

## Problem Set 4

Due Oct 31, before class

### 1. Equivalence of $F$ and $\bar{R}^2$

Consider testing  $\beta_2 = 0$  for the two nested Normal linear models:

$$Y = 1_n\beta_0 + X_1\beta_1 + \epsilon$$

$$Y = 1_n\beta_0 + X_1\beta_1 + X_2\beta_2 + \epsilon$$

We can use the standard  $F$  statistic. We can also compare the adjusted  $R^2$ 's from these two models:  $\bar{R}_1^2$  and  $\bar{R}_2^2$ . Show that

$$F > 1 \iff \bar{R}_1^2 < \bar{R}_2^2.$$

### 2. Best subset selection in *lalonge* data

Produce the figure similar to the best subset selection example in class using the *lalonge* data in the *Matching* package. Report the AIC, BIC, PRESS, and GCV of the selected model.

### 3. Derivative of the MSE of the ridge regression

Show that

$$\left. \frac{\partial \text{MSE}(\lambda)}{\partial \lambda} \right|_{\lambda=0} < 0.$$

*Remark:* This result ensures that the ridge estimator must have smaller MSE than OLS in a neighborhood of  $\lambda = 0$ .

#### 4. Ridge as OLS with augmented data

Show that  $\hat{\beta}^{\text{ridge}}(\lambda)$  equals the OLS coefficient of  $\tilde{Y}$  on  $\tilde{X}$  with augmented data

$$\tilde{Y} = \begin{pmatrix} Y \\ 0_p \end{pmatrix}, \quad \tilde{X} = \begin{pmatrix} X \\ \sqrt{\lambda} I_p \end{pmatrix},$$

where  $\tilde{Y}$  is an  $n + p$  dimensional vector and  $\tilde{X}$  is an  $(n + p) \times p$  matrix.

#### 5. Leave-one-out formulas for ridge

Define  $\hat{\beta}(\lambda) = (X^T X + \lambda I_p)^{-1} X^T Y$  as the ridge coefficient (dropping the superscript “ridge” for simplicity),  $\hat{\epsilon}(\lambda) = Y - X \hat{\beta}(\lambda)$  as the residual vector using the full data, and  $h_{ii}(\lambda) = x_i (X^T X + \lambda I_p)^{-1} x_i^T$  as the  $(i, i)$ -th diagonal element of  $H(\lambda) = X (X^T X + \lambda I_p)^{-1} X^T$ . Define  $\hat{\beta}_{[-i]}(\lambda)$  as the ridge coefficient without observation  $i$ , and  $\hat{\epsilon}_{[-i]}(\lambda) = y_i - x_i^T \hat{\beta}_{[-i]}(\lambda)$  as the predicted residual. Prove the following leave-one-out formulas for ridge regression

$$\begin{aligned} \hat{\beta}_{[-i]}(\lambda) &= \hat{\beta}(\lambda) - \{1 - h_{ii}(\lambda)\}^{-1} (X^T X + \lambda I_p)^{-1} x_i \hat{\epsilon}_i(\lambda), \\ \hat{\epsilon}_{[-i]}(\lambda) &= \frac{\hat{\epsilon}_i(\lambda)}{1 - h_{ii}(\lambda)}. \end{aligned}$$

*Hint:* You can use the result in the last problem and apply the leave-one-out formulas for OLS.

#### 6. An equivalent form of ridge coefficient

Show that the ridge coefficient has two equivalent forms: for  $\lambda > 0$

$$(X^T X + \lambda I_p)^{-1} X^T Y = X^T (X X^T + \lambda I_n)^{-1} Y.$$

*Hint:* Use the Woodbury formula  $(A + BDC)^{-1} = A^{-1} - A^{-1}B(D^{-1} + CA^{-1}B)^{-1}CA^{-1}$  by setting  $A = \lambda I_p$ ,  $B = X^T$ ,  $D = I_n$ , and  $C = X$ .

*Remark:* This formula is more useful when  $p > n$ ; If  $p > n$  and  $XX^T$  is invertible, then we can let  $\lambda$  go to zero on the right-hand side, yielding

$$\hat{\beta}^{\text{ridge}}(0) = X^T(XX^T)^{-1}Y.$$

## 7. Penalized OLS with an orthogonal design matrix

Consider the special case with standardized and orthogonal design matrix:

$$X^T 1_n = 0, \quad X^T X = I_p.$$

For the fixed  $\lambda \geq 0$ , find the explicit formulas of the  $j$ -th coordinate of the following estimators in terms of the corresponding  $j$ -th coordinate of the OLS estimator  $\hat{\beta}_j$  and  $\lambda$  ( $j = 1, \dots, p$ ):

$$\begin{aligned} \hat{\beta}^{\text{ridge}}(\lambda) &= \arg \min_{b \in \mathbb{R}^p} \{ \|Y - Xb\|_2^2 + \lambda \|b\|_2^2 \}, \\ \hat{\beta}^{\text{lasso}}(\lambda) &= \arg \min_{b \in \mathbb{R}^p} \{ \|Y - Xb\|_2^2 + \lambda \|b\|_1 \}, \\ \hat{\beta}^{\text{enet}}(\lambda) &= \arg \min_{b \in \mathbb{R}^p} \{ \|Y - Xb\|_2^2 + \lambda (\alpha \|b\|_2^2 + (1 - \alpha) \|b\|_1) \}, \\ \hat{\beta}^{\text{subset}}(\lambda) &= \arg \min_{b \in \mathbb{R}^p} \{ \|Y - Xb\|_2^2 + \lambda \|b\|_0 \}, \end{aligned}$$

where

$$\|b\|_2^2 = \|b\|^2 = \sum_{j=1}^p b_j^2, \quad \|b\|_1 = \sum_{j=1}^p |b_j|, \quad \|b\|_0 = \sum_{j=1}^p I(b_j \neq 0)$$

with  $I(\cdot)$  being the indicator function.

## 8. Coordinate descent for the elastic net

Give the detailed coordinate descent algorithm for the elastic net.

## 9. More noise in the Boston housing data

The Boston housing data have  $n = 506$  observations. Add  $p = n$  columns of covariates of random noise and randomly split the data into the training set and testing set with ratio 8 : 2. Compare ridge and lasso in terms of the estimated coefficients and the mean squared predicted error in the testing dataset.