

Exact simulation scheme for the Ornstein–Uhlenbeck
driven stochastic volatility (OUSV) model with the
Karhunen–Loève (KL) expansions
Applied Stochastic Processes (FIN 514)

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Ornstein–Uhlenbeck driven stochastic volatility (OUSV) model

$$\frac{dS_t}{S_t} = rdt + \sigma_t dW_t \quad \text{and} \quad d\sigma_t = \kappa(\theta - \sigma_t)dt + \xi dZ_t,$$

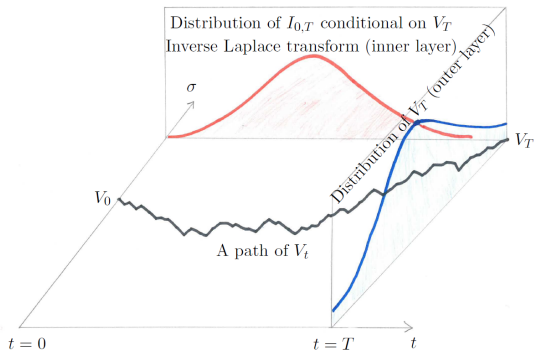
where W_t and Z_t are the standard BMs correlated by ρ , κ is the mean reversion speed, ξ is the volatility of volatility (vol-of-vol), and θ is the long-term volatility.

- Volatility σ_t is driven by the Ornstein–Uhlenbeck (OU) process.
- One of the earliest stochastic volatility (SV) models:
Uncorrelated OUSV ([Stein and Stein, 1991](#))
v.s. [Hull and White \(1987\)](#) and [Heston \(1993\)](#) models.
- Vanilla option can be efficiently priced with inverse Fourier transform ([Schöbel and Zhu, 1999](#)).
- Efficient and accurate Monte–Carlo (MC) simulation (i.e., sampling S_T) under the OUSV model is recently studied.

Introduction: Exact simulation

Exact simulation for various SV models

- **Heston (1993)** model: **Broadie and Kaya (2006)** and **Glasserman and Kim (2011, Gamma RV series)**
 - **3/2 volatility model**: **Baldeaux (2012)**
 - **Stochastic-alpha-beta-rho (SABR)**: **Cai et al. (2017)**, **Choi et al. (2019, $\beta = 0$ SABR)**, and **Cui et al. (2021, CTMC)**
 - **OUSV model**: **Li and Wu (2019)**
-
- Directly sample S_T from S_0 for any time step T without time-discretization.
 - Avoid possible discretization bias from the Euler/Milstein scheme.
 - Typical procedures:
 - ① Sample terminal volatility σ_T from a known distribution (SDE)
 - ② Sample integrated variance $V_{0,T}$ by numerically inverting its Laplace transform.
 - ③ Sample price S_T from a log-normal distribution.
 - **Heavy computation cost** from the numerical inversion, step (2).
Not necessarily more efficient than the simple Euler/Milstein scheme.



Contribution summary

Existing studies

- [Li and Wu \(2019, EJOR\)](#) propose the first exact simulation scheme for the OUSV model.
- As in other exact simulation schemes, [Li and Wu \(2019\)](#) suffer from costly numerical inversion.

Contribution of this study

- Using the [Karhunen–Loève \(KL\)](#) expansions, the volatility path (σ_t , OU process) is expressed as the infinite sine series.
- The time integrals of σ_t and σ_t^2 are analytically derived as the sum of independent normal variates.
- Only $N \leq 10$ sine terms required after handling the truncation error.
- The new simulation is **several hundred times faster than** [Li and Wu \(2019\)](#) by avoiding the transform inversion approach.

