

# Stochastic Finance (FIN 519)

## Midterm Exam

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**BM** stands for Brownian motion. Assume that  $B_t$  is a standard **BM**. **RN** and **RV** stand for random number and random variable, respectively. The PDF and CDF of the standard normal distribution are denoted by  $n(z)$  and  $N(z)$  respectively. You can use  $n(z)$  and  $N(z)$  in your answers without further evaluation.

1. (5 points) [**Lognormal distribution**] A lognormal random variables with parameters  $(\mu, \sigma)$  is given by

$$Y = \mu \exp(\sigma Z - \sigma^2/2) \quad \text{for a standard normal variable } Z.$$

- (a) (2 points) Obtain the mean and variance of  $Y$ .
- (b) (3 points) Suppose that two lognormal random variables are give by

$$Y_1 = \mu_1 \exp(\sigma_1 Z_1 - \sigma_1^2/2) \quad \text{and} \quad Y_2 = \mu_2 \exp(\sigma_2 Z_2 - \sigma_2^2/2),$$

and that the two standard normals,  $Z_1$  and  $Z_2$ , are correlated by  $\rho$  (i.e.,  $E(Z_1 Z_2) = \rho$ ). Obtain the covariance and correlation between  $Y_1$  and  $Y_2$ .

### Solution:

- (a) The mean and variance are given by  $\mu$  and  $\mu^2(e^{\sigma^2} - 1)$  respectively.

$$\begin{aligned} E(Y) &= \mu E(\exp(\sigma Z - \sigma^2/2)) = \mu. \\ E(Y^2) &= \mu^2 E(\exp(2\sigma Z - \sigma^2)) = \mu^2 \exp(\sigma^2) E(\exp(2\sigma Z - (2\sigma)^2/2)) = \mu^2 \exp(\sigma^2). \\ \text{Var}(Y) &= E(Y^2) - E(Y)^2 = \mu^2 (\exp(\sigma^2) - 1). \end{aligned}$$

- (b) Using  $\sigma_1 Z_1 + \sigma_2 Z_2 \sim N(0, \sigma_1^2 + \sigma_2^2 + 2\rho\sigma_1\sigma_2)$ ,

$$\begin{aligned} E(Y_1 Y_2) &= \mu_1 \mu_2 E(\exp(\sigma_1 Z_1 + \sigma_2 Z_2 - (\sigma_1^2 + \sigma_2^2)/2)) \\ &= \mu_1 \mu_2 \exp(\rho\sigma_1\sigma_2) E(\exp(\sigma_1 Z_1 + \sigma_2 Z_2 - (\sigma_1^2 + \sigma_2^2 + 2\rho\sigma_1\sigma_2)/2)) \\ &= \mu_1 \mu_2 \exp(\rho\sigma_1\sigma_2) \\ \text{Cov}(Y_1, Y_2) &= E(Y_1 Y_2) - E(Y_1)E(Y_2) = \mu_1 \mu_2 (\exp(\rho\sigma_1\sigma_2) - 1) \\ \text{Corr}(Y_1, Y_2) &= \frac{\text{Cov}(Y_1, Y_2)}{\sqrt{\text{Var}(Y_1)\text{Var}(Y_2)}} = \frac{\exp(\rho\sigma_1\sigma_2) - 1}{\sqrt{(\exp(\sigma_1^2) - 1)(\exp(\sigma_2^2) - 1)}}. \end{aligned}$$

2. (2×4 points) [**Martingale related to BM**] If  $B_t$  is a standard BM, determine whether the following is a martingale or not. Give a brief reason.
- (a)  $Y_t = B_{\lambda t}^2 - \lambda^2 t$
  - (b)  $Y_t = \exp(-B_{at} - a^2 t/2)$
  - (c)  $Y_t = \begin{cases} 2A - B_t & \text{if } t \leq \tau \\ B_t & \text{if } t > \tau \end{cases}$ , where  $\tau$  is the first time  $B_t$  hits  $A$  ( $\tau = \min\{t : B_t = A\}$ )
  - (d)  $S_t = S_0 + \sigma B_{t \wedge \tau}$  where  $\tau = \min\{t : S_{t+1} - S_t < -A\}$  for some  $A > 0$ .

**Solution:**

- (a) No.  $Y_t = B_{\lambda t}^2 - \lambda t$  is a martingale.
- (b) No.  $Y_t = \exp(-B_{at} - at/2)$  is a martingale.
- (c) Yes.  $Y_t$  is a BM starting with  $2A$  from the reflection principle. Although  $Y_t$  is not a standard BM, it is still a martingale.
- (d) No.  $\tau$  is not a proper stopping time because it is based on forward-looking information (i.e.,  $S_{t+1}$ ).

3. (3 points) [**Forward-starting option**] A forward-starting option with expiry  $T$  is an option whose strike price is set relative to the stock price at time  $t = T'$  ( $< T$ ) not at time  $t = 0$ . Suppose that the strike will be set as  $K = S_{T'} + \Delta$  at  $t = T'$ . Therefore, the payout of the forward-starting call option at expiry  $T$  is given by

$$\text{Payout} = \max(S_T - K, 0) = \max(S_T - S_{T'} - \Delta, 0).$$

Assume that the underlying stock price follows a BM,  $S_t = S_0 + \sigma B_t$  (and  $r = q = 0$ ). Derive the price of the forward-starting call option. You may use the Bachelier model option formula without proof.

