbusters” question has substantially stronger explanatory power for investor trading.¹⁵

Realization utility. [Shefrin and Statman \(1985\)](#), [Odean \(1999\)](#), [Grinblatt and Keloharju \(2001\)](#) and

¹⁴ [Kumar \(2009\)](#) and [Boyer et al. \(2010\)](#) provide evidence that supports the presence of such a gambling preference. [Barberis et al. \(2020\)](#) study how prospect theory can explain stock market anomalies.

¹⁵ An alternative explanation for the difference between these two questions is that the “blockbusters” question helps to identify the “impatient” gamblers. As the literature does not offer any link between trading volume and the discount rate, we attribute the question’s better explanatory power to incorrect probability assessment rather than to impatience.

Grinblatt and Han (2005) argue that trading can arise as a result of the widely observed disposition effect. Barberis and Xiong (2009, 2012) and Ingersoll and Jin (2013) formalize theories of realization utility, in which realization utility contributes to excessive trading.¹⁶ In Table 5, Panel B reports the summary statistics for the two questions on realization utility. Similar to the questions on extrapolative beliefs, these two questions ask respondents to make investment decisions under two hypothetical scenarios. In the first scenario, the respondent is given a stock whose price has gone up since purchase and is asked which of two actions would bring more personal happiness: selling the stock or holding it. In the second scenario, the respondent instead faces a stock whose price has gone down since purchase and is asked which action would be more painful. According to realization utility, selling winners is more pleasing than holding winners, while selling losers is more painful than holding losers. Survey responses for the two questions are mixed. In the first question, consistent with realization utility, more respondents say selling winners makes them happier. In the second question, however, more respondents report that holding losers is more painful than selling them. In what follows, we differentiate these two questions by labeling the first question as realization utility for winners and the second question as realization utility for losers.

Sensation seeking. Sensation seeking, a measurable psychological trait linked to gambling, risky driving, drug abuse, and a host of other behaviors, has been shown to be an important motivation for trading (Grinblatt and Keloharju, 2009; Dorn and Sengmueller, 2009).¹⁷ We designed two questions to capture two distinct dimensions of sensation seeking: novelty seeking, which says that people derive utility from doing something new, and volatility seeking, which says that people derive utility from doing something risky. In Table 5, Panel C reports the summary statistics for these two questions. Overall, answers to these questions exhibit a similar distribution, but the respondents in general do not exhibit a strong tendency toward sensation seeking.

Information. Economists have long argued that access to private information is a key reason why investors trade in financial markets. However, the classic no-trade theorem posits that when all investors are rational and share the same prior beliefs, asymmetric information can cause them not to trade, due to the concern of adverse selection (Milgrom and Stokey, 1982). Theories of financial market trading with asymmetric information (e.g., Grossman and Stiglitz, 1980; Kyle, 1985) typically involve the presence of noise traders, who trade at losses so that rational traders trade despite the potential concern of adverse selection.

Are retail investors in China rational investors with a genuine information advantage or noise traders who believe they hold superior information even though they do

not? We included two questions in the survey to elicit a respondent's perception of their information advantage. The first question measures one's belief in having an information advantage by asking how often they believe they know stocks better than other investors. A positive response to this question may be associated with a genuine information advantage, but it could also reflect a misperceived information advantage due to overconfidence. This latter possibility potentially reflects a tendency to exaggerate one's own information but not the information of others. Various theoretical models have used this tendency to specify investor overconfidence, the third form of overconfidence we mentioned earlier.¹⁸ Later, we differentiate a genuine information advantage from a misperceived one by examining whether the respondent actually performs better.

The second question measures one's potential adverse selection concerns by asking investors how often they worry that others know stocks better than they do. This question measures dismissiveness of others' information, a form of investor bias that offers distinct implications from overconfidence for equilibrium prices and trading volume (Eyster et al., 2019). Panel A of Table 6 shows that about 18% of the respondents say they often or always believe they have an information advantage, and 47% say they never or rarely believe they face an information disadvantage. Despite the relatively small fraction of respondents who indicate a perceived information advantage, these respondents indeed trade more than others, as we will show later.

Social interaction. Shiller (1984) argues that investing in speculative assets is a social activity because investors enjoy discussing investments and gossiping about others' investment successes or failures. As a result, social influences could lead to excessive trading.¹⁹ We designed two questions to capture social interactions, one about the influence of family, friends, and other acquaintances, and the other about the influence of investment advisors. Panel B of Table 6 shows that around 14% of the respondents say that they are often or always influenced by family, friends, or other acquaintances, while 8% say their investment advisors often or always influence their trading.

Other trading motives. In Table 6, Panel C reports the responses to the two questions related to liquidity needs and rebalancing motives. Overall, only about 11% of the respondents say portfolio rebalancing often or always affects their trading, whereas about 17% say liquidity needs often or always affect their trading. Consistent with prior literature, retail investors do not appear to be considering these rational trading motives in their day-to-day trading activities.

Panel D of Table 6 reports on three standard questions for measuring risk aversion. We elicit investors' risk at-

¹⁶ Frydman et al. (2014) provide neural evidence to support realization utility in financial decision making.

¹⁷ Brown et al. (2018) further argue that sensation seeking may even affect the trading of hedge fund managers.

¹⁸ For example, Kyle and Wang (1997), Odean (1998), Gervais and Odean (2001), and Scheinkman and Xiong (2003) all model overconfidence as stemming from a perceived information advantage.

¹⁹ Hong et al. (2004) provide evidence that stock market participation is influenced by social interaction. Han et al. (2020) develop a model to show that social interaction exacerbates excessive trading among investors.

Table 6

Summary statistics for responses to questions, part III

This table tabulates the distribution of investors' answers to questions related to information (Q24, Q25), social interaction (Q28, Q29), others (Q30, Q31), and risk aversion (Q32, Q33, Q34).

Panel A: Information						
<i>Perceived information advantage</i>						
1. When you decide to trade a stock, how often do you believe that you know the stock better than others?	Never 8.7%	Rarely 27.9%	Sometimes 40.3%	Often 14.5%	Always 3.2%	N/A 5.4%
<i>Dismissive of others' information</i>						
2. When you decide to trade a stock, how often do you worry that other investors know about the stock better than you do?	Never 18.2%	Rarely 28.9%	Sometimes 32.3%	Often 12.6%	Always 2.5%	N/A 5.6%
Panel B: Social Interaction						
<i>Social influence</i>						
1. When you decide to trade a stock, how often are you influenced by your family members, friends, or other acquaintances?	Never 11.6%	Rarely 31.2%	Sometimes 40.0%	Often 11.8%	Always 1.7%	N/A 3.8%
<i>Advisor influence</i>						
2. When you decide to trade a stock, how often are you influenced by your investment advisors?	Never 17.8%	Rarely 35.0%	Sometimes 35.8%	Often 7.2%	Always 1.2%	N/A 3.1%
Panel C: Others						
<i>Portfolio rebalancing needs</i>						
1. When you decide to trade a stock, how often is it that you need to rebalance your portfolio?	Never 9.6%	Rarely 30.5%	Sometimes 44.5%	Often 9.5%	Always 1.7%	N/A 4.2%
<i>Liquidity needs</i>						
2. When you decide to trade a stock, how often is it because you need money somewhere else?	Never 7.0%	Rarely 25.9%	Sometimes 45.0%	Often 14.4%	Always 2.6%	N/A 5.1%
Panel D: Risk Aversion						
1. Suppose you are the only income earner in the family, and you have a good job guaranteed to give you your current income every year for life. You are given the opportunity to take a new, equally good job. With a 50% chance it will double your income, and with a 50% chance, it will cut your income by 20%. Would you take the new job?	Yes 51.6%	No 34.1%	Don't Know 11.3%	N/A 3.0%		
2. Suppose the chances were 50% that it would double your income and 50% that it would cut it by 1/3. Would you take the new job?	Yes 45.3%	No 37.5%	Don't Know 13.8%	N/A 3.4%		
3. Suppose the chances were 50% that it would double your income and 50% that it would cut it by 1/2. Would you take the new job?	Yes 26.0%	No 57.4%	Don't Know 13.2%	N/A 3.5%		

titude by asking whether they would be willing to give up their current stable jobs for other jobs with higher expected income but also higher uncertainty in three hypothetical scenarios. About 34% of the investors were unwilling to take the job with the smallest risk, and 26% were willing to take the riskiest job.²⁰

3. A horse race based on survey responses

In this section, we use survey responses to differentiate various explanations for the excessive trading puzzle. We start by merging the respondents' survey responses with their transaction data. We demonstrate the external validity of survey responses by showing their ability to capture actual trading behaviors. We examine the explanatory power of each trading motive alone for turnover, followed by a horse race among all survey-based trading motives. We also provide some robustness checks and additional evidence for several key motives at the end of the section.

