

L11 Q.P. 1

6. PSD का नियम 2 का प्रमाणण PSD का $A \geq 0$ है ④

$M_{1,2,3,4} \geq 0$ का नियम 2 का प्रमाणण PSD का $A \geq 0$ है ④

$b_{ij}, s_{ij} \geq 0$ का नियम 2 का प्रमाणण PSD का $A \geq 0$ है ④

$M_{1,2,3,4} \geq 0$ का नियम 2 का प्रमाणण PSD का $A \geq 0$ है ④

$$A = Q D Q^T = Q \cdot \begin{pmatrix} M & 0 \\ 0 & \lambda_2 \end{pmatrix} \cdot Q^T = Q \cdot \boxed{\begin{pmatrix} \sqrt{M} & 0 \\ 0 & \sqrt{\lambda_2} \end{pmatrix}} \cdot \begin{pmatrix} \sqrt{M} & 0 \\ 0 & \sqrt{\lambda_2} \end{pmatrix} \cdot Q^T = X \cdot X^T$$

$$X^T = Q \cdot \begin{pmatrix} \sqrt{M} & 0 \\ 0 & \sqrt{\lambda_2} \end{pmatrix} = \begin{pmatrix} \sqrt{M} & 0 \\ 0 & \sqrt{\lambda_2} \end{pmatrix} \cdot Q^T = \boxed{\begin{pmatrix} \sqrt{M} & 0 \\ 0 & \sqrt{\lambda_2} \end{pmatrix} \cdot Q^T}$$

~~प्रमाणण करने की कोशिश~~

13) $A = X X^T$ का प्रमाणण करने की कोशिश

प्रमाणण $V^T A V = V^T X X^T V = (X^T V)^T \cdot (X^T V) = \|X^T V\|^2 \geq 0$ ④

$$V^T A V = V^T X X^T V = (X^T V)^T \cdot (X^T V) = \|X^T V\|^2 \geq 0$$

PSD का नियम 2 का प्रमाणण $V^T A V \geq 0$ ④

□

• Für $V \in \mathbb{R}^n$ gilt $\lambda A + \beta B \geq 0 \Leftrightarrow \lambda \geq 0 \wedge \text{PSD } A, \text{ PSD } B$ (b)

$$V^\top (\lambda A + \beta B) V = V^\top \lambda A V + V^\top \beta B V = \sum_{i=1}^n \lambda \cdot V_i^\top A V_i + \beta \cdot \sum_{i=1}^n V_i^\top B V_i \geq 0$$

PSD $\Leftrightarrow \lambda A + \beta B \geq 0 \Leftrightarrow \lambda \geq 0, \text{ PSD } B$ psl

$\forall \lambda \geq 0 \exists \beta \geq 0 \ L \lambda A + \beta B \geq 0 \Leftrightarrow \text{PSD } A, \text{ PSD } B$ psl

$\Rightarrow \text{PSD } A \Leftrightarrow \exists \lambda \geq 0 \ L \lambda A \geq 0 \Leftrightarrow \text{PSD } A$ psl

(max von X_1, \dots, X_n)

$y = \max(X_1, \dots, X_n) \quad x_i \in \mathbb{R}, \forall i \in \{1, \dots, n\} \Rightarrow \text{i.i.d. r.v. } X_1, \dots, X_n \text{ psl}$ (9)

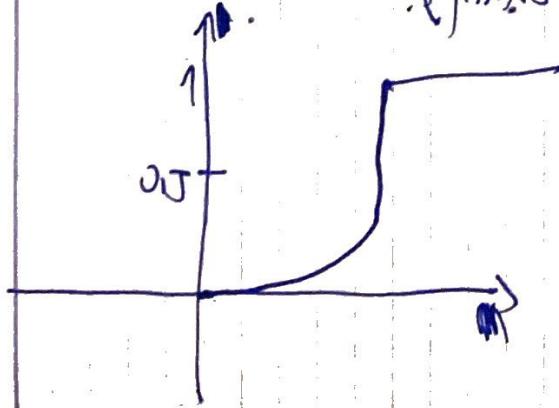
W.S.U. $\Rightarrow y$ def. W.S.U. \Rightarrow max ist s.o. s.o. s.o.