**Solution:** Since  $S_T - S_{T'} = \sigma(B_T - B_{T'}) \sim \sigma\sqrt{T - T'} Z$  for a standard normal  $Z$ , we can use the Bachelier option price formula with time-to-expiry  $T - T'$ , spot price 0, and strike price  $\Delta$ :

$$C = -\Delta N(d_N) + \sigma\sqrt{T - T'} n(d_N), \quad d_N = -\frac{\Delta}{\sigma\sqrt{T - T'}}.$$

4. (7 points) [**Stochastic integral**] Based on the highschool calculus,  $\int_0^x e^{-x} dx = 1 - e^{-x}$ , I make the statement on the following stochastic integral:

$$\int_0^T e^{-B_t} dB_t = 1 - e^{-B_T}.$$

We will check if this is true or false.

- (a) (3 points) What is the mean and variance of the left-hand side (LHS)?  
 (b) (3 points) What is the mean and variance of the right-hand side (RHS)?  
 (c) (1 point) Is the statement true or false?

**Solution:**

- (a) The mean of LHS is zero by symmetry. Using the Itô's isometry,

$$\text{Var}(\text{LHS}) = E \left( \int_0^T e^{-2B_t} dt \right) = \int_0^T E(e^{-2B_t}) dt = \int_0^T e^{2t} dt = \frac{e^{2T} - 1}{2}$$

- (b) Regarding RHS,

$$\begin{aligned} E(\text{RHS}) &= E(1 - e^{-B_T}) = 1 - E(e^{-B_T}) = 1 - e^{T/2}, \\ \text{Var}(\text{RHS}) &= E \left( (e^{-B_T} - e^{T/2})^2 \right) = E(e^{-2B_T}) - 2e^{T/2}E(e^{-B_T}) + e^T \\ &= e^{2T} - 2e^{T/2} \cdot e^{T/2} + e^T = e^T(e^T - 1). \end{aligned}$$

- (c) Because the means and variance are not same, the statement is false.

5. (3 points) [**Geometric BM**] Assume that a stock price follows a geometric BM with volatility  $\sigma$  and initial price  $S_0$ ,

$$S_t = S_0 \exp \left( \sigma B_t - \frac{\sigma^2 t}{2} \right)$$

When you observe the stock price ever year ( $t = 1, 2, \dots$ ), find the probability that the annual stock return is positive for the following 3 years,

$$P(S_1 > S_0 \text{ and } S_2 > S_1 \text{ and } S_3 > S_2).$$

Express the answer with  $n(\cdot)$  or  $N(\cdot)$ .

**Solution:** For  $t = n$ ,

$$\begin{aligned} P(S_{n+1} > S_n) &= P(S_{n+1}/S_n > 1) = P(\sigma(B_{n+1} - B_n) - \sigma^2/2 > 0) \\ &= P(\sigma Z - \sigma^2/2 > 0) = P(Z > \sigma/2) \quad \text{for a standard normal } Z \\ &= 1 - N(\sigma/2) \quad \text{or} \quad N(-\sigma/2) \end{aligned}$$

Thanks to the independent increments of BM,  $S_1 > S_0$ ,  $S_2 > S_1$ , and  $S_3 > S_2$  are independent events. Therefore,

$$P(S_1 > S_0 \text{ and } S_2 > S_1 \text{ and } S_3 > S_2) = (1 - N(\sigma/2))^3$$

For example, if  $\sigma = 20\%$  ( $= 0.2$ ), the probability is  $(1 - N(0.1))^3 \approx (0.460)^3 = 9.74\%$ .

6. (4 points) [**Last hitting time of BM**] In class, we know that the **first** hitting time of BM to  $\delta$ ,  $\tau = \min \{t : B_t = \delta\}$ , has the following CDF:

$$P(\tau \leq t) = 2 - 2N(\delta/\sqrt{t}).$$

Instead, let us consider the **last** hitting time  $\tau'$  as

$$\tau' = \max \{t : B_t = \delta\}.$$

Derive the CDF (i.e.,  $P(\tau' \leq t)$ ) and PDF of  $\tau'$ . **Hint:**  $Y_t = t B_{1/t}$  ( $Y_0 = 0$ ) is also a standard BM.

**Solution:** If  $\tau$  is the first hitting time,  $1/t$  is the last exit time of BM. Using that  $Y_t = t B_{1/t}$  is also a standard BM,

$$\begin{aligned}\tau' &= \max \{t : Y_t = \delta\} = \max \{t : t B_{1/t} = \delta\} = \max \{t : B_{1/t} = \delta/t\} \\ &= \max \{1/t : B_t = \delta t\} = 1 / \min \{t : B_t = \delta t\} = 1/\tau.\end{aligned}$$

Therefore, the CDF and PDF are respectively given by

$$\begin{aligned}P(\tau' \leq t) &= P(\tau \geq 1/t) = 2N(\delta\sqrt{t}) - 1, \\ f_{\tau'}(t) &= \frac{1}{dt}P(\tau' \leq t) = \frac{\delta}{\sqrt{t}} n(\delta\sqrt{t}).\end{aligned}$$

Note that the PDF is similar to, but different from, the PDF of the first hitting time  $\tau$ :

$$f_{\tau}(t) = \frac{\delta}{\sqrt{t^3}} n\left(\frac{\delta}{\sqrt{t}}\right).$$

This question was inspired by a question on [Math StackExchange](#).