

## Lecture Notes 6: Wishart Distribution

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## 6.1 Multivariable Normal Distribution

**Definition 6.1.** Let  $X \sim N(\mu, \Sigma)$  be a  $p \times 1$  random vector, where  $\mu$  is a  $p \times 1$  vector and  $\Sigma$  is a positive definite matrix, with probability density function:

$$p(X) = \frac{1}{(2\pi)^{\frac{p}{2}} |\Sigma|^{\frac{1}{2}}} \exp\left(-\frac{1}{2}(X - \mu)^T \Sigma^{-1}(X - \mu)\right)$$

where  $\mathbb{E}(X) = \mu$ ,  $Cov(X) = \Sigma$ .

Take  $X$  into two parts. i.e.  $X = \begin{pmatrix} X^{(1)} \\ X^{(2)} \end{pmatrix}$ ,  $\mu = \begin{pmatrix} \mu^{(1)} \\ \mu^{(2)} \end{pmatrix}$ ,  $\Sigma = \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix}$

where  $X^{(1)}$  is  $q \times 1$  and  $X^{(2)}$  is  $(p - q) \times 1$ , so as  $\mu^{(1)}, \mu^{(2)}$

Let's define  $X_{2,1} = X^{(2)} - \Sigma_{21} \Sigma_{11}^{-1} X^{(1)}$

**Theorem 6.1.** If  $X \sim N_p(\mu, \Sigma)$  then

1.  $X^{(1)} \sim N_q(\mu^{(1)}, \Sigma_{11})$ ,  $X^{(2)} \sim N_{p-q}(\mu^{(2)}, \Sigma_{22})$   
 $X_{2,1} \sim N_{p-q}(\mu_{2,1}, \Sigma_{22,1})$
2.  $X^{(1)}$  and  $X_{2,1}$  are independent
3.  $X^{(2)} | X^{(1)} \sim N_{p-q}(\mu^{(2)} + \Sigma_{21} \Sigma_{11}^{-1}(x^{(1)} - \mu^{(1)}), \Sigma_{22,1})$

where

$$\mu_{2,1} = \mu^{(2)} - \Sigma_{21} \Sigma_{11}^{-1} \mu^{(1)}$$

and

$$\Sigma_{22,1} = \Sigma_{22} - \Sigma_{21} \Sigma_{11}^{-1} \Sigma_{12}$$

is the Schur Complement of  $\Sigma_{11}$ .

**Remarks:** The Jacobian of transform  $(X^{(1)}, X^{(2)}) \rightarrow (X^{(1)}, X_{2,1})$  is 1.

*Proof.*

$$Z = \begin{pmatrix} X^{(1)} \\ X_{2,1} \end{pmatrix} = \begin{pmatrix} \mathbf{I} & 0 \\ -\Sigma_{21} \Sigma_{11}^{-1} & \mathbf{I} \end{pmatrix} \begin{pmatrix} X^{(1)} \\ X^{(2)} \end{pmatrix}$$

That makes  $dZ = \det \begin{pmatrix} \mathbf{I} & 0 \\ -\Sigma_{21} \Sigma_{11}^{-1} & \mathbf{I} \end{pmatrix} dX$

obviously,  $\det \begin{pmatrix} \mathbf{I} & 0 \\ -\Sigma_{21} \Sigma_{11}^{-1} & \mathbf{I} \end{pmatrix} = 1$ . □

Let  $\mathbf{B} = \begin{pmatrix} \mathbf{I} & 0 \\ -\boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1} & \mathbf{I} \end{pmatrix}$ , then  $X = \mathbf{B}^{-1}Z$ .

