

# Review Chap 2: Loss Distributions

Fall 2023 Actuarial Modelling, Nankai University

**TA: Yuan Zhuang** (Instructor: Prof. Lianzeng Zhang)

November 5, 2023

# 目录

- ▶ Structure of the Chapter
- ▶ Classical Loss Distributions
- ▶ Fitting loss distributions
- ▶ Mixture Distributions
- ▶ Loss Distributions and Reinsurance

# Distribution in Actuarial Modelling

- Fit a probability distribution with reasonably tractable mathematical properties
- Concern about the chances and sizes of large claims

## What's each section about?

- **Section 2.2:** What are the basic properties of some commonly used and classic loss distributions? How to fit data to a loss distribution? (**Classical loss distributions**)
- **Section 2.3:** How well does our selected distribution fit the data? (**Analyzing fit**)
- **Section 2.4:** Well, what's going on if the parameter is not a constant, but follows a certain distribution? (**Mixture distribution**)
- **Section 2.5:** What will change under reinsurance contract? (**Loss distributions and reinsurance**)

# 目录

- ▶ Structure of the Chapter
- ▶ **Classical Loss Distributions**
- ▶ Fitting loss distributions
- ▶ Mixture Distributions
- ▶ Loss Distributions and Reinsurance

## Exponential Distribution (Page 36 - 39)

- PDF, CDF, MGF, Mean, Variance (Have a look at the form of CDF!)
- Memoryless property
- Thinner tail than the dataset Theft
- Failure (or hazard) rate function

$$r_X(x) = \lim_{h \rightarrow 0} \frac{F_X(x+h) - F_X(x)}{h} \frac{1}{\bar{F}_X(x)} = \frac{f_X(x)}{\bar{F}_X(x)} = \frac{\lambda e^{-\lambda x}}{e^{-\lambda x}} = \lambda.$$

- Method of moments and maximum likelihood
- $\bar{F}_X(x) = e^{-\lambda x}$  converges to 0 really fast

## Pareto Distribution (Page 39 - 42)

- PDF, CDF, Mean, Variance

$$\bar{F}_X(x) = \left( \frac{\lambda}{\lambda + x} \right)^\alpha$$

- Dealing with inflation in claims and deductibles
- Method of moments and maximum likelihood (numerical methods needed)
- Much closer to observed relative frequencies of these events for the Theft claim data than the ML fitted exponential distribution

## Gamma Distribution (Page 43 - 44)

- Shape and scale parameter
- Method of moments and maximum likelihood (Reparametrization needed)



## Weibull Distribution (Page 45 - 46)

- Shape and scale parameter  $c$  and  $\gamma$ 
  - $\gamma = 1$ , exponential
  - $\gamma < 1$ , tail is fatter (heavier) than that of any exponential distribution, but not as heavy as that of a Pareto

- 

$$r_X(x) = \frac{f_X(x)}{\bar{F}_X(x)} = \frac{c\gamma x^{\gamma-1} e^{-cx^\gamma}}{e^{-cx^\gamma}} = c\gamma x^{\gamma-1}$$

- Moments (Gamma function needed)
- Maximum Likelihood and method of percentiles

## Lognormal Distribution (Page 47 - 51)

- Normal after Logged
- PDF, Mean, Variance
- Method of moments or Maximum likelihood

# 目录

- ▶ Structure of the Chapter
- ▶ Classical Loss Distributions
- ▶ **Fitting loss distributions**
- ▶ Mixture Distributions
- ▶ Loss Distributions and Reinsurance

# Tests on Goodness of Fit

- Kolmogorov–Smirnov test
- Anderson–Darling test
- Chi-square goodness-of-fit tests (\*Important)

## Chi-square goodness-of-fit tests (Page 54 - 56)

- Maximum likelihood on the grouped data

- 

$$\chi_{GF}^2 = \sum_1^k (O_i - E_i)^2 / E_i$$

- Degrees of Freedom:  $d = k - 1 - r$ 
  - $k$ : Number of groups
  - $r$ : Number of parameters to be estimated

# 目录

- ▶ Structure of the Chapter
- ▶ Classical Loss Distributions
- ▶ Fitting loss distributions
- ▶ **Mixture Distributions**
- ▶ Loss Distributions and Reinsurance

## Mixture Distributions (Page 58 - 60)

### Simple Case

$$F(x) = pF_1(x) + qF_2(x)$$

### General Case

$$F(x) = \int_{\Theta} F_{\theta}(x) dG(\theta) = \int_{\Theta} F_{\theta}(x) g(\theta) d\theta$$

Focus on Example 2.5

# 目录

- ▶ Structure of the Chapter
- ▶ Classical Loss Distributions
- ▶ Fitting loss distributions
- ▶ Mixture Distributions
- ▶ Loss Distributions and Reinsurance



## Assumptions (Page 61)

- Claim-by-claim Based

- 

$$X = Y + Z = h_I(X) + h_R(X)$$

## Proportional reinsurance (Page 62)

- $$h_I(X) = Y = \alpha X, \quad h_R(X) = Z = (1 - \alpha)X$$
- Closed under multiplication by a positive scalar (exponential, Pareto, gamma, Weibull and lognormal)

## Excess of Loss Reinsurance (Page 62 - 63)

- 

$$Y = \begin{cases} X & \text{if } X \leq M \\ M & \text{if } X > M \end{cases}$$

$$E(Y) = \int_0^M x f_X(x) dx + M \bar{F}_X(M) = E(X) - \int_0^\infty y f_X(y + M) dy$$

- Maximum Likelihood:

$$\mathbf{x} = x_1, x_2, M, x_4, M, x_6, x_7, M, \dots$$

$$L(\theta) = \prod_1^n f_X(x_i, \theta) \prod_1^m \bar{F}_X(M, \theta)$$

## The Reinsurer's View of Excess of Loss Reinsurance (Page 63 - 64)

- 

$$E(Z) = 0 \cdot F_X(M) + E(Z_R) \bar{F}_X(M) = E(Z_R) P(X > M)$$

- 

$$\bar{F}_{Z_R}(z) = P(X > M + z \mid X > M) = \frac{\bar{F}_X(M + z)}{\bar{F}_X(M)}$$

## Dealing with Claims Inflation (Page 64 - 65)

- 

$$E(Y^*) = \int_0^{M/k} kx f_X(x) dx + \int_{M/k}^{\infty} M f_X(x) dx$$

- 

$$E(Y^*) = k \left[ E(X) - \int_0^{\infty} y f_X(y + M/k) dy \right]$$

- The increase overall in payment by the (ceding) insurer is less than  $k$

## Policy Excess and Deductibles (Page 66 - 68)

- 

$$Y = \begin{cases} 0 & \text{if } X \leq D \\ X - D & \text{if } X > D \end{cases}$$

- 

$$E(Y) = \int_D^{\infty} (x - D) f_X(x) dx = \int_0^{\infty} y f_X(y + D) dy$$

Focus on Example 2.7