hw06

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8.1

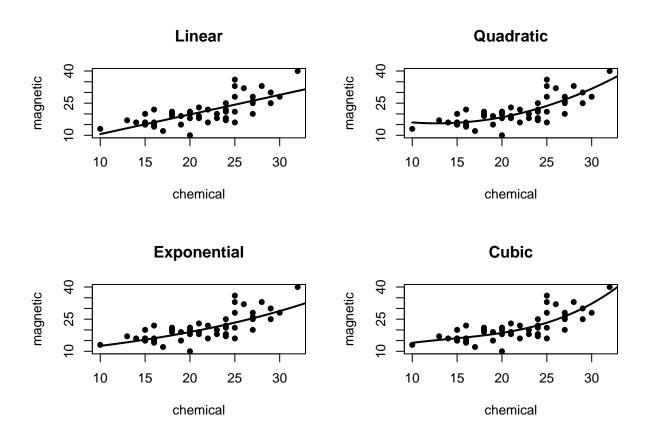
```
library(bootstrap)
compute_correlation_stats <- function(data, x_col, y_col) {</pre>
  n <- nrow(data)</pre>
  theta.hat <- cor(data[[x_col]], data[[y_col]])</pre>
  theta.jack <- sapply(1:n, function(i) cor(data[-i, x_col], data[-i, y_col]))</pre>
  bias <- (n - 1) * (mean(theta.jack) - theta.hat)</pre>
  se <- sqrt((n - 1) * mean((theta.jack - mean(theta.jack))^2))</pre>
  list(est = theta.hat, bias = bias, se = se)
}
data("law")
results <- compute_correlation_stats(law, "LSAT", "GPA")</pre>
print(results)
## $est
## [1] 0.7763745
##
## $bias
## [1] -0.006473623
##
## $se
## [1] 0.1425186
rm(law)
detach(package:bootstrap)
8.3
library(boot)
library(bootstrap)
data("law")
attach(law)
```

```
cor.stat <- function(data, indices = 1:nrow(data)) {</pre>
  cor(data[indices, 1], data[indices, 2])
# 定义 Bootstrap t 区间计算函数
boot.t.ci <- function(data, B = 500, R = 100, level = 0.95, statistic) {
  # 将数据转换为矩阵
  data matrix <- as.matrix(data)</pre>
 n <- nrow(data_matrix)</pre>
  # 初始化统计量和标准误向量
  stats <- numeric(B)</pre>
  ses <- numeric(B)</pre>
  # 定义 Bootstrap 标准误计算函数
  boot.se <- function(data, R, statistic) {</pre>
    m <- nrow(data)</pre>
    bootstrap_samples <- replicate(R, {</pre>
      indices <- sample(1:m, size = m, replace = TRUE)</pre>
      statistic(data[indices, ])
    })
    sd(bootstrap_samples)
  # 进行 Bootstrap 抽样
  for (b in 1:B) {
    indices <- sample(1:n, size = n, replace = TRUE)</pre>
    bootstrap_sample <- data_matrix[indices, ]</pre>
    stats[b] <- statistic(bootstrap_sample)</pre>
    ses[b] <- boot.se(bootstrap_sample, R = R, statistic)</pre>
  # 计算原始统计量
  original_stat <- statistic(data_matrix)</pre>
  # 计算 t 统计量
  t_stats <- (stats - original_stat) / ses
  # 计算标准误
  se0 <- sd(stats)
  # 计算置信水平
  alpha <- 1 - level
  # 计算 t 分布的分位数
  q_t <- quantile(t_stats, c(alpha / 2, 1 - alpha / 2), type = 1)</pre>
  names(q_t) <- rev(names(q_t))</pre>
  # 计算置信区间
  ci <- rev(original_stat - q_t * se0)</pre>
  return(ci)
}
```

```
result <- boot.t.ci(law, B = 1000, R = 25, statistic = cor.stat)</pre>
print(result)
         2.5%
                   97.5%
## -0.2594146 0.9782805
detach(law)
detach(package:bootstrap)
detach(package:boot)
8.5
library(boot)
x <- aircondit[1]</pre>
meant <- function(x, i) return(mean(as.matrix(x[i, ])))</pre>
b <- boot(x, statistic = meant, R = 1000)</pre>
## ORDINARY NONPARAMETRIC BOOTSTRAP
##
##
## Call:
## boot(data = x, statistic = meant, R = 1000)
## Bootstrap Statistics :
       original bias std. error
## t1* 108.0833 0.17775 37.49542
boot.ci(b, type = c("norm", "perc", "basic"))
## BOOTSTRAP CONFIDENCE INTERVAL CALCULATIONS
## Based on 1000 bootstrap replicates
##
## boot.ci(boot.out = b, type = c("norm", "perc", "basic"))
##
## Intervals :
## Level
              Normal
                                   Basic
                                                      Percentile
       (34.4, 181.4) (27.1, 168.9) (47.2, 189.1)
## Calculations and Intervals on Original Scale
detach(package:boot)
8.10
library(DAAG, warn.conflict = FALSE)
```