The above derivations are also established if replacing  $X$  with  $X - \mu$ . Hence, we have  $X - \mu = \mathbf{B}^{-1}Z$ . Since the Jacobian from  $X$  to  $Z$  is 1, we can derive the p.d.f of  $Z$  easily (just ignoring the constants):

$$\begin{aligned} & (X - \mu)^T \boldsymbol{\Sigma}^{-1} (X - \mu) \\ &= Z^T (\mathbf{B}^{-1})^T \boldsymbol{\Sigma}^{-1} \mathbf{B}^{-1} Z \\ &= Z^T (\mathbf{B} \boldsymbol{\Sigma} \mathbf{B}^T)^{-1} Z \\ &= Z^T \left[ \begin{pmatrix} \mathbf{I} & 0 \\ -\boldsymbol{\Sigma}_{21}\boldsymbol{\Sigma}_{11}^{-1} & \mathbf{I} \end{pmatrix} \begin{pmatrix} \boldsymbol{\Sigma}_{11} & \boldsymbol{\Sigma}_{12} \\ \boldsymbol{\Sigma}_{21} & \boldsymbol{\Sigma}_{22} \end{pmatrix} \begin{pmatrix} \mathbf{I} & -\boldsymbol{\Sigma}_{11}^{-1}\boldsymbol{\Sigma}_{12} \\ 0 & \mathbf{I} \end{pmatrix} \right]^{-1} Z \\ &= Z^T \begin{pmatrix} \boldsymbol{\Sigma}_{11}^{-1} & 0 \\ 0 & \boldsymbol{\Sigma}_{22.1}^{-1} \end{pmatrix} Z \end{aligned}$$

So  $Z$  forms a Gaussian Distribution with variance matrix  $\begin{pmatrix} \boldsymbol{\Sigma}_{11} & 0 \\ 0 & \boldsymbol{\Sigma}_{22.1}^{-1} \end{pmatrix}$ .

Since the covariance is 0, so  $X^{(1)}$  and  $X_{2.1}$  are independent. Now we have proved the proposition 1 and 2 in theorem 1.

Let consider the constant part  $|\boldsymbol{\Sigma}|^{\frac{1}{2}}$  to confirm our conclusion.

$$\begin{aligned} |\mathbf{B} \boldsymbol{\Sigma} \mathbf{B}^T| &= \begin{vmatrix} \boldsymbol{\Sigma}_{11} & 0 \\ 0 & \boldsymbol{\Sigma}_{22.1} \end{vmatrix} = |\boldsymbol{\Sigma}_{11}| |\boldsymbol{\Sigma}_{22.1}| = |\mathbf{B}|^2 |\boldsymbol{\Sigma}| \\ \Rightarrow |\boldsymbol{\Sigma}| &= |\boldsymbol{\Sigma}_{11}| |\boldsymbol{\Sigma}_{22.1}| \\ \Rightarrow |\boldsymbol{\Sigma}|^{\frac{1}{2}} &= |\boldsymbol{\Sigma}_{11}|^{\frac{1}{2}} |\boldsymbol{\Sigma}_{22.1}|^{\frac{1}{2}} \end{aligned}$$

So, the p.d.f of  $Z$  is

$$p(Z) = \frac{1}{(2\pi)^{\frac{p}{2}} |\boldsymbol{\Sigma}_{11}|^{\frac{1}{2}} |\boldsymbol{\Sigma}_{22.1}|^{\frac{1}{2}}} \exp\left(-\frac{1}{2} Z^{(1)} \boldsymbol{\Sigma}_{11}^{-1} Z^{(1)}\right) \exp\left(-\frac{1}{2} Z^{(2)} \boldsymbol{\Sigma}_{22.1}^{-1} Z^{(2)}\right)$$

**Corollary 6.1.**  $\boldsymbol{\Sigma}$  is positive definite  $\iff \boldsymbol{\Sigma}_{11}, \boldsymbol{\Sigma}_{22.1}$  is positive definite.

Now let's prove the proposition 3 in theorem 1. Since  $X^{(1)}$  is a constant in conditional probability, we have

$$\begin{aligned} X^{(2)} &= X_{2.1} + \boldsymbol{\Sigma}_{21} \boldsymbol{\Sigma}_{11}^{-1} X^{(1)} \\ \Rightarrow \mathbb{E}(X^{(2)} | X^{(1)}) &= \mu_{2.1} + \boldsymbol{\Sigma}_{21} \boldsymbol{\Sigma}_{11}^{-1} X^{(1)} \\ \Rightarrow \text{Cov}(X^{(2)} | X^{(1)}) &= \boldsymbol{\Sigma}_{22.1} \end{aligned}$$

That's all of the proving of theorem 6.1.

