

The paper I am going to present is “Trends in Extreme Value Indices”. This paper is published on Journal of the American Statistical Association in 2020. It is written by Laurens de Haan and Chen Zhou.

This is the outline of My presentation. First, I will give a introduction to this study. And then I will present the main methodology. And Finally, a simulation study and a real data application will be presented.

Now, we turn to the first part.

Classic extreme value analysis assumes that the observations are independent and identically distributed. Recent studies aim at dealing with case when observations are drawn from different distribution. This paper considers a continuously changing extreme value index and try to estimate the functional extreme value index accurately.

Mathematically, consider a set of distributions $F_s(x)$ for $s \in [0, 1]$. Now we have independent random variables X_1 to X_n and the distribution function of X_i is $F_{i/n}$.

Here, assume that the function $F_s(x)$ is continuous with respect to s and x . Also, assume that F_s belongs to the maximum domain of attraction with extreme value index $\gamma(s)$.

This paper considers the case that the function γ is positive. This means that for each s , the function F_s is a Frechet distribution. Also, the function γ is assumed to be continuous.

The goal of this paper is to estimate the function γ and test the hypothesis that $\gamma = \gamma_0$ for some given function γ_0 . In particular, it can be applied to test whether the extreme value index remains at a constant level across all observations.

We first discuss how to estimate the function γ locally. And then I will present how to obtain a global estimator and how to use this to do the hypothesis testing that $\gamma = \gamma_0$ for some given function γ_0 .

The idea for estimating $\gamma(s)$ locally is similar to the kernel density estimation. More specifically, use only observations X_i in the h -neighborhood of s . And the mathematical definition of the h -neighborhood is displayed as this, where h is bandwidth and h satisfies that as n goes to infinity, h will go to infinity and n times h will go to infinity.

Correspondingly there will be $[2nh]$ observations in the h -neighborhood for each $s \in [h, 1 - h]$. To apply any known estimators for extreme value index, such as the Hill estimator, we choose an intermediate sequence $k = k(n)$ such that as $n \rightarrow \infty, k \rightarrow \infty$ and $k/n \rightarrow 0$. This choice of k is very standard in the extreme value analysis.

Then, we will use the top $[2kh]$ order statistics among the $[2nh]$ local observation in the h -neighborhood of s to estimate $\gamma(s)$. Rank the $[2nh]$ observations in the h -neighborhood of s as this. Then we can apply the Hill estimator to estimate

$\gamma(s)$ locally. And the mathematical definition of the local Hill estimator is defined as this.

This is the main idea of how to estimate the function γ locally. To obtain the asymptotic properties of this estimator, we need some conditions.

The first condition is second order condition. Suppose there exist a continuous negative function $\rho(s)$ and a set of function $A_s(t)$, such that the equation (3) holds. Assume a second order condition is quite standard and sometimes necessary for the extreme value analysis. What is different here is that this is a functional form of the second order condition.

The approach can be regarded as a combination of kernel density estimation and extreme value statistics. To prove the local and global asymptotic normality, we need to combine two limiting procedures as the number of observations tending to infinity. First, the observations used are from a h -neighborhood that is shrinking. Second, within each h -neighborhood, we need to apply a threshold to all observations that is increasing toward infinity. If the h -neighborhood shrinks too fast, there will not be sufficient observations in each neighborhood for statistical inference. If it shrinks too slowly, we would have involved too many observations with very different extreme value indices such that the estimation is distorted. Therefore, the two limiting procedures have to be balanced such that the resulting estimators possess proper asymptotic behavior. Then, we also need some conditions for k and h . They are displayed as condition 4-7. Condition (4) ensures that the number of high order statistics used in each local interval tends to infinity. Condition (5) is the one usually required for extreme value analysis to guarantee to have no asymptotic bias in the estimator. Condition (6) assumes that k_n is compatible with the h_n -variation in the γ function. Condition (7) states that $(1 - k_n)$ -quantiles of distributions are sufficiently smooth in short-interval.

Under the above conditions, the estimation for $\gamma(s)$ possess the asymptotical normality. This result is displayed in Theorem 2.1. This result is comparable with the asymptotical normality of Hill estimator, but now the estimation is based on the observations with different extreme value indices. The speed of convergence is $\sqrt{2kh}$ because only the top $[2kh]$ order statistics are used in the estimation.

Now, consider testing the hypothesis that $\gamma(s) = \gamma_0(s)$ for all $s \in [0, 1]$. Although we are able to estimate the function γ locally, since the local estimators use only local observations, their asymptotic limits are obviously independent. That prevents us from constructing a hypothesis testing procedure. In addition, the local estimators converge with a slow speed of convergence $1/\sqrt{2kh}$. To achieve the hypothesis testing goal, consider the estimation $\Gamma(s)$, which is the integral of γ function from 0 to s . $\Gamma = \Gamma_0$ is the equivalent hypothesis.

The function Γ is estimated by aggregating the local estimators of $\gamma(s)$ to a global estimator. Consider a discretized version of the local Hill estimator. And then define the estimator of estimator of $\Gamma(s)$ as the integral of the discretized version of local Hill estimator. However, the local Hill estimator is only defined

for s less than $1 - h$. So, we need to extend the range of local Hill estimator to accomodate the case when $s > 1 - h$. This is just some accomodations.

Note that $\hat{\Gamma}(s)$ is a stochastic process defined on $[0,1]$. Then, assume the same conditions as in Theorem 2.1, There exist a series of Brownian motions $W_n(s)$ such that $\hat{\Gamma}(s)$ can be approximated by this process.

Firstly, the convergence is uniformly for all $s \in [0,1]$. Secondly, the speed of convergence for the estimators $\hat{\Gamma}(s)$ is $1/\sqrt{k}$.

Now, the above Theorem, Theorem 2.2 provides