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STAT 488-001

14 June 2022

1 3.3A) Posterior distribution: $\prod_{i=1}^{37} N(y_{ci} | \mu_c, \sigma_c^2) \prod_{i=1}^{22} N(y_{ti} | \mu_t, \sigma_t^2)$

control distributions Treatment distributions

From Section 3.2, we see the marginal posterior distribution of μ_t is $t_{n_t-1}(\mu_t, \frac{\sigma_t^2}{n_t}) = t_{(301)-1}(1.173, \frac{(0.2)^2}{(35)}) = t_{35}(1.173, \frac{1}{900})$.

3.3B) See next page

2 3.4A) From Section 3.7, we see that a reasonable prior is $\text{Beta}(\frac{1}{2}, \frac{1}{2})$.
 $\frac{\Gamma(\frac{1}{2})\Gamma(\frac{1}{2})}{\Gamma(1)} \propto p^{\frac{1}{2}}(1-p)^{\frac{1}{2}}$. The posterior for sampling would then be \propto
 $\frac{P_0^{39}(1-P_0)^{674-39} P_1^{22}(1-P_1)^{602-22}}{P_0^{38.5}(1-P_0)^{634.5} P_1^{21.5}(1-P_1)^{657.5}}$ $\alpha = \text{deaths}$
 $\beta = \text{total-deaths}$

3.4B) We can see the distribution is right-skew with a mean of 0.56.

3.4C) We can see from the contour plots in problem 3.4A that a lot of the simulations fall outside the outermost contour level. Although the choice of prior appears to be good, better choices of the α and β parameters could improve the contour plot.

3 3.6A) This noninformative prior distribution is easier to work with. It is improper because $\int_0^\infty \frac{1}{\lambda} d\lambda$ diverges. $p(N, \theta) = \int_0^\infty p(N, \lambda, \theta) d\lambda = \int_0^\infty p(N | \lambda, \theta) p(\lambda, \theta) d\lambda = \int_0^\infty \frac{(\lambda)^{N-1} e^{-\lambda}}{\Gamma(N)} \cdot \frac{1}{\lambda} d\lambda = \frac{1}{\Gamma(N)} \int_0^\infty \lambda^{N-2} e^{-\lambda} d\lambda = \frac{1}{\Gamma(N)} \Gamma(N-1) = \frac{1}{N}$

3.6B) $p(N, \theta | y) = \frac{p(y | N, \theta) p(N, \theta)}{\int_0^\infty \int_0^1 p(y | N, \lambda, \theta) p(\lambda, \theta) d\lambda d\theta}$
 $\left(\prod_{i=1}^N \binom{N}{y_i} \theta^{y_i} (1-\theta)^{N-y_i} \right) \cdot \frac{1}{N} \propto \frac{\theta^{\sum y_i} (1-\theta)^{N-\sum y_i}}{\theta^{\sum y_i} (1-\theta)^{N-\sum y_i}} \sim \text{Beta}(\sum y_i + 1, N - \sum y_i + 1)$

$p(N | y) = \frac{\frac{1}{N} \prod_{i=1}^N \binom{N}{y_i} \int_0^1 \theta^{\sum y_i} (1-\theta)^{N-\sum y_i} d\theta}{\int_0^1 \frac{1}{N} \prod_{i=1}^N \binom{N}{y_i} \cdot \text{Beta}(\sum y_i + 1, N - \sum y_i + 1) d\theta}$ (for any size n)

3.6C) We could not use a Poisson distribution with fixed μ as a prior because we don't know what μ is and don't have any reasonable estimate.

