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Modu	de <b>4M24</b>	Tit	le of report	High-Dime	Dimensional MCMC									
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C O N T E N T	Completeness, quantity of content: Has the report covered all aspects of the lab? Has the analysis been carried out thoroughly?													
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	Overall assessment (circle grade)  A*			A*		A		В		С	D			
	Guideli	ne standard	I	>75%		65-7	5%	55-65%	40-55%		<	<40%		
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# 4M24 CW - High-Dimensional MCMC

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#### Abstract

This report outlines the result of the 4M24 coursework on high-dimensional Markov Chain Monte Carlo (MCMC).

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## 1 Introduction

### 2 Simulation

#### a Sampling from a Gaussian Process

We begin with a Gaussian Process (GP) defined on a 2D domain  $\mathbf{x} \in [0, 1]^2$ . The realisations from this process are denoted  $\mathbf{u} \sim \mathcal{N}(0, C)$  where  $C_{ij} = k(\mathbf{x}_i, \mathbf{x}_j)$  and k is the Squared Exponential (SE) covariance function with length parameter l:

$$k(\boldsymbol{x}, \boldsymbol{x}') = \exp\left(\frac{-||\boldsymbol{x} - \boldsymbol{x}'||^2}{2l^2}\right) \tag{1}$$

If we specify the latent variables  $\{x_i\}_{i=1}^N$ , then we can compute C and hence fully specify the prior on u. We choose to place  $\{x_i\}_{i=1}^N$  on a  $D \times D$  grid with equal spacing, starting at (0,0) and ending at (1,1). Obviously, we require  $N=D^2$ .

We can now plot the *u*-surface atop this grid by ensuring that for each  $i \in \{1...N\}$ ,  $u_i$  denotes the Z-position and  $x_i$  the X-Y-position. We can then investigate the effect of varying the length-scale parameter l. Three settings of l and the associated plots are given in figure 1 for D = 16.

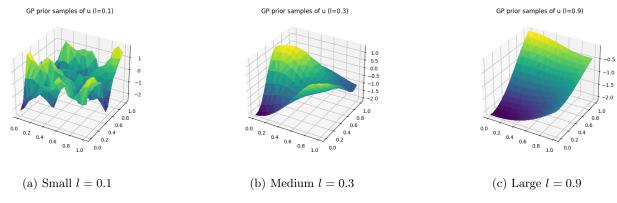


Figure 1: Samples from GP prior for varying length scales (D = 16)

We now proceed to make M random draws (denoted by the  $M \times 1$  vector  $\mathbf{v}$ ) from these samples  $\mathbf{u}$  with additive Gaussian noise  $\boldsymbol{\epsilon} \sim \mathcal{N}(0, I)$ . The subsampling factor f is defined as f := N/M. The draws can be computed as follows:

$$\boldsymbol{v} = G\boldsymbol{u} + \boldsymbol{\epsilon} \tag{2}$$

Where G is an  $M \times N$  matrix with a single one in each row in a random location (without repetition) and rest zeros. The result is that the observations  $\boldsymbol{v}$  are a jumbled subsample of  $\boldsymbol{u}$  with additive noise  $\boldsymbol{\epsilon}$ . We can plot the data overlaid on the original prior samples by simply matching each entry of  $\boldsymbol{v}$  back to the coordinate it was selected from. The result is plotted on figure 2.

#### Simulated data v overlaid onto u surface (I=0.3)

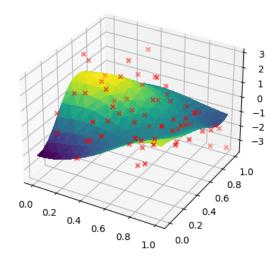


Figure 2: v (red crosses) overlaid on u-surface (f = 4, D = 16, l = 0.3)

We observe  $M = N/f = 16^2/4 = 64$  samples contained in the  $\boldsymbol{v}$  vector. These are equally likely to appear above or below the  $\boldsymbol{u}$ -surface as the noise has zero mean. The noise variance for each data-point is of similar magnitude to the variation in the surface ( $\sigma^2 = 1$ ) so some crosses appear relatively far away from the surface. Moreover, as the subsampling is random, the  $\boldsymbol{v}$ -points appear at randomly chosen (but distinct) locations X-Y plane.

#### b Log probabilities and MCMC

The log prior can be calculated simply, through manipulation of the Gaussian pdf:

$$\ln p(\boldsymbol{u}) = \ln \mathcal{N}(\boldsymbol{u}; 0, C)$$

$$= \ln \frac{1}{(2\pi)^{N/2} |C|^{1/2}} \exp\left(-\frac{1}{2}\boldsymbol{u}^T C^{-1}\boldsymbol{u}\right)$$

$$= -\left(\frac{N}{2} \ln 2\pi + \frac{1}{2} \ln |C| + \frac{1}{2}\boldsymbol{u}^T C^{-1}\boldsymbol{u}\right)$$
(3)

Likewise, v|u is also a Gaussian such that  $v|u \sim \mathcal{N}(Gu, I)$  (see equation 2). By comparison with the form of equation 3, we can jump straight to the log-likelihood, noting that  $\ln |I| = 0$ :

$$\ln p(\boldsymbol{v}|\boldsymbol{u}) = -\left(\frac{M}{2}\ln 2\pi + \frac{1}{2}\left(\boldsymbol{v} - G\boldsymbol{u}\right)^{T}\left(\boldsymbol{v} - G\boldsymbol{u}\right)\right)$$
(4)

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