# Data Science II with python (Class notes)

**STAT 303-2** 

Arvind Krishna

1/3/23

## Table of contents

Pr	face	4
I	Linear regression	5
1	Simple Linear Regression	6
2	Multiple Linear Regression	12
3	Variable interactions and transformations  3.0.1 Variable interaction between continuous predictors	17 17 20
	3.1 Variable transformations	26 27 29
4	Model assumptions 4.1 Non-linearity of data	33 35 36
Αļ	pendices	41
Α	Assignment A  A.1 Regression vs Classification; Prediction vs Inference	42 43 44 44 45
В	Assignment B  B.1 Multiple linear regression	50 50
	B.1.3 Coefficient interpretation	50

	B.1.4	Variable significance	50
	B.1.5	Variable significance from confidence interval	50
	B.1.6	<i>p</i> -value	50
	B.1.7	Predictor significance in presence / absence of other predictors	51
	B.1.8	Prediction	51
	B.1.9	Variable selection	51
B.2	Using	MLR coefficients and variable transformation $\ldots \ldots \ldots \ldots$	51
	B.2.1	Data visualisation	52
	B.2.2	Removing effect of predictor from response	52
	B.2.3	Data visualisation after removing effect of predictor from response	53
В.3	Variab	ple transformations and interactions	53
	B.3.1	Training SLR	53
	B.3.2	Linearity in relationship	54
	B.3.3	Variable transformation	54
	B.3.4	Model visualisation with transformed predictor	54
	B.3.5	Training MLR with qualitative predictor	54
	B.3.6	Variable interaction	55
	B.3.7	Model visualisation with qualitative predictor	55
	B.3.8	Model interpretation	55
C Da	tasets, a	assignment and project files	56
Refere	nces		57

## **Preface**

These are class notes for the course STAT303-2. This is not the course text-book. You are required to read the relevant sections of the book as mentioned on the course website.

The course notes are currently being written, and will continue to being developed as the course progresses (just like the course textbook last quarter). Please report any typos / mistakes / inconsistencies / issues with the class notes / class presentations in your comments here. Thank you!

# Part I Linear regression

## 1 Simple Linear Regression

```
import pandas as pd
import numpy as np
import statsmodels.formula.api as smf
import seaborn as sns
import matplotlib.pyplot as plt
```

Develop a simple linear regression model that predicts car price based on engine size. Datasets to be used: Car\_features\_train.csv, Car\_prices\_train.csv

```
trainf = pd.read_csv('./Datasets/Car_features_train.csv')
trainp = pd.read_csv('./Datasets/Car_prices_train.csv')
train = pd.merge(trainf,trainp)
train.head()
```

	carID	brand	model	year	transmission	mileage	fuelType	tax	mpg	engineSize	price
0	18473	bmw	6 Series	2020	Semi-Auto	11	Diesel	145	53.3282	3.0	37980
1	15064	bmw	6 Series	2019	Semi-Auto	10813	Diesel	145	53.0430	3.0	33980
2	18268	bmw	6 Series	2020	Semi-Auto	6	Diesel	145	53.4379	3.0	36850
3	18480	bmw	6 Series	2017	Semi-Auto	18895	Diesel	145	51.5140	3.0	25998
4	18492	bmw	6 Series	2015	Automatic	62953	Diesel	160	51.4903	3.0	18990

```
#Using the ols function to create an ols object. 'ols' stands for 'Ordinary least squares'
ols_object = smf.ols(formula = 'price~engineSize', data = train)

#Using the fit() function of the 'ols' class to fit the model
model = ols_object.fit()

#Printing model summary which contains among other things, the model coefficients
model.summary()
```

Table 1.2: OLS Regression Results

Dep. Variable:	price	R-squared:	0.390
Model:	OLS	Adj. R-squared:	0.390
Method:	Least Squares	F-statistic:	3177.
Date:	Thu, 19 Jan 2023	Prob (F-statistic):	0.00
Time:	16:44:04	Log-Likelihood:	-53949.
No. Observations:	4960	AIC:	1.079e + 05
Df Residuals:	4958	BIC:	1.079e + 05
Df Model:	1		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[0.025	0.975]
Intercept	-4122.0357	522.260	-7.893	0.000	-5145.896	-3098.176
${\it engine Size}$	1.299e + 04	230.450	56.361	0.000	1.25e + 04	1.34e + 04

Omnibus:	1271.986	Durbin-Watson:	0.517
Prob(Omnibus):	0.000	Jarque-Bera (JB):	6490.719
Skew:	1.137	Prob(JB):	0.00
Kurtosis:	8.122	Cond. No.	7.64

The model equation is: car price = -4122.0357 + 12990 \* engine Size

#### Visualize the regression line

```
sns.regplot(x = 'engineSize', y = 'price', data = train, color = 'orange', line_kws={"color plt.xlim(-1,7)}
#Note that some of the engineSize values are 0. They are incorrect, and should ideally be
```

(-1.0, 7.0)



Predict the car price for the cars in the test dataset. Datasets to be used:  $Car\_features\_test.csv,\ Car\_prices\_test.csv$ 

```
testf = pd.read_csv('./Datasets/Car_features_test.csv')
testp = pd.read_csv('./Datasets/Car_prices_test.csv')

#Using the predict() function associated with the 'model' object to make predictions of capred_price = model.predict(testf)#Note that the predict() function finds the predictor 'en
```

# Make a visualization that compares the predicted car prices with the actual car prices

```
sns.scatterplot(x = testp.price, y = pred_price)
#In case of a perfect prediction, all the points must lie on the line x = y.
sns.lineplot(x = [0,testp.price.max()], y = [0,testp.price.max()],color='orange') #Plottin
plt.xlabel('Actual price')
plt.ylabel('Predicted price')
```

Text(0, 0.5, 'Predicted price')



The prediction doesn't look too good. This is because we are just using one predictor - engine size. We can probably improve the model by adding more predictors when we learn multiple linear regression.

#### What is the RMSE of the predicted car price?

```
np.sqrt(((testp.price - pred_price)**2).mean())
```

12995.1064515487

The root mean squared error in predicting car price is around \$13k.

#### What is the residual standard error based on the training data?

```
np.sqrt(model.mse_resid)
```

#### 12810.109175214136

The residual standard error on the training data is close to the RMSE on the test data. This shows that the performance of the model on unknown data is comparable to its performance

on known data. This implies that the model is not overfitting, which is good! In case we overfit a model on the training data, it's performance on unknown data is likely to be worse than that on the training data.

#### Find the confidence and prediction intervals of the predicted car price

#Using the get\_prediction() function associated with the 'model' object to get the interval
intervals = model.get\_prediction(testf)

#The function requires specifying alpha (probability of Type 1 error) instead of the confiintervals.summary\_frame(alpha=0.05)

	mean	mean_se	$mean\_ci\_lower$	$mean\_ci\_upper$	$obs\_ci\_lower$	$obs\_ci\_upper$
0	34842.807319	271.666459	34310.220826	35375.393812	9723.677232	59961.937406
1	34842.807319	271.666459	34310.220826	35375.393812	9723.677232	59961.937406
2	34842.807319	271.666459	34310.220826	35375.393812	9723.677232	59961.937406
3	8866.245277	316.580850	8245.606701	9486.883853	-16254.905974	33987.396528
4	47831.088340	468.949360	46911.740050	48750.436631	22700.782946	72961.393735
	•••	•••		•••	•••	•••
2667	47831.088340	468.949360	46911.740050	48750.436631	22700.782946	72961.393735
2668	34842.807319	271.666459	34310.220826	35375.393812	9723.677232	59961.937406
2669	8866.245277	316.580850	8245.606701	9486.883853	-16254.905974	33987.396528
2670	21854.526298	184.135754	21493.538727	22215.513869	-3261.551421	46970.604017
2671	21854.526298	184.135754	21493.538727	22215.513869	-3261.551421	46970.604017

Show the regression line predicting car price based on engine size for test data. Also show the confidence and prediction intervals for the car price.

```
interval_table = intervals.summary_frame(alpha=0.05)

sns.scatterplot(x = testf.engineSize, y = pred_price,color = 'orange', s = 10)
sns.lineplot(x = testf.engineSize, y = pred_price, color = 'red')
sns.lineplot(x = testf.engineSize, y = interval_table.mean_ci_lower, color = 'blue')
sns.lineplot(x = testf.engineSize, y = interval_table.mean_ci_upper, color = 'blue',label=
sns.lineplot(x = testf.engineSize, y = interval_table.obs_ci_lower, color = 'green')
sns.lineplot(x = testf.engineSize, y = interval_table.obs_ci_upper, color = 'green')
plt.legend(labels=["Regression line", "Confidence interval", "Prediction interval"])
```