Homework 2

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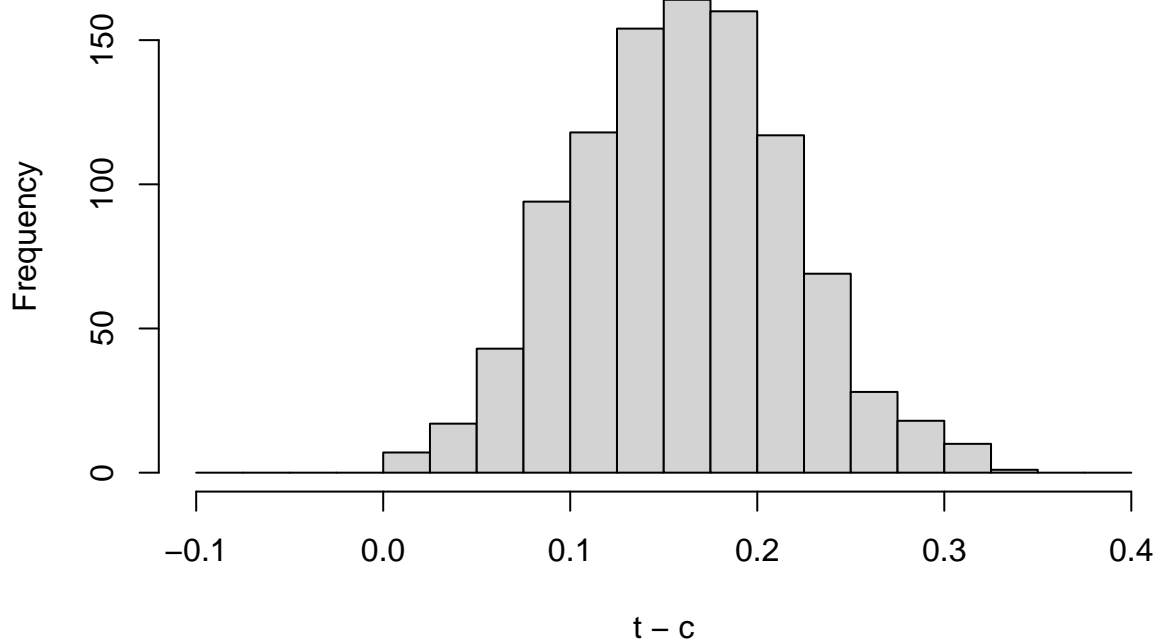
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Problem 3.3

Problem 3.3(b)

```
rm(list=ls())  
set.seed(1406)  
c<-1.013+0.24/sqrt(32)*rt(1000,32-1)  
t<-1.173+0.20/sqrt(36)*rt(1000,36-1)  
hist(t-c,breaks=seq(-0.1,0.4,0.025))
```

Histogram of $t - c$



```
quantile(t-c,c(0.05/2,1-0.05/2))
```

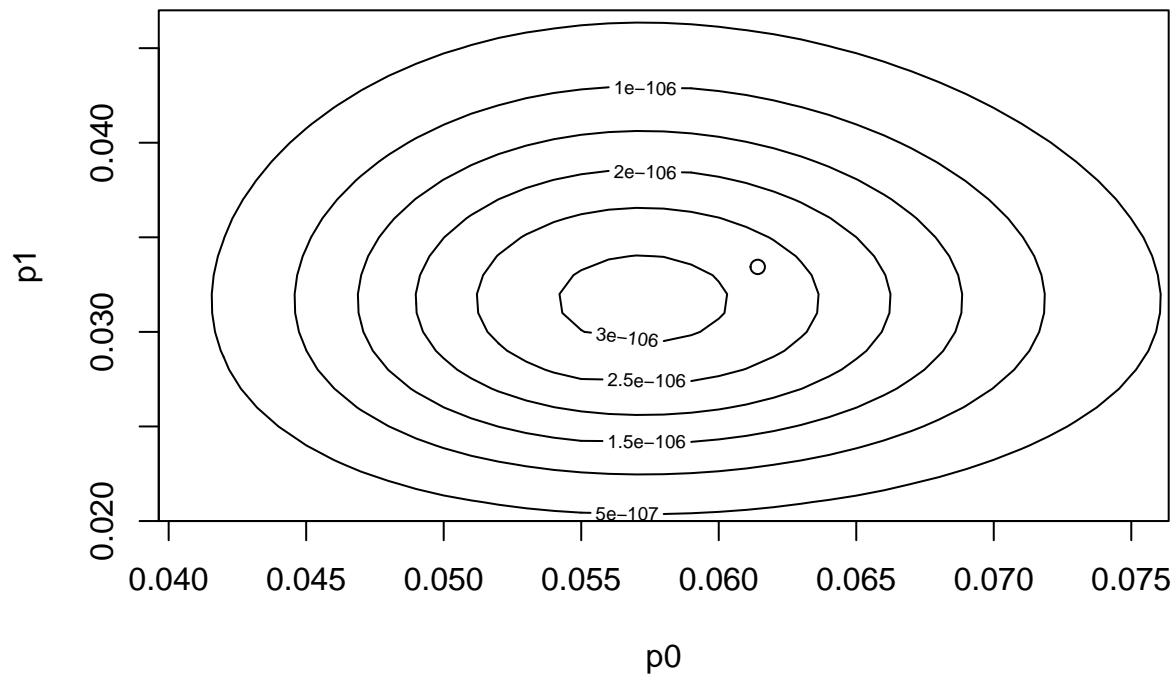
```
##          2.5%      97.5%
## 0.05193314 0.27704710
```