The average volatility ($U_{0,T}$) and variance ($V_{0,T}$)

- Define the time averages of volatility and variance, respectively:

$$U_{0,T} = \frac{1}{T} \int_0^T \sigma_t dt \quad \text{and} \quad V_{0,T} = \frac{1}{T} \int_0^T \sigma_t^2 dt.$$

- The final price, S_T , conditional on $(\sigma_T, U_{0,T}, V_{0,T})$, follows a log-normal distribution (Li and Wu, 2019, Proposition 1):

$$\begin{aligned} \log(S_T/S_0) &= \mu_{0,T} + \Sigma_{0,T} Z \quad (Z : \text{standard normal}), \\ \mu_{0,T} &= rT + \frac{\rho}{2\xi} [(-\xi^2 - 2\kappa\theta U_{0,T} + (2\kappa - \xi/\rho) V_{0,T}) T + (\sigma_T^2 - \sigma_0^2)] \\ \Sigma_{0,T}^2 &= \rho_*^2 T V_{0,T} \quad (\rho_* = \sqrt{1 - \rho^2}). \end{aligned}$$

Li and Wu (2019)'s sampling

- 1 $(\sigma_T, U_{0,T})$ from a bi-variate normal distribution.
- 2 $V_{0,T}$ from the inverse transform of $V_{0,T}$ conditional on $(\sigma_T, U_{0,T})$.
- 3 S_T from a lognormal distribution.

Derivation

- Let $U_{0,T} = \frac{1}{T} \int_0^T \sigma_t dt$ and $V_{0,T} = \frac{1}{T} \int_0^T \sigma_t^2 dt$.
- Derive SDE for σ_t^2 :

$$d\sigma_t = \kappa(\theta - \sigma_t)dt + \xi dZ_t.$$

$$d\sigma_t^2 = 2\sigma_t d\sigma_t + (d\sigma_t)^2 = (\xi^2 + 2\kappa(\theta\sigma_t - \sigma_t^2))dt + 2\xi\sigma_t dZ_t$$

$$\sigma_T^2 - \sigma_0^2 = \xi^2 T + 2\kappa T(\theta U_{0,T} - V_{0,T}) + 2\xi \int_0^T \sigma_t dZ_t$$

- Stochastic integral of σ_t :

$$\int_0^T \sigma_t dZ_t = \frac{1}{2\xi}(\sigma_T^2 - \sigma_0^2) - T \left(\frac{\xi}{2} + \frac{\kappa\theta}{\xi} U_{0,T} - \frac{\kappa}{\xi} V_{0,T} \right)$$

- S_T is expressed by σ_T and $V_{0,T}$!

$$\begin{aligned} \log \left(\frac{S_T}{S_0} \right) &= \rho \int_0^T \sigma_t dZ_t + \rho_* \int_0^T \sigma_t dX_t - \frac{1}{2} T V_{0,T} \\ &= \rho \left[\dots \right] - \frac{1}{2} T V_{0,T} + \rho_* \sqrt{T V_{0,T}} Z \end{aligned}$$

Sampling the terminal volatility

- Define the centered (de-measured) OU process $\hat{\sigma}_T$:

$$\hat{\sigma}_T = \sigma_T - \underbrace{\theta - (\sigma_0 - \theta)e^{-\kappa T}}_{E(\sigma_T)} = \xi e^{-\kappa T} \int_0^T e^{\kappa t} dZ_t \quad (\hat{\sigma}_0 = 0),$$

- $\hat{\sigma}_T$ follows a Gaussian process with zero mean and the covariance given by

$$\text{Cov}(\hat{\sigma}_t, \hat{\sigma}_T) = \frac{\xi^2}{2\kappa} \left(e^{-\kappa(T-t)} - e^{-\kappa(T+t)} \right) = \xi^2 \frac{\sinh(\kappa t)}{\kappa e^{\kappa T}} \quad \text{for } 0 \leq t \leq T.$$

- $\hat{\sigma}_T$ can be sampled by

$$\hat{\sigma}_T \stackrel{d}{=} \xi \sqrt{\frac{\sinh(\kappa T)}{\kappa e^{\kappa T}}} Z_0 \quad \text{for } Z_0 \sim N(0, 1).$$

Auxiliary processes (1/2)

- Remove the long-term volatility θ :

$$\bar{\sigma}_t = \sigma_t - \theta, \quad \bar{U}_{0,T} = \frac{1}{T} \int_0^T \bar{\sigma}_t dt \quad \text{and} \quad \bar{V}_{0,T} = \frac{1}{T} \int_0^T \bar{\sigma}_t^2 dt.$$

