

10.2A) A Type I error is concluding that the drug induces sleep in less than 80 percent of insomniacs when it actually induces sleep in 80 percent of insomniacs.

$$10.2B) \alpha = P(Y \leq 12 | p = 0.8) = \boxed{0.032} \quad (n=20) \quad \text{Appendix III, Table 1D}$$

10.2C) A Type II error is concluding that the drug induces sleep in 80 percent of insomniacs when it actually induces sleep in less than 80 percent of insomniacs.

$$10.2D) \beta_{p=0.6} = P(Y > 12 | p = 0.6) = 1 - P(Y \leq 12 | p = 0.6) = 1 - 0.584 = \boxed{0.416}$$

$n=20, \alpha=12, p=0.6$ Appendix III, Table 1D

$$10.2E) \beta_{p=0.4} = P(Y > 12 | p = 0.4) = 1 - P(Y \leq 12 | p = 0.4) = 1 - 0.979 = \boxed{0.021}$$

$n=20, \alpha=12, p=0.4$ Appendix III, Table 1D

10.3A) Since $n=20$ and $p=0.8$, we can check Appendix III, Table 1D to see that $\alpha = \boxed{11} = \boxed{C}$ when $\alpha = 0.01$.

$$10.3B) \beta_{p=0.6} = P(Y > 11 | p = 0.6) = 1 - P(Y \leq 11 | p = 0.6) = 1 - 0.404 = \boxed{0.596}$$

$n=20, \alpha=11, p=0.6$ Appendix III, Table 1D

$$10.3C) \beta_{p=0.4} = P(Y > 11 | p = \overset{0.4}{0.4}) = 1 - P(Y \leq 11 | p = 0.4) = 1 - 0.943 = \boxed{0.057}$$

$n=20, \alpha=11, p=0.4$ Appendix III, Table 1D

10.19) $H_0: \mu = 130, H_A: \mu < 130, \alpha = 0.05$ $n=40, \bar{y}=128.6, s=2.1$
 Because $n=40 > 30$, the Central Limit Theorem dictates that we can use the normal distribution.

$$z = \frac{\bar{y} - \mu_0}{\frac{s}{\sqrt{n}}} = \frac{128.6 - 130}{\frac{2.1}{\sqrt{40}}} = \frac{-1.4}{0.33} = -4.22$$

We reject H_0 at the $\alpha = 0.05$ level. There is sufficient evidence that the mean output voltage is less than 130.

130

$$10.37) \bar{Y} + Z_{0.05} \left(\frac{s}{\sqrt{n}} \right) = 128.6 + (1.645) \left(\frac{2.1}{\sqrt{40}} \right) = 128.6 + \frac{10.2910}{5000} \approx 129.25 < \mu = 130$$

$$\beta_{\mu=128} = P(\bar{Y} \geq 129.454 | \mu = 128) = P\left(\frac{\bar{Y} - \mu_0}{\frac{s}{\sqrt{n}}} \geq \frac{129.454 - 128}{\frac{2.1}{\sqrt{40}}}\right) = P(Z \geq 4.3784) < 0.00001$$

$$10.49A) \bar{Y} + Z_{0.025} \left(\frac{s}{\sqrt{n}} \right) = 128.6 + (1.96) \left(\frac{2.1}{\sqrt{40}} \right) = 128.6 + \frac{10.2910}{5000} \approx 129.25 < \mu = 130$$

The value $\mu = 130$ is greater than the 95 percent upper confidence bound.

$$10.49B) \text{Yes} \quad 10.49C) \text{No} \quad 10.24) H_0: p = 0.15, H_1: p < 0.15, \alpha = 0.05 \quad n = 100, \hat{p} = 0.15, q_0 = 1 - p_0 = 0.85$$

$$Z = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0(1-p_0)}{n}}} = \frac{0.13 - 0.15}{\sqrt{\frac{(0.15)(0.85)}{100}}} \approx -0.5601 \quad |-0.5601| < Z_{0.05} = 1.645$$

We fail to reject H_0 at the $\alpha = 0.05$ level. There is insufficient evidence that the percentage reported by Children's Hospital is too high.

