Bjørnar’s assignment in Gretl

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Table of contents

[Bjørnar’s assignment](https://hiof.instructure.com/courses/9596/files/folder/Assignment?preview=1695176)

## Exercise 1

﻿ Mean Median S.D. Min Max  
price 1406 1080 1027 495.2 4884  
WinterRain 608.4 600.0 129.0 376.0 830.0  
temp 16.48 16.42 0.6592 14.98 17.65  
HarvestRain 144.8 123.0 73.07 38.00 292.0  
Age 16.19 16.00 8.246 3.000 31.00

A case of wine has an average market **price** of 1,406, a median of 1,080, and a standard deviation of 1,027. 495.2 is the lowest price that has been determined, while 4884 is the highest. This comparatively high standard deviation indicates significant price volatility for wines, which is probably caused by a mix of aging, climate, and vintage quality.

Before the growing season, the average amount of **WinterRain** (October–March) is 608.4 mm, with a median of 600 mm and a standard deviation of 129 mm. It varies slightly between vintages, with values ranging from 376 mm to 830 mm.

With a mean **temperature** of 16.48°C and a standard deviation of 0.66°C, the growth season temperature (April–September) is mainly constant throughout the sample. The estimated temperatures range from a minimum of 14.98°C to a maximum of 17.65°C.

There is a higher volatility in **HarvestRain** (August–September), with a mean of 144.8 mm, a standard deviation of 73.07 mm, and a range of 38 mm to 292 mm. This variation raises the possibility that some vintages had unusually wet harvest seasons, which could affect the quality of the grapes.

With a mean of 16.19 years and a standard deviation of 8.25 years, the wine’s **age** in 1983 varied from 3 to 31 years. With a median age of 16, the distribution appears to be pretty symmetrical.

In summary, the data show that wine pricing and environment variables fluctuate significantly, while age and temperature are more evenly balanced.

## Exercise 2

*Model 1*

**a.**

﻿Model 1: OLS, using observations 1952-1980 (T = 27)  
Missing or incomplete observations dropped: 2  
Dependent variable: price  
  
 coefficient std. error t-ratio p-value  
 -------------------------------------------------------  
 const 494.278 402.112 1.229 0.2304   
 Age 56.3183 22.2218 2.534 0.0179 \*\*  
  
Mean dependent var 1405.800 S.D. dependent var 1027.226  
Sum squared resid 21827151 S.E. of regression 934.3907  
R-squared 0.204406 Adjusted R-squared 0.172582  
F(1, 25) 6.423065 P-value(F) 0.017902  
Log-likelihood −221.9495 Akaike criterion 447.8990  
Schwarz criterion 450.4907 Hannan-Quinn 448.6697  
  
Warning: The version of this datafile (1.5) is higher than that fully  
supported by this build of gretl (1.4). Some features may not be  
correctly recognized.

**b.**

If the age increases by one year, we expect the price of the wine which is 494.278 to increase by 56.3183 per year.

**c.**

Test wheather Age has a significant effect on Price by looking at:

* Null hypothesis (): (no effect)
* Alternative hypothesis (): (significant effect)

At the level, we reject if:

Since the p-value is 0.0179 > 0.01, we fail to reject the null hypothesis at the 1% level.

The Age coefficient does not achieve significance at the more restrictive 1% level, despite being significant at the 5% level. As a result, we make the conclusion that there is not enough data to support that wine age significantly affects pricing at the 1% level.

**d.**

We use :

We calculate predicted wine prices for three values of the variable Age: the minimum of 3, the mean of 16.19, and the maximum of 31, based on the output from gretl in exercise 1.

When Age = 3, the predicted wine price is approximately 663.24.

When Age = 16.19, the predicted price is approximately 1406.1, which is very close to the mean price in the output which is 1405.8.

When Age = 31, the predicted price increases to approximately 2240.2.

These results demonstrate the positive correlation between price and age that the regression model suggests, that the wine’s expected market value increases linearly with age.

**e.**

Since the prices were measured in 1983, the wine from 1961 is:

Then we will calculate the predicted price for Age = 22:

Then we look at the degrees of freedom:

From a standard t-table, with 25 degrees of freedom, the critical t-value at 95% confidence is 2.060.