<matplotlib.legend.Legend at 0x26a3a32c550>



## 2 Multiple Linear Regression

```
import pandas as pd
import numpy as np
import statsmodels.formula.api as smf
import seaborn as sns
import matplotlib.pyplot as plt
```

Develop a multiple linear regression model that predicts car price based on engine size, year, mileage, and mpg. Datasets to be used: Car\_features\_train.csv, Car\_prices\_train.csv

```
trainf = pd.read_csv('./Datasets/Car_features_train.csv')
trainp = pd.read_csv('./Datasets/Car_prices_train.csv')
train = pd.merge(trainf,trainp)
train.head()
```

	carID	brand	model	year	transmission	mileage	fuelType	tax	mpg	engineSize	price
0	18473	bmw	6 Series	2020	Semi-Auto	11	Diesel	145	53.3282	3.0	37980
1	15064	bmw	6 Series	2019	Semi-Auto	10813	Diesel	145	53.0430	3.0	33980
2	18268	bmw	6 Series	2020	Semi-Auto	6	Diesel	145	53.4379	3.0	36850
3	18480	bmw	6 Series	2017	Semi-Auto	18895	Diesel	145	51.5140	3.0	25998
4	18492	bmw	6 Series	2015	Automatic	62953	Diesel	160	51.4903	3.0	18990

```
#Using the ols function to create an ols object. 'ols' stands for 'Ordinary least squares'
ols_object = smf.ols(formula = 'price~year+mileage+mpg+engineSize', data = train)
model = ols_object.fit()
model.summary()
```

Table 2.2: OLS Regression Results

Dep. Variable:	price	R-squared:	0.660
Model:	OLS	Adj. R-squared:	0.660
Method:	Least Squares	F-statistic:	2410.

Table 2.2: OLS Regression Results

Date:	Tue, 27 Dec 2022	Prob (F-statistic):	0.00
Time:	01:07:25	Log-Likelihood:	-52497.
No. Observations:	4960	AIC:	1.050e + 05
Df Residuals:	4955	BIC:	1.050e + 05
Df Model:	4		
Covariance Type:	nonrobust		

	coef	$\operatorname{std}$ err	$\mathbf{t}$	P> t	[0.025]	0.975]
Intercept	-3.661e + 06	1.49e + 05	-24.593	0.000	-3.95e + 06	-3.37e + 06
year	1817.7366	73.751	24.647	0.000	1673.151	1962.322
$_{ m mileage}$	-0.1474	0.009	-16.817	0.000	-0.165	-0.130
mpg	-79.3126	9.338	-8.493	0.000	-97.620	-61.006
engine Size	1.218e + 04	189.969	64.107	0.000	1.18e + 04	1.26e + 04

Omnibus:	2450.973	Durbin-Watson:	0.541
Prob(Omnibus):	0.000	Jarque-Bera (JB):	31060.548
Skew:	2.045	Prob(JB):	0.00
Kurtosis:	14.557	Cond. No.	3.83e + 07

The model equation is: estimated car price = -3.661e6 + 1818 \* year -0.15 \* mileage - 79.31 \* mpg + 12180 \* engineSize

Predict the car price for the cars in the test dataset. Datasets to be used:  $Car\_features\_test.csv, Car\_prices\_test.csv$ 

```
testf = pd.read_csv('./Datasets/Car_features_test.csv')
testp = pd.read_csv('./Datasets/Car_prices_test.csv')

#Using the predict() function associated with the 'model' object to make predictions of capred_price = model.predict(testf)#Note that the predict() function finds the predictor 'en
```

# Make a visualization that compares the predicted car prices with the actual car prices

```
sns.scatterplot(x = testp.price, y = pred_price)
#In case of a perfect prediction, all the points must lie on the line x = y.
sns.lineplot(x = [0,testp.price.max()], y = [0,testp.price.max()],color='orange') #Plotting
```

```
plt.xlabel('Actual price')
plt.ylabel('Predicted price')
```

Text(0, 0.5, 'Predicted price')



The prediction looks better as compared to the one with simple linear regression. This is because we have four predictors to help explain the variation in car price, instead of just one in the case of simple linear regression. Also, all the predictors have a significant relationship with price as evident from their p-values. Thus, all four of them are contributing in explaining the variation. Note the higher values of R2 as compared to the one in the case of simple linear regression.

What is the RMSE of the predicted car price?

```
np.sqrt(((testp.price - pred_price)**2).mean())
```

9956.82497993548

What is the residual standard error based on the training data?

```
np.sqrt(model.mse_resid)
```

#### 9563.74782917604

```
sns.scatterplot(x = model.fittedvalues, y=model.resid,color = 'orange')
sns.lineplot(x = [pred_price.min(),pred_price.max()],y = [0,0],color = 'blue')
plt.xlabel('Predicted price')
plt.ylabel('Residual')
```

Text(0, 0.5, 'Residual')



Will the explained variation (R-squared) in car price always increase if we add a variable?

Should we keep on adding variables as long as the explained variation (R-squared) is increasing?

```
#Using the ols function to create an ols object. 'ols' stands for 'Ordinary least squares' np.random.seed(1)
```

```
train['rand_col'] = np.random.rand(train.shape[0])
ols_object = smf.ols(formula = 'price~year+mileage+mpg+engineSize+rand_col', data = train)
model = ols_object.fit()
model.summary()
```

Table 2.5: OLS Regression Results

Dep. Variable:	price	R-squared:	0.661
Model:	OLS	Adj. R-squared:	0.660
Method:	Least Squares	F-statistic:	1928.
Date:	Tue, 27 Dec 2022	Prob (F-statistic):	0.00
Time:	01:07:38	Log-Likelihood:	-52497.
No. Observations:	4960	AIC:	1.050e + 05
Df Residuals:	4954	BIC:	1.050e + 05
Df Model:	5		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[0.025]	0.975]
Intercept	-3.662e+06	1.49e + 05	-24.600	0.000	-3.95e + 06	-3.37e + 06
year	1818.1672	73.753	24.652	0.000	1673.578	1962.756
$_{ m mileage}$	-0.1474	0.009	-16.809	0.000	-0.165	-0.130
mpg	-79.2837	9.338	-8.490	0.000	-97.591	-60.976
engine Size	1.218e + 04	189.972	64.109	0.000	1.18e + 04	1.26e + 04
$rand\_col$	451.1226	471.897	0.956	0.339	-474.004	1376.249

Omnibus:	2451.728	Durbin-Watson:	0.541
Prob(Omnibus):	0.000	Jarque-Bera (JB):	31040.331
Skew:	2.046	Prob(JB):	0.00
Kurtosis:	14.552	Cond. No.	3.83e + 07

Adding a variable with random values to the model ( $rand\_col$ ) increased the explained variation (R-squared). This is because the model has one more parameter to tune to reduce the residual squared error (RSS). However, the p-value of  $rand\_col$  suggests that its coefficient is zero. Thus, using the model with  $rand\_col$  may give poorer performance on unknown data, as compared to the model without  $rand\_col$ . This implies that it is not a good idea to blindly add variables in the model to increase R-squared.

### 3 Variable interactions and transformations

```
import pandas as pd
import numpy as np
import statsmodels.formula.api as smf
import seaborn as sns
import matplotlib.pyplot as plt

trainf = pd.read_csv('./Datasets/Car_features_train.csv')
trainp = pd.read_csv('./Datasets/Car_prices_train.csv')
testf = pd.read_csv('./Datasets/Car_features_test.csv')
testp = pd.read_csv('./Datasets/Car_prices_test.csv')
train = pd.merge(trainf,trainp)
train.head()
```

	carID	brand	model	year	transmission	mileage	fuelType	tax	mpg	engineSize	price
0	18473	bmw	6 Series	2020	Semi-Auto	11	Diesel	145	53.3282	3.0	37980
1	15064	bmw	6 Series	2019	Semi-Auto	10813	Diesel	145	53.0430	3.0	33980
2	18268	bmw	6 Series	2020	Semi-Auto	6	Diesel	145	53.4379	3.0	36850
3	18480	bmw	6 Series	2017	Semi-Auto	18895	Diesel	145	51.5140	3.0	25998
4	18492	bmw	6 Series	2015	Automatic	62953	Diesel	160	51.4903	3.0	18990

Until now, we have assumed that the association between a predictor  $X_j$  and response Y does not depend on the value of other predictors. For example, the muliple linear regression model that we developed in Chapter 2 assumes that the average increase in price associated with a unit increase in engineSize is always \$12,180, regardless of the value of other predictors. However, this assumption may be incorrect.