3.1. Merging surveys with transactions

In the third part of our survey, we asked respondents to provide information on various demographic variables, including name, date of birth, broker name, and branch name. This allows us to uniquely identify a substantial fraction of the respondents in the transaction database of the Shenzhen Stock Exchange. Specifically, out of the 11,268 respondents that remain in our sample, we can uniquely identify 6,013 investors. Our transaction data cover January 2018 through June 2019, which nicely straddles our survey date of September 2018. We further require an investor to have held at least one stock in the Shenzhen Stock Exchange during the two-year window before the survey.²¹ This further reduces the sample size to 4,671, which is our *main sample*. Table 2 shows that investor characteristics are comparable between all the re-

²⁰ We offer a comparison between Chinese and U.S. investors in Table A3 of the Internet Appendix.

²¹ An investor may have been invited to take our survey without any stockholdings in the Shenzhen Stock Exchange due to various reasons, including holding mutual funds or ETFs, or holding stocks listed on the Shanghai Stock Exchange.

Table 7

Summary statistics of turnover and portfolio returns

Panel A shows the summary statistics of the monthly turnover, raw return, and net return for investors in the main sample between October 2018 and June 2019. The main sample includes 4,671 survey respondents that (1) can be identified in the Shenzhen Stock Exchange centralized database, and (2) held at least one SZSE stock during the two-year window before the survey. Panel B shows the correlation coefficients among the three variables. See Table A2 in the Internet Appendix for more details about variable definitions.

Panel A: Summary Statistics							
	Min	P25	Median	P75	Max	Mean	Std Dev
Turnover	0.0%	4.7%	35.5%	109.8%	650.6%	84.8%	123.4%
Raw returns	−12.6%	−1.4%	0.0%	2.0%	10.0%	0.0%	3.5%
Net returns	−12.9%	−1.6%	0.0%	1.8%	9.6%	−0.2%	3.6%

Panel B: Correlation Matrix			
	Turnover	Raw returns	Net returns
Turnover	1		
Raw returns	−0.07***	1	
Net returns	−0.16***	0.99***	1

*** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

spondents and those in the main sample, suggesting that the merging process does not induce further biases.

Is the excessive trading prevalent among investors in our sample? Table 7 reports the summary statistics of the monthly turnover and portfolio return for the post-survey sample from October 2018 through June 2019, the nine-month window after the survey. When needed, we also extend the window to cover the nine months before the survey. Table 7 confirms the existence of excessive trading. First, investors trade intensively: the median monthly turnover rate is almost one, suggesting that they fully reshuffle their portfolios almost once every month. Second, their performance is poor: while the monthly return of the Shenzhen Composite Index was about 0.6% from October 2018 through June 2019, the median net return in our sample is 0.0%. Third, those who trade more perform worse: the correlation between turnover and raw returns is −0.07 while the correlation between turnover and net returns is −0.16. These negative correlations are statistically significant and confirm the key findings of Odean (1999) and Barber and Odean (2000).

3.2. Encoding survey-based variables

To make different survey-based variables comparable, we encode them into dummy variables. A detailed description of the construction of these dummy variables is in Table A2 of the Internet Appendix. In a nutshell, for the agreement questions, we code “strongly agree” and “agree” as 1 and other answers as 0; for the frequency questions, we code “always” and “often” as 1 and other answers as 0; and for quantitative questions, we typically use zero as the cut-off value.²² Table 8 reports the summary statistics of these dummy variables and their pairwise correlations. Note that for the multiple questions targeting the same trading motive, their pairwise correlation, highlighted in bold, is generally high, which suggests that their responses are internally consistent.

²² The only exception is that, when we code the question of dismissiveness, we code “never” or “rarely” as 1 and others as 0.

A high supporting rate in the survey for a certain trading motive does *not* necessarily mean that this motive is a key determinant of excess trading, as question-specific biases could be induced by the survey. We filter out such biases by examining the cross-sectional explanatory power of survey responses for actual turnover in the aforementioned regression framework. Only when variation in the survey responses for a given motive explains the cross-sectional variation in turnover can we conclude that the motive is relevant to excessive trading. Column (1) of Table 8 shows the degree to which each trading motive is supported by the respondents in our survey. Several motives, such as overplacement of performance, miscalibration, and underestimation of transaction costs, have strong supporting rates (above 60%). Interestingly, as we will show, these motives do not have the strongest explanatory power for turnover in the cross-section, possibly because these survey questions are easier for the respondents to understand. In contrast, certain other motives, such as gambling preference for blockbusters and perceived information advantage, have substantially stronger explanatory power despite having lower rankings, as indicated by the values in column (1).

3.3. Validating survey responses

There are several widely held concerns about the use of survey responses in testing economic hypotheses. First, respondents may not take the survey seriously and may not truthfully report what they think or believe. Second, even if their responses are truthful, they may not act in a way that is consistent with their responses. Indeed, because most existing papers are limited to the use of either survey data or transaction data only, the literature is still missing a systematic test of the external validity of subjective survey responses from investors.²³

²³ Several earlier examples of such validation exercises are worth noting. Using survey and administrative data from Denmark and Sweden, respectively, Koijen et al. (2014) and Kreiner et al. (2015) show that, while survey-based consumption is noisy at the individual level, it is consistent with actual consumption measured from administrative data. More recently, Giglio et al. (2021a) examine the relation between survey expect-

Table 8

Summary statistics and pairwise correlation coefficients of dummy variables based on survey responses

This table shows the mean value of dummy variables based on survey responses and their pair-wise correlation coefficients. See Table A2 in the Internet Appendix for more details about variable definitions. The bold fonts highlight correlation coefficients for survey responses that capture different aspects of the same mechanism.

Variable	Mean	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18	19	20	21
1 Overplacement, performance	0.67	1.00																				
2 Overplacement, literacy	0.24	0.03	1.00																			
3 Miscalibration	0.69	0.08	0.02	1.00																		
4 Underestimation of transaction costs	0.69	−0.02	0.02	0.00	1.00																	
5 Do not consider transaction costs	0.53	0.03	−0.01	0.01	0.11	1.00																
6 Do not think bid-ask spread is a cost	0.33	−0.01	−0.01	−0.05	−0.08	−0.06	1.00															
7 Extrapolation, up	0.32	−0.01	0.04	0.02	0.00	0.08	−0.09	1.00														
8 Extrapolation, down	0.34	0.00	0.04	0.02	0.00	0.07	−0.10	0.62	1.00													
9 Gambling preference, blockbusters	0.37	−0.01	0.04	−0.01	−0.02	0.05	−0.09	0.25	0.21	1.00												
10 Gambling preference, lotteries	0.30	−0.01	0.04	0.01	0.02	0.07	−0.10	0.24	0.21	0.40	1.00											
11 Realization utility, winners	0.36	−0.03	0.02	0.05	0.07	0.01	−0.09	−0.01	0.05	0.04	0.07	1.00										
12 Realization utility, losers	0.22	0.01	0.02	−0.01	0.03	0.04	−0.08	0.06	0.06	0.04	0.04	0.22	1.00									
13 Sensation seeking, novelty	0.24	−0.03	0.03	0.00	0.03	0.08	−0.12	0.19	0.18	0.18	0.24	0.07	0.12	1.00								
14 Sensation seeking, volatility	0.29	0.00	0.04	0.03	0.03	0.05	−0.12	0.22	0.23	0.22	0.26	0.09	0.13	0.42	1.00							
15 Perceived information advantage	0.18	0.06	0.07	0.01	−0.02	−0.03	−0.03	−0.01	0.01	−0.06	−0.09	−0.02	−0.02	0.01	0.02	1.00						
16 Dismissive of others' information	0.14	−0.02	0.03	−0.03	−0.05	−0.11	0.08	−0.03	0.01	0.02	−0.01	−0.01	−0.04	−0.02	−0.03	0.14	1.00					
17 Social influence	0.13	−0.01	0.02	−0.04	−0.04	−0.02	0.07	−0.01	−0.01	0.06	0.05	0.01	−0.02	0.00	−0.03	0.01	0.22	1.00				
18 Advisor influence	0.07	−0.01	0.01	−0.01	−0.02	0.00	0.03	0.01	0.00	0.00	0.02	0.02	−0.02	0.03	0.01	0.03	0.15	0.32	1.00			
19 Portfolio rebalancing needs	0.17	0.01	0.02	−0.02	−0.07	−0.07	0.07	−0.07	−0.06	−0.06	−0.07	−0.07	−0.02	−0.01	−0.02	0.20	0.17	0.12	0.08	1.00		
20 Liquidity needs	0.10	0.00	0.03	−0.07	−0.07	−0.10	0.08	−0.04	−0.03	0.05	−0.01	−0.07	−0.02	−0.03	−0.02	0.09	0.22	0.21	0.10	0.29	1.00	
21 Risk aversion	0.34	0.02	−0.01	0.01	0.01	0.00	0.06	0.02	0.02	0.00	−0.01	−0.03	−0.01	−0.03	−0.02	−0.01	−0.02	0.00	−0.03	−0.05	−0.01	1.00