Problem 3.4

Problem 3.4(a)

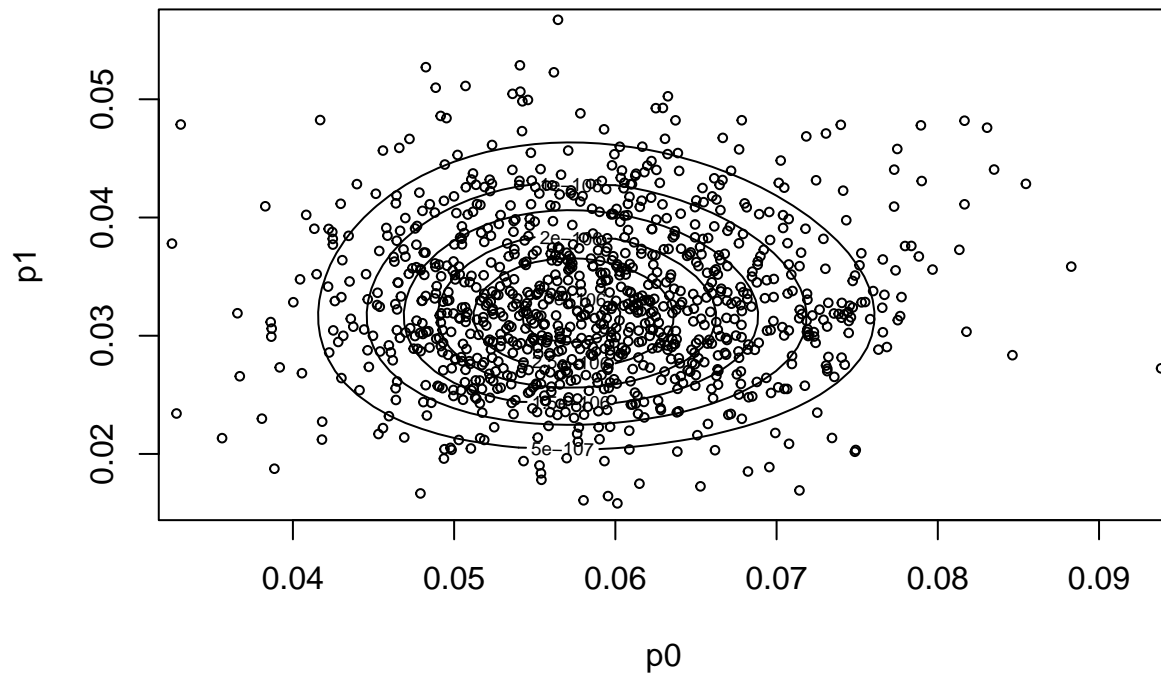
```
library(ggplot2)
library(metR)
p0<-p1<-seq(0,1,0.001)
p0p1<-expand.grid(p0,p1)
zim<-function(p0,p1){return(p0^38.5*(1-p0)^634.5*p1^21.5*(1-p1)^657.5)}
z<-vector(mode="numeric",length=nrow(p0p1))
for(i in 1:nrow(p0p1)){z[i]<-zim(p0p1[i,"Var1"],p0p1[i,"Var2"])}
contour(p0,p1,matrix(z,1001,1001),xlim=c(0.041,0.075),ylim=c(0.021,0.046),xlab="p0",ylab="p1",main="Problem 3.4(a) - Contour Plot with MLE")
points(39/(674-39),22/(680-22))
```