#### 3.0.1 Variable interaction between continuous predictors

We can relax this assumption by considering another predictor, called an interaction term. Let us assume that the average increase in price associated with a one-unit increase in engineSize depends on the model year of the car. In other words, there is an interaction between engineSize and year. This interaction can be included as a predictor, which is the

product of engineSize and year. Note that there are several possible interactions that we can consider. Here the interaction between engineSize and year is just an example.

```
#Considering interaction between engineSize and year
ols_object = smf.ols(formula = 'price~year*engineSize+mileage+mpg', data = train)
model = ols_object.fit()
model.summary()
```

Table 3.2: OLS Regression Results

Dep. Variable:	price	R-squared:	0.682
Model:	OLS	Adj. R-squared:	0.681
Method:	Least Squares	F-statistic:	2121.
Date:	Tue, 24 Jan 2023	Prob (F-statistic):	0.00
Time:	15:28:11	Log-Likelihood:	-52338.
No. Observations:	4960	AIC:	1.047e + 05
Df Residuals:	4954	BIC:	1.047e + 05
Df Model:	5		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[0.025	0.975]
Intercept	5.606e + 05	2.74e + 05	2.048	0.041	2.4e + 04	1.1e + 06
year	-275.3833	135.695	-2.029	0.042	-541.405	-9.361
engineSize	-1.796e + 06	9.97e + 04	-18.019	0.000	-1.99e + 06	-1.6e + 06
year:engineSize	896.7687	49.431	18.142	0.000	799.861	993.676
mileage	-0.1525	0.008	-17.954	0.000	-0.169	-0.136
mpg	-84.3417	9.048	-9.322	0.000	-102.079	-66.604

Omnibus:	2330.413	Durbin-Watson:	0.524
Prob(Omnibus):	0.000	Jarque-Bera (JB):	29977.437
Skew:	1.908	Prob(JB):	0.00
Kurtosis:	14.423	Cond. No.	7.66e + 07

Note that the R-squared has increased as compared to the model in Chapter 2 since we added a predictor.

The model equation is:

```
price = \beta_0 + \beta_1 * year + \beta_2 * engineSize + \beta_3 * (year * engineSize) + \beta_4 * mileage + \beta_5 * mpg, \ (3.1)
```

$$price = \beta_0 + \beta_1 * year + (\beta_2 + \beta_3 * year) * engineSize + \beta_4 * mileage + \beta_5 * mpg, \quad (3.2)$$

or

$$price = \beta_0 + \beta_1 * year + \tilde{\beta} * engineSize + \beta_4 * mileage + \beta_5 * mpg, \tag{3.3}$$

Since  $\tilde{\beta}$  is a function of year, the association between engineSize and price is no longer a constant. A change in the value of year will change the association between price and engineSize.

Substituting the values of the coefficients:

```
price = 5.606e5 - 275.3833*year + (-1.796e6 + 896.7687*year)*engineSize - 0.1525*mileage - 84.3417*mpg \end{substitute} (3.4)
```

Thus, for cars launched in the year 2010, the average increase in price for one liter increase in engine size is -1.796e6 + 896.7687 \* 2010  $\approx$  \\$6,500, assuming all the other predictors are constant. However, for cars launched in the year 2020, the average increase in price for one liter increase in engine size is -1.796e6 + 896.7687\*2020  $\approx$  \\$15,500 , assuming all the other predictors are constant.

Similarly, the equation can be re-arranged as:

```
price = 5.606e5 + (-275.3833 + 896.7687*engineSize)*year - 1.796e6*engineSize - 0.1525*mileage - 84.3417*mprice = 5.606e5 + (-275.3833 + 896.7687*engineSize)*year - 1.796e6*engineSize - 0.1525*mileage - 84.3417*mprice = 5.606e5 + (-275.3833 + 896.7687*engineSize)*year - 1.796e6*engineSize - 0.1525*mileage - 84.3417*mprice = 5.606e5 + (-275.3833 + 896.7687*engineSize)*year - 1.796e6*engineSize - 0.1525*mileage - 84.3417*mprice = 5.606e5 + (-275.3833 + 896.7687*engineSize)*year - 1.796e6*engineSize - 0.1525*mileage - 84.3417*mprice = 5.606e5*mileage - 86.606*mileage - 86.606*mil
```

Thus, for cars with an engine size of 2 litres, the average increase in price for a one year newer model is -275.3833+896.7687 \* 2  $\approx$  \\$1500, assuming all the other predictors are constant. However, for cars with an engine size of 3 litres, the average increase in price for a one year newer model is -275.3833+896.7687 \* 3  $\approx$  \\$2400, assuming all the other predictors are constant.

```
#Computing the RMSE of the model with the interaction term
pred_price = model.predict(testf)
np.sqrt(((testp.price - pred_price)**2).mean())
```

9423.598872501092

Note that the RMSE reduced as compared to that of the model in Chapter 2. This is because the interaction term between engineSize and year is significant and relaxes the assumption of constant association between price and engine size, and between price and year. This added flexibility makes the model better fit the data. Caution: Too much flexibility may lead to overfitting!

Note that interaction terms corresponding to other variable pairs, and higher order interaction terms (such as those containing 3 or 4 variables) may also be significant and improve the model fit & thereby the prediction accuracy of the model.

#### 3.0.2 Including qualitative predictors in the model

Let us develop a model for predicting price based on engineSize and the qualitative predictor transmission.

```
#checking the distribution of values of transmission
train.transmission.value_counts()
```

Manual 1948 Automatic 1660 Semi-Auto 1351 Other 1

Name: transmission, dtype: int64

Note that the *Other* category of the variable *transmission* contains only a single observation, which is likely to be insufficient to train the model. We'll remove that observation from the training data. Another option may be to combine the observation in the *Other* category with the nearest category, and keep it in the data.

```
train_updated = train[train.transmission!='Other']

ols_object = smf.ols(formula = 'price~engineSize+transmission', data = train_updated)
model = ols_object.fit()
model.summary()
```

Table 3.5: OLS Regression Results

Dep. Variable:	price	R-squared:	0.459
Model:	OLS	Adj. R-squared:	0.458
Method:	Least Squares	F-statistic:	1400.

Table 3.5: OLS Regression Results

Date:	Tue, 24 Jan 2023	Prob (F-statistic):	0.00
Time:	15:28:21	Log-Likelihood:	-53644.
No. Observations:	4959	AIC:	1.073e + 05
Df Residuals:	4955	BIC:	1.073e + 05
Df Model:	3		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[0.025	0.975]
Intercept	3042.6765	661.190	4.602	0.000	1746.451	4338.902
${\bf transmission}[{\bf T.Manual}]$	-6770.6165	442.116	-15.314	0.000	-7637.360	-5903.873
transmission[T.Semi-Auto]	4994.3112	442.989	11.274	0.000	4125.857	5862.765
engineSize	1.023e + 04	247.485	41.323	0.000	9741.581	1.07e + 04

Omnibus:	1575.518	Durbin-Watson:	0.579
Prob(Omnibus):	0.000	Jarque-Bera (JB):	11006.609
Skew:	1.334	Prob(JB):	0.00
Kurtosis:	9.793	Cond. No.	11.4

Note that there is no coefficient for the *Automatic* level of the variable Transmission. If a car doesn't have *Manual* or *Semi-Automatic* transmission, then it has an *Automatic* transmission. Thus, the coefficient of *Automatic* will be redundant, and the dummy variable corresponding to *Automatic* transmission is dropped from the model.

The level of the categorical variable that is dropped from the model is called the baseline level. Here *Automatic* transmission is the baseline level. The coefficients of other levels of transission should be interpreted with respect to the baseline level.

#### **Q:** Interpret the intercept term

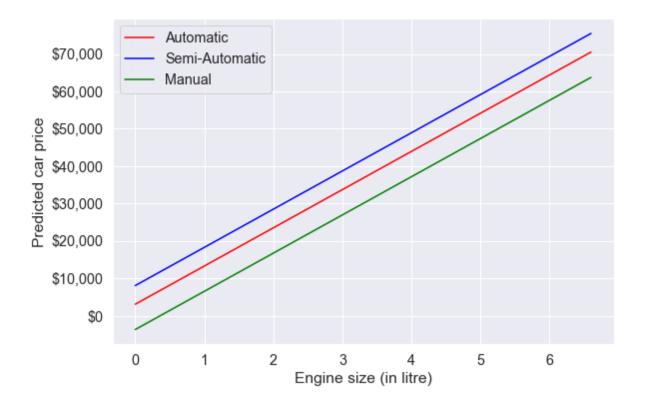
**Ans:** For the hypothetical scenario of a car with zero engine size and *Automatic* transmission, the estimated mean car price is  $\approx \$3042$ .

#### Q: Interpret the coefficient of transmission [T.Manual]

**Ans:** The estimated mean price of a car with manual transmission is  $\approx \$6770$  less than that of a car with *Automatic* transmission.