Ideally, we would like to validate responses to all the questions asked in the survey, but this is not plausible. For instance, although the survey has several questions regarding the sources of information and the influence of social interactions, it is difficult, if not impossible, to infer these aspects from observational data without collecting additional data and making strong assumptions. Given these limitations, we validate survey responses for a set of four questions with natural empirical counterparts that can be directly constructed from transactions. These questions concern extrapolation, gambling preference, risk aversion, and return expectation. In addition to having straightforward implications about trading behaviors, these questions span a wide range of trading motives: belief formation, preferences, and return expectations. For brevity, we focus on gambling preference in the main text. We briefly talk about other validation exercises and their details are included in Section 8 of the Internet Appendix.²⁴

Gambling preference. We start by measuring gambling behavior from transaction data. Gambling preference motivates investors to buy assets with positively skewed returns. While it seems straightforward to measure gambling behavior based on return skewness, the literature, including Kumar (2009), argues that return skewness is difficult to compute and is not a metric that is sufficiently intuitive to investors. Instead, salient stock characteristics, such as realizations of extreme returns, would attract investors with a gambling preference. This argument is particularly compelling as it connects well with our earlier discussion of gambling preference that is driven by investors' overweighting of tail outcomes (Barberis and Huang, 2008; Bordalo et al., 2012). Motivated by this argument, we take advantage of a unique regulation in the Chinese stock market: the daily price limits rule. This rule states that daily stock returns of individual stocks cannot exceed 10%. We use the total count of up-limit hits (i.e., the number of days with prices hitting the up limit) in a preceding period to proxy for a stock's positive return skewness. As hitting the daily up limit puts a stock in the headlines of the stock exchange, this event is highly salient and attracts attention from investors. Thus, we measure an investor's gambling behavior by the volume-weighted count of up-limit hits over either a month or a quarter, based on all the stocks they added to the portfolio.

Table 9 reports the results when regressing transaction-based gambling behavior on survey-based gambling preference. Panel A uses the total count of up-limit hits over the preceding one-month horizon, while Panel B uses one quarter as the horizon. Recall that we included two survey

questions regarding gambling preference, one about the desire to pick blockbusters to get rich and the other about a conscious perception of stocks being lotterylike. Indeed, responses to the first question significantly explain gambling behavior with a positive sign. On average, the stocks purchased by investors who answered affirmatively to this question have a larger count of up-limit hits by around 0.1 (0.2) times in the preceding month (quarter), and this relation holds in both the pre-survey and post-survey periods. Interestingly, responses to the second question do not explain gambling behavior. We document a similar pattern about their explanatory power on turnover later.

Extrapolation, risk aversion and survey expectations. We perform three additional exercises to validate survey-based measures of extrapolative beliefs, risk aversion, and return expectations, using a method similar to the one before. The results are reported in Tables A5–A7 of the Internet Appendix. First, investors who report having extrapolative beliefs exhibit stronger extrapolative behavior: on average, the stocks they purchase experience 1% higher returns in the preceding month and more than 2% higher returns in the preceding quarter, and this holds in both pre-survey and post-survey samples. Second, consistent with Dorn and Huberman (2005), survey-based measures of risk aversion are negatively associated with holding more-volatile stocks. Third, consistent with Giglio et al. (2021a), survey-based expectations about future stock market returns are positively associated with an increase in stock holdings, but the magnitude, as noted by Giglio et al. (2021a), is relatively small.

Finally, we note that throughout the validation exercises, although the coefficient between the survey response and the targeted trading behavior is highly significant, the *R*-squared is generally small. For instance, in Table 9, across all specifications, the *t*-statistic for gambling preference (blockbusters) remains around 4, but the *R*-squared is consistently below 2.5%. This suggests that although survey responses are consistent with the targeted behavior, much of the variation in the targeted behavior is unexplained. This low *R*-squared could be due to measurement errors in survey responses, but it could also be that the behavior itself is driven simultaneously by multiple factors. We will discuss this important issue further in Section 4.

3.4. Baseline results on turnover

After validating the survey responses, we proceed to examine the relation between survey-based trading motives and turnover. We primarily focus on using survey responses to explain post-survey turnover.²⁵ Table 10 presents the baseline results. In each column, we regress turnover on a particular survey-based trading motive. Most regressions are univariate, except for a few instances where we need to control for additional characteristics.

tations and mutual fund holdings and find that survey expectations are consistent with respondents' mutual fund holdings. Unlike these earlier papers, which study consumption and expectation, our main interest is to validate whether survey-based trading motives reflect investors' actual trading behavior.

²⁴ Note that while we demonstrate consistency between survey responses and trading behaviors, we do not claim that the targeted trading behavior is solely captured by the designed question. Indeed, as we will show later in Section 4, one type of observed behavior (such as purchase of gambling stocks) can be driven by multiple motives. Therefore, the purpose of our validation exercise is simply to demonstrate the relevance and usefulness of survey responses.

²⁵ If we measure turnover at the time of or before the survey, then the exercise is subject to the concern that some common shocks may have affected both survey responses and trading behavior. For instance, a positive shock to one's recent return may lead one to report a higher self-assessed performance (resulting in more overplacement of performance) and to trade more.

Table 9

Validating gambling preferences using gambling behavior

This table studies the relation between survey-based gambling preference and transaction-based gambling behavior. Gambling behavior is measured by the buy-volume (in RMB) weighted average of the past one-month (Panel A) or one-quarter (Panel B) number of up-limit hits based on the stocks an investor purchases in a given sample period. A purchase is considered as an initial buy if the investor holds zero share of the stock before the purchase. Each panel presents OLS regression results based on three sample periods: full (January 2018–June 2019), pre-survey (January 2018–September 2018), and post-survey (October 2018–June 2019). Gambling preference (blockbusters) equals one if an investor answers “Strongly agree” or “Agree” when asked if they aim to make a lot of money quickly through stock investment and zero otherwise. Gambling preference (lotteries) equals one if an investor answers “Strongly agree” or “Agree” when asked if they often think of stocks as lotteries and zero otherwise. See Table 5 for the exact phrasing of the survey questions. Control variables include age, gender, net worth, income, trading experience, account size, and education. *T*-statistics are based on robust standard errors and are reported in parentheses.

Panel A: Volume-Weighted Past One-Month Count of Up-Limit Hits Based on Initial Buys												
	Full sample (2018:01–2019:06)				Pre-survey (2018:01–2018:09)				Post-survey (2018:10–2019:06)			
Gambling preference, blockbusters	0.112*** (3.875)	0.109*** (3.768)			0.087*** (3.640)	0.086*** (3.608)			0.142*** (3.660)	0.139*** (3.573)		
Gambling preference, lotteries			0.038 (1.257)	0.019 (0.653)			0.025 (1.013)	0.018 (0.727)			0.051 (1.237)	0.029 (0.698)
Male		−0.034 (−1.164)		−0.033 (−1.140)		−0.011 (−0.444)		−0.01 (−0.403)		−0.035 (−0.884)		−0.034 (−0.866)
Controls	NO	YES	NO	YES	NO	YES	NO	YES	NO	YES	NO	YES
R2	0.004	0.023	0.000	0.019	0.004	0.017	0.000	0.014	0.004	0.020	0.000	0.016
N	4,145	4,145	4,145	4,145	3,435	3,435	3,435	3,435	3,550	3,550	3,550	3,550
Panel B: Volume-Weighted Past One-Quarter Count of Up-Limit Hits Based on Initial Buys												
	Full sample (2018:01–2019:06)				Pre-survey (2018:01–2018:09)				Post-survey (2018:10–2019:06)			
Gambling preference, blockbusters	0.209*** (4.550)	0.199*** (4.299)			0.174*** (4.354)	0.169*** (4.240)			0.256*** (4.066)	0.239*** (3.774)		
Gambling preference, lotteries			0.091* (1.897)	0.055 (1.144)			0.103** (2.389)	0.086** (1.994)			0.071 (1.107)	0.024 (0.373)
Male		−0.051 (−1.084)		−0.049 (−1.051)		−0.04 (−0.996)		−0.039 (−0.949)		−0.051 (−0.798)		−0.05 (−0.784)
Controls	NO	YES	NO	YES	NO	YES	NO	YES	NO	YES	NO	YES
R2	0.005	0.025	0.001	0.021	0.006	0.017	0.002	0.013	0.005	0.021	0.000	0.017
N	4,145	4,145	4,145	4,145	3,435	3,435	3,435	3,435	3,550	3,550	3,550	3,550

t-statistics in parentheses; *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

Table 10

Univariate regression results on turnover

In this table, we run univariate cross-sectional regressions of each investor's turnover (%) on survey-based trading motives. T-statistics are based on robust standard errors and are reported in parentheses. See Table A2 in the Internet Appendix for more details about variable definitions.