$$F_Y(y) = P(Y \leq y) = P(\max(X_1, \dots, X_n) \leq y) = P(X_1 \leq y, \dots, X_n \leq y) = (P(X_1 \leq y))^n = (F_{X_1}(y))^n$$

$$F_{X_1}(x) = \begin{cases} 0 & x < x_1 \\ \frac{x - x_1}{x_2 - x_1} & x \in [x_1, x_2] \\ 1 & x > x_2 \end{cases} \Rightarrow f_{X_1}(x) = \begin{cases} 0 & x \notin [x_1, x_2] \\ \frac{1}{x_2 - x_1} & x \in [x_1, x_2] \end{cases}$$

$$f_Y(y) = (F_Y(y))' = \frac{d}{dy} (F_{X_1}(y))^n = n \cdot (F_{X_1}(y))^{n-1} \cdot f_{X_1}(y) = n y^{n-1} \cdot I_{[x_1, x_2]}$$

: l. r. s. d. y. (c. w. g.)



$$E[y] = \int_0^1 x \cdot f_{X,Y}(x,y) dx = \int_0^1 x \cdot x^k y^k dy = k \int_0^1 x^k y^k dy = \frac{x^{k+1}}{k+1}$$

$y \neq 0$ \Rightarrow $f_{X,Y}(x,y) = 0$

$$\lim_{n \rightarrow \infty} \frac{1}{n} = 1 \Rightarrow$$

$$E[y] \xrightarrow{n \rightarrow \infty} 1$$

\therefore $N \rightarrow \infty$ \Rightarrow $E[y] \approx 1$

$E[x^2] = \int_0^1 x^2 f_X(x) dx = \int_0^1 x^2 x^k dx = \frac{x^{k+3}}{k+3}$

$$E[y^2] = \int_0^1 x^2 f_Y(y) dy = k \int_0^1 x^2 y^k dy = \frac{1}{k+2}$$

$\therefore E[y^2] \approx 1$

$$Vy = E[y^2] - (E[y])^2 = \frac{1}{k+2} - \frac{1}{(k+1)^2}$$

$\therefore V[y]$

$$\lim_{n \rightarrow \infty} Vy = \frac{1}{k+2} - \frac{1}{k^2} = 1 - 1 = 0$$

\therefore $V[y] \approx 0$

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$$h(u) = E[X-u] = \int_{-\infty}^{\infty} (x-u) f_X(x) dx = \int_{-\infty}^{\infty} (x-u) f_X(x) dx + \int_u^{\infty} (x-u) f_X(x) dx$$

\therefore $E[X-u] = \int_u^{\infty} (x-u) f_X(x) dx$

$$\frac{d}{du} (h(u)) = (u-k) \cdot \frac{d}{du} \int_{-\infty}^u f_X(x) dx + \int_u^{\infty} f_X(x) dx + (k-u) \cdot \frac{d}{du} \int_u^{\infty} f_X(x) dx =$$

\therefore $0 = f_X(u)$

$$= (u-k+k-u) \cdot f_X(u) + \int_u^{\infty} f_X(x) dx - \int_u^{\infty} f_X(x) dx = 0 \Rightarrow \int_u^{\infty} f_X(x) dx = \int_u^{\infty} f_X(x) dx$$

$$P(X \leq u) = P(X < u) = \int_{-\infty}^u f_X(x) dx$$

\therefore $P(X \leq u) = P(X < u) = \int_{-\infty}^u f_X(x) dx$

(JULY 2012)