Let us visualize the developed model.

```
#Visualizing the developed model
plt.rcParams["figure.figsize"] = (9,6)
sns.set(font_scale = 1.3)
x = np.linspace(train_updated.engineSize.min(),train_updated.engineSize.max(),100)
ax = sns.lineplot(x = x, y = model.params['engineSize']*x+model.params['Intercept'], color
sns.lineplot(x = x, y = model.params['engineSize']*x+model.params['Intercept']+model.param
sns.lineplot(x = x, y = model.params['engineSize']*x+model.params['Intercept']+model.param
plt.legend(labels=["Automatic", "Semi-Automatic", "Manual"])
plt.xlabel('Engine size (in litre)')
plt.ylabel('Predicted car price')
ax.yaxis.set_major_formatter('${x:,.0f}')
```



Based on the developed model, for a given engine size, the car with a semi-automatic transmission is estimated to the most expensive on an average, while the car with a manual transmission is estimated to be the least expensive on an average.

Changing the baseline level: By default, the baseline level is chosen as the one that comes first if the levels are arranged in an alphabetical order. However, you can change the baseline level by specifying one explicitly.

Internally, statsmodels uses the patsy package to convert formulas and data to the matrices that are used in model fitting. You may refer to this section in the patsy documentation to specify a particular level of the categorical variable as the baseline.

For example, suppose we wish to change the baseline level to *Manual* transission. We can specify this in the formula as follows:

```
ols_object = smf.ols(formula = 'price~engineSize+C(transmission, Treatment("Manual"))', da
model = ols_object.fit()
model.summary()
```

Table 3.8: OLS Regression Results

Dep. Variable:	price	R-squared:	0.459
Model:	OLS	Adj. R-squared:	0.458
Method:	Least Squares	F-statistic:	1400.
Date:	Tue, 24 Jan 2023	Prob (F-statistic):	0.00
Time:	15:28:39	Log-Likelihood:	-53644.
No. Observations:	4959	AIC:	1.073e + 05
Df Residuals:	4955	BIC:	1.073e + 05
Df Model:	3		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[0.025]	0.9
Intercept	-3727.9400	492.917	-7.563	0.000	-4694.275	-27
C(transmission, Treatment("Manual"))[T.Automatic]	6770.6165	442.116	15.314	0.000	5903.873	763
C(transmission, Treatment("Manual"))[T.Semi-Auto]	1.176e + 04	473.110	24.867	0.000	1.08e + 04	1.2
engineSize	1.023e + 04	247.485	41.323	0.000	9741.581	1.0

Omnibus:	1575.518	Durbin-Watson:	0.579
Prob(Omnibus):	0.000	Jarque-Bera (JB):	11006.609
Skew:	1.334	Prob(JB):	0.00
Kurtosis:	9.793	Cond. No.	8.62

# 3.0.3 Including qualitative predictors and its interaction with continuous predictor in the model

Note that the qualitative predictor leads to fitting 3 parallel lines to the data, as there are 3 categories.

However, note that we have made the constant association assumption. The fact that the lines are parallel means that the average increase in car price for one litre increase in engine size does not depend on the type of transmission. This represents a potentially serious limitation of the model, since in fact a change in engine size may have a very different association on the price of an automatic car versus a semi-automatic or manual car.

This limitation can be addressed by adding an interaction variable, which is the product of engineSize and the dummy variables for semi-automatic and manual transmissions.

```
#Using the ols function to create an ols object. 'ols' stands for 'Ordinary least squares'
ols_object = smf.ols(formula = 'price~engineSize*transmission', data = train_updated)
model = ols_object.fit()
model.summary()
```

Table 3.11: OLS Regression Results

Dep. Variable:	price	R-squared:	0.479
Model:	OLS	Adj. R-squared:	0.478
Method:	Least Squares	F-statistic:	909.9
Date:	Sun, 22 Jan 2023	Prob (F-statistic):	0.00
Time:	22:55:55	Log-Likelihood:	-53550.
No. Observations:	4959	AIC:	1.071e + 05
Df Residuals:	4953	BIC:	1.072e + 05
Df Model:	5		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[0.025]	0.975]
Intercept	3754.7238	895.221	4.194	0.000	1999.695	5509.753
transmission[T.Manual]	1768.5856	1294.071	1.367	0.172	-768.366	4305.538
transmission[T.Semi-Auto]	-5282.7164	1416.472	-3.729	0.000	-8059.628	-2505.805
engineSize	9928.6082	354.511	28.006	0.000	9233.610	1.06e + 04
engineSize:transmission[T.Manual]	-5285.9059	646.175	-8.180	0.000	-6552.695	-4019.117
engine Size: transmission [T. Semi-Auto]	4162.2428	552.597	7.532	0.000	3078.908	5245.578

Omnibus:	1379.846	Durbin-Watson:	0.622
Prob(Omnibus):	0.000	Jarque-Bera (JB):	9799.471
Skew:	1.139	Prob(JB):	0.00
Kurtosis:	9.499	Cond. No.	30.8

The model equation for the model with interactions is:

```
Automatic transmission: price = 3754.7238 + 9928.6082 * engineSize, Semi-Automatic transmission: price = 3754.7238 + 9928.6082 * engineSize + (-5282.7164 + 4162.2428 * engineSize), Manual transmission: price = 3754.7238 + 9928.6082 * engineSize + (1768.5856 - 5285.9059 * engineSize), or Automatic transmission: price = 3754.7238 + 9928.6082 * engineSize, Semi-Automatic transmission: price = 3754.7238 + 9928.6082 * engineSize, Manual transmission: price = -1527 + 7046 * engineSize, Manual transmission: price = 5523 + 4642 * engineSize,
```

Q: Interpert the coefficient of manual transmission, i.e., the coefficient of transmission [T.Manual].

A: For a given engine size, the estimated mean price of a car with *Manual* transmission is  $\approx$  \\$1768 more than the estimated mean price of a car with *Automatic* transmission.

**Q:** Interpret the coefficient of the interaction between engine size and manual transmission, i.e., the coefficient of engineSize:transmission[T.Manual].

A: For a unit (or a litre) increase in engineSize, the increase in estimated mean price of a car with *Manual* transmission is  $\approx \$5285$  less than the increase in estimated mean price of a car with *Automatic* transmission.

```
#Visualizing the developed model with interaction terms
plt.rcParams["figure.figsize"] = (9,6)
sns.set(font_scale = 1.3)
x = np.linspace(train_updated.engineSize.min(),train_updated.engineSize.max(),100)
ax = sns.lineplot(x = x, y = model.params['engineSize']*x+model.params['Intercept'], label
plt.plot(x, (model.params['engineSize']+model.params['engineSize:transmission[T.Semi-Auto]
plt.plot(x, (model.params['engineSize']+model.params['engineSize:transmission[T.Manual]'])
plt.legend(loc='upper left')
plt.xlabel('Engine size (in litre)')
plt.ylabel('Predicted car price')
ax.yaxis.set_major_formatter('${x:,.0f}')
```



Note the interaction term adds flexibility to the model.

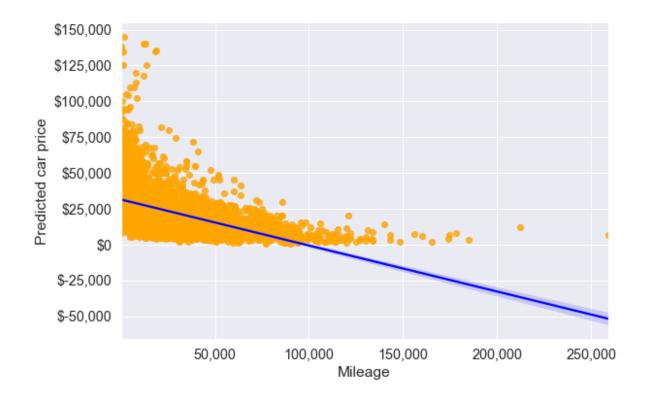
The slope of the regression line for semi-automatic cars is the largest. This suggests that increase in engine size is associated with a higher increase in car price for semi-automatic cars, as compared to other cars.

#### 3.1 Variable transformations

So far we have considered only a linear relationship between the predictors and the response. However, the relationship may be non-linear.

Consider the regression plot of price on mileage.

```
ax = sns.regplot(x = train_updated.mileage, y =train_updated.price,color = 'orange', line_
plt.xlabel('Mileage')
plt.ylabel('Predicted car price')
ax.yaxis.set_major_formatter('${x:,.0f}')
ax.xaxis.set_major_formatter('{x:,.0f}')
```



```
#R-squared of the model with just mileage
model = smf.ols('price~mileage', data = train_updated).fit()
model.rsquared
```

#### 0.22928048993376182

From the first scatterplot, we see that the relationship between price and mileage doesn't seem to be linear, as the points do not lie on a straight line. Also, we see the regression line (or the curve), which is the best fit line doesn't seem to fit the points well. However, price on average seems to decrease with mileage, albeit in a non-linear manner.