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Actual performance in 2017	4.104*** (5.332)							
Overplacement, performance	15.695*** (2.760)							
Financial literacy, dummy		11.922*** (3.127)						
Overplacement, literacy		1.729 (0.400)						
Miscalibration			1.116 (0.289)					
Underestimation of trading costs				−3.549 (−0.980)				
Do not consider trading costs					−2.143 (−0.548)			
Do not think bid-ask spread is a cost						−15.135*** (−4.254)		
Extrapolation, up							4.379 (1.110)	
Extrapolation, down								3.810 (1.005)
R2	0.007	0.002	0.000	0.000	0.000	0.004	0.000	0.000
	(9)	(10)	(11)	(12)	(13)	(14)	(15)	(16)
Gambling preference, blockbusters	10.924*** (2.878)							
Gambling preference, lotteries		2.750 (0.684)						
Realization utility, winners			7.188* (1.874)					
Realization utility, losers				0.409 (0.093)				
Sensation seeking, novelty					10.184** (2.270)			
Sensation seeking, volatility						11.984*** (2.885)		
Perceived information advantage							21.747*** (4.254)	
Dismissive of others' information								4.778 (1.318)
R2	0.002	0.000	0.001	0.000	0.001	0.002	0.005	0.000
	(17)	(18)	(19)	(20)				
Social influence	−15.647*** (−3.317)							
Advisor influence		−16.469** (−2.708)						
Portfolio rebalancing needs			12.652** (2.423)					
Liquidity needs				−9.974* (−1.853)				
R2	0.002	0.001	0.001	0.001				

t-statistics in parentheses; *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

Columns (1)–(3) report the results on three measures of overconfidence: overplacement of performance, overplacement of literacy, and miscalibration of uncertainty. Out of these three measures of overconfidence, the only one that is significantly and positively related to turnover is overplacement of performance: in column (1), conditional on having the same past performance, investors who self-report having higher performance tend to trade more subsequently. In column (3), miscalibration of uncertainty does not significantly predict future turnover. These results

are consistent with [Glaser and Weber \(2007\)](#), who show that overplacement predicts excess trading while miscalibration of uncertainty does not.

Column (1) also shows that past performance positively predicts future turnover. In column (2), financial literacy positively predicts future turnover. This finding is in sharp contrast to a view that excessive trading may be driven by a lack of financial knowledge. Therefore, improving investors' financial literacy, a policy often advocated in emerging economies such as China, may not be effective.

tive in reducing excessive trading. Furthermore, column (2) shows that overplacement of literacy does not predict future turnover.

Columns (4)–(6) report the results on neglect of trading costs. Surprisingly, none of the three measures we constructed significantly predict future turnover with the predicted sign: in columns (4) and (5), the coefficients are close to zero and insignificant; in column (6), investors who do not understand the bid-ask spread as a form of trading cost trade *less*. The result in column (4) is particularly puzzling, because the measure is constructed using direct estimates of transaction fees in a round-trip trade and should clearly identify investors who underestimate trading costs.²⁶ That we cannot find any supporting evidence despite having constructed three measures for neglect of trading costs gives us pause about its role in explaining investor trading. It is possible that investors who underestimate trading costs are more naïve about financial markets, and that such naivety results in passivity in trading and low volume. It is also possible that this pattern reflects a reverse selection in which investors who trade more incur more total costs and are more aware of their existence. We further examine these issues in Section 3.7.

Columns (7) and (8) report the results on extrapolative beliefs. For the two measures of extrapolation of positive and negative returns, we do not find a strong relation between extrapolative beliefs and turnover. One possibility is that extrapolation generates trading only in a bullish market (Barberis et al., 2018; Liao et al., 2021), but the period we examine is relatively quiet, with the market increasing by just a few percentage points during the nine-month window. Another possibility is that extrapolation alone cannot explain volume and must be combined with some additional forces to generate a trading frenzy (Liao et al., 2021).

Columns (9) and (10) report the results on gambling preference. We find that, consistent with the conjecture in Barber and Odean (2000) and the implications of Barberis and Huang (2008) and Bordalo et al. (2012), investors who are subject to gambling preference trade significantly more. Again, the question about “blockbusters” is much more powerful than the “lotteries” question. This is consistent with the pattern in Table 9, which shows that gambling behavior can be explained by answers to the “blockbusters” question but not by answers to the “lotteries” question.

Columns (11) and (12) report the results on realization utility and show an asymmetry. The first measure (the one that proxies for taking pleasure in selling winners) positively predicts future turnover, whereas the second measure (the one that proxies for feeling pain when selling losers) does not predict future turnover. This pattern is consistent with the implications of realization util-

ity (Barberis and Xiong, 2012), as investors who exhibit realization utility are more willing to let go of stocks in gains and to hold on to stocks in losses.

Columns (13) and (14) report the results on sensation seeking. Both the “novelty-seeking” and the “volatility-seeking” measures positively predict future turnover with a large coefficient. These results are consistent with the finding, in Grinblatt and Keloharju (2009) and Dorn and Sengmueller (2009), that investors most prone to sensation seeking trade more frequently.

Columns (15) and (16) report the results on perceived information advantage and dismissiveness of others' information. Column (15) shows that those who believe they have an information advantage trade more, whereas column (16) shows that those who are more dismissive do *not* trade more. As we discussed earlier, the first measure could capture a particular form of overconfidence if we show that these investors do not deliver better returns; indeed, we do show this later in Section 3.7. The second measure captures the dismissiveness modelled by Eyster et al. (2019). Thus, we find supportive evidence for perceived information advantage in explaining excessive trading, but not for dismissiveness.

Finally, columns (17) and (18) concern two measures of social influence. Interestingly, investors who are more influenced by their family, friends, and investment advisors tend to trade *less*, not more. This pattern does not lend support to the aforementioned literature, which argues that social interaction contributes to the spread of investor sentiment and excessive trading.²⁷ Columns (19) and (20) show that rational trading motives such as portfolio rebalancing needs and liquidity needs can only explain a small part of the variation in turnover across investors.

In sum, Table 10 confirms several of the existing explanations for trading volume: for example, overplacement of performance, gambling preference, sensation seeking, realization utility, and perceived information advantage. Table 10 also shows a number of “null” results for some prominent explanations of excessive trading: for example, lack of financial literacy, neglect of trading costs, dismissiveness, and social interaction.

3.5. Horse race results on turnover

Although the baseline results confirm several of the previous explanations for trading volume, it remains unclear whether their explanatory power will survive once we included all trading motives in the same regression. Such a horse race has not been run before. Table 11 presents the full regression results. In addition to including all the survey-based trading motives, we also include: (1) basic demographic characteristics such as gender, income, net worth, and education; (2) return expectations, to control for differences in optimism and pessimism; and (3) re-

²⁶ Transaction fees are standard and almost homogeneous across different brokers. While some variation across brokers still remains, in our construction we use a rather conservative bound to identify those who underestimate trading costs. In addition, we control for differences in fees across brokers with branch fixed effects.

²⁷ However, we note that recent models of social interactions such as Han et al. (2020) are inherently conditional: social interactions lead to more trading when the market is going up and people are making money. Our tests rely on a period of quiet market reactions and therefore does not test these models directly.

Table 11

Regression results using the full set of trading motives

In this table, we run a multivariate cross-sectional regression of each investor's turnover on all survey-based measures of trading motives. Control variables include age, gender, wealth, income, trading experience, account size, and education. *T*-statistics are based on robust standard errors and are reported in parentheses. See Table A2 in the Internet Appendix for more details about variable definitions.

Dependent Variable: Average Monthly Turnover Ratio (%) (October 2018–June 2019)			
Actual performance in 2017	4.198*** (5.219)	Gambling preference, blockbusters	11.764*** (2.920)
Overplacement, performance	11.549** (2.063)	Gambling preference, lotteries	−1.159 (−0.263)
Financial literacy, dummy	7.065* (1.800)	Sensation seeking, novelty	6.598 (1.360)
Overplacement, literacy	−2.621 (−0.625)	Sensation seeking, volatility	3.632 (0.824)
Miscalibration of uncertainty	−2.989 (−0.764)	Perceived information advantage	15.660*** (2.988)
Do not consider trading costs	−3.989 (−1.071)	Dismissive of others' information	2.942 (0.805)
Underestimation of trading costs	−4.029 (−1.052)	Social influence	−7.839 (−1.616)
Do not think bid-ask spread is a cost	−9.456*** (−2.650)	Advisor influence	−12.089* (−1.943)
Extrapolation, up	−1.255 (−0.254)	Portfolio rebalancing needs	12.571** (2.280)
Extrapolation, down	−1.208 (−0.262)	Liquidity needs	−7.651 (−1.335)
Realization utility, winners	7.049* (1.848)	Risk Aversion	−2.943 (−0.692)
Realization utility, losers	−2.321 (−0.538)	Expected 1-year market return	0.709* (1.901)
Gender: male	21.488*** (6.124)	Controls	YES
		N	4,648
		R ²	0.089

t-statistics in parentheses; *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

cent performance, to control for “mood.”²⁸ Table 11 reveals a number of notable observations.