2.11 BC

Given: $f(x) = \begin{cases} 1 & 0 < x < 4 \\ 0 & \text{otherwise} \end{cases}$

$$h(u) = E[|x-u|] = \int_{-\infty}^{\infty} |x-u| f_x(x) dx = \int_{-\infty}^{\infty} (x-u) f_x(x) dx + \int_u^{\infty} (u-x) f_x(x) dx =$$
$$= u \int_{-\infty}^u f_x(x) dx - \int_{-\infty}^u x f_x(x) dx + \int_u^{\infty} x f_x(x) dx - u \int_u^{\infty} f_x(x) dx =$$
~~$$= u F_x(u) - u \int_{-\infty}^u x f_x(x) dx - \int_u^{\infty} x f_x(x) dx + u \int_u^{\infty} F_x(x) dx$$~~
$$= u F_x(u) - u(1 - F_x(u)) + \int_u^{\infty} x f_x(x) dx - \int_u^{\infty} x f_x(x) dx =$$
$$= 2u F_x(u) - u + \int_u^{\infty} x f_x(x) dx - \int_{\infty}^u x f_x(x) dx$$
$$h(u) = (2u F_x(u))' - u f_x(u) - u f_x(u) - 1 = 2u f_x(u) - 2u f_x(u) - 1 + 2f_x(u) =$$
$$= 2f_x(u) - 1$$

Final answer: $F_x(u) = \frac{u}{4}$ for $0 < u < 4$

Final answer: $h(u) = 2u - 1$

□

$(\lambda(G)) \cdot f(\lambda) = \lambda^2 S_n(\lambda) W(\lambda) + \lambda f(\lambda)$

$$L(h) = E[\text{log}(Y_i | h(x))] = \sum_{i=1}^L b_{i-1}(i, h(x)) \cdot P(Y_i = i | h(x))$$

$$P(X=x, Y=y) = P(X=x) \cdot P(Y=y|X=x)$$

$\int_{-1}^1 f(x) dx = \int_{-1}^1 g(x) dx$

$$L(h) = p(k=x) \cdot \sum_{i=1}^k b_{i-1}(i|h(x)) \cdot p(y=i|k=x)$$

$$(o_{\alpha} \circ h)(x) = \begin{cases} 0 & \text{if } \alpha = i \\ 1 & \text{else.} \end{cases} : h(x) \in \{1, -1\} = C \quad \text{proof}$$

$$P(X=x) = \sum_{y=1}^L P(X=x, Y=y) = \sum_{y=1}^L P(Y=y|X=x) \cdot P(X=x)$$

∴ $\lim_{x \rightarrow 0} f(x) = 1$ (as $f(0) = 1$)

$$h(x) = \sum_{i=1}^k P(Y=i | X=x)$$

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$$\begin{aligned} \text{LHS} &= \mathbb{E}[h(Y, f(x))] = P(X=x, Y=j) \cdot h(j, h(x)) + P(X=x, Y=i) \cdot h(i, h(x)) \\ &= P(X=x) \cdot P(Y=j|X=x) \cdot h(j, h(x)) + P(X=x) \cdot P(Y=i|X=x) \cdot h(i, h(x)) = \\ &\stackrel{(a)}{=} P(X=x) \cdot (P(Y=j|X=x) h(j, h(x)) + P(Y=i|X=x) h(i, h(x))) \end{aligned}$$

$P(X=x) = P(\text{number of } 1's = x) = \binom{n}{x} p^x (1-p)^{n-x}$

$$q \cdot P(Y=0|X=x) : \text{fN} \text{ SK } \text{bmm} \rightarrow \text{N}$$

$$b \cdot p(Y=1 | X=x) = f_2(x) \cdot \Pr_{\substack{h(X)=1 \\ h(X) \geq 0}}[Y=1 | X=x]$$

$$b \cdot P(Y=1|k=x) \geq 4 \cdot P(Y=0|k=x) \quad ; h(x) = \frac{1}{4} \quad \text{if } P(Y=1|k=x) > \frac{1}{4}$$

$$h(x) = 0 \quad \text{if } P(Y=1|k=x) \leq \frac{1}{4} \quad \text{else} \quad \text{and} \quad b = \frac{P(Y=1|k=x)}{P(Y=0|k=x)}$$

Now we can write for given x and y $b = P(Y=1|k=x)$ $\text{and} \quad h(x) = \underline{b}$

$$h(x) = \begin{cases} 0 & \text{if } P(Y=1|k=x) \leq \frac{1}{4} \\ 1 & \text{else} \end{cases} \quad ; \quad \text{if } P(Y=1|k=x) < \frac{P(Y=0|k=x)}{b}$$

or if $x \in \{1, 0\}$ $P(Y=1|k=x)$ \rightarrow $h(x) = \underline{b}$ $\text{and} \quad b = P(Y=1|k=x)$.