- The process $\bar{\sigma}_t$ satisfies a simpler form of the OU process:

$$d\bar{\sigma}_t = -\kappa \bar{\sigma}_t dt + \xi dZ_t \quad (\bar{\sigma}_0 = \sigma_0 - \theta).$$

- Its solution is well-known as

$$\bar{\sigma}_T = \bar{\sigma}_0 e^{-\kappa T} + \xi e^{-\kappa T} \int_0^T e^{\kappa t} dZ_t.$$

- The introduction of $(\bar{\sigma}_T, \bar{U}_{0,T}, \bar{V}_{0,T})$ will simplify algebra. The original triplets $(\sigma_T, U_{0,T}, V_{0,T})$ and the new triplet $(\bar{\sigma}_T, \bar{U}_{0,T}, \bar{V}_{0,T})$ are interchangeable by

$$\sigma_T = \theta + \bar{\sigma}_T, \quad U_{0,T} = \theta + \bar{U}_{0,T}, \quad \text{and} \quad V_{0,T} = \theta^2 + 2\theta \bar{U}_{0,T} + \bar{V}_{0,T}.$$

Auxiliary processes (2/2)

- Define the centered OU process $\hat{\sigma}_T$ by removing its mean from $\bar{\sigma}_T$,

$$\hat{\sigma}_T = \bar{\sigma}_T - \bar{\sigma}_0 e^{-\kappa T} = \xi e^{-\kappa T} \int_0^T e^{\kappa t} dZ_t \quad (\hat{\sigma}_0 = 0).$$

- $\hat{\sigma}_T$ is a Gaussian process with zero mean and the covariance given by

$$\text{Cov}(\hat{\sigma}_t, \hat{\sigma}_T) = \frac{\xi^2}{2\kappa} \left(e^{-\kappa(T-t)} - e^{-\kappa(T+t)} \right) \quad \text{for } 0 \leq t \leq T.$$

- The terminal value $\hat{\sigma}_T$ can be sampled by

$$\hat{\sigma}_T \sim \xi \sqrt{\frac{1 - e^{-2\kappa T}}{2\kappa}} Z_0 = \xi \sqrt{T\phi(2\kappa T)} Z_0,$$

$$\text{where } Z_0 \sim N(0, 1) \quad \text{and} \quad \phi(x) = \frac{1 - e^{-x}}{x} \quad (\phi(0) = 1).$$

- $\hat{\sigma}_T$ is interchangeably used with σ_T or $\bar{\sigma}_T$:

$$\sigma_T - \theta = \bar{\sigma}_T = \hat{\sigma}_T + \bar{\sigma}_0 e^{-\kappa T}.$$

The KL expansion of the OU bridge

- Given $\hat{\sigma}_T$, construct the OU bridge process, B_t , of $\hat{\sigma}_t$ for $0 \leq t \leq T$:

$$B_t = \hat{\sigma}_t - \frac{\text{Cov}(\hat{\sigma}_t, \hat{\sigma}_T)}{\text{Cov}(\hat{\sigma}_T, \hat{\sigma}_T)} \hat{\sigma}_T = \hat{\sigma}_t - \frac{\sinh(\kappa t)}{\sinh(\kappa T)} \hat{\sigma}_T \quad (B_0 = B_T = 0).$$

- The OU bridge can be decomposed into the KL expansions ([Daniluk and Muchorski, 2016](#), Theorem 2.3)

$$B_t = \xi \sum_{n=1}^{\infty} a_n \sqrt{T} \sin\left(\frac{n\pi t}{T}\right) Z_n \quad \text{for} \quad a_n = \sqrt{\frac{2}{(\kappa T)^2 + (n\pi)^2}},$$