$$10.27) H_0: p_R - p_{PS} = 0, H_A: p_R - p_{PS} \neq 0, \alpha = 0.01 \quad n_R = 6124, n_{PS} = 5512, \hat{p}_R = 0.4, \hat{p}_{PS} = 0.37$$

$$\hat{p} = \frac{n_R \hat{p}_R + n_{PS} \hat{p}_{PS}}{n_R + n_{PS}} = \frac{(6124)(0.4) + (5512)(0.37)}{6124 + 5512} \approx 0.386 \quad (\text{Pooled samples}) \quad \hat{q} = 1 - \hat{p} \approx 0.614$$

$$Z = \frac{\hat{p}_R - \hat{p}_{PS}}{\sqrt{\hat{p}\hat{q}\left(\frac{1}{n_R} + \frac{1}{n_{PS}}\right)}} = \frac{0.4 - 0.37}{\sqrt{(0.386)(0.614)\left(\frac{1}{6124} + \frac{1}{5512}\right)}} \approx 3.319 \quad |3.319| > Z_{0.005} = 2.579$$

We reject H_0 at the $\alpha = 0.01$ level. There is sufficient evidence that there is a difference between the two proportions.

$$10.43A) H_0: \mu_S - \mu_{NS} = 0, H_A: \mu_S - \mu_{NS} > 0, \alpha = 0.05 \quad \bar{y}_S = 32.19, s_S = 4.34, n_S = 31$$

$$\bar{y}_{NS} = 31.68, s_{NS} = 4.56, n_{NS} = 31$$

$$Z = \frac{\bar{y}_S - \bar{y}_{NS}}{\sqrt{\frac{s_S^2}{n_S} + \frac{s_{NS}^2}{n_{NS}}}} = \frac{32.19 - 31.68}{\sqrt{\frac{(4.34)^2}{31} + \frac{(4.56)^2}{31}}} \approx 0.4928 \quad |0.4928| < Z_{0.05} = 1.645$$

We fail to reject H_0 at the $\alpha = 0.05$ level. There is insufficient evidence that second graders who participate in sports have a higher mean dexterity score.

$$10.43B) \bar{y}_S - \bar{y}_{NS} > 1.645 \sqrt{\frac{(4.34)^2}{(31)} + \frac{(4.56)^2}{(31)}} \rightarrow \bar{y}_S - \bar{y}_{NS} > 1.702$$

$$\beta_{\mu_S - \mu_{NS} = 3} = P(\bar{y}_S - \bar{y}_{NS} \leq 1.702 | \mu_S - \mu_{NS} = 3) = P\left(Z \leq \frac{1.702 - 3}{\sqrt{\frac{(4.34)^2}{(31)} + \frac{(4.56)^2}{(31)}}}\right) = P(Z \leq -1.25) \approx 0.1056$$

$$10.44) \alpha = \beta = 0.05 = P(Z \leq Z_{0.05} | \mu_S - \mu_{NS} = 3) = P\left(\frac{\bar{y}_S - \bar{y}_{NS} - 0}{\sqrt{\frac{s_S^2}{n} + \frac{s_{NS}^2}{n}}} \leq Z_{0.05} | \mu_S - \mu_{NS} = 3\right) = P\left(\frac{\bar{y}_S - \bar{y}_{NS} - 3}{\sqrt{\frac{s_S^2}{n} + \frac{s_{NS}^2}{n}}} \leq Z_{0.05} - \frac{3}{\sqrt{\frac{s_S^2}{n} + \frac{s_{NS}^2}{n}}}\right)$$

$$Z_{0.05} - \frac{3}{\sqrt{\frac{s_S^2}{n} + \frac{s_{NS}^2}{n}}} = -Z_{0.05} \rightarrow \frac{3}{\sqrt{\frac{s_S^2}{n} + \frac{s_{NS}^2}{n}}} = 2Z_{0.05} \rightarrow \frac{3}{2Z_{0.05}} = \sqrt{\frac{s_S^2}{n} + \frac{s_{NS}^2}{n}}$$

$$\frac{3}{4(1.645)^2} = \frac{4.34^2 + 4.56^2}{n} \rightarrow n = \frac{4(1.645)^2(4.34^2 + 4.56^2)}{9} \approx 47.66 \rightarrow n = 48$$