Problem 3.4(a) – Contour Plot with MLE



```
set.seed(1406)
x<-rbeta(1000,39.5,635.5)
y<-rbeta(1000,22.5,658.5)
contour(p0,p1,matrix(z,1001,1001),xlim=c(0.034,0.092),ylim=c(0.016,0.056),xlab="p0",ylab="p1",main="Problem 3.4(a) - Contour Plot with MLE")
points(x,y,cex=.6)
```

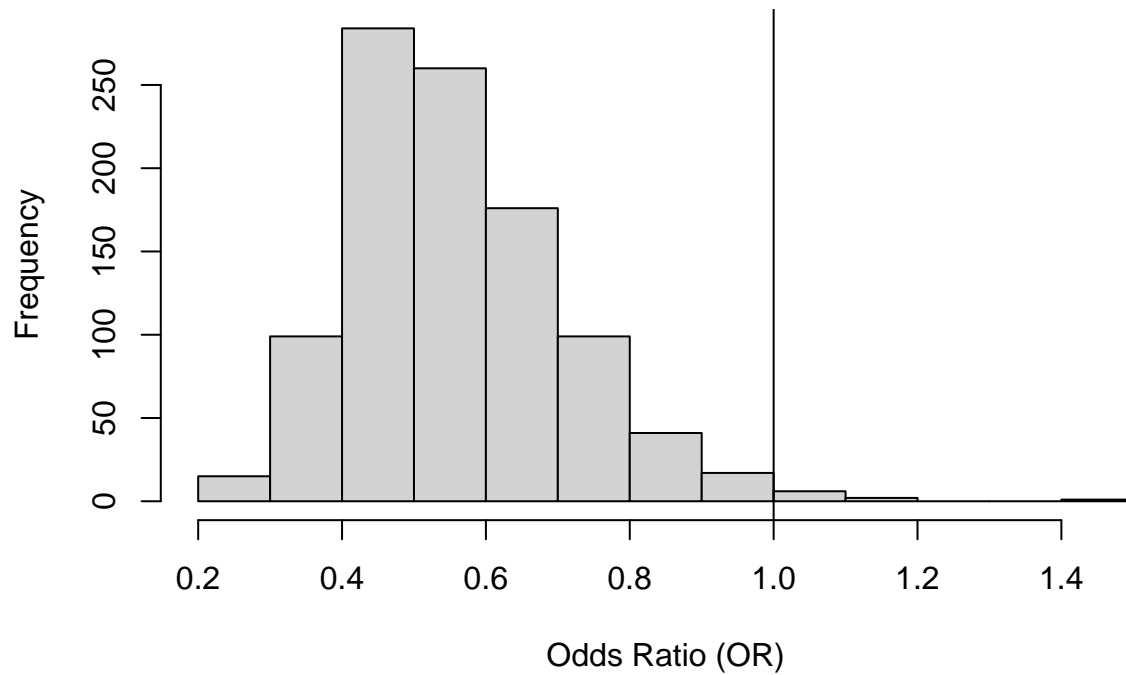
Problem 3.4(a) – Contour Plot with 1,000 Simulations