**Theorem 6.2.** If  $\mathbf{C} = \boldsymbol{\Sigma}^{-1}$ , i.e.  $\mathbf{C} = \begin{pmatrix} \mathbf{C}_{11} & \mathbf{C}_{12} \\ \mathbf{C}_{21} & \mathbf{C}_{22} \end{pmatrix} = \begin{pmatrix} \boldsymbol{\Sigma}_{11} & \boldsymbol{\Sigma}_{12} \\ \boldsymbol{\Sigma}_{21} & \boldsymbol{\Sigma}_{22} \end{pmatrix}^{-1}$ , then

$$1. \mathbf{C}_{22}^{-1} = \boldsymbol{\Sigma}_{22.1}$$

$$2. \mathbf{C}_{11}^{-1} \mathbf{C}_{12} = -\mathbf{\Sigma}_{12} \mathbf{\Sigma}_{22}^{-1}$$

*Proof.*

$$\begin{aligned} & \begin{pmatrix} \mathbf{I} & 0 \\ -\mathbf{\Sigma}_{21} \mathbf{\Sigma}_{11}^{-1} & \mathbf{I} \end{pmatrix} \begin{pmatrix} \mathbf{\Sigma}_{11} & \mathbf{\Sigma}_{12} \\ \mathbf{\Sigma}_{21} & \mathbf{\Sigma}_{22} \end{pmatrix} \begin{pmatrix} \mathbf{I} & -\mathbf{\Sigma}_{11}^{-1} \mathbf{\Sigma}_{12} \\ 0 & \mathbf{I} \end{pmatrix} = \begin{pmatrix} \mathbf{\Sigma}_{11} & 0 \\ 0 & \mathbf{\Sigma}_{22.1} \end{pmatrix} \\ \Rightarrow & \begin{pmatrix} \mathbf{\Sigma}_{11} & \mathbf{\Sigma}_{12} \\ \mathbf{\Sigma}_{21} & \mathbf{\Sigma}_{22} \end{pmatrix} = \begin{pmatrix} \mathbf{I} & 0 \\ \mathbf{\Sigma}_{21} \mathbf{\Sigma}_{11}^{-1} & \mathbf{I} \end{pmatrix} \begin{pmatrix} \mathbf{\Sigma}_{11} & 0 \\ 0 & \mathbf{\Sigma}_{22.1} \end{pmatrix} \begin{pmatrix} \mathbf{I} & \mathbf{\Sigma}_{11}^{-1} \mathbf{\Sigma}_{12} \\ 0 & \mathbf{I} \end{pmatrix} \\ \Rightarrow & \begin{pmatrix} \mathbf{C}_{11} & \mathbf{C}_{12} \\ \mathbf{C}_{21} & \mathbf{C}_{22} \end{pmatrix} = \begin{pmatrix} \mathbf{I} & -\mathbf{\Sigma}_{11}^{-1} \mathbf{\Sigma}_{12} \\ 0 & \mathbf{I} \end{pmatrix} \begin{pmatrix} \mathbf{\Sigma}_{11}^{-1} & 0 \\ 0 & \mathbf{\Sigma}_{22.1}^{-1} \end{pmatrix} \begin{pmatrix} \mathbf{I} & 0 \\ -\mathbf{\Sigma}_{21} \mathbf{\Sigma}_{11}^{-1} & \mathbf{I} \end{pmatrix} \\ \Rightarrow & \begin{pmatrix} \mathbf{C}_{11} & \mathbf{C}_{12} \\ \mathbf{C}_{21} & \mathbf{C}_{22} \end{pmatrix} = \begin{pmatrix} \mathbf{\Sigma}_{11}^{-1} + \mathbf{\Sigma}_{11}^{-1} \mathbf{\Sigma}_{12} \mathbf{\Sigma}_{22.1}^{-1} \mathbf{\Sigma}_{21} \mathbf{\Sigma}_{11}^{-1} & -\mathbf{\Sigma}_{11} \mathbf{\Sigma}_{12} \mathbf{\Sigma}_{22.1}^{-1} \\ -\mathbf{\Sigma}_{22.1}^{-1} \mathbf{\Sigma}_{21} \mathbf{\Sigma}_{11}^{-1} & \mathbf{\Sigma}_{22.1}^{-1} \end{pmatrix} \end{aligned}$$

