

一、  
(1)  $X \sim \text{Nb}(2, \theta)$   
 $E(X) = \frac{2}{\theta}$

$\therefore \hat{\theta} = \frac{2}{\bar{x}}$

(2)  
 $E(X) = \int_0^1 x (\theta+1) x^{\theta} dx = \int_0^1 (\theta+1) x^{\theta+1} dx = (\theta+1) \frac{x^{\theta+2}}{\theta+2} \Big|_0^1$   
 $\hat{\theta} = \frac{1-2\bar{x}}{\bar{x}-1} = \frac{(\theta+1)}{\theta+2}$

二、  
 $\hat{\mu} = \bar{x}$  假设  $X \sim N(\mu, 1)$  则  $Y = \frac{X-\mu}{1} \sim N(0, 1)$

$f = \frac{k}{n} = P(X > 0)$

$P(X < 0) = 1 - \frac{k}{n}$

$\therefore P(Y < -\mu) = 1 - \frac{k}{n} = \Phi(-\mu) = 1 - \Phi(\mu)$

$\therefore \Phi(\hat{\mu}) = \frac{k}{n}$

$\therefore \hat{\mu} = u_{k/n}$

3、  
 $L(\theta) = \left(\frac{1}{2\theta}\right)^n \exp\left\{-\frac{\sum_{i=1}^n |x_i|}{\theta}\right\} \quad \theta > 0$

取对数

$\ln L(\theta) = n \ln \frac{1}{2\theta} + \frac{-\sum_{i=1}^n |x_i|}{\theta}$

$\frac{d \ln L(\theta)}{d\theta} = n \cdot \left(-\frac{1}{\theta}\right) + \frac{-\sum_{i=1}^n |x_i|}{\theta^2} \cdot (-1) = -\frac{n}{\theta} + \frac{\sum_{i=1}^n |x_i|}{\theta^2} = 0$

$\therefore \hat{\theta} = \frac{\sum_{i=1}^n |x_i|}{n}$

验证

$\frac{d^2 \ln L(\theta)}{d\theta^2} = \frac{n}{\theta^2} + \frac{-2 \sum_{i=1}^n |x_i|}{\theta^3} \Big|_{\theta=\hat{\theta}} = \frac{n^3}{\left(\sum_{i=1}^n |x_i|\right)^2} + \frac{-2n^3}{\left(\sum_{i=1}^n |x_i|\right)^2} = \frac{-n^3}{\left(\sum_{i=1}^n |x_i|\right)^2} < 0$

$\therefore \hat{\theta} = \frac{\sum_{i=1}^n |x_i|}{n}$



(b)

$$L(\theta) = \frac{1}{(\theta_2 - \theta_1)^n} I\{\theta_1 < X_i < \theta_2, i=1, 2, \dots, n\}$$

$$= \frac{1}{(\theta_2 - \theta_1)^n} I\{\theta_1 < X_{(1)} < X_{(n)} < \theta_2\} \uparrow$$

$\theta_2 - \theta_1$  小 但在  $X_{(1)}$  和  $X_{(n)}$  两边

$$\therefore \hat{\theta}_1 = X_{(1)} \quad \hat{\theta}_2 = X_{(n)}$$

IV、

$$L(\theta) = e^{-\sum_{i=1}^n X_i + n\theta} \quad X > \theta$$

$$\ln L(\theta) = -\sum_{i=1}^n X_i + n\theta \quad X > \theta$$

$$L(\theta) \uparrow \quad \theta \uparrow \quad \therefore \hat{\theta}_1 = X_{(1)}$$

$$p(x) = e^{-(x-\theta)} \quad x > \theta$$

$$F(x) = \int_{-\infty}^x e^{-x+\theta} dx + \theta$$

$$= - (e^{-x+\theta} \big|_{\theta}^x)$$

$$= - (e^{-x+\theta} - e^0)$$

$$= -e^{-x+\theta} + 1$$

VE

$$E(\hat{\theta}_1) = E(X_{(1)})$$

$$p_1(x) = n \cdot L(x) e^{-x+\theta} \cdot n^{n-1} e^{-x+\theta}$$

$$= n (e^{-x+\theta})^{n-1} \cdot e^{-x+\theta}$$

$$= n (e^{-x+\theta})^n \quad x > \theta$$

$$E(X_{(1)}) = n \int_{\theta}^{+\infty} x \cdot n e^{-nx+n\theta} dx = \int_{\theta}^{+\infty} x e^{-nx+n\theta} d(-nx+n\theta)$$

$$= - \int_{\theta}^{+\infty} x d(e^{-nx+n\theta}) = - \left( x \cdot e^{-nx+n\theta} \big|_{\theta}^{+\infty} - \int_{\theta}^{+\infty} e^{-nx+n\theta} dx \right)$$

$$= - \left( 0 - \theta \cdot e^{-\theta+\theta} - \frac{1}{n} \right) = - \left( -\frac{1}{n} - \theta \right) = \theta + \frac{1}{n}$$