where $Z_n \sim N(0, 1)$ *i.i.d.*

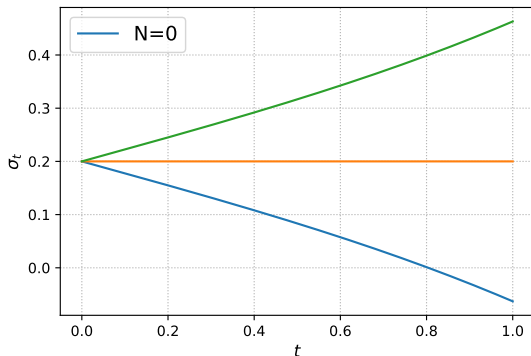
- Effectively, **KL expansions are the PCA on B_t** (∞ -dimensional data).
- When $\kappa = 0$, the expansions are nested to the KL expansions of the Brownian bridge, $a_n = \sqrt{2}/(n\pi)$.
- Finally, the volatility (i.e., OU process) path is represented by

$$\bar{\sigma}_t = \bar{\sigma}_0 e^{-\kappa t} + \hat{\sigma}_T \frac{\sinh(\kappa t)}{\sinh(\kappa T)} + \xi \sqrt{T} \sum_{n=1}^{\infty} a_n \sin\left(\frac{n\pi t}{T}\right) Z_n \quad (0 \leq t \leq T).$$

Illustration of σ_t paths from the KL expansions

Three sample volatility (OU process) paths generated by KL expansions with increasing $N = 0, 2, \dots, 64$ sine terms for

$$\sigma_0 = \theta = \xi = 0.2, \quad \text{and} \quad \kappa = 1.$$



Exact simulation scheme

$\bar{U}_{0,T}$ and $\bar{V}_{0,T}$ are analytically integrated to,

$$\bar{U}_{0,T} = \underbrace{\left[\bar{\sigma}_0 + \frac{\hat{\sigma}_T}{1 + e^{-\kappa T}} \right] \phi(\kappa T)}_{:=E(\bar{U}_{0,T} | \hat{\sigma}_T)} + 2\xi\sqrt{T} \sum_{\substack{n=1 \\ n: \text{ odd}}}^{\infty} \frac{a_n}{n\pi} Z_n,$$

$$\bar{V}_{0,T} = \underbrace{\bar{\sigma}_0^2 \phi(2\kappa T) + \hat{\sigma}_T^2 \frac{\sinh(2\kappa T) - 2\kappa T}{4\kappa T \sinh^2(\kappa T)} + \frac{\xi^2}{2\kappa} \left[\coth(\kappa T) - \frac{1}{\kappa T} \right] + \bar{\sigma}_0 \hat{\sigma}_T \frac{e^{-\kappa T}}{\kappa T} \left[\frac{1}{\phi(2\kappa T)} - 1 \right]}_{:=E(\bar{V}_{0,T} | \hat{\sigma}_T)}$$

$$+ \xi\sqrt{T} \left[\bar{\sigma}_0 \sum_{n=1}^{\infty} n\pi a_n^3 Z_n + \bar{\sigma}_T \sum_{n=1}^{\infty} (-1)^{n-1} n\pi a_n^3 Z_n \right] + \frac{\xi^2 T}{2} \sum_{n=1}^{\infty} a_n^2 (Z_n^2 - 1).$$

$$\bar{U}_{0,T} = \underbrace{\left[\bar{\sigma}_0 + \frac{\hat{\sigma}_T}{1 + e^{-\kappa T}} \right] \phi(\kappa T)}_{:=E(\bar{U}_{0,T} | \hat{\sigma}_T)} + 2\xi\sqrt{T} \sum_{\substack{n=1 \\ n: \text{ odd}}}^L \frac{a_n}{n\pi} Z_n,$$

$$\bar{V}_{0,T} = \underbrace{\bar{\sigma}_0^2 \phi(2\kappa T) + \hat{\sigma}_T^2 \frac{\sinh(2\kappa T) - 2\kappa T}{4\kappa T \sinh^2(\kappa T)} + \frac{\xi^2}{2\kappa} \left[\coth(\kappa T) - \frac{1}{\kappa T} \right] + \bar{\sigma}_0 \hat{\sigma}_T \frac{e^{-\kappa T}}{\kappa T} \left[\frac{1}{\phi(2\kappa T)} - 1 \right]}_{:=E(\bar{V}_{0,T} | \hat{\sigma}_T)}$$

$$+ \xi\sqrt{T} \left[\bar{\sigma}_0 \sum_{n=1}^L n\pi a_n^3 Z_n + \bar{\sigma}_T \sum_{n=1}^L (-1)^{n-1} n\pi a_n^3 Z_n \right] + \frac{\xi^2 T}{2} \sum_{n=1}^L a_n^2 (Z_n^2 - 1).$$