So,  $\mathbf{C}_{22}^{-1} = \mathbf{\Sigma}_{22.1}$ . And

$$\begin{aligned} \mathbf{C}_{11}^{-1} \mathbf{C}_{12} &= (\mathbf{\Sigma}_{11}^{-1} + \mathbf{\Sigma}_{11}^{-1} \mathbf{\Sigma}_{12} \mathbf{\Sigma}_{22.1}^{-1} \mathbf{\Sigma}_{21} \mathbf{\Sigma}_{11}^{-1}) (-\mathbf{\Sigma}_{11} \mathbf{\Sigma}_{12} \mathbf{\Sigma}_{22.1}^{-1}) \\ &= -\mathbf{\Sigma}_{11} \mathbf{\Sigma}_{21}^{-1} - \mathbf{\Sigma}_{12} \mathbf{\Sigma}_{22.1}^{-1} \\ &= -\mathbf{\Sigma}_{11} \mathbf{\Sigma}_{21}^{-1} - \mathbf{\Sigma}_{12} (\mathbf{\Sigma}_{22} - \mathbf{\Sigma}_{21} \mathbf{\Sigma}_{11}^{-1} \mathbf{\Sigma}_{12})^{-1} \\ &= -\mathbf{\Sigma}_{12} \mathbf{\Sigma}_{22}^{-1} \end{aligned}$$

□

## 6.2 Matrix Variate Distribution

Let  $\mathbf{X} = (X_1, X_2, \dots, X_n)^T$ ,  $X_i \in \mathbb{R}^p$  and  $X_i \sim N(\mu_i, \mathbf{\Sigma})$ . If

$$\begin{aligned} p(\mathbf{X}) &= \prod_{i=1}^n p(X_i) \\ &= \prod_{i=1}^n \frac{1}{(2\pi)^{\frac{p}{2}} |\mathbf{\Sigma}|^{\frac{1}{2}}} \exp\left(-\frac{1}{2} (X_i - \mu_i)^T \mathbf{\Sigma}^{-1} (X_i - \mu_i)\right) \\ &= \frac{1}{(2\pi)^{\frac{p}{2}} |\mathbf{\Sigma}|^{\frac{1}{2}}} \exp\left(-\frac{1}{2} \text{tr}\left(\mathbf{\Sigma}^{-1} \sum_{i=1}^n (X_i - \mu_i)(X_i - \mu_i)^T\right)\right) \end{aligned}$$

Suppose  $\boldsymbol{\mu} = (\mu_1, \mu_2, \dots, \mu_n)^T$ , then

$$\begin{aligned} p(\mathbf{X}) &= \frac{1}{(2\pi)^{\frac{np}{2}} |\mathbf{\Sigma}|^{\frac{n}{2}}} \exp\left(-\frac{1}{2} \text{tr}(\mathbf{\Sigma}^{-1} (\mathbf{X} - \boldsymbol{\mu})^T \mathbf{I} (\mathbf{X} - \boldsymbol{\mu}))\right) \\ &= \frac{1}{(2\pi)^{\frac{np}{2}} |\mathbf{\Sigma}|^{\frac{n}{2}}} \text{etr}\left(-\frac{1}{2} \mathbf{\Sigma}^{-1} (\mathbf{X} - \boldsymbol{\mu})^T \mathbf{I} (\mathbf{X} - \boldsymbol{\mu})\right) \end{aligned}$$

We call  $\mathbf{X}$  is Matrix-variate normal distributed.