First, two trading motives stand out in the horse race: gambling preference (“blockbusters”) and overconfidence in the form of perceived information advantage. Both coefficients are quantitatively large and significant at the 1% level. The finding of overconfidence as a key driver of turnover supports the large volume of prior studies emphasizing the roles of overconfidence. Even more interesting, our finding highlights that a particular form of overconfidence through perceived information advantage—rather than other forms such as overplacement of literacy and miscalibration of uncertainty—is most relevant in explaining trading. This form of overconfidence also confirms the specification adopted by Kyle and Wang (1997), Daniel et al. (1998, 2001), Odean (1998), Gervais and Odean (2001), and Scheinkman and Xiong (2003) in modeling investor overconfidence in financial markets.

Our finding of gambling preference as a key driver of investor trading is surprising. Earlier literature tends to treat gambling preference as an important mechanism for understanding demand for lotterylike stocks but hasn't fully established its link with excessive trading. Our finding suggests that gambling preference may also lead investors to trade more. Barber and Odean (2000) conjecture

a mechanism that works as follows. As individual stocks fluctuate in their volatility and tail distribution, the set of lotterylike stocks changes over time. Consequently, investors subject to gambling preference chase one lotterylike stock after another, leading to large trading volume.

We note several interesting aspects of our findings of perceived information advantage and gambling preference as the most powerful factors in explaining turnover. First, in Table 8, perceived information advantage is supported by only 18% of the respondents, and gambling preference (for blockbusters) is supported by 37%. Both are substantially lower than some other factors (with over 60% supporting rates). Therefore, although the two motives affect a smaller fraction of the population, their explanatory power is greater. This contrast echoes our earlier point that question-specific biases could make it challenging to rely on a simple ranking of survey responses to compare the importance of different trading motives in explaining actual turnover.

Second, in Table 8, the correlation coefficient between perceived information advantage and gambling preference for blockbusters is −0.06. The small correlation suggests that overconfidence and gambling preference contribute to trading volume through two orthogonal channels. Below, we present additional evidence to support these trading motives as key drivers of excessive trading.

Third, several trading motives that are significant in the baseline regressions become insignificant or marginally significant in the horse race. They include financial literacy,

²⁸ We also have a specification that includes branch fixed effects to control for clustering at the branch level. Results are essentially unchanged and reported in Table A8 of the Internet Appendix.

sensation seeking for novelty, sensation seeking for volatility, social influence, and advisor influence. The results for the two sensation-seeking measures are particularly striking: while both measures are highly significant in univariate regressions, their significance largely disappears after controlling for other factors, suggesting that their explanatory power is subsumed by other factors. The contrast between sensation seeking and gambling preference is also worth noting, given that the literature sometimes mixes the two. Sensation seeking suggests that investors like to gamble because they derive utility from gambling activities independent of the final payoffs, while gambling preference suggests that the appeal of gambles is ultimately driven by the potential of a large payoff. Our analysis suggests that, while sensation seeking and gambling preference are correlated, gambling preference is the more relevant factor for the observed trading.

Finally, consistent with the finding of Barber and Odean (2001), we report a significant gender effect: on average, the monthly turnover of male investors is 21% higher than that of female investors. Barber and Odean (2001) attribute this difference to overconfidence: men trade more because they are more overconfident. Interestingly, the gender effect in Table 11 persists even after controlling for various forms of overconfidence, suggesting the gender effect may go beyond overconfidence.

3.6. Robustness and subsample analysis

As robustness checks, we report the results from alternative regressions in Section 9 of the Internet Appendix, including specifications in which we bootstrap standard errors, add branch fixed effects as control variables, use a larger sample that includes investors that have not traded for more than two years before the survey, and use a smaller sample that only includes investors who are active around the time of the survey. We also consider alternative measures of turnover, including an equal-weighted version (as opposed to the value-weighted one we use throughout the paper) and a version measured in the nine-month window before the survey (as opposed to in the nine-month window after the survey). Throughout all these specifications, gambling preference and perceived information advantage remain the most powerful factors for explaining turnover.

We also perform two sets of subsample analyses and report the results in Section 9 of the Internet Appendix. In the first one, we split the full sample based on account size and compare the behaviors of small and large investors. Overall, consistent with the notion that small investors are more affected by behavioral biases, we find that the results are slightly stronger among small investors. In the second subsample, we split the full sample based on the fraction of wealth invested in the stock market. In both subsamples, gambling preference and perceived information advantage remain significant factors. However, for investors whose wealth is more invested in the stock market, portfolio rebalancing needs become a more pronounced factor to their trading.

We discuss two limitations of our horse race. First, it is possible that the importance of each mechanism is

time-varying. Without a panel of survey responses, we can only capture a snapshot of their relative importance. For instance, realization utility (Barberis and Xiong, 2012; Liao et al., 2021) and social interactions (Han et al., 2020) may contribute to excessive trading more in a market boom than in a downturn. However, we show, in Table A17 of the Internet Appendix, that the explanatory power of each motive remains stable during the 9-month window before the survey, suggesting relatively persistent importance in the time series. Second, and relatedly, it is also possible that some retail investors learn to debias themselves from past mistakes, so the importance of certain mechanisms may decay over time (Seru et al., 2010). While our cross-sectional setting does not allow us to directly speak to the issue of learning, we note that some recent evidence suggests that retail investors do not appear to learn from their prior mistakes (e.g., Anagol et al., 2021).

3.7. Additional evidence of excessive trading

Trading is not necessarily excessive if more trading is associated with better returns. We further examine the portfolio returns based on investors' responses to the gambling preference and perceived information advantage questions to show that indeed the associated trading is excessive.

Panel A of Table 12 sorts investors into five groups based on their answers to the “blockbusters” question and reports each group's monthly turnover and portfolio return. While this single-sorting approach ignores the correlations of gambling preference with other trading motives, it provides a more granular look at the explanatory power of gambling preference.²⁹ For turnover, there is a monotonically increasing pattern from the least gambling-prone to the most gambling-prone group. This monotonic pattern is present not just in the mean and the median of the monthly turnover rate, but also across various percentiles in the distribution, indicating that this pattern is not driven by outliers. On average, the difference between the “strongly agree” group and the “strongly disagree” group is about 21%, suggesting sizable economic significance; a monthly turnover rate of 21% translates into an annualized transaction fee of 0.6%.

Is the trading associated with gambling preference excessive? The result shows that this is the case: the five groups exhibit similar raw returns before fees. In fact, the “strongly agree” group on average earns –0.35% lower monthly returns than the “strongly disagree” group, although the difference is not statistically significant. Together, the lack of superior performance and the large transaction costs suggest that the trading by the “strongly agree” group is excessive.

We also examine the characteristics of stocks purchased by the five groups of investors in Table A18 of the Internet Appendix. Investors with a survey-based gambling preference tend to buy stocks that are smaller and have a larger market beta, larger counts of daily up-limit hits, and higher

²⁹ Note that the coefficient of gambling preference is virtually unchanged from the univariate regression to the horse race, suggesting that the effect is not affected by other trading motives.

Table 12

Additional analysis of gambling preference, blockbusters and perceived information advantage

In Panel A, we sort investors into five groups based on their answers to the question, “Do you agree with the following statement? When I trade stocks, I often wish to select those stocks whose price would rise sharply in a short period of time so that I can make a lot of money quickly.” In Panel B, we sort investors into five groups based on their answers to the question, “When you decide to trade a stock, how often do you believe that you know the stock better than others?” In each panel, we tabulate the summary statistics of monthly turnover ratios (monthly raw returns) for investors in each group. The last one or two rows report the differences between the bottom and top groups. When testing for the significance of the differences, we use robust standard errors.

