

Q1:

$\forall j \in \{1, 2, \dots, p\}$, 将第 j 列与前 $j-1$ 列交换, 则有

$$T_j = \begin{pmatrix} 1 & & & \\ & \ddots & & \\ & & 1 & \\ & & & \ddots \\ & & & & 1 \end{pmatrix} \quad T_j^{-1} = T_j^T = \begin{pmatrix} 1 & & & \\ & \ddots & & \\ & & 1 & \\ & & & \ddots \\ & & & & 1 \end{pmatrix}$$

$$\text{令 } X_t = (X_j, \underbrace{X_1, \dots, X_{j-1}, X_{j+1}, \dots, X_p}_{X_0})$$

$$X_t = X_s T_j \Rightarrow X_s = X_t T_j^{-1} = X_t T_j^T$$

$$\begin{aligned} \therefore C &= (X_s^T X_s)^{-1} = \{c_{ij}\} = (X_t T_j^T)^T (X_t T_j^T)^{-1} \\ &= (T_j X_t^T X_t T_j^T)^{-1} = T_j (X_t^T X_t)^{-1} T_j^T \end{aligned}$$

$$X_t = (X_j, X_0) \quad X_t^T X_t = \begin{pmatrix} X_j^T \\ X_0^T \end{pmatrix} (X_j, X_0) = \begin{pmatrix} X_j^T X_j & X_j^T X_0 \\ X_0^T X_j & X_0^T X_0 \end{pmatrix}$$

$$(X_t^T X_t)^{-1} = \begin{pmatrix} \chi_{11}^* & \chi_{12}^* \\ \chi_{21}^* & \chi_{22}^* \end{pmatrix}$$

$$\chi_{11}^* = (X_j^T X_j - X_j^T X_0 (X_0^T X_0)^{-1} X_0^T X_j)^{-1}$$

$$\chi_{12}^* = - (X_j^T X_j - X_j^T X_0 (X_0^T X_0)^{-1} X_0^T X_j)^{-1} X_j^T X_0 (X_0^T X_0)^{-1}$$

$$\chi_{21}^* = - (X_0^T X_0)^{-1} X_0^T X_j (X_j^T X_j - X_j^T X_0 (X_0^T X_0)^{-1} X_0^T X_j)^{-1}$$

$$\chi_{22}^* = (X_0^T X_0)^{-1} + (X_0^T X_0)^{-1} X_0^T X_j (X_j^T X_j - X_j^T X_0 (X_0^T X_0)^{-1} X_0^T X_j)^{-1} X_j^T X_0 (X_0^T X_0)^{-1}$$

X_j 是经标准化后的

$$\therefore \sum_{i=1}^n X_{ij}^2 = 1 \Rightarrow X_j^T X_j = 1$$

$$\begin{aligned} C_{jj} &= X_{jj}^* = (X_j^T X_j - X_j^T X_0 (X_0^T X_0)^{-1} X_0^T X_j)^{-1} \\ &= (1 - X_j^T X_0 (X_0^T X_0)^{-1} X_0^T X_j)^{-1} \end{aligned}$$

$$\therefore \bar{X}_j = \frac{1}{n} \sum_{k=1}^n X_{jk} = 0$$

$$\begin{aligned} \therefore SS_R^j &= \sum_{k=1}^n (\hat{X}_{jk} - \bar{X}_j)^2 = \sum_{k=1}^n \hat{X}_{jk}^2 = \hat{X}_j^T \hat{X}_j = (X_0 \hat{\beta})^T X_0 \hat{\beta} \\ &= X_j^T X_0 (X_0^T X_0)^{-1} X_0^T X_0 (X_0^T X_0)^{-1} X_0^T X_j \\ &= X_j^T X_0 (X_0^T X_0)^{-1} X_0^T X_j \end{aligned}$$

$$SS_T^j = \sum_{k=1}^n (X_{jk} - \bar{X}_j)^2 = \sum_{k=1}^n X_{jk}^2 = 1$$

$$R_j^2 = \frac{SS_R^j}{SS_T^j} = X_j^T X_0 (X_0^T X_0)^{-1} X_0^T X_j$$

$$\therefore C_{jj} = (1 - X_j^T X_0 (X_0^T X_0)^{-1} X_0^T X_j)^{-1} = \frac{1}{1 - R_j^2}$$