**Homework 1** If  $\text{vec}(\mathbf{X}^T) \sim N_{np}(\text{vec}(\boldsymbol{\mu}^T), \mathbf{B} \otimes \mathbf{A})$ , show the p.d.f of  $\mathbf{X}$  is

$$\frac{1}{(2\pi)^{\frac{np}{2}} |\mathbf{A}|^{\frac{n}{2}} |\mathbf{B}|^{\frac{p}{2}}} \exp\left(-\frac{1}{2} \text{tr}(\mathbf{A}^{-1} (\mathbf{X} - \boldsymbol{\mu})^T \mathbf{B}^{-1} (\mathbf{X} - \boldsymbol{\mu}))\right)$$

**Definition 6.2** (Wishart Distribution). If  $\mathbf{S} = \mathbf{X}^T \mathbf{X}$ , where the  $n \times p$  matrix  $\mathbf{X}$  is  $N(0, \mathbf{I}_n \otimes \mathbf{\Sigma})$ , then  $\mathbf{S}$  is positive definite and is said to have the Wishart distribution with  $n$  degrees of freedom and covariance matrix  $\mathbf{\Sigma}$ . We will write that  $\mathbf{S}$  is  $W_p(\mathbf{\Sigma}, n)$ , the subscript on  $W$  denoting the size of the matrix  $\mathbf{S}$ .

**Theorem 6.3.** If  $\mathbf{S}$  is  $W_p(\mathbf{\Sigma}, r)$  with  $r \geq p$  then the density function of  $\mathbf{S}$  is

$$p(\mathbf{S}) = \frac{|\mathbf{S}|^{\frac{r-p-1}{2}} \exp(-\frac{1}{2}\text{tr}(\mathbf{\Sigma}^{-1}\mathbf{S}))}{2^{\frac{rp}{2}} \pi^{\frac{p(p-1)}{4}} |\mathbf{\Sigma}|^{\frac{r}{2}} \prod_{i=1}^p \Gamma(\frac{r+1-i}{2})}, r \geq p$$

In Bayesian statistics, in the context of the multivariate normal distribution, the Wishart distribution is the conjugate prior to the precision matrix  $\mathbf{\Omega} = \mathbf{\Sigma}^{-1}$ , where  $\mathbf{\Sigma}$  is the covariance matrix.

Splitting  $\mathbf{S}$  into parts of  $q$  and  $p - q$ , i.e.

$$\mathbf{S} = \begin{pmatrix} \mathbf{S}_{11} & \mathbf{S}_{12} \\ \mathbf{S}_{21} & \mathbf{S}_{22} \end{pmatrix}$$

where  $\mathbf{S}_{11}$  is  $q \times q$  and  $\mathbf{S}_{22}$  is  $(p - q) \times (p - q)$ . So as  $\mathbf{\Sigma}$ .

**Theorem 6.4.** Let  $\mathbf{S} \sim W_p(\mathbf{\Sigma}, r)$ ,  $\mathbf{S}_{11.2} = \mathbf{S}_{11} - \mathbf{S}_{12}\mathbf{S}_{22}^{-1}\mathbf{S}_{21}$ ,  $\mathbf{\Sigma}_{11.2} = \mathbf{\Sigma}_{11} - \mathbf{\Sigma}_{12}\mathbf{\Sigma}_{22}^{-1}\mathbf{\Sigma}_{21}$  then

1.  $\mathbf{S}_{11} \sim W_q(\mathbf{\Sigma}_{11}, r)$ ,  $\mathbf{S}_{22} \sim W_{p-q}(\mathbf{\Sigma}_{22}, r)$
2.  $\mathbf{S}_{11.2} \sim W_q(\mathbf{\Sigma}_{11.2}, r - (p - q))$
3.  $\mathbf{S}_{11.2}$  and  $(\mathbf{S}_{12}, \mathbf{S}_{22})$  are independent.
4.  $\mathbf{S}_{12}|\mathbf{S}_{22} \sim N_{q,p-q}(\mathbf{\Sigma}_{12}\mathbf{\Sigma}_{22}^{-1}\mathbf{S}_{22}, \mathbf{\Sigma}_{11.2} \otimes \mathbf{S}_{22})$