#### 3.1.1 Quadratic transformation

So, we guess that if we model price as a quadratic function of mileage, the model may better fit the points (or the curve may better fit the points). Let us transform the predictor mileage to include  $mileage^2$  (i.e., perform a quadratic transformation on the predictor).

```
#Including mileage squared as a predictor and developing the model
ols_object = smf.ols(formula = 'price~mileage+I(mileage**2)', data = train_updated)
model = ols_object.fit()
model.summary()
```

Table 3.14: OLS Regression Results

Dep. Variable:	price	R-squared:	0.271
Model:	OLS	Adj. R-squared:	0.271
Method:	Least Squares	F-statistic:	920.6
Date:	Sun, 22 Jan 2023	Prob (F-statistic):	0.00
Time:	23:26:05	Log-Likelihood:	-54382.
No. Observations:	4959	AIC:	1.088e + 05
Df Residuals:	4956	BIC:	1.088e + 05
Df Model:	2		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[0.025	0.975]
Intercept	3.44e + 04	332.710	103.382	0.000	3.37e + 04	3.5e + 04
mileage	-0.5662	0.017	-33.940	0.000	-0.599	-0.534
I(mileage ** 2)	2.629 e-06	1.56e-07	16.813	0.000	2.32e-06	2.94e-06

Omnibus:	2362.973	Durbin-Watson:	0.325
Prob(Omnibus):	0.000	Jarque-Bera (JB):	22427.952
Skew:	2.052	Prob(JB):	0.00
Kurtosis:	12.576	Cond. No.	4.81e + 09

Note that in the formula specified within the ols() function, the I() operator isolates or insulates the contents within I(...) from the regular formula operators. Without the I() operator, mileage\*\*2 will be treated as the interaction of mileage with itself, which is mileage. Thus, to add the square of mileage as a separate predictor, we need to use the I() operator.

Let us visualize the model fit with the quadratic transformation of the predictor - mileage.

```
#Visualizing the regression line with the model consisting of the quadratic transformation
pred_price = model.predict(train_updated)
ax = sns.scatterplot(x = 'mileage', y = 'price', data = train_updated, color = 'orange')
sns.lineplot(x = train_updated.mileage, y = pred_price, color = 'blue')
plt.xlabel('Mileage')
```

```
plt.ylabel('Predicted car price')
ax.yaxis.set_major_formatter('${x:,.0f}')
ax.xaxis.set_major_formatter('{x:,.0f}')
```



The above model seems to better fit the data (as compared to the model without transformation) at least upto mileage around 125,000. The  $R^2$  of the model with the quadratic transformation of mileage is also higher than that of the model without transformation indicating a better fit.

#### 3.1.2 Cubic transformation

Let us see if a cubic transformation of mileage can further improve the model fit.

```
#Including mileage squared and mileage cube as predictors and developing the model
ols_object = smf.ols(formula = 'price~mileage+I(mileage**2)+I(mileage**3)', data = train_u
model = ols_object.fit()
model.summary()
```

Table 3.17: OLS Regression Results

Dep. Variable:	price	R-squared:	0.283
Model:	OLS	Adj. R-squared:	0.283
Method:	Least Squares	F-statistic:	652.3
Date:	Sun, 22 Jan 2023	Prob (F-statistic):	0.00
Time:	23:33:27	Log-Likelihood:	-54340.
No. Observations:	4959	AIC:	1.087e + 05
Df Residuals:	4955	BIC:	1.087e + 05
Df Model:	3		
Covariance Type:	nonrobust		

	coef	$\operatorname{std}$ err	t	P> t	[0.025]	0.975]
Intercept	3.598e + 04	371.926	96.727	0.000	3.52e + 04	3.67e + 04
mileage	-0.7742	0.028	-27.634	0.000	-0.829	-0.719
I(mileage ** 2)	6.875 e-06	4.87e-07	14.119	0.000	5.92e-06	7.83e-06
I(mileage ** 3)	-1.823e-11	1.98e-12	-9.199	0.000	-2.21e-11	-1.43e-11

Omnibus:	2380.788	Durbin-Watson:	0.321
Prob(Omnibus):	0.000	Jarque-Bera (JB):	23039.307
Skew:	2.065	Prob(JB):	0.00
Kurtosis:	12.719	Cond. No.	7.73e + 14

```
#Visualizing the model with the cubic transformation of mileage
pred_price = model.predict(train_updated)
ax = sns.scatterplot(x = 'mileage', y = 'price', data = train_updated, color = 'orange')
sns.lineplot(x = train_updated.mileage, y = pred_price, color = 'blue')
plt.xlabel('Mileage')
plt.ylabel('Predicted car price')
ax.yaxis.set_major_formatter('${x:,.0f}')
ax.xaxis.set_major_formatter('{x:,.0f}')
```



Note that the model fit with cubic transformation of mileage seems slighty better as compared to the models with the quadratic transformation, and no transformation of mileage, for mileage upto 180k. However, the model should not be used to predict car prices of cars with a mileage higher than 180k.

Let's update the model created earlier (in the beginning of this chapter) to include the transformed predictor.

```
#Model with an interaction term and a variable transformation term
ols_object = smf.ols(formula = 'price~year*engineSize+mileage+mpg+I(mileage**2)', data = t
model = ols_object.fit()
model.summary()
```

Table 3.20: OLS Regression Results

Dep. Variable:	price	R-squared:	0.702
Model:	OLS	Adj. R-squared:	0.702
Method:	Least Squares	F-statistic:	1947.
Date:	Sun, 22 Jan 2023	Prob (F-statistic):	0.00
Time:	23:42:13	Log-Likelihood:	-52162.

Table 3.20: OLS Regression Results

No. Observations:	4959	AIC:	1.043e + 05
Df Residuals:	4952	BIC:	1.044e + 05
Df Model:	6		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[0.025	0.975]
Intercept	1.53e + 06	2.7e + 05	5.671	0.000	1e + 06	2.06e + 06
year	-755.7419	133.791	-5.649	0.000	-1018.031	-493.453
engineSize	-2.022e+06	9.72e + 04	-20.803	0.000	-2.21e+06	-1.83e + 06
year:engineSize	1008.6993	48.196	20.929	0.000	914.215	1103.184
mileage	-0.3548	0.014	-25.973	0.000	-0.382	-0.328
mpg	-54.7450	8.896	-6.154	0.000	-72.185	-37.305
I(mileage ** 2)	1.926e-06	1.04e-07	18.536	0.000	1.72e-06	2.13e-06

Omnibus: Prob(Omnibus):		Durbin-Watson: Jarque-Bera (JB):	0.562 38317.404
Skew:	1.857	Prob(JB):	0.00
Kurtosis:	16.101	Cond. No.	6.40e + 12

Note that the R-squared has increased as compared to the model with just the interaction term.

```
#Computing RMSE on test data
pred_price = model.predict(testf)
np.sqrt(((testp.price - pred_price)**2).mean())
```

#### 9074.494088619422

Note that the prediction accuracy of the model has further increased, as the RMSE has reduced. The transformed predictor is statistically significant and provides additional flexibility to better capture the trend in the data, leading to an increase in prediction accuracy.

## 4 Model assumptions

Note that in this course, lecture notes are not sufficient, you must read the book for better understanding. Lecture notes are just implementing the concepts of the book on a dataset, but not explaining the concepts elaborately.