Problem 3.4(b)

```
OR<-function(p0,p1){return(p1/(1-p1)/(p0/(1-p0)))}  
pOR<-apply(data.frame(p0=x,p1=y),1,function(row) OR(row[1],row[2]))  
hist(pOR,xlab="Odds Ratio (OR)",main="Problem 3.4(b) - Posterior Distribution of Odds Ratio")  
abline(v=1)
```

Problem 3.4(b) – Posterior Distribution of Odds Ratio



```
summary(pOR)
```

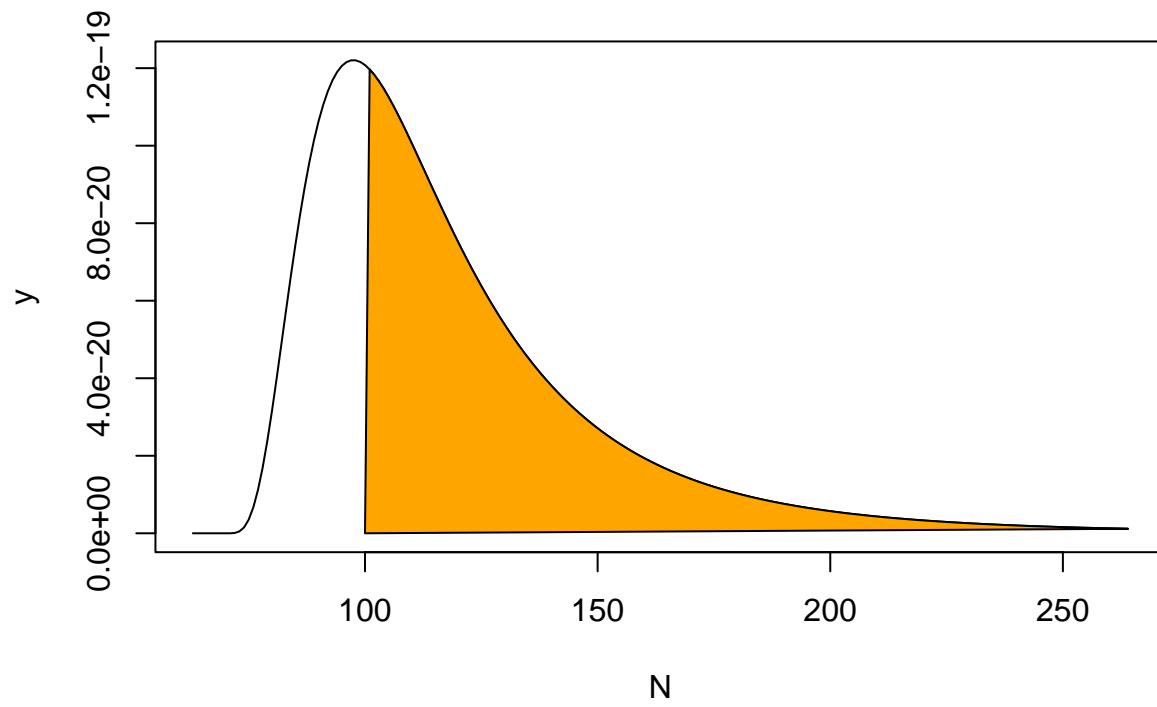
```
##      Min. 1st Qu.  Median    Mean 3rd Qu.    Max.
## 0.2237 0.4544 0.5359 0.5589 0.6384 1.4712
```

We can see the distribution is right-skew with a mean of 0.5589333.

Problem 3.6(b)

```
w<-c(53,57,66,67,72)
b<-function(N,w){return(prod(sapply(w,function(x) choose(N,x)/N))*beta(sum(w)+1,5*N-sum(w)+1))}
N<-seq(63,264,1) # Testing different bounds
y<-sapply(N,function(x) b(x,w))
plot(N,y,type="l",main="Problem 3.6(b) - Marginal Posterior of N")
polygon(N[N>=100],c(0,y[(100-min(N)+1+1):length(y)]),col="orange")
```

Problem 3.6(b) – Marginal Posterior of N



```
sum(y[(N>100)]) / sum(y)
```

```
## [1] 0.6729019
```