*Proof.* Making the transformation

$$\begin{cases} \mathbf{S}_{11.2} = & \mathbf{S}_{11} - \mathbf{S}_{12}\mathbf{S}_{22}^{-1}\mathbf{S}_{21} \\ \mathbf{B}_{12} = & \mathbf{S}_{12} \\ \mathbf{B}_{22} = & \mathbf{S}_{22} \end{cases}$$

i.e.  $(\mathbf{S}_{11}, \mathbf{S}_{12}, \mathbf{S}_{22}) \rightarrow (\mathbf{S}_{11.2}, \mathbf{B}_{12}, \mathbf{B}_{22})$ . Since

$$\begin{aligned} (d(\mathbf{S}_{11.2}, \mathbf{B}_{12}, \mathbf{B}_{22})) &= (d(\mathbf{S}_{11.2}, \mathbf{S}_{12}, \mathbf{S}_{22})) \\ &= (d(\mathbf{S}_{11} - \mathbf{S}_{12}\mathbf{S}_{22}^{-1}\mathbf{S}_{21}, \mathbf{B}_{12}, \mathbf{B}_{22})) \\ &= (d(\mathbf{S}_{11}, \mathbf{B}_{12}, \mathbf{B}_{22})) \\ &= (d(\mathbf{S}_{11}, \mathbf{S}_{12}, \mathbf{S}_{22})) \end{aligned}$$

So the Jacobian is 1. Hence, we can replace  $\mathbf{S}$  with  $\mathbf{S}_{11.2}, \mathbf{B}_{12}, \mathbf{B}_{22}$ .

First, we have  $|\mathbf{S}| = |\mathbf{S}_{11.2}||\mathbf{S}_{22}| = |\mathbf{S}_{11.2}||\mathbf{B}_{22}|$ . Second, in the trace part, we have

$$\begin{aligned}
\text{tr}(\boldsymbol{\Sigma}^{-1}\mathbf{S}) &= \text{tr}\left(\begin{pmatrix} \mathbf{C}_{11} & \mathbf{C}_{12} \\ \mathbf{C}_{21} & \mathbf{C}_{22} \end{pmatrix} \begin{pmatrix} \mathbf{S}_{11} & \mathbf{S}_{12} \\ \mathbf{S}_{21} & \mathbf{S}_{22} \end{pmatrix}\right) \\
&= \text{tr}(\mathbf{C}_{11}\mathbf{S}_{11}) + 2\text{tr}(\mathbf{C}_{21}\mathbf{S}_{12}) + \text{tr}(\mathbf{C}_{22}\mathbf{S}_{22}) \\
&= \text{tr}(\mathbf{C}_{11}\mathbf{S}_{11}) + 2\text{tr}(\mathbf{C}_{12}\mathbf{B}_{21}) + \text{tr}(\mathbf{C}_{22}\mathbf{B}_{22}) \\
&\quad (\text{since } \mathbf{S}_{11} = \mathbf{S}_{11.2} + \mathbf{S}_{12}\mathbf{S}_{22}^{-1}\mathbf{S}_{21}) \\
&= \text{tr}(\mathbf{C}_{11}\mathbf{S}_{11.2}) + \text{tr}(\mathbf{C}_{11}\mathbf{B}_{12}\mathbf{B}_{22}^{-1}\mathbf{B}_{21}) + 2\text{tr}(\mathbf{C}_{12}\mathbf{B}_{21}) + \text{tr}(\mathbf{C}_{22}\mathbf{B}_{22}) \\
&\quad (\text{using theorem 6.2}) \\
&= \text{tr}(\boldsymbol{\Sigma}_{11.2}^{-1}\mathbf{S}_{11.2}) + \text{tr}(\mathbf{C}_{11}\mathbf{B}_{12}\mathbf{B}_{22}^{-1}\mathbf{B}_{21}) + 2\text{tr}(\mathbf{C}_{12}\mathbf{B}_{21}) + \text{tr}(\mathbf{C}_{22}\mathbf{B}_{22}) \\
&\quad (\text{using } \boldsymbol{\Sigma}_{22}^{-1} = \mathbf{C}_{22.1} = \mathbf{C}_{22} - \mathbf{C}_{21}\mathbf{C}_{11}^{-1}\mathbf{C}_{12}) \\
&= \text{tr}(\boldsymbol{\Sigma}_{11.2}^{-1}\mathbf{S}_{11.2}) + \text{tr}(\mathbf{C}_{11}\mathbf{B}_{12}\mathbf{B}_{22}^{-1}\mathbf{B}_{21}) + 2\text{tr}(\mathbf{C}_{12}\mathbf{B}_{21}) + \text{tr}(\boldsymbol{\Sigma}_{22}^{-1}\mathbf{B}_{22}) \\
&\quad + \text{tr}((\mathbf{C}_{21}\mathbf{C}_{11}^{-1}\mathbf{C}_{12})\mathbf{B}_{22})
\end{aligned}$$