Consider the model with interactions and transformation developed previously.

```
import pandas as pd
import numpy as np
import statsmodels.formula.api as smf
import seaborn as sns
import matplotlib.pyplot as plt

trainf = pd.read_csv('./Datasets/Car_features_train.csv')
trainp = pd.read_csv('./Datasets/Car_prices_train.csv')
testf = pd.read_csv('./Datasets/Car_features_test.csv')
testp = pd.read_csv('./Datasets/Car_prices_test.csv')
train = pd.merge(trainf,trainp)
train.head()
```

	carID	brand	model	year	transmission	$_{ m mileage}$	fuelType	tax	mpg	engineSize	price
0	18473	bmw	6 Series	2020	Semi-Auto	11	Diesel	145	53.3282	3.0	37980
1	15064	bmw	6 Series	2019	Semi-Auto	10813	Diesel	145	53.0430	3.0	33980
2	18268	bmw	6 Series	2020	Semi-Auto	6	Diesel	145	53.4379	3.0	36850
3	18480	bmw	6 Series	2017	Semi-Auto	18895	Diesel	145	51.5140	3.0	25998
4	18492	bmw	6 Series	2015	Automatic	62953	Diesel	160	51.4903	3.0	18990

```
ols_object = smf.ols(formula = 'price~(year+engineSize+mileage+mpg)**2+I(mileage**2)', dat
model = ols_object.fit()
model.summary()
```

Table 4.2: OLS Regression Results

Dep. Variable:	price	R-squared:	0.732
Model:	OLS	Adj. R-squared:	0.731
Method:	Least Squares	F-statistic:	1229.
Date:	Wed, $25$ Jan $2023$	Prob (F-statistic):	0.00
Time:	11:36:00	Log-Likelihood:	-51911.
No. Observations:	4960	AIC:	1.038e + 05
Df Residuals:	4948	BIC:	1.039e + 05
Df Model:	11		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[0.025	0.975]
Intercept	-1.282e+06	7.14e + 05	-1.795	0.073	-2.68e + 06	1.18e + 05
year	632.3954	353.865	1.787	0.074	-61.338	1326.128
engineSize	-1.465e + 06	1.61e + 05	-9.129	0.000	-1.78e + 06	-1.15e + 06
mileage	56.4581	3.811	14.815	0.000	48.987	63.929
mpg	-2.951e + 04	9550.775	-3.089	0.002	-4.82e+04	-1.08e + 04
year:engineSize	735.8074	79.532	9.252	0.000	579.890	891.725
year:mileage	-0.0281	0.002	-14.898	0.000	-0.032	-0.024
year:mpg	14.6915	4.731	3.105	0.002	5.417	23.966
engineSize:mileage	-0.0808	0.011	-7.143	0.000	-0.103	-0.059
engineSize:mpg	-120.5780	11.384	-10.592	0.000	-142.896	-98.260
mileage:mpg	0.0026	0.000	5.173	0.000	0.002	0.004
I(mileage ** 2)	3.495 e-07	1.56e-07	2.236	0.025	4.31e-08	6.56 e - 07

Omnibus:	1958.631	Durbin-Watson:	0.542
Prob(Omnibus):	0.000	Jarque-Bera (JB):	44560.042
Skew:	1.349	Prob(JB):	0.00
Kurtosis:	17.434	Cond. No.	1.73e + 13

```
np.sqrt(model.mse_resid)
```

#### 8502.851955843495

```
pred_price = model.predict(testf)
np.sqrt(((testp.price - pred_price)**2).mean())
```

```
#Computing MAE on test data
pred_price = model.predict(testf)
(np.abs(testp.price - pred_price)).mean()
```

5395.006622253402

Let us check if this model satisfies the assumptions of the linear regression model

#### 4.1 Non-linearity of data

We have assumed that there is a linear relationship between the predictors and the response. Residual plots, which are scatter plots of residuals vs fitted values, can be used to identify non-linearity. Fitted values are the values estimated by the model on training data, denoted by  $\hat{y}_i$ , and residuals are given by  $e_i = y_i - \hat{y}_i$ .

```
#Plotting residuals vs fitted values
plt.rcParams["figure.figsize"] = (9,6)
sns.set(font_scale=1.25)
ax = sns.scatterplot(x = model.fittedvalues, y=model.resid,color = 'orange')
sns.lineplot(x = [pred_price.min(),pred_price.max()],y = [0,0],color = 'blue')
plt.xlabel('Fitted values')
plt.ylabel('Residuals')
ax.yaxis.set_major_formatter('${x:,.0f}')
ax.xaxis.set_major_formatter('${x:,.0f}')
```



The model seems to satisfy this assumption, as we do not observe a strong pattern in the residuals around the line Residuals = 0. Residuals are distributed more or less in a similar manner on both sides of the blue line for all fitted values.

For the model to satisfy the linearity assumption perfectly, the points above the line (Residuals = 0), should be mirror age of the points below the line, i.e., the blue line in the above plot should act as a mirror.

What to do if there is non-linear association (page 94 of book): If the residual plot indicates that there are non-linear associations in the data, then a simple approach is to use non-linear transformations of the predictors, such as log X,  $\sqrt{X}$ , and  $X^2$ , in the regression model.

#### 4.2 Non-constant variance of error terms

The variance of the error terms is assumed to be constant, i.e.,  $Var(\epsilon_i) = \sigma^2$ , and this assumption is used while deriving the standard errors of the regression coefficients. The standard errors in turn are used to test the significant of the predictors, and obtain their confidence interval. Thus, violation of this assumption may lead to incorrect inference. Non-constant variance of error terms, or violation of the constant variance assumption, is called *heteroscedasticity*.

This assumption can be checked by plotting the residuals against fitted values.

```
#Plotting residuals vs fitted values
ax = sns.scatterplot(x = model.fittedvalues, y=model.resid,color = 'orange')
sns.lineplot(x = [pred_price.min(),pred_price.max()],y = [0,0],color = 'blue')
plt.xlabel('Fitted values')
plt.ylabel('Residuals')
ax.yaxis.set_major_formatter('${x:,.0f}')
ax.xaxis.set_major_formatter('${x:,.0f}')
```



We see that the variance of errors seems to increase with increase in the fitted values. In such a case a log transformation of the response can resolve the issue to some extent. This is because a log transform will result in a higher shrinkage of larger values.

```
#Model with an interaction term and a variable transformation term
ols_object = smf.ols(formula = 'np.log(price)~(year+engineSize+mileage+mpg)**2+I(mileage**
model_log = ols_object.fit()
model_log.summary()
```

Table 4.5: OLS Regression Results

Dep. Variable:	np.log(price)	R-squared:	0.803
Model:	OLS	Adj. R-squared:	0.803
Method:	Least Squares	F-statistic:	1834.
Date:	Wed, $25$ Jan $2023$	Prob (F-statistic):	0.00
Time:	11:37:55	Log-Likelihood:	-1173.8
No. Observations:	4960	AIC:	2372.
Df Residuals:	4948	BIC:	2450.
Df Model:	11		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[0.025	0.975]
Intercept	-238.2125	25.790	-9.237	0.000	-288.773	-187.652
year	0.1227	0.013	9.608	0.000	0.098	0.148
engineSize	13.8349	5.795	2.387	0.017	2.475	25.195
mileage	0.0005	0.000	3.837	0.000	0.000	0.001
mpg	-1.2446	0.345	-3.610	0.000	-1.921	-0.569
year:engineSize	-0.0067	0.003	-2.324	0.020	-0.012	-0.001
year:mileage	-2.67e-07	6.8e-08	-3.923	0.000	-4e-07	-1.34e-07
year:mpg	0.0006	0.000	3.591	0.000	0.000	0.001
engineSize:mileage	-2.668e-07	4.08e-07	-0.654	0.513	-1.07e-06	5.33e-07
engineSize:mpg	0.0028	0.000	6.842	0.000	0.002	0.004
mileage:mpg	7.235e-08	1.79e-08	4.036	0.000	3.72e-08	1.08e-07
I(mileage ** 2)	1.828e-11	5.64e-12	3.240	0.001	7.22e-12	2.93e-11

Omnibus:	711.515	Durbin-Watson:	0.498
Prob(Omnibus):	0.000	Jarque-Bera (JB):	2545.807
Skew:	0.699	Prob(JB):	0.00
Kurtosis:	6.220	Cond. No.	1.73e + 13

Note that the coefficient of year turns out to be significant (at 5% significance level), unlike in the previous model. Intuitively, the coefficient of year should have been significant, as year has the highest linear correlation of 50% with car price.

Although the R-squared has increased as compared to the previous model, violation of this assumption does not cause bias in the regression coefficients. Thus, there may not be a large improvement in the model fit, unless we add predictor(s) to address heteroscedasticity.