□

$$(N \cdot b) \cdot (1 - b) + (1 - N) \cdot b^2 \leq b^2$$

~~$$\frac{1}{N} f(x_i, \mu_i, \Sigma_{i,i}) = \frac{1}{N} \frac{1}{\sqrt{2\pi \sum_{i,i}}} e^{-\frac{(x_i - \mu_i)^2}{2\Sigma_{i,i}}} = \frac{1}{(2\pi)^{\frac{N}{2}} \prod_{i=1}^N \Sigma_{i,i}^{\frac{1}{2}}} e^{-\frac{\sum_{i=1}^N (x_i - \mu_i)^2}{2\Sigma_{i,i}}}$$~~

$$\prod_{i=1}^N f(x_i, \mu_i, \Sigma_{i,i}) = \prod_{i=1}^N \frac{1}{(2\pi)^{\frac{N}{2}} \prod_{i=1}^N \Sigma_{i,i}^{\frac{1}{2}}} e^{-\frac{\sum_{i=1}^N (x_i - \mu_i)^2}{2\Sigma_{i,i}}}$$

$$\left(\prod_{i=1}^N f(x_i, \mu_i, \Sigma_{i,i}) \right)^{\frac{1}{N}} = \left(\frac{1}{(2\pi)^{\frac{N}{2}} \prod_{i=1}^N \Sigma_{i,i}^{\frac{1}{2}}} e^{-\frac{\sum_{i=1}^N (x_i - \mu_i)^2}{2\Sigma_{i,i}}} \right)^{\frac{1}{N}} = \frac{1}{(2\pi)^{\frac{N}{2}} \prod_{i=1}^N \Sigma_{i,i}^{\frac{1}{2}}} e^{-\frac{\sum_{i=1}^N (x_i - \mu_i)^2}{2N}}$$

$$P(Y=1|X=x) > P(Y=0|X=x) \quad \text{if } \lambda_1 > \lambda_0 \quad \text{and} \quad \lambda_1 > \lambda_0$$

$$\frac{f_X(Y=1|x)p(Y=1)}{f_X(x)} > \frac{f_X(Y=0|x)p(Y=0)}{f_X(x)} \quad \text{if } p(Y=1) = p \quad \text{and} \quad \sum_{i=1}^d x_i > \sum_{i=1}^d \lambda_i$$

$$P \cdot f_1(x, \mu_1, \Sigma) > (1-p) f_0(x, \mu_0, \Sigma)$$

$$\frac{1-p}{p} < \frac{\frac{1}{2} \cdot (x-\mu_0)^T \Sigma^{-1} (x-\mu_0)}{\frac{1}{2} \cdot (x-\mu_1)^T \Sigma^{-1} (x-\mu_1)}$$

$$\frac{1-p}{p} < e^{-\frac{1}{2} \cdot (x-\mu_0)^T \Sigma^{-1} (x-\mu_0) + \frac{1}{2} \cdot (x-\mu_1)^T \Sigma^{-1} (x-\mu_1)}$$

$$2 \ln \left(\frac{1-p}{p} \right) < (x - \mu_0)^T \Sigma^{-1} (x - \mu_0) - (x - \mu_1)^T \Sigma^{-1} (x - \mu_1)$$

$$RHS = \sum_{i=1}^d \frac{(x_i - \mu_{0,i})^2}{\Sigma_{ii}} - \sum_{i=1}^d \frac{(x_i - \mu_{1,i})^2}{\Sigma_{ii}} = \sum_{i=1}^d \frac{x_i^2 - 2x_i \mu_{0,i} + \mu_{0,i}^2 - x_i^2 + 2x_i \mu_{1,i} - \mu_{1,i}^2}{\Sigma_{ii}} =$$