The margin of error:

Interval:

The 95% prediction interval is [−191.53, 3658.17]

## Exercise 3

*Model 2*

**a.**

﻿Model 2: OLS, using observations 1952-1980 (T = 27)  
Missing or incomplete observations dropped: 2  
Dependent variable: price  
  
 coefficient std. error t-ratio p-value   
 -------------------------------------------------------------  
 const −15509.0 3379.87 −4.589 0.0001 \*\*\*  
 Age 39.2126 14.3490 2.733 0.0121 \*\*  
 WinterRain 2.75098 0.965119 2.850 0.0093 \*\*\*  
 temp 930.787 190.557 4.885 6.97e-05 \*\*\*  
 HarvestRain −5.04694 1.61682 −3.122 0.0050 \*\*\*  
  
Mean dependent var 1405.800 S.D. dependent var 1027.226  
Sum squared resid 7238994 S.E. of regression 573.6246  
R-squared 0.736141 Adjusted R-squared 0.688166  
F(4, 22) 15.34443 P-value(F) 3.93e-06  
Log-likelihood −207.0499 Akaike criterion 424.0999  
Schwarz criterion 430.5791 Hannan-Quinn 426.0265

**b.**

- If the **Age** of the wine increases by one year, we expect the price of the wine, which is initially predicted to be −15,509.0, to increase by **39.21 per year**.

* If **WinterRain** increases by 1, we expect the price of the wine to increase by **2.75**.
* If the **temperature** during the growing season increases by 1°C, we expect the price of the wine to increase by **930.79**.
* If **HarvestRain** increases by 1, we expect the price of the wine to **decrease by 5.05**.

**c.**

Exercise 3 multiple regression model provides a much better fit than Exercise 2 basic regression model. The expanded model explains a greater amount of the variation in wine prices, as shown by the R-squared rising from 0.204 to 0.736. The regression’s standard error drops from 934.39 to 573.62, indicating that predictions are now more accurate.

When the weather variables are included, the coefficient for age drops from 56.32 to 39.21, but it is still statistically significant. This decrease raises the possibility that the basic model overemphasized the impact of age since some variables were left out. The multiple regression model offers a more precise and trustworthy examination of the elements influencing wine pricing by incorporating the other variables.

## Exercise 4

*Model 3*

**a.**

﻿Model 3: OLS, using observations 1952-1980 (T = 27)  
Missing or incomplete observations dropped: 2  
Dependent variable: price  
  
 coefficient std. error t-ratio p-value  
 ----------------------------------------------------------  
 const −16289.5 4395.18 −3.706 0.0012 \*\*\*  
 Dheavyrain 11634.1 8990.38 1.294 0.2085   
 temp 1082.95 266.391 4.065 0.0005 \*\*\*  
 temp\_Dheavy −756.902 546.064 −1.386 0.1790   
  
Mean dependent var 1405.800 S.D. dependent var 1027.226  
Sum squared resid 14038659 S.E. of regression 781.2660  
R-squared 0.488295 Adjusted R-squared 0.421550  
F(3, 23) 7.315911 P-value(F) 0.001294  
Log-likelihood −215.9914 Akaike criterion 439.9829  
Schwarz criterion 445.1662 Hannan-Quinn 441.5242  
  
Excluding the constant, p-value was highest for variable 6 (Dheavyrain)

**b.**

The consistent effect of heavy HarvestRain is captured by the variable Dheavyrain. Although its coefficient 11634.1 indicates that the expected base price is greater in years with extremely high HarvestRain, this effect is not statistically significant (p = 0.2085), therefore we are unable to draw firm conclusions about the impact of rain alone.

With a coefficient of -756.90, the interaction term temp\_Dheavyrain shows that the price-boosting impact of temperature is less pronounced in years with high rain. This interaction implies that heavy rain lessens the value of warm growing seasons, most likely because of lower grape quality at harvest, even if it is also not statistically significant (p = 0.1790).

## Exercise 5

* Next we determine each model’s expected price and residual.

*Model 1* – Simple regression

*Model 2* – Multiple regression

*Model 3* -Dummy regression

**b.**

With a residual of 1306.023, Model 2 (the multiple regression incorporating age and weather factors) appears to offer the best accurate prediction for the 1961 vintage, as opposed to 2402.303 for Model 3 and 3150.58 for Model 1. This suggests that Model 2 is the one that most closely resembles the 1961 vintage’s actual market price.

With an R-squared of 0.736, Model 2 has the strongest explanatory power outside of residual analysis, accounting for roughly 73.6% of the variation in wine prices. Compared to Model 1’s R-squared of 0.204 and Model 3’s R-squared of 0.488, this is much higher.

The Model 2’s standard error was the lowest of the three explanatory factors, and all four of them—Age, HarvestRain, WinterRain, and temperature—were statistically significant.

In conclusion, Model 2 is the best model for forecasting Bordeaux wine prices when taking into account both the predicted accuracy for the 1961 vintage and the overall model performance.