Let us check the constant variance assumption again.

```
#Plotting residuals vs fitted values
sns.scatterplot(x = (model_log.fittedvalues), y=(model_log.resid),color = 'orange')
sns.lineplot(x = [model_log.fittedvalues.min(),model_log.fittedvalues.max()],y = [0,0],col
plt.xlabel('Fitted values')
plt.ylabel('Residuals')
```

Text(0, 0.5, 'Residuals')



Now we observe that the constant variance assumption is satisfied. Let us see the RMSE of this model on test data.

```
#Computing RMSE on test data
pred_price_log = model_log.predict(testf)
np.sqrt(((testp.price - np.exp(pred_price_log))**2).mean())
```

9094.209503063496

Note that the RMSE of the log-transformed model has increased as compared to the model without transformation. Does it mean the log-transformed model is less accurate?

```
#Computing MAE on test data
pred_price_log = model_log.predict(testf)
((np.abs(testp.price - np.exp(pred_price_log))).mean())
```

#### 5268.398904745121

Although the RMSE has increased a bit for the log-transformed model, the MAE has reduced. This means the log-transformed model does a bit worse on reducing relatively large errors, but does better in reducing the absolute errors on an average.

```
#Comparing errors of the log-transformed model with the previous model
err = np.abs(testp.price - pred_price)
err_log = np.abs(testp.price - np.exp(pred_price_log))
sns.scatterplot(x = err,y = err_log, color = 'orange')
sns.lineplot(x = [0,100000], y = [0,100000], color = 'blue')
np.sum(err_log<err)/len(err)</pre>
```

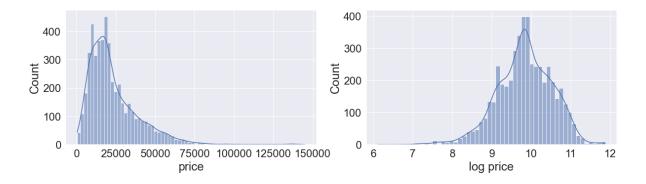
### 0.5572604790419161



For 56% of the cars, the log transformed makes a more accurate prediction than the previous model, which is another criterion based on which the log-transformed model is more accurate. However, the conclusion based on RMSE is different. This is because RMSE can be influenced by a few large errors. Thus, RMSE, though sometimes appropriate than other criteria, should not be used as the sole measure to compare the accuracy of models.

```
#Visualizing the distribution of price and log(price)
fig = plt.figure()
fig.subplots_adjust(hspace=0.4, wspace=0.2)
sns.set(rc = {'figure.figsize':(20,12)})
sns.set(font_scale = 2)
ax = fig.add_subplot(2, 2, 1)
sns.histplot(train.price,kde=True)
ax.set(xlabel='price', ylabel='Count')
ax = fig.add_subplot(2, 2, 2)
sns.histplot(np.log(train.price),kde=True)
ax.set(xlabel='log price', ylabel='Count')
```

[Text(0.5, 0, 'log price'), Text(0, 0.5, 'Count')]



We can see that the log transformation shrinked the higher values of price, making its distribution closer to normal.

Note that heterscedasticity can also occur due to model misspecification, i.e., in case of missing predictor(s). Some of the cars are too expensive, which makes the *price* distribution skewed. Perhaps, the price of expensive cars be better explained by the car *model*, a predictor that is missing in the current model.

# A Assignment A

- 1. You may talk to a friend, discuss the questions and potential directions for solving them. However, you need to write your own solutions and code separately, and not as a group activity.
- 2. Do not write your name on the assignment.
- 3. Write your code in the *Code* cells and your answer in the *Markdown* cells of the Jupyter notebook. Ensure that the solution is written neatly enough to understand and grade.
- 4. Use Quarto to print the *.ipynb* file as HTML. You will need to open the command prompt, navigate to the directory containing the file, and use the command: quarto render filename.ipynb --to html. Submit the HTML file.
- 5. The assignment is worth 100 points, and is due on **Tuesday**, 17th January 2023 at 11:59 pm.
- 6. There is a **bonus** question worth 5 points.
- 7. **Five points are for properly formatting the assignment**. The breakdown is as follows:
- Must be an HTML file rendered using Quarto (1 pt); If you have a Quarto issue, you must mention the issue & quote the error you get when rendering using Quarto in the comments section of Canvas, and submit the ipynb file.
- No name can be written on the assignment, nor can there be any indicator of the student's identity—e.g., printouts of the working directory should not be included in the final submission (1 pt).
- There aren't excessively long outputs of extraneous information (e.g. no printouts of entire data frames without good reason, there aren't long printouts of which iteration a loop is on, there aren't long sections of commented-out code, etc.) (1 pt).
- Final answers of each question are written in Markdown cells (1 pt).
- There is no piece of unnecessary / redundant code, and no unnecessary / redundant text (1 pt).
- 8. The maximum possible score in the assignment is 95 + 5 (formatting) + 5 (bonus question) = 105 out of 100. There is no partial credit for the bonus question.

### A.1 Regression vs Classification; Prediction vs Inference

Explain (1) whether each scenario is a classification or regression problem, and (2) whether we are most interested in inference or prediction. Answers to both parts must be supported by a justification.

### A.1.1

Consider a company that is interested in conducting a marketing campaign. The goal is to identify individuals who are likely to respond positively to a marketing campaign, based on observations of demographic variables (such as age, gender, income, etc.) measured on each individual.

(2+2 points)

### A.1.2

Consider that the company mentioned in the previous question is interested in understanding the impact of advertising promotions in different media types on the company sales. For example, the company is interested in the question, 'how large of an increase in sales is associated with a given increase in radio vis-a-vis TV advertising?'

(2+2 points)

### A.1.3

Consider a company selling furniture is interested in the finding the association between demographic characterisits of customers (such as age, gender, income, etc.) and their probability of purchase of a particular company product.

(2+2 points)

### A.1.4

We are interested in predicting the % change in the USD/Euro exchange rate in relation to the weekly changes in the world stock markets. Hence we collect weekly data for all of 2022. For each week we record the % change in the USD/Euro, the % change in the US market, the % change in the British market, and the % change in the German market.

(2+2 points)

### A.2 RMSE vs MAE

### A.2.1

Describe a regression problem, where it will be more appropriate to assess the model accuracy using the root mean squared error (RMSE) metric as compared to the mean absolute error (MAE) metric.

**Note:** Don't use the examples presented in class

(4 points)

### A.2.2

Describe a regression problem, where it will be more appropriate to assess the model accuracy using the mean absolute error (MAE) metric as compared to the root mean squared error (RMSE) metric.

**Note:** Don't use the examples presented in class

(4 points)

### A.3 FNR vs FPR

### A.3.1

A classification model is developed to predict those customers who will respond positively to a company's tele-marketing campaign. All those customers that are predicted to respond positively to the campaign will be called by phone to buy the product being marketed. If the customer being called purchases the product (y=1), the company will get a profit of \$100. On the other hand, if they are called and they don't purchase (y=0), the company will have a loss of \$1. Among FPR (False positive rate) and FNR (False negative rate), which metric is more important to be minimized to reduce the loss associated with misclassification? Justify your answer.

In your justification, you must clearly interpret False Negatives (FN) and False Postives (FP) first.

**Assumption:** Assume that based on the past marketing campaigns, around 50% of the customers will actually respond positively to the campaign.

(4 points)

### A.3.2

Can the answer to the previous question change if the assumption stated in the question is false? Justify your answer.

(6 points)

### A.4 Petrol consumption

Read the dataset petrol\_consumption\_train.csv. It contains the following five columns:

Petrol\_tax: Petrol tax (cents per gallon)

Per\_capita\_income: Average income (dollars)

Paved\_highways: Paved Highways (miles)

Prop\_license: Proportion of population with driver's licenses

Petrol\_consumption: Consumption of petrol (millions of gallons)

### A.4.1

Make a pairwise plot of all the variables in the dataset. Which variable seems to have the highest linear correlation with Petrol\_consumption? Let this variable be predictor P. Note: If you cannot figure out P by looking at the visualization, you may find the pairwise linear correlation coefficient to identify P.

(4 points)

### A.4.2

Fit a simple linear regression model to predict  $Petrol\_consumption$  based on predictor P (identified in the previous part). Print the model summary.

(4 points)

### A.4.3

Interpret the coefficient of P. What is the increase in petrol consumption for an increase of 0.05 in P?

(2+2 points)

### A.4.4

Does petrol consumption have a statistically significant relationship with the predictor P? Justify your answer.

```
(4 points)
```

### A.4.5

What is the R-squared? Interpret its value.

```
(4 points)
```

### A.4.6

Use the model developed above to estimate the petrol consumption for a state in which 50% of the population has a driver's license. What are the confidence and prediction intervals for your estimate? Which interval includes the irreducible error?

```
(4+3+3+2 = 12 points)
```

### A.4.7

Use the model developed above to estimate the petrol consumption for a state in which 10% of the population has a driver's license. Are you getting a reasonable estimate? Why or why not?

```
(5 points)
```

### A.4.8

What is the residual standard error of the model?

```
(4 points)
```

### A.4.9

Using the model developed above, predict the petrol consumption for the observations in petrol\_consumption\_test.csv. Find the RMSE (Root mean squared error). Include the units of RMSE in your answer.

```
(5 points)
```

### A.4.10

Based on the answers to the previous two questions, do you think the model is overfitting? Justify your answer.

```
(4 points)
```

Make a scatterplot of Petrol\_consumption vs Prop\_license using petrol\_consumption\_test.csv. Over the scatterplot, plot the regression line, the prediction interval, and the confidence interval. Distinguish the regression line, prediction interval lines, and confidence interval lines with the following colors. Include the legend as well.

• Regression line: red

Confidence interval lines: bluePrediction interval lines: green

(4 points)

Among the confidence and prediction intervals, which interval is wider, and why?