Panel A: Sort investors by their answers to the statement about gambling preference, blockbusters								
	Monthly Turnover						Monthly Raw Returns	
	P10	P25	P75	P90	Median	Mean	Median	Mean
1. Strongly disagree	0%	4%	99%	206%	25%	74%	0.19%	0.15%
2. Disagree	0%	3%	100%	222%	31%	77%	0.00%	0.04%
3. Neutral	0%	5%	112%	238%	33%	84%	0.01%	0.11%
4. Agree	0%	7%	117%	248%	42%	90%	0.03%	−0.04%
5. Strongly agree	0%	5%	119%	274%	42%	95%	0.00%	−0.20%
5 – 1	0%	0%	20%	68%	17%	21%**	−0.19%	−0.35%
Annual transaction fee (5–1)	0.00%	0.00%	0.60%	1.96%	0.51%	0.63%		
Panel B: Sort investors by their perceived information advantages								
	Monthly Turnover						Monthly Raw Returns	
	P10	P25	P75	P90	Median	Mean	Median	Mean
1. Never	0%	4%	102%	232%	30%	76%	0.10%	0.12%
2. Rarely	0%	3%	100%	218%	32%	76%	0.07%	0.06%
3. Sometimes	0%	5%	109%	244%	34%	86%	0.00%	0.08%
4. Often	0%	11%	139%	286%	46%	103%	0.00%	−0.13%
5. Always	0%	10%	139%	253%	44%	100%	0.00%	−0.01%
5 – 1	0%	6%	37%	21%	14%**	24%**	−0.10%	−0.13%
Annual transaction fee (5–1)	0.00%	0.18%	1.11%	0.63%	0.42%	0.72%		

*** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

past volatility and past returns. These stocks also perform worse subsequently, confirming that investors with a gambling preference trade in the wrong direction and trade excessively.

Panel B of Table 12 performs a similar exercise by sorting investors into five groups based on their answers to the “information advantage” question. Similar to before, investors who “always” think they have an information advantage exhibit higher turnover than those who “never” think so, for almost all the distribution percentiles we look at. The magnitude is also similar: the difference in the monthly turnover rate between the “always” and “never” groups is about 24%, implying an annual transaction fee of 0.7%.

Is the perceived information advantage supported by superior performance in portfolio returns? This is not the case: the five groups exhibit similar performance before fees, indicating that those who report having an information advantage do not outperform others in selecting better stocks. Accounting for trading fees would make their net performance clearly worse. Thus, the perceived information advantage reflects a form of overconfidence rather than genuinely better information.

3.8. Neglect of trading costs

In both the baseline and the horse race, none of the survey variables for neglect of trading costs can explain turnover in the right direction. This contradicts the popular view that Chinese retail investors trade so much because they neglect trading costs. In Tables 10 and 11, some measures even suggest an opposite pattern, in which in-

vestors with more awareness of trading costs trade more. To further isolate the effect of awareness of trading costs, we have also implemented a randomized experiment.

Among the 500 brokerage branches we distributed the survey to, we randomly selected 250 branches to include an additional “nudge.” The nudge asked the respondent to read a one-page article that highlighted the negative consequences of excessive trading. As shown in Fig. A3 in the Internet Appendix, the article contained a detailed calculation of how much investors lose from frequent trading, along with a quote from Warren Buffett advising investors to buy and hold. Instead of presenting trading costs as a fraction of total transaction value, we made them more salient by presenting the annualized fee rate for a frequent trader. We also included a validation question after the article, asking the respondent to compute the total trading cost for a given level of turnover. Answers to this question help identify those who actually read the article and therefore were treated. We study the effect of this nudge in a difference-in-difference framework and report the results in Table A19 of the Internet Appendix. Overall, the nudge had no effect on reducing trading. One might argue that the nudge was not sufficiently strong, and the treated group may not have read the article carefully. However, we identify an investor as treated only if they were in the treated group and answered the validation question correctly.

Taken together, our analysis suggests that neglect of trading costs is not a key driver of excessive trading. This finding has an important policy implication. Policy makers across the world, including China’s stock market regulator, the China Securities Regulatory Commission (CSRC),

Table 13

Trading characteristics for investors sorted on transaction-based gambling behavior

We construct a measure for transaction-based gambling behavior for each investor in two steps. First, for each of the nine months prior to the survey (January 2018–September 2018), we calculate the past one-month count of up-limit hits of the stock for each buy transaction and then take the transaction value-weighted average across all buy orders. Then, we take the time-series average value weighted by monthly buy values. We then sort investors into five groups according to transaction-based gambling behavior and compare their behaviors after the survey, from October 2018–June 2019. In Panel A (B), we tabulate the summary statistics of monthly turnover ratios (characteristics of stocks bought) for investors in each group. In the last row of each panel, we report the differences between the bottom and top groups. When testing for the significance of the differences, standard errors are adjusted for heteroscedasticity.

	Panel A: Monthly Turnover		Panel B: Characteristics of Stocks Bought						
	Mean	Median	Past 30-day # of Up-limit Hits	Past 30-day Return Volatility (%)	Past 30-day Return (%)	Size (Billion RMB)	Beta	B/M	Future 30-day Return (%)
1 (lowest)	60.37	29.43	0.70	3.30	10.65	36.46	0.94	0.66	−0.91
2	80.76	38.69	0.67	3.36	10.28	35.14	0.95	0.62	−0.91
3	71.91	29.49	0.80	3.41	11.18	29.79	0.99	0.61	−0.81
4	92.69	43.92	0.74	3.48	10.13	23.37	1.04	0.58	−0.88
5 (highest)	157.29	98.45	1.12	3.78	14.63	20.13	1.02	0.59	−2.02
5 – 1	96.92***	69.02***	0.42***	0.48***	3.97***	−16.34***	0.09***	−0.07***	−1.11**

*** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

frequently use Tobin taxes as a policy tool to curb speculative trading in stock markets. To the extent that investors may engage in excessive trading despite their awareness of the trading costs, our finding casts doubt on the effectiveness of Tobin taxes.³⁰

4. Comparing survey-based and transaction-based measures

In our analysis so far, we have taken survey responses as direct measures of trading motives and used them to study why investors trade so much. These survey-based measures have some clear advantages over transaction-based measures. First, well-designed survey questions can measure trading motives in a way that is closer to textbook definitions. Second, survey responses allow researchers to measure a large set of trading motives from the perspectives of the respondents at the same time, including those that are hard to measure from administrative data. However, there are also various concerns about survey data. The primary concern, the one we have already addressed through various validation exercises, is that survey responses may not capture actual trading behavior. A second concern is that survey responses are noisy: perhaps respondents, on average, do answer truthfully, but their responses at the individual level may be noisy. This is a concern that also arises in our setting. For instance, in Table 9, while the relation between survey-based gambling preference and transaction-based gambling behavior is statistically significant, the R -squared is rather small across all specifications.

This concern about noise in survey responses raises a follow-up question: do transaction-based behavioral measures have stronger power than survey-based measures?

We address this question by comparing survey-based and transaction-based measures of gambling behavior. Table 13 reports the results when we sort investors into different groups based on their gambling behavior directly measured from transaction data in the pre-survey sample period. This transaction-based measure turns out to be much more powerful in explaining turnover in the post-survey sample: the difference in the monthly turnover rate between the top and bottom groups is 97%, quadrupling the magnitude of 21% reported in Table 12 based on the survey-based measure of gambling behavior.

While the transaction-based measure of gambling behavior appears to work well, we demonstrate that this measure may capture multiple forces at the same time. We regress the transaction-based measure of gambling behavior on all survey-based trading motives and report the results in Table 14. It is reassuring to see that the survey-based measure of gambling preference is indeed the most powerful explanatory variable in this regression. However, a number of other survey-based trading motives are also significantly correlated with the transaction-based measure of gambling behavior. For instance, investors with a perceived information advantage also gamble more. Therefore, although the transaction-based measure of gambling behavior is more powerful in explaining trading, this measure is partially correlated with other trading motives, and its explanatory power may not come solely from gambling preference.³¹

Taken together, our results show a trade-off between survey-based and transaction-based measures of trading motives. Survey-based measures have stronger power from the economic perspective of qualitatively testing different trading motives, even though they may contain more noise

³⁰ There is mixed evidence on the effects of Tobin taxes in reducing speculative trading and price volatility. See Song and Xiong (2018) for a detailed review of the CSRC's policy interventions in the stock market and Deng et al. (2018) and Cai et al. (2020) for studies of effects of increasing the stamp tax for stock trading in China.

³¹ The transaction-based measure of gambling behavior may also contain effects from other omitted variables. For example, one possible omitted variable is investor attention: investors who pay more attention to the stock market are more likely to be drawn to lotterylike stocks, as those stocks appear more often in the news. While these investors may exhibit gambling-like behavior, their frequent trading is explained by their attention to the stock market.