Truncated terms (1/3)

We also need to simulate the truncated terms in $\bar{U}_{0,T}$ and $\bar{V}_{0,T}$:

$$G_L = \sum_{\substack{n=L+1 \\ n \text{ odd}}}^{\infty} \frac{a_n}{n\pi} Z_n, \quad P_L = \sum_{\substack{n=L+1 \\ n \text{ odd}}}^{\infty} n\pi a_n^3 Z_n, \quad Q_L = \sum_{\substack{n=L+1 \\ n \text{ even}}}^{\infty} n\pi a_n^3 Z_n.$$

- G_L , P_L , and Q_L follow a multivariate normal distribution with

$$\mu = 0, \quad \Sigma = \begin{pmatrix} f_L^{\text{odd}} & c_L^{\text{odd}} & 0 \\ c_L^{\text{odd}} & g_L^{\text{odd}} & 0 \\ 0 & 0 & g_L^{\text{even}} \end{pmatrix},$$

where

$$c_L = \sum_{n=L+1}^{\infty} a_n^4, \quad f_L = \sum_{n=L+1}^{\infty} \frac{a_n^2}{(n\pi)^2}, \quad g_L = \sum_{n=L+1}^{\infty} (n\pi)^2 a_n^6.$$

and **odd** and **even** superscripts are defined by

$$c_L^{\text{odd}} = \sum_{\substack{n=L+1 \\ n \text{ odd}}}^{\infty} a_n^4 \quad \text{and} \quad c_L^{\text{even}} = \sum_{\substack{n=L+1 \\ n \text{ even}}}^{\infty} a_n^4.$$

Truncated terms (2/3)

- We can exactly sample G_L , P_L , and Q_L :

$$G_L = \sqrt{f_L^{\text{odd}}} \left(\sqrt{1 - \rho_{GP}^2} W_1 + \rho_{GP} W_2 \right) \quad \text{with} \quad \rho_{GP} = c_L^{\text{odd}} / \sqrt{f_L^{\text{odd}} g_L^{\text{odd}}},$$
$$P_L = \sqrt{g_L^{\text{odd}}} W_2 \quad \text{and} \quad Q_L = \sqrt{g_L^{\text{even}}} W_3,$$

where W_1 , W_2 , and W_3 are independent normal random variates.

- Even more, the infinite sums have analytic expressions: for $\lambda = \kappa T$,

$$c_0 = \frac{1}{\lambda^4} \left(\frac{\lambda}{\tanh(\lambda)} + \frac{\lambda^2}{\sinh(\lambda)^2} - 2 \right),$$
$$f_0 = \frac{b_0(0) - b_0(\lambda)}{\lambda^2}, \quad \text{and} \quad g_0(\lambda) = 2c_0 - \lambda^2 d_0.$$

- No approximation so far.

Truncated terms (3/3)

$$R_L = \sum_{n=L+1}^{\infty} a_n^2 (Z_n^2 - 1) \approx \sqrt{c_L} (W_4^2 - 1)$$

- Mean and variance of R_L are

$$E(R_L) = 0 \quad \text{and} \quad \text{Var}(R_L) = 2c_L = 2 \sum_{n=L+1}^{\infty} a_n^4.$$

- Not independent from, but zero-correlation with G_L , P_L , and Q_L .
- W_4 is a normal variate independent from W_1 , W_2 , and W_3 .
- Approximation of R_L is the only source of error.