We can see that  $\text{tr}(\boldsymbol{\Sigma}_{11.2}^{-1}\mathbf{S}_{11.2})$  is corresponding to  $p(\mathbf{S}_{11.2})$ ,  $\text{tr}(\boldsymbol{\Sigma}_{22}^{-1}\mathbf{B}_{22})$  is corresponding to  $p(\mathbf{B}_{22})$ . And to prove  $\mathbf{S}_{11.2}$  and  $(\mathbf{B}_{12}, \mathbf{B}_{22})$  are independent, we should have

$$p(\mathbf{S}_{11.2}, \mathbf{B}_{12}, \mathbf{B}_{22}) = p(\mathbf{S}_{11.2})p(\mathbf{B}_{12}, \mathbf{B}_{22}) = p(\mathbf{S}_{11.2})p(\mathbf{B}_{12}|\mathbf{B}_{22})p(\mathbf{B}_{22})$$

So, the residue terms should be corresponding to  $p(\mathbf{B}_{12}, \mathbf{B}_{22})$ , which is the 4th proposition in theorem 6.4. Now we rewritten them to show that they are corresponding to  $N_{q,p-q}(\boldsymbol{\Sigma}_{12}\boldsymbol{\Sigma}_{22}^{-1}\mathbf{S}_{22}, \boldsymbol{\Sigma}_{11.2} \otimes \mathbf{S}_{22})$ .

$$\begin{aligned}
&\text{tr}(\mathbf{C}_{21}\mathbf{C}_{11}^{-1}\mathbf{C}_{12}\mathbf{B}_{22}) + \text{tr}(\mathbf{C}_{11}\mathbf{B}_{12}\mathbf{B}_{22}^{-1}\mathbf{B}_{21}) + 2\text{tr}(\mathbf{C}_{12}\mathbf{B}_{21}) \\
&= \text{tr}(\mathbf{C}_{11}(\mathbf{B}_{12} + \mathbf{C}_{11}^{-1}\mathbf{C}_{12}\mathbf{B}_{22})\mathbf{B}_{22}^{-1}(\mathbf{B}_{12} + \mathbf{C}_{11}^{-1}\mathbf{C}_{12}\mathbf{B}_{22})^T) \\
&= \text{tr}(\boldsymbol{\Sigma}_{11.2}^{-1}(\mathbf{B}_{12} - \boldsymbol{\Sigma}_{12}\boldsymbol{\Sigma}_{22}^{-1}\mathbf{B}_{22})\mathbf{B}_{22}^{-1}(\mathbf{B}_{12} - \boldsymbol{\Sigma}_{12}\boldsymbol{\Sigma}_{22}^{-1}\mathbf{B}_{22})^T)
\end{aligned}$$

Finally, we have  $|\boldsymbol{\Sigma}| = |\boldsymbol{\Sigma}_{11.2}||\boldsymbol{\Sigma}_{22}|$ .

Now we have proved that the p.d.f of  $\mathbf{S}$  can be decomposed into terms  $\mathbf{S}_{22}$ ,  $\mathbf{S}_{11.2}$  and  $\mathbf{S}_{12}|\mathbf{S}_{22}$ .  $\square$

The following theorem is used to solve the problem: how to sample from a Wishart distribution.