Table 14

Regressing transaction-based gambling behavior on survey-based trading motives

In this table, we run a multivariate cross-sectional regression of each investor's transaction-based gambling behavior on survey-based trading motives. We construct a measure for transaction-based gambling behavior for each investor in two steps. First, for each of the nine months prior to the survey (January 2018–September 2018), we first calculate the past one-month count of up-limit hits of the stock for each buy transaction and then take the transaction value weighted average across all buy orders. Second, we take the time-series average value weighted by monthly buy values. Control variables include age, gender, wealth, income, trading experience, account size, and education. *T*-statistics are based on robust standard errors and are reported in parentheses. See Table A2 in the Internet Appendix for more details about variable definitions.

Dependent Variable: Volume-Weighted Past One-Month Count of Up-Limit Hits Based on Initial Buys (January 2018–September 2018)			
Actual performance in 2017	−0.009** (−2.533)	Gambling preference, blockbusters	0.071*** (3.598)
Overplacement, performance	0.002 (0.071)	Gambling preference, lotteries	−0.011 (−0.482)
Financial literacy, dummy	−0.031 (−1.478)	Sensation seeking, novelty	−0.032 (−1.518)
Overplacement, literacy	−0.014 (−0.633)	Sensation seeking, volatility	0.022 (1.030)
Miscalibration of uncertainty	0.017 (0.942)	Perceived information advantage	0.049** (2.097)
Do not consider trading costs	0.040** (2.221)	Dismissive of others' information	−0.001 (−0.031)
Underestimation of trading costs	−0.005 (−0.276)	Social influence	−0.005 (−0.178)
Do not think bid-ask spread is a cost	−0.043** (−2.436)	Advisor influence	0.025 (0.647)
Extrapolation, up	0.003 (0.133)	Portfolio rebalancing needs	−0.039* (−1.741)
Extrapolation, down	−0.001 (−0.045)	Liquidity needs	0.021 (0.679)
Realization utility, winners	0.015 (0.843)	Risk Aversion	0.004 (0.205)
Realization utility, losers	0.009 (0.409)	Expected 1-year market return	0.000 (0.266)
Gender: male	0.011 (0.623)	Controls	YES
		N	3,528
		R ²	0.031

t-statistics in parentheses; *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

and thus have weaker statistical power. Transaction-based measures have stronger statistical power, but one measure may reflect multiple mechanisms and the economic interpretations are thus not as sharp as survey-based measures.

5. Conclusion

We propose a new approach to consolidate the zoo of behavioral biases by examining the cross-sectional explanatory power of survey-based measures of different biases for certain observed investor behavior. This approach allows researchers to examine multiple mechanisms at the same time and addresses some econometric issues that arise in a purely survey-based approach. We illustrate this approach by designing and administering a nationwide survey to study why investors trade so much. We highlight a number of new findings. First, survey responses are consistent with actual trading behaviors. Second, overconfidence (in having an information advantage) and gambling preference dominate other trading motives in explaining observed turnover, despite their relatively low supporting rates in the survey. Third, other explanations, such as neglect of trading costs and low financial literacy, do not contribute to excessive trading. Finally, by analyzing the pros and cons of survey-based and transaction-based approaches, we argue that our integrated approach can help

mitigate the concerns faced by each of these approaches alone.

The analysis in this paper has focused on one of the fundamental puzzles in financial markets: why do retail investors trade so much? A similar exercise can be carried out to compare competing mechanisms for other anomalies. It is possible that a few mechanisms are most relevant for explaining multiple anomalies, but it is also possible that different mechanisms are driving different anomalies. In this regard, our exercise takes a necessary first step to eventually tame the bias zoo.

References

- Allen, F., Qian, J., Shan, C., Zhu, J., 2020. The development of the Chinese stock market. In: Amstad, M., Sun, G., Xiong, W. (Eds.), *Handbook of China's Financial System*. Princeton University Press, pp. 283–313.
- Alpert, M., Raiffa, H., 1982. A progress report on the training of probability assessors. In: Kahneman, D., Slovic, P., Tversky, A. (Eds.), *Judgment Under Uncertainty: Heuristics and Biases*. Cambridge University Press, pp. 294–305.
- Anagol, S., Balasubramaniam, V., Ramadorai, T., 2021. Learning from noise: evidence from India's IPO lotteries. *J. Financ. Econ.* 140, 965–986.
- Barber, B.M., Odean, T., 2000. Trading is hazardous to your wealth: the common stock investment performance of individual investors. *J. Financ.* 55, 773–806.
- Barber, B.M., Odean, T., 2001. Boys will be boys: gender, overconfidence, and common stock investment. *Q. J. Econ.* 116, 261–292.
- Barber, B.M., Odean, T., Constantinides, G.M., Harris, M., Stulz, R.M., 2013. The behavior of individual investors. In: *Handbook of the Economics of Finance*, 2. Elsevier, pp. 1533–1570.