Exact MC procedure

The procedure for simulating $\bar{\sigma}_T$, $\bar{U}_{0,T}$, and $\bar{V}_{0,T}$

$$\bar{\sigma}_T = \bar{\sigma}_0 e^{-\kappa T} + \hat{\sigma}_T \quad \text{where} \quad \hat{\sigma}_T = \xi \sqrt{T \phi(2\kappa T)} Z_0,$$

$$\bar{U}_{0,T} = E(\bar{U}_{0,T} | \hat{\sigma}_T) + 2\xi\sqrt{T} \left(\sum_{\substack{n=1 \\ n: \text{ odd}}}^L \frac{a_n}{n\pi} Z_n + G_L \right),$$

$$\begin{aligned} \bar{V}_{0,T} \approx & E(\bar{V}_{0,T} | \hat{\sigma}_T) + \xi\sqrt{T} \left[\bar{\sigma}_0 \left(\sum_{n=1}^L n\pi a_n^3 Z_n + P_L + Q_L \right) \right. \\ & \left. + \bar{\sigma}_T \left(\sum_{n=1}^L (-1)^{n-1} n\pi a_n^3 Z_n + P_L - Q_L \right) \right] + \frac{\xi^2 T}{2} \left(\sum_{n=1}^L a_n^2 (Z_n^2 - 1) + R_L \right), \end{aligned}$$

- We use $L + 5$ ($Z_0, \dots, Z_L, W_1, W_2, W_3, W_4$) independent normal variates for simulating from $t = 0$ to T .
- R_T is the only source of approximation error.

Numerical test

- We test the new MC scheme against the spot and vanilla option prices as non-trivial examples.
- Exact option values are available from the inverse Fourier transform ([Schöbel and Zhu, 1999](#)).

Parameter set to test

- For direct comparison, we test the parameter set from [Li and Wu \(2019, Table 2\)](#):

$$S_0 = K = 100, \sigma_0 = \theta = 0.2, \kappa = 4, \xi = 0.1, \rho = -0.7, \text{ and } r = 0.09531, \\ T = 1, \quad 5, \quad \text{and} \quad 10$$

- As in Fourier series, we expect more sine terms (L) is required for bigger T .

MC variance reduction

Conditional MC (Willard, 1997)

Instead of sampling S_T , MC-average conditional BS price over $(\sigma_T, U_{0,T}, V_{0,T})$:

$$C_{\text{OUSV}} = E_{\text{MC}} \left\{ C_{\text{BS}} \left(K, F_T, \rho_* \sqrt{V_{0,T}} \right) \right\}, \quad \text{where}$$

$$\begin{aligned} F_T &= E\{S_T \mid \sigma_T, U_{0,T}, V_{0,T}\} = S_0 \exp(\mu_{0,T} + \Sigma_{0,T}^2/2) \\ &= S_0 \exp \left(rT + \frac{\rho}{2\xi} [(-\xi^2 - 2\kappa\theta U_{0,T} + (2\kappa - \rho\xi) V_{0,T}) T + (\sigma_T^2 - \sigma_0^2)] \right). \end{aligned}$$

- Cai et al. (2017) reports 99% variance reduction (in the SABR model).
- With the reduced variance, we aim to measure the bias accurately.

Martingale-correcting control variate

Because $S_0 = e^{-rT} E_{\text{MC}} \{F_T\}$, correct F_T by

$$F_T^{\text{cv}} = \mu F_T \quad \text{for} \quad \mu = S_0 e^{rT} / E_{\text{MC}} \{F_T\}.$$

Numerical results: $T = 1$

Table: The simulation result for $T = 1$. “Spot Price” and “Option Price” evaluate \bar{S}_T and C_{OUSV} , respectively. “Option Price with CV” uses the control variate \bar{S}_T^{cv} . The true option price is [13.21492](#).

