Graphical Models for Extremes

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August 11, 2020

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Introduction

Extreme Value theory in multivariate case. Assume $X \in \mathbb{R}^d$.

- Max-stable distributions arise as limits of normalized maxima of independent copies of X.
- Multivariate Pareto distributions describe the random vector X
 conditioned on the event that at least one component exceeds a high
 threshold.

Sparse Multivariate Models

- Sparse multivariate models require the notion of conditional independence.
- The problem is how to define the conditional independence for tail dependence.
- If (Z_1, Z_2, Z_3) is a max-stable random vector with positive continuous density, then the conditional independence of $Z_1 \perp \!\!\! \perp Z_3 \mid Z_2$ already implies the independence $Z_1 \perp \!\!\! \perp Z_3$.
- Meaningful conditional independence structures can thus only be obtained for max-stable distributions with discrete spectral measure.

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Graphical Models

We now consider the Multivariate Pareto distribution $\mathbf{Y} = (Y_1, \dots, Y_d)$.

For an undirected graph $\mathcal{G}=(V,E)$ with nodes $V=\{1,2,\ldots,d\}$ and edge set E, we say that \mathbf{Y} is an extremal graphical model if it satisfies the pairwise Markov property

$$\mathbf{Y}_i \perp_{\mathrm{e}} \mathbf{Y}_j \mid \mathbf{Y}_{\setminus \{i,j\}}, \quad (i,j) \notin E,$$

where we use $\perp_{\rm e}$ to stress that it is designed for extremes. (And we will define this later.)

The main advantage of conditional independence and graphical models is that they imply a simple probabilistic structure and possibly sparse patterns in multivariate random vectors.

- Let $X_i = (X_{i1}, \dots, X_{id})$, $i = 1, \dots, n$ be independent copies of X and denote the componentwise maximum by $M_n = (M_{1n}, \dots, M_{dn}) = (\max_{i=1}^n X_{i1}, \dots, \max_{i=1}^n X_{id})$.
- Assume there are sequences of normalizing constants $b_{jn} \in \mathbb{R}$, $a_{jn} > 0$, such that

$$\lim_{n\to\infty}\mathbb{P}\left(\frac{M_{jn}-b_{jn}}{a_{jn}}\leqslant x\right)=G_{j}(x)=\exp\left\{-\left(1+\xi_{j}x\right)_{+}^{-1/\xi_{j}}\right\},\quad x\in\mathbb{R}.$$

Assume

$$\lim_{n\to\infty}\mathbb{P}\left(\max_{i=1,\ldots,n}X_{i1}\leqslant nz_1,\ldots,\max_{i=1,\ldots,n}X_{id}\leqslant nz_d\right)=\mathbb{P}(\boldsymbol{Z}\leqslant\boldsymbol{z})$$

In this case, Z is max-stable with standard Fréchet marginals $\mathbb{P}(Z_j \leq z) = \exp(-1/z)$ and we may write

$$\mathbb{P}(z \leqslant z) = \exp\{-\Lambda(z)\}, \quad z \in \mathcal{E}$$

where the exponent measure Λ is a Radon measure on the cone $\mathcal{E} = [0, \infty)^d \setminus \{0\}$ and $\Lambda(z)$ is shorthand for $\Lambda(\mathcal{E} \setminus [0, z])$.

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If Λ is absolutely continuous with respect to Lebesgue measure on \mathcal{E} , its Radon–Nikodym derivative, denoted by λ has the following properties

- homogeneity of order -(d+1), i.e. $\lambda(t\mathbf{y}) = t^{-(d+1)}\lambda(\mathbf{y})$.
- normalized marginals, i.e. for any i = 1, ..., d,

$$\int_{\mathbf{y}\in\mathcal{E}:y_i>1}\lambda(\boldsymbol{y})\mathrm{d}\boldsymbol{y}=1$$

Another perspective on multivariate extremes is through threshold exceedances

$$\lim_{u\to\infty} u\{1-\mathbb{P}(\boldsymbol{X}\leqslant u\boldsymbol{z})\} = \Lambda(\boldsymbol{z}), \quad \boldsymbol{z}\in\mathcal{E}$$

Consequently, the multivariate distribution of the threshold exceedances of \boldsymbol{X} satisfies

$$\mathbb{P}(\mathsf{Y} \leqslant \mathbf{z}) = \lim_{u \to \infty} \mathbb{P}\left(\frac{\mathbf{X}}{u} \leqslant \mathbf{z} \mid ||\mathbf{X}||_{\infty} > u\right) = \frac{\Lambda(\mathbf{z} \wedge 1) - \Lambda(\mathbf{z})}{\Lambda(1)}.$$

The distribution of the limiting random vector \mathbf{Y} is called amultivariate Pareto distribution.

It is defined through the exponent measure Λ of the max-stable distribution \boldsymbol{Z} , with support on the L-shaped space $\mathcal{L} = \{x \in \mathcal{E} : \|\boldsymbol{x}\|_{\infty} > 1\}$.