Q2:

$$\hat{\beta} = (X'X)^{-1} X'y \quad E(\hat{\beta}) = \beta \quad \text{Var}(\hat{\beta}) = \sigma^2 (X'X)^{-1}$$

$$MSE(\hat{\beta}) = E(\hat{\beta} - \beta)'(\hat{\beta} - \beta) = E \left[\underbrace{(\hat{\beta} - E(\hat{\beta}))'(\hat{\beta} - E(\hat{\beta}))}_{-\uparrow = \text{次型}} \right]$$

$$= \text{tr}(E[(\hat{\beta} - E(\hat{\beta}))'(\hat{\beta} - E(\hat{\beta}))])$$

$$= \text{tr}(\text{Var}(\hat{\beta})) = \text{tr}(\sigma^2 (X'X)^{-1}) = \sigma^2 \text{tr}[(X'X)^{-1}]$$

假定 $X'X$ 的特征值为 $\lambda_1, \lambda_2, \dots, \lambda_p$.

则 $(X'X)^{-1}$ 的特征值有 $\frac{1}{\lambda_1}, \frac{1}{\lambda_2}, \dots, \frac{1}{\lambda_p}$

$$\text{tr}[(X'X)^{-1}] = \sum_{i=1}^p \frac{1}{\lambda_i}$$

$$\therefore MSE(\hat{\beta}) = \sigma^2 \text{tr}[(X'X)^{-1}] = \sigma^2 \sum_{i=1}^p \frac{1}{\lambda_i}$$

从贝叶斯统计的角度解释岭回归
最大化后验估计

$$J(w) = L(w) + \lambda P(w) \quad \text{正则化框架}$$

\downarrow \downarrow
 loss penalty

$$\hat{w} = \operatorname{argmin}_w J(w), \quad \lambda > 0$$

岭回归: $P(w) = w^T w$

$$J(w) = \sum_{i=1}^n (w^T x_i - y_i)^2 + \lambda w^T w$$

$$\hat{w} = \operatorname{argmin}_w J(w) = (X^T X + \lambda I)^{-1} X^T y$$

贝叶斯角度

$$\epsilon \sim N(0, \sigma^2)$$

$$Y = w^T X + \epsilon = f(w) + \epsilon \quad Y \sim N(w^T X, \sigma^2)$$

假设参数 w 也符合高斯分布

$$w \sim N(0, \sigma_1^2)$$

$$P(w) = \frac{1}{\sqrt{2\pi} \sigma_1} \exp \left\{ -\frac{w^T w}{2\sigma_1^2} \right\}$$

$$P(Y|w; X) = \frac{1}{\sqrt{2\pi} \sigma} \exp \left\{ -\frac{(Y - w^T X)^T (Y - w^T X)}{2\sigma^2} \right\}$$

最大后验概率估计

$$P(w|Y) = \frac{P(w) P(Y|w)}{P(Y)}$$

$$\hat{w} = \operatorname{argmax}_w P(w|Y) = \operatorname{argmax}_w \frac{P(w) P(Y|w)}{P(Y)} = \operatorname{argmax}_w P(w) P(Y|w) = \operatorname{argmax}_w \log(P(w) P(Y|w))$$

$$= \operatorname{argmax}_w \log \frac{1}{\sqrt{2\pi} \sigma_1} - \frac{w^T w}{2\sigma_1^2} + \log \frac{1}{\sqrt{2\pi} \sigma} - \frac{(Y - w^T X)^T (Y - w^T X)}{2\sigma^2}$$

$$= \operatorname{argmin}_w (Y - w^T X)^2 + \frac{\sigma^2}{\sigma_1^2} w^T w = \operatorname{argmin}_w \sum_{i=1}^n (y_i - w^T x_i)^2 + \lambda w^T w$$

$$\therefore \hat{w} = \operatorname{argmin}_w \underbrace{\sum_{i=1}^n (y_i - w^T x_i)^2}_{\text{loss}} + \underbrace{\left(\frac{\sigma^2}{\sigma_1^2} \right) w^T w}_{\lambda \text{ penalty}}$$

(最大后验)

一回事