(1+2 points)

### A.4.11

Find the correlation between Petrol\_consumption and the rest of the variables in petrol\_consumption\_train.csv. Based on the correlations, a simple linear regression model with which predictor will have the least R-squared value for predicting Petrol\_consumption. Don't develop any linear regression models.

(4 points)

### Bonus point question

(5 points - no partial credit)

### A.4.12

Fit a simple linear regression model to predict  $Petrol\_consumption$  based on predictor P, but without an intercept term.

(you must answer this correctly to qualify for earning bonus points)

### A.4.13

Estimate the petrol consumption for the observations in *petrol\_consumption\_test.csv* using the model in developed in the previous question. Find the RMSE.

(you must answer this correctly to qualify for earning bonus points)

### A.4.14

The RMSE for the models with and without the intercept are similar, which indicates that both models are almost equally good. However, the R-squared for the model without intercept is much higher than the R-squared for the model with the intercept. Why? Justify your answer.

(5 points)

## B Assignment B

- 1. You may talk to a friend, discuss the questions and potential directions for solving them. However, you need to write your own solutions and code separately, and not as a group activity.
- 2. Do not write your name on the assignment.
- 3. Write your code in the *Code* cells and your answer in the *Markdown* cells of the Jupyter notebook. Ensure that the solution is written neatly enough to understand and grade.
- 4. Use Quarto to print the *.ipynb* file as HTML. You will need to open the command prompt, navigate to the directory containing the file, and use the command: quarto render filename.ipynb --to html. Submit the HTML file.
- 5. The assignment is worth 100 points, and is due on Thursday, 26th January 2023 at 11:59 pm.
- 6. Five points are properly formatting the assignment. The breakdown is as follows:
- Must be an HTML file rendered using Quarto (1 pt). If you have a Quarto issue, you must mention the issue & quote the error you get when rendering using Quarto in the comments section of Canvas, and submit the ipynb file.
- No name can be written on the assignment, nor can there be any indicator of the student's identity—e.g. printouts of the working directory should not be included in the final submission (1 pt)
- There aren't excessively long outputs of extraneous information (e.g. no printouts of entire data frames without good reason, there aren't long printouts of which iteration a loop is on, there aren't long sections of commented-out code, etc.) (1 pt)
- Final answers of each question are written in Markdown cells (1 pt).
- There is no piece of unnecessary / redundant code, and no unnecessary / redundant text (1 pt)

### **B.1** Multiple linear regression

A study was conducted on 97 men with prostate cancer who were due to receive a radical prostatectomy. The dataset *prostate.csv* contains data on 9 measurements made on these 97 men. The description of variables can be found here:

### **B.1.1 Training MLR**

Fit a linear regression model with lpsa as the response and all the other variables as predictors. Write down the equation to predict lpsa based on the other eight variables.

(2+2 points)

### **B.1.2** Model significance

Is the overall regression significant at 5% level? Justify your answer.

(2 points)

### **B.1.3 Coefficient interpretation**

Interpret the coefficient of svi.

(2 points)

### **B.1.4 Variable significance**

Report the *p*-values for gleason and age. What do you conclude about the significance of these variables?

(2+2 points)

### B.1.5 Variable significance from confidence interval

What is the 95% confidence interval for the coefficient of age? Can you conclude anything about its significance based on the confidence interval?

(2+2 points)

### **B.1.6** *p*-value

Fit a simple linear regression on lpsa against gleason. What is the *p*-value for gleason?

(1+1 points)

### B.1.7 Predictor significance in presence / absence of other predictors

Is the predictor gleason statistically significant in the model developed in the previous question (B.1.6)?

Was gleason statistically significant in the model developed in the first question (B.1.1) with multiple predictors?

Did the statistical significance of gleason change in the absence of other predictors? Why or why not?

(1+1+4 points)

### **B.1.8 Prediction**

Predict lpsa of a 65-year old man with lcavol = 1.35, lweight = 3.65, lbph = 0.1, svi = 0.22, lcp = -0.18, gleason = 6.75, and pgg45 = 25 and find 95% prediction intervals.

(2 points)

### **B.1.9 Variable selection**

Find the largest subset of predictors in the model developed in the first question (B.1.1), such that their coefficients are zero, i.e., none of the predictors in the subset are statistically significant.

Does the model R-squared change a lot if you remove the set of predictors identified above from the model in the first question (B.1.1)?

Hint: You may use the f\_test() method to test hypotheses.
(4+1 points)

### **B.2 Using MLR coefficients and variable transformation**

The dataset *infmort.csv* gives the infant mortality of different countries in the world. The column mortality contains the infant mortality in deaths per 1000 births.

### **B.2.1** Data visualisation

Make the following plots:

- 1. a boxplot of log(mortality) against region (note that a plot of log(mortality) against region better distinguishes the mortality among regions as compared to a plot of mortality against region,
- 2. a boxplot of income against region, and
- 3. a scatter plot of mortality against income.

What trends do you see in these plots? Mention the trend separately for each plot. (3+2 points)

### B.2.2 Removing effect of predictor from response

Europe seems to have the lowest infant mortality, but it also has the highest per capita annual income. We want to see if Europe still has the lowest mortality if we remove the effect of income from the mortality. We will answer this question with the following steps.

### **B.2.2.1** Variable transformation

Plot:

- 1. mortality against income,
- 2. log(mortality) against income,
- 3. mortality against log(income), and
- 4. log(mortality) against log(income).

Based on the plots, postulate an appropriate model to predict mortality as a function of income. *Print the model summary.* 

(2+4 points)

### B.2.2.2 Model update

Update the model developed in the previous question by adding **region** as a predictor. Print the model summary.

(2 points)

Use the model developed in the previous question to compute adjusted\_mortality for each observation in the data, where adjusted mortality is the mortality after removing the estimated effect of income. Make a boxplot of log(adjusted\_mortality) against region.

(4+2 points)

### B.2.3 Data visualisation after removing effect of predictor from response

From the plot in the previous question:

- 1. Does Europe still seem to have the lowest mortality as compared to other regions after removing the effect of income from mortality?
- 2. After adjusting for income, is there any change in the mortality comparison among different regions. Compare the plot developed in the previous question to the plot of log(mortality) against region developed earlier (B.2.1) to answer this question.

**Hint:** Do any African / Asian / American countries seem to do better than all the European countries with regard to mortality after adjusting for income?

(1+3 points)

### **B.3** Variable transformations and interactions

The dataset  $soc\_ind.csv$  contains the GDP per capita of some countries along with several social indicators.

### **B.3.1 Training SLR**

For a simple linear regression model predicting gdpPerCapita. Which predictor will provide the best model fit (ignore categorical predictors)? Let that predictor be P.

(2 points)

### **B.3.2** Linearity in relationship

Make a scatterplot of gdpPerCapita vs P. Does the relationship between gdpPerCapita and P seem linear or non-linear?

```
(1 + 2 points)
```

### **B.3.3 Variable transformation**

If the relationship identified in the previous question is non-linear, identify and include transformation(s) of the predictor P in the model to improve the model fit.

Mention the predictors of the transformed model, and report the change in the R-squared value of the transformed model as compared to the simple linear regression model with only P.

(4+4 points)

### B.3.4 Model visualisation with transformed predictor

Plot the regression curve of the transformed model (developed in the previous question) over the scatterplot in (b) to visualize model fit. Also make the regression line of the simple linear regression model with only P on the same plot.

(3 + 1 points)

### B.3.5 Training MLR with qualitative predictor

Develop a model to predict gdpPerCapita with P and continent as predictors.

- 1. Interpert the intercept term.
- 2. For a given value of P, are there any continents that **do not** have a signficant difference between their mean <code>gdpPerCapita</code> and that of Africa? If yes, then which ones, and why? If no, then why not? Consider a significance level of 5%.

(4 + 4 points)

### **B.3.6** Variable interaction

The model developed in the previous question has a limitation. It assumes that the increase in mean gdpPerCapita with a unit increase in P does not depend on the continent.

- 1. Eliminate this limitation by including interaction of **continent** with P in the model developed in the previous question. Print the model summary of the model with interactions.
- 2. Interpret the coefficient of any one of the interaction terms.

```
(4 + 4 points)
```

### B.3.7 Model visualisation with qualitative predictor

Use the model developed in the previous question to plot the regression lines for Africa, Asia, and Europe. Put gdpPerCapita on the vertical axis and P on the horizontal axis. Use a legend to distinguish among the regression lines of the three continents.

(4 points)

### **B.3.8 Model interpretation**

Based on the plot develop in the previous question, which continent has the highest increase in mean gdpPerCapita for a unit increase in P, and which one has the least? Justify your answer.

(2+2 points)

# C Datasets, assignment and project files

Datasets used in the book, assignment files, and project files can be found here

# References