$$= \sum_{i=1}^d \frac{2x_i(\mu_{1,i} - \mu_{0,i}) - (\mu_{1,i} - \mu_{0,i})(\mu_{1,i} + \mu_{0,i})}{\Sigma_{ii}} = \sum_{i=1}^d \frac{(\mu_{1,i} - \mu_{0,i})(2x_i - \mu_{1,i} - \mu_{0,i})}{\Sigma_{ii}} =$$

$$= (\mu_1^T - \mu_0^T) \Sigma^{-1} (2x - \mu_0 - \mu_1)$$

$$2 \ln \left(\frac{1-p}{p} \right) \rightarrow \text{if } 2 \ln \left(\frac{1-p}{p} \right) > 0 \text{ then } \mu_1^T - \mu_0^T > 0 \text{ if } 2 \ln \left(\frac{1-p}{p} \right) < 0 \text{ then } \mu_1^T - \mu_0^T < 0$$

□

$$0 \neq P(Y=1) \neq p \neq 1 - P(Y=0) \neq 1 - p \neq \lambda_1 - \lambda_0 \quad \text{and} \quad 0 = P(Y=1) \neq p \neq 1 - p$$

where $\lambda_1 > \lambda_0$

$$\text{rank}(V) \leq n-k, V = \sum_{i=1}^k v_i x_i \Rightarrow \text{rank } V = (n-k) \leq n \quad \text{per } \textcircled{D}$$

$$\text{Q.E.D.} \quad \text{rank}(V) = (n-k) \leq n \quad \text{by } \text{rank } V \leq n$$

$$\text{P.P.} \quad \sum_{i=1}^k v_i x_i - u = 0 \quad \text{for all } u \in \mathbb{R}^n$$

$$(R^J \rightarrow R^J - R^J u) \text{ with } J-1 \text{ linear functions}$$

□

①

$$\text{Gauss's Law} \quad S(x) = \frac{1}{2\pi b_0^2} \quad P(Y=1|X=x) \quad P(Y=0|X=x)$$

$$P(Y=1|X=x) > P(Y=0|X=x)$$

$$\frac{f_X(x|Y=1) P(Y=1)}{f_X(x|Y=0)} > \frac{P(X|Y=0) P(Y=0)}{f_X(x|Y=0)}$$

$$\frac{f_X(x|Y=1)}{f_X(x|Y=0)} > \frac{P(Y=1)}{P(Y=0)}$$

$$\therefore P := P(Y=1)$$

$$\frac{\frac{1}{\sqrt{2\pi b_0^2}} e^{-\frac{(x-\mu_1)^2}{2b_0^2}}}{\frac{1}{\sqrt{2\pi b_1^2}} e^{-\frac{(x-\mu_0)^2}{2b_1^2}}} > \frac{1-p}{p}$$

$$\frac{e^{-\frac{(x-\mu_1)^2}{2b_0^2}}}{e^{-\frac{(x-\mu_0)^2}{2b_1^2}}} > \frac{1-p}{p}$$

$$e^{-\frac{(x-\mu_1)^2}{2b_0^2}} - e^{-\frac{(x-\mu_0)^2}{2b_1^2}} > \frac{1-p}{p} \cdot \frac{b_1}{b_0}$$

$$e^{-\frac{(x-\mu_1)^2}{2b_0^2}} - e^{-\frac{(x-\mu_0)^2}{2b_1^2}} > \frac{1-p}{p} \cdot \frac{b_1}{b_0} \quad \mu_0 = \mu_1 = \mu$$

$$\frac{(x-\mu)^2}{b_0^2} \left(\frac{1}{b_0^2} - \frac{1}{b_1^2} \right) > \ln \left(\frac{1-p}{p} \cdot \frac{b_1}{b_0} \right)$$

$$(x-\mu)^2 > \frac{2b_0^2 b_1^2 \ln \left(\frac{1-p}{p} \cdot \frac{b_1}{b_0} \right)}{b_1^2 - b_0^2}$$

$$0 - \int_{-\infty}^{\mu} L_{\mu, b_0} \lambda d\lambda + \int_{\mu}^{\infty} L_{\mu, b_1} \lambda d\lambda > \mu - \frac{1}{2} \ln \left(\frac{1-p}{p} \cdot \frac{b_1}{b_0} \right) \quad \boxed{C}$$

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