$$\therefore E(\hat{\theta}_1) = \theta + \frac{1}{n} \neq \theta \quad \therefore \hat{\theta}_1 \text{ 不是 } \theta \text{ 的 UE}$$

$$\lim_{n \rightarrow \infty} E(\hat{\theta}_1) = \lim_{n \rightarrow \infty} \left( \theta + \frac{1}{n} \right) = \theta$$

$$\lim_{n \rightarrow \infty} \text{Var}(\hat{\theta}_1) = \lim_{n \rightarrow \infty} E(\hat{\theta}_1^2) - \theta^2 = \lim_{n \rightarrow \infty} E(X_{(1)}^2) - \theta^2$$

$$E(X_{(1)}^2) = n \int_{\theta}^{+\infty} x^2 e^{-nx+n\theta} dx = - \left( x^2 \cdot e^{-nx+n\theta} \big|_{\theta}^{+\infty} - 2 \int_{\theta}^{+\infty} x \cdot e^{-nx+n\theta} dx \right)$$

$$= - \left( 0 - \theta^2 \cdot e^{-\theta+\theta} - 2 \frac{\theta + \frac{1}{n}}{n} \right) = \theta^2 + \frac{2\theta + \frac{2}{n}}{n} = \theta^2 + \frac{2\theta}{n} + \frac{2}{n^2}$$

$$\therefore \lim_{n \rightarrow \infty} \text{Var}(\hat{\theta}_1) = \theta^2 - \theta^2 = 0$$

$\therefore \hat{\theta}_1$  是  $\theta$  的相合估计



(b)

$$E(X) = \int_0^{+\infty} x e^{-(x-\theta)} dx = - \int_0^{+\infty} x e^{-x+\theta} d(-x) = - \int_0^{+\infty} x d e^{-x+\theta}$$

$$= - (x \cdot e^{-x+\theta} \Big|_0^{+\infty} - \int_0^{+\infty} e^{-x+\theta} dx) = - (0 - \theta - 1) = 1 + \theta$$

$$\therefore \hat{\theta}_2 = \bar{X} - 1$$

$$UE: E(\hat{\theta}_2) = E(\bar{X} - 1) = E(\bar{X}) - 1 = E(X) - 1 = 1 + \theta - 1 = \theta$$

$\therefore \hat{\theta}_2$  是  $\theta$  的 UE

$$\lim_{n \rightarrow \infty} E(\hat{\theta}_2) = \theta$$

$$\lim_{n \rightarrow \infty} \text{Var}(\hat{\theta}_2) = \lim_{n \rightarrow \infty} \text{Var}(\bar{X} - 1) = \lim_{n \rightarrow \infty} \text{Var}\left(\frac{\sum_{i=1}^n X_i}{n}\right) = \lim_{n \rightarrow \infty} \frac{n \text{Var}(X)}{n^2}$$

$$= \lim_{n \rightarrow \infty} \frac{\text{Var}(X)}{n}$$

$$\text{Var}(X) = E(X^2) - (1+\theta)^2 = -(0 - \theta^2 - (2 + 2\theta)) - 1 - 2\theta - \theta^2$$

$$= \cancel{\theta^2} + \cancel{2} + \cancel{2\theta} - 1 - \cancel{2\theta} - \cancel{\theta^2} = 1$$

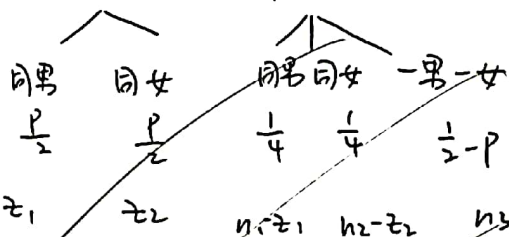
$$\therefore \lim_{n \rightarrow \infty} \text{Var}(\hat{\theta}_2) = \lim_{n \rightarrow \infty} \frac{1}{n} = 0$$

$\therefore \hat{\theta}_2$  是  $\theta$  的相合估计

5.

同 异

P 1-P

\*  $n_1 + n_2 + n_3 > 1$ 

$$L(p; n_1, n_2, n_3, z_1, z_2) = \left(\frac{p}{2}\right)^{z_1} \left(\frac{p}{2}\right)^{z_2} \cdot \left(\frac{1}{4}\right)^{n_1 - z_1} \left(\frac{1}{4}\right)^{n_2 - z_2} \left(\frac{1}{2} - p\right)^{n_3}$$

$$l(p; n_1, n_2, n_3, z_1, z_2) = (z_1 + z_2) \ln \frac{p}{2} + (n_1 + n_2 - z_1 - z_2) \ln \frac{1}{4} + n_3 \ln \left(\frac{1}{2} - p\right)$$

$$\frac{\partial l(p; n_1, n_2, n_3, z_1, z_2)}{\partial p} = \frac{(z_1 + z_2)}{p} + \frac{2n_3}{2p - 1} = 0$$

$$E(z_1 | p) = \frac{p}{2}$$

$$E(z_2 | p) =$$