N	n_{path} (number of paths)	Spot Price		Option Price		Option Price with CV		
		Bias $\times 10^{-4}$	RMSE $\times 10^{-2}$	Bias $\times 10^{-4}$	RMSE $\times 10^{-2}$	Bias $\times 10^{-4}$	RMSE $\times 10^{-2}$	CPU Time (Seconds)
2	10,000	0.3	1.74	-0.4	4.04	-0.7	3.48	0.006
	40,000	0.3	0.87	-0.4	2.33	-0.6	1.74	0.024
	160,000	0.3	0.44	-0.4	1.17	-0.6	0.88	0.096
4	10,000	-0.2	1.73	-0.9	4.03	-0.7	3.47	0.008
	40,000	-0.2	0.87	-0.9	2.33	-0.7	1.74	0.028
	160,000	-0.2	0.44	-0.9	1.17	-0.7	0.87	0.109
6	10,000	-0.2	1.74	-0.4	4.04	-0.3	3.48	0.008
	40,000	-0.2	0.87	-0.4	2.33	-0.2	1.74	0.030
	160,000	-0.2	0.44	-0.4	1.17	-0.2	0.87	0.122

Numerical results: $T = 5$

Table: The simulation result for $T = 5$. “Spot Price” and “Option Price” evaluate \bar{S}_T and C_{OUSV} , respectively. “Option Price with CV” uses the control variate \bar{S}_T^{cv} . The true option price is 40.797689.

N	n_{path} (number of paths)	Spot Price		Option Price		Option Price with CV		
		Bias $\times 10^{-4}$	RMSE $\times 10^{-2}$	Bias $\times 10^{-4}$	RMSE $\times 10^{-2}$	Bias $\times 10^{-4}$	RMSE $\times 10^{-2}$	CPU Time (Seconds)
4	10,000	19.7	9.56	10.7	12.13	-6.8	4.11	0.006
	40,000	19.7	4.78	10.7	6.35	-6.9	2.06	0.026
	160,000	19.7	2.41	10.7	3.19	-7.0	1.04	0.103
6	10,000	5.4	9.57	3.4	12.13	-1.2	4.12	0.007
	40,000	5.4	4.79	3.4	6.37	-1.4	2.06	0.028
	160,000	5.4	2.39	3.4	3.18	-1.4	1.03	0.110
8	10,000	-0.2	9.54	-1.2	12.11	-0.7	4.11	0.007
	40,000	-0.2	4.76	-1.2	6.33	-0.9	2.05	0.028
	160,000	-0.2	2.37	-1.2	3.15	-0.9	1.02	0.111

Numerical results: $T = 10$

Table: The simulation result for $T = 10$. “Spot Price” and “Option Price” evaluate \bar{S}_T and C_{OUSV} , respectively. “Option Price with CV” uses the control variate \bar{S}_T^{cv} . The true option price is 62.76312.

N	n_{path} (number of paths)	Spot Price		Option Price		Option Price with CV		
		Bias $\times 10^{-4}$	RMSE $\times 10^{-2}$	Bias $\times 10^{-4}$	RMSE $\times 10^{-2}$	Bias $\times 10^{-4}$	RMSE $\times 10^{-2}$	CPU Time (Seconds)
6	10,000	19.3	19.14	16.8	20.68	-1.3	2.65	0.007
	40,000	19.3	9.59	16.8	10.56	-1.6	1.33	0.028
	160,000	19.3	4.78	16.8	5.27	-1.7	0.66	0.111
8	10,000	-1.5	19.06	-1.9	20.63	0.0	2.64	0.007
	40,000	-1.5	9.50	-1.9	10.47	-0.3	1.32	0.029
	160,000	-1.5	4.73	-1.9	5.20	-0.4	0.65	0.117
10	10,000	-4.6	19.10	-4.5	20.65	0.5	2.65	0.008
	40,000	-4.6	9.57	-4.5	10.54	0.2	1.33	0.033
	160,000	-4.6	4.82	-4.5	5.31	0.1	0.67	0.132

Contribution of this study

- Using the [Karhunen–Loève \(KL\)](#) expansions, the volatility path (OU process) is expressed as the infinite sine series.
- The time integrals of σ_t and σ_t^2 are analytically derived as the sum of independent normal variates.
- Only $L \leq 10$ sine terms required after handling the truncation error.
- The new simulation is **several hundred times faster than Li and Wu (2019)** by avoiding the transform inversion approach.

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