**Theorem 6.5.** *Let  $\mathbf{S} \sim W_p(\mathbf{I}_p, r)$  and  $\mathbf{S} = \mathbf{T}^T\mathbf{T}$  where  $\mathbf{T} = (t_{i,j})$  is a upper triangle matrix,  $t_{i,i} > 0$  then*

1.  $t_{i,j} \quad 1 \leq j \leq i \leq p$  are independently distributed.
2.  $t_{i,i}^2 \sim \chi_{r-i+1}^2$
3.  $t_{i,j} \sim N(0, 1) \quad 1 \leq j < i \leq p$

*Proof.* First, we have

$$|\mathbf{S}|^{\frac{1}{2}(r-p-1)} \text{etr}\left(-\frac{1}{2}\mathbf{S}\right) = \left(\prod_{i=1}^p t_{ii}^{\frac{2(r-p-1)}{2}}\right) \text{etr}\left(-\frac{1}{2} \sum_{1 \leq j \leq i \leq p} t_{ij}^2\right)$$

According to Theorem 5.6, we have  $J(\mathbf{S} \rightarrow \mathbf{T}) = 2^p \prod_{i=1}^p t_{ii}^{p-i+1}$ . Also, we have  $\text{tr}(\mathbf{S}) = \text{tr}(\mathbf{T}^T \mathbf{T})$ . Thus,

$$\begin{aligned} p(\mathbf{T}) &\propto \prod_{1 \leq j \leq i \leq p} \exp\left(-\frac{1}{2}t_{ij}^2\right) \prod_{i=1}^p (t_{ii})^{\frac{r-p-1}{2}} |J(\mathbf{S} \rightarrow \mathbf{T})| \\ &\propto \prod_{1 \leq j < i \leq p} \exp\left(-\frac{1}{2}t_{ij}^2\right) \prod_{i=1}^p (t_{ii})^{\frac{r-i+1}{2}} \exp\left(-\frac{1}{2}t_{ii}^2\right) \end{aligned}$$

$\prod_{1 \leq j < i \leq p} \exp\left(-\frac{1}{2}t_{ij}^2\right)$  denote the independent standard normal distributions of  $t_{i,j}$ .

$\prod_{i=1}^p (t_{ii})^{\frac{r-i+1}{2}} \exp\left(-\frac{1}{2}t_{ii}^2\right)$  denote the independent distributions  $\chi_{r-i+1}^2$ .  $\square$

Wishart distribution is a generalization to multiple dimensions of the chi-squared distribution, If  $p = 1$  and  $\Sigma = 1$  then this distribution is a chi-squared distribution with  $r$  degrees of freedom.

**Definition 6.3.**  $\mathbf{S}^{-1}$  is said to have an inverse Wishart Distribution  $W_p^{-1}(\Sigma, r)$  if its p.d.f. ( $\mathbf{M} = \mathbf{S}^{-1}$ )

$$f(\mathbf{M}) = \frac{|\mathbf{M}|^{-\frac{r+p+1}{2}} \text{etr}\left(-\frac{1}{2}\Sigma^{-1}\mathbf{M}^{-1}\right)}{2^{\frac{rp}{2}} \pi^{\frac{p(p-1)}{4}} |\Sigma|^{\frac{r}{2}} \prod_{i=1}^p \Gamma\left(\frac{r+1-i}{2}\right)} \quad (1)$$

**Theorem 6.6.** Let  $\mathbf{A}$  and  $\mathbf{B}$  be independent where  $\mathbf{A} \sim W_p(\Sigma, r_1)$ ,  $\mathbf{B} \sim W_p(\Sigma, r_2)$ , with  $r_1 \geq p$ ,  $r_2 \geq p$ . Put  $\mathbf{A} + \mathbf{B} = \mathbf{T}^T \mathbf{T}$ .  $\mathbf{T}$  is upper triangular. And  $\mathbf{A} = \mathbf{T}^T \mathbf{U} \mathbf{T}$ . Let  $\mathbf{U}$  be an  $m \times m$  symmetric matrix. then  $0 \prec \mathbf{U} \prec \mathbf{I}$ , and

1.  $\mathbf{A} + \mathbf{B}$  and  $\mathbf{U}$  are independent

2.  $\mathbf{A} + \mathbf{B} \sim W_p(\Sigma, r_1 + r_2)$

3.  $p(u) \propto |\mathbf{U}|^{\frac{r_1-p-1}{2}} |\mathbf{I} - \mathbf{U}|^{\frac{r_2-p-1}{2}}$

$p(\mathbf{U})$  is called matrix-variate Beta Distribution.

**Homework 2** Prove theorem 6.6.