Heteroskedasticity, Part 2

EC 421, Set 5

Luciana Etcheverry October 15, 2019

Prologue

Schedule

Last Time

Heteroskedasticity: Issues and tests

Today

Living with heteroskedasticity

EC 421

Goals

- Develop intuition for econometrics.
- Learn how to *apply* econometrics—strengths, weaknessed, etc.
- Learn R.

R does the calculations and has already memorized the formulas.

I want you to know what the formulas mean, when/why we use them, and when they fail/work.

This course has the potential to be one of the most useful/valuable/applicable/marketable classes that you take at UO.

Review

Review

Three review questions

Question 1: What is the difference between u_i and e_i ?

Question 2: We spend a lot of time discussing u_i^2 . Why?

Question 3: We also spend a lot of time discussing e_i^2 . Why?

Review

Question 1: What is the difference between u_i and e_i ?

Answer 1:

 u_i gives the **population disturbance** for the i^{th} observation. u_i measures how far the i^{th} observation is from the **population** line, *i.e.*,

$$u_i = y_i - \underbrace{(eta_0 + eta_1 x_i)}_{ ext{Population line}}$$

 e_i gives the **regression residual (error)** for the i^{th} observation. e_i measures how far the i^{th} observation is from the **regression** line, *i.e.*,

$$e_i = y_i - \underbrace{\left(\hat{eta}_0 + \hat{eta}_1 x_i
ight)}_{ ext{Regression line} = \hat{y}} = y_i - \hat{y}_i$$

Review

Question 2: We spend a lot of time discussing u_i^2 . Why?

Answer 2:

One of major assumptions is that our disturbances (the u_i 's) are homoskedastic (they have constant variance), i.e., $\mathrm{Var}(u_i|x_i)=\sigma^2$.

We also assume that the mean of these disturbances is zero, $[u_i|x_i]=0$.

By definition,
$$\operatorname{Var}(u_i|x_i) = \begin{bmatrix} u_i^2 - \underbrace{\begin{bmatrix} u_i|x_i\end{bmatrix}^2} |x_i\end{bmatrix} = \begin{bmatrix} u_i^2|x_i\end{bmatrix}$$

Thus, if we want to learn about the variance of u_i , we can focus on u_i^2 .

Review

Question 3: We also spend a lot of time discussing e_i^2 . Why?

Answer 3:

We cannot observe u_i (or u_i^2).

But u_i^2 tells us about the variance of u_i .

We use e_i^2 to learn about u_i^2 and, consequently, σ_i^2 .

Review: Current assumptions

- 1. Our sample (the x_k 's and y_i) was **randomly drawn** from the population.
- 2. y is a **linear function** of the β_k 's and u_i .
- 3. There is no perfect **multicollinearity** in our sample.
- 4. The explanatory variables are **exogenous**: [u|X] = 0 (\Longrightarrow [u] = 0).
- 5. The disurbances have **constant variance** σ^2 and **zero covariance**, *i.e.*,
 - $egin{array}{ccc} \circ & \left[u_i^2 | X_i
 ight] = \mathrm{Var}(u_i | X_i) = \sigma^2 \implies \mathrm{Var}(u_i) = \sigma^2 \end{array}$
 - $\circ \ \operatorname{Cov}ig(u_i,\, u_j|X_i,\, X_jig) = \ ig[u_iu_j|X_i,\, X_jig] = 0$ for i
 eq j
- 6. The disturbances come from a **Normal** distribution, *i.e.*, $u_i \stackrel{ ext{iid}}{\sim} \mathrm{N}(0,\sigma^2)$.

Review

Today we're focusing on assumption #5:

5. The disurbances have **constant variance** σ^2 and **zero covariance**, *i.e.*,

$$ullet \left[u_i^2 | X_i
ight] = \mathrm{Var}(u_i | X) = \sigma^2 \implies \mathrm{Var}(u_i) = \sigma^2$$

$$ullet$$
 $\operatorname{Cov}ig(u_i,\,u_j|X_i,\,X_jig) = ig[u_iu_j|X_i,\,X_jig] = 0$ for $i
eq j$

Specifically, we will focus on the assumption of **constant variance** (also known as *homoskedasticity*).

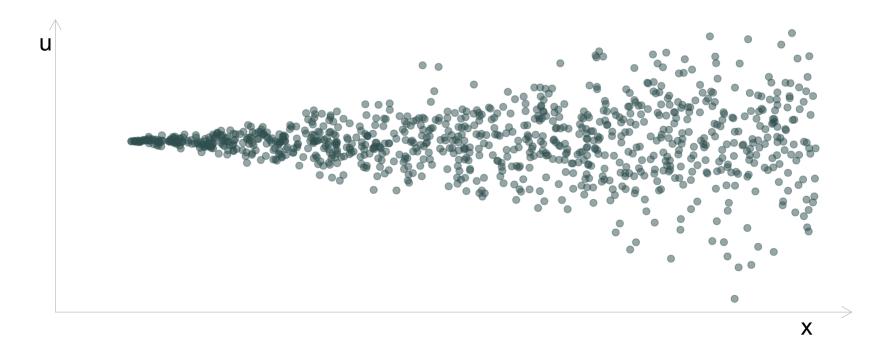
Violation of this assumption: Our disturbances have different variances.

Heteroskedasticity: $\mathrm{Var}(u_i) = \sigma_i^2$ and $\sigma_i^2 \neq \sigma_i^2$ for some $i \neq j$.

Review

Classic example of heteroskedasticity: The funnel