Warning: 程序包'DAAG'是用R版本4.4.2 来建造的

```
attach(ironslag)
a \leftarrow seq(10, 40, 0.1)
par(mfrow = c(2, 2))
L1 <- lm(magnetic ~ chemical)
plot(chemical, magnetic, main = "Linear", pch = 16)
yhat1 < -L1$coef[1] + L1$coef[2] *a
lines(a,yhat1,lwd=2)
L2<-lm(magnetic~chemical+I(chemical^2))</pre>
plot(chemical, magnetic, main="Quadratic", pch=16)
yhat2 < -L2\$coef[1] + L2\$coef[2]*a + L2\$coef[3]*a^2
lines(a,yhat2,lwd=2)
L3<-lm(log(magnetic)~chemical)
plot(chemical, magnetic, main="Exponential", pch=16)
logyhat3 < -L3$coef[1] + L3$coef[2]*a
yhat3<-exp(logyhat3)</pre>
lines(a,yhat3,lwd=2)
c2<-chemical^2
c3<-chemical^3
L4<-lm(magnetic~chemical+c2+c3)
plot(chemical, magnetic, main="Cubic", pch=16)
yhat4 < -L4$coef[1] + L4$coef[2] *a + L4$coef[3] *a^2 + L4$coef[4] *a^3
lines(a,yhat4,lwd=2)
```



```
par(mfrow=c(1,1))
Rsq<-numeric(4)
Rsq[1] <- summary(L1) $adj.r.squared
Rsq[2] <- summary(L2) $adj.r.squared
Rsq[3] <- summary(L3) $adj.r.squared
Rsq[4] <- summary(L4) $adj.r.squared
Rsq
## [1] 0.5281545 0.5768151 0.5280556 0.5740396
8.11
library(DAAG, warn.conflict = FALSE)
attach(ironslag)
## The following objects are masked from ironslag (pos = 3):
##
##
        chemical, magnetic
n <- length(magnetic)</pre>
N <- choose(n, 2)
e1 <- e2 <- e3 <- e4 <- e5 <- numeric(N)
ij<-1
for (i in 1:(n-1)) for (j in (i + 1):n) {
k \leftarrow c(i, j)
y <- magnetic[-k]
x <- chemical[-k]</pre>
 J1 \leftarrow lm(y \sim x)
yhat1 <- J1$coef[1] + J1$coef[2] * chemical[k]</pre>
 e1[ij] <- sum((magnetic[k] - yhat1)^2)
 J2 \leftarrow lm(y \sim x + I(x^2))
yhat2 <- J2$coef[1] + J2$coef[2] * chemical[k] +</pre>
 J2$coef[3] * chemical[k]^2
 e2[ij] <- sum((magnetic[k] - yhat2)^2)
 J3 \leftarrow lm(log(y) \sim x)
logyhat3 <- J3$coef[1] + J3$coef[2] * chemical[k]</pre>
 yhat3 <- exp(logyhat3)</pre>
 e3[ij] <- sum((magnetic[k] - yhat3)^2)
 J4 \leftarrow lm(log(y) \sim log(x))
 logyhat4 <- J4$coef[1] + J4$coef[2] * log(chemical[k])</pre>
 yhat4 <- exp(logyhat4)</pre>
 e4[ij] <- sum((magnetic[k]- yhat4)^2)
 c2 <- x^2
 c3 <- x<sup>3</sup>
 J5 \leftarrow lm(y \sim x + c2 + c3)
yhat5 <- J5$coef[1] + J5$coef[2] * chemical[k] +</pre>
 J5\$coef[3] * chemical[k]^2 + J5\$coef[4] * chemical[k]^3
 e5[ij] <- sum((magnetic[k] - yhat5)^2)
 ij<-ij+1
c(sum(e1), sum(e2), sum(e3), sum(e4), sum(e5))/N
```

[1] 39.14455 35.74037 36.90983 40.93436 36.46996

detach(ironslag)
detach(package:DAAG)