- Barber, B.M., Lee, Y.T., Liu, Y.J., Odean, T., 2009. Just how much do individual investors lose by trading? *Rev. Financ. Stud.* 22, 609–632.
- Barber, B.M., Huang, X., Ko, K.J., Odean, T., 2020. Leveraging overconfidence. Unpublished working paper. UC Berkeley and Washington University in St. Louis.
- Barberis, N., Huang, M., 2008. Stocks as lotteries: the implications of probability weighting for security prices. *Am. Econ. Rev.* 98, 2066–2100.
- Barberis, N., Xiong, W., 2009. What drives the disposition effect? an analysis of a long-standing preference-based explanation. *J. Financ.* 64, 751–784.
- Barberis, N., Xiong, W., 2012. Realization utility. *J. Financ. Econ.* 104, 251–271.
- Barberis, N., Shleifer, A., Vishny, R., 1998. A model of investor sentiment. *J. Financ. Econ.* 49, 307–343.
- Barberis, N., Greenwood, R., Jin, L.J., Shleifer, A., 2015. X-CAPM: an extrapolative capital asset pricing model. *J. Financ. Econ.* 115, 1–24.
- Barberis, N., Greenwood, R., Jin, L.J., Shleifer, A., 2018. Extrapolation and bubbles. *J. Financ. Econ.* 129, 203–227.
- Barberis, N., Jin, L., Wang, B., 2020. Prospect theory and stock market anomalies. *J. Financ.* forthcoming.
- Barberis, N., Bernheim, B.D., DellaVigna, S., Laibson, D., 2018. Psychology-based models of asset prices and trading volume. In: *Handbook of Behavioral Economics: Applications and Foundations 1*, I. North-Holland, pp. 79–175.
- Ben-David, I., Graham, J.R., Harvey, C.R., 2013. Managerial miscalibration. *Q. J. Econ.* 128, 1547–1584.
- Benos, A.V., 1998. Aggressiveness and survival of overconfident traders. *J. Financ. Mark.* 1, 353–383.
- Bertrand, M., Mullainathan, S., 2001. Do people mean what they say? implications for subjective survey data. *Am. Econ. Rev.* 91, 67–72.
- Birru, J., 2015. Confusion of confusions: a test of the disposition effect and momentum. *Rev. Financ. Stud.* 28, 1849–1873.
- Bordalo, P., Gennaioli, N., Shleifer, A., 2012. Salience theory of choice under risk. *Q. J. Econ.* 127, 1243–1285.
- Boyer, B., Mitton, T., Vorkink, K., 2010. Expected idiosyncratic skewness. *Rev. Financ. Stud.* 23, 169–202.
- Brown, S., Lu, Y., Ray, S., Teo, M., 2018. Sensation seeking and hedge funds. *J. Financ.* 73, 2871–2914.
- Cai, J., He, J., Jiang, W., Xiong, W., 2020. The Whack-a-mole game: tobin taxes and trading frenzy. *Rev. Financ. Stud.* forthcoming.
- Chetty, R., Looney, A., Kroft, K., 2009. Salience and taxation: theory and evidence. *Am. Econ. Rev.* 99, 1145–1177.
- Chinco, A., Hartzmark, S.M., Sussman, A.B., 2021. Risk-factor irrelevance. Unpublished working paper.
- Choi, J.J., Robertson, A.Z., 2020. What matters to individual investors? evidence from the horse's mouth. *J. Financ.* 75, 1965–2020.
- Da, Z., Huang, X., Jin, L.J., 2021. Extrapolative beliefs in the cross-section: what can we learn from the crowds? *J. Financ. Econ.* 140, 175–196.
- Daniel, K., Hirshleifer, D., Subrahmanyam, A., 1998. Investor psychology and security market under- and overreactions. *J. Financ.* 53, 1839–1885.
- Daniel, K., Hirshleifer, D., Subrahmanyam, A., 2001. Overconfidence, arbitrage, and equilibrium asset pricing. *J. Financ.* 56, 921–965.
- DellaVigna, S., Linos, E., 2020. RCTs to scale: comprehensive evidence from two nudge units. Unpublished working paper. UC Berkeley.
- Deng, Y., Liu, X., Wei, S.J., 2018. One fundamental and two taxes: when does a Tobin tax reduce financial price volatility? *J. Financ. Econ.* 130, 663–692.
- Dorn, D., Huberman, G., 2005. Talk and action: what individual investors say and what they do. *Rev. Financ.* 9, 437–481.
- Dorn, D., Sengmueller, P., 2009. Trading as entertainment? *Manag. Sci.* 55, 591–603.
- Epper, T., Fehr, E., Fehr-Duda, H., Kreiner, C.T., Lassen, D.D., Leth-Petersen, S., Rasmussen, G.N., 2020. Time discounting and wealth inequality. *Am. Econ. Rev.* 110, 1177–1205.
- Eyster, E., Rabin, M., Vayanos, D., 2019. Financial markets where traders neglect the informational content of prices. *J. Financ.* 74, 371–399.
- Fama, E.F., 1998. Market efficiency, long-term returns, and behavioral finance. *J. Financ. Econ.* 49, 283–306.
- Friedman, M., 1953. The methodology of positive economics. *Essays Posit. Econ.* 3, 145–178.
- Frydman, C., Rangel, A., 2014. Debiasing the disposition effect by reducing the saliency of information about a stock's purchase price. *J. Econ. Behav. Organ.* 107, 541–552.
- Frydman, C., Wang, B., 2020. The impact of salience on investor behavior: evidence from a natural experiment. *J. Financ.* 75, 229–276.
- Frydman, C., Barberis, N., Camerer, C., Bossaerts, P., Rangel, A., 2014. Using neural data to test a theory of investor behavior: an application to realization utility. *J. Financ.* 69, 907–946.
- Gao, X., Lin, T.C., 2015. Do individual investors treat trading as a fun and exciting gambling activity? evidence from repeated natural experiments. *Rev. Financ. Stud.* 28, 2128–2166.
- Gervais, S., Odean, T., 2001. Learning to be overconfident. *Rev. Financ. Stud.* 14, 1–27.
- Giglio, S., Maggiori, M., Stroebe, J., Utkus, S., 2021b. The joint dynamics of investor beliefs and trading during the COVID-19 crash. *Proc. Natl. Acad. Sci.* 118 (4).
- Giglio, S., Maggiori, M., Stroebe, J., Utkus, S., 2021a. Five facts about beliefs and portfolios. *Am. Econ. Rev.* 111, 1481–1522.
- Glaser, M., Weber, M., 2007. Overconfidence and trading volume. *Geneva Risk Insur. Rev.* 32, 1–36.
- Graham, J.R., Harvey, C.R., Huang, H., 2009. Investor competence, trading frequency, and home bias. *Manag. Sci.* 55, 1094–1106.
- Gramp, W.D., 1948. Adam Smith and the economic man. *J. Political Econ.* 56 (4), 315–336.
- Greenwood, R., Shleifer, A., 2014. Expectations of returns and expected returns. *Rev. Financ. Stud.* 27, 714–746.
- Grinblatt, M., Han, B., 2005. Prospect theory, mental accounting, and momentum. *J. Financ. Econ.* 78, 311–339.
- Grinblatt, M., Keloharju, M., 2001. How distance, language, and culture influence stockholdings and trades. *J. Financ.* 56, 1053–1073.
- Grinblatt, M., Keloharju, M., 2009. Sensation seeking, overconfidence, and trading activity. *J. Financ.* 64 (2), 549–578.
- Grinblatt, M., Keloharju, M., Linnainmaa, J., 2011. IQ and stock market participation. *J. Financ.* 66, 2121–2164.
- Grossman, S.J., Stiglitz, J.E., 1980. On the impossibility of informationally efficient markets. *Am. Econ. Rev.* 70, 393–408.
- Han, B., Hirshleifer, D., Walden, J., 2020. Social transmission bias and investor behavior. *J. Financ. Quant. Anal.* forthcoming.
- Hirshleifer, D., 2015. Behavioral finance. *Annu. Rev. Financ. Econ.* 7, 133–159.
- Hong, H., Stein, J.C., 1999. A unified theory of underreaction, momentum trading, and overreaction in asset markets. *J. Financ.* 54, 2143–2184.
- Hong, H., Kubik, J.D., Stein, J.C., 2004. Social interaction and stock-market participation. *J. Financ.* 59, 137–163.
- Hong, H., Scheinkman, J., Xiong, W., 2008. Advisors and asset prices: a model of the origins of bubbles. *J. Financ. Econ.* 89, 268–287.
- Ingersoll, J.E., Jin, L.J., 2013. Realization utility with reference-dependent preferences. *Rev. Financ. Stud.* 26, 723–767.
- Jin, L.J., Sui, P., 2021. Asset pricing with return extrapolation. Unpublished working paper. California Institute of Technology.
- Kahneman, D., Frederick, S., 2002. Representativeness revisited: attribute substitution in intuitive judgment. *Heuristics Biases Psychol. Intuitive Judgm.* 49, 81.
- Kelly, M., O Grada, C., 2000. Market contagion: evidence from the panics of 1854 and 1857. *Am. Econ. Rev.* 90, 1110–1124.
- Koijen, R., Van Nieuwerburgh, S., Vestman, R., 2014. Judging the quality of survey data by comparison with “truth” as measured by administrative records evidence from Sweden. In: Carroll, C.D., Crossley, T.F., Sabelhaus, J. (Eds.), *Improving the Measurement of Consumer Expenditures*. University of Chicago Press, pp. 308–346.
- Kreiner, C.T., Lassen, D.D., Leth-Petersen, S., 2015. Measuring the accuracy of survey responses using administrative register data: evidence from Denmark. In: Carroll, C.D., Crossley, T.F., Sabelhaus, J. (Eds.), *Improving the Measurement of Consumer Expenditures*. University of Chicago Press, pp. 289–307.
- Kumar, A., 2009. Who gambles in the stock market? *J. Financ.* 64, 1889–1933.
- Kyle, A.S., Wang, F.A., 1997. Speculation duopoly with agreement to disagree: can overconfidence survive the market test? *J. Financ.* 52, 2073–2090.
- Kyle, A.S., 1985. Continuous auctions and insider trading. *Econometrica* 53, 1315–1335.
- Liao, J., Peng, C., Zhu, N., 2021. Extrapolative bubbles and trading volume. *Rev. Financ. Stud.* forthcoming.
- Lusardi, A., Mitchell, O.S., 2007. Baby boomer retirement security: the roles of planning, financial literacy, and housing wealth. *J. Monet. Econ.* 54, 205–224.
- Lusardi, A., Mitchell, O.S., 2011. Financial literacy around the world: an overview. *J. Pension Econ. Financ.* 10, 497–508.
- Lusardi, A., Mitchell, O.S., 2014. The economic importance of financial literacy: theory and evidence. *J. Econ. Lit.* 52, 5–44.
- Milgrom, P., Stokey, N., 1982. Information, trade and common knowledge. *J. Econ. Theory* 26, 17–27.
- Odean, T., 1998. Volume, volatility, price, and profit when all traders are above average. *J. Financ.* 53, 1887–1934.
- Odean, T., 1999. Do investors trade too much? *Am. Econ. Rev.* 89, 1279–1298.

- Pool, V.K., Stoffman, N., Yonker, S.E., 2015. The people in your neighborhood: social interactions and mutual fund portfolios. *J. Financ.* 70, 2679–2732.
- Scheinkman, J.A., Xiong, W., 2003. Overconfidence and speculative bubbles. *J. Political Econ.* 111, 1183–1219.
- Seru, A., Shumway, T., Stoffman, N., 2010. Learning by trading. *Rev. Financ. Stud.* 23, 705–739.
- Shefrin, H., Statman, M., 1985. The disposition to sell winners too early and ride losers too long: theory and evidence. *J. Financ.* 40, 777–790.
- Shefrin, H., Statman, M., 2000. Behavioral portfolio theory. *J. Financ. Quant. Anal.* 35, 127–151.
- Shiller, R.J., 1984. Stock prices and social dynamics. *Brook. Pap. Econ. Act.* 2, 457–498.
- Shiller, R.J., 1989. Comovements in stock prices and comovements in dividends. *J. Financ.* 44, 719–729.
- Song, Z., Xiong, W., 2018. Risks in china's financial system. *Annu. Rev. Financ. Econ.* 10, 261–286.
- Taubinsky, D., Rees-Jones, A., 2018. Attention variation and welfare: theory and evidence from a tax salience experiment. *Rev. Econ. Stud.* 85, 2462–2496.
- Tversky, A., Kahneman, D., 1992. Advances in prospect theory: cumulative representation of uncertainty. *J. Risk Uncertain.* 5, 297–323.
- Van Rooij, M., Lusardi, A., Alessie, R., 2011. Financial literacy and stock market participation. *J. Financ. Econ.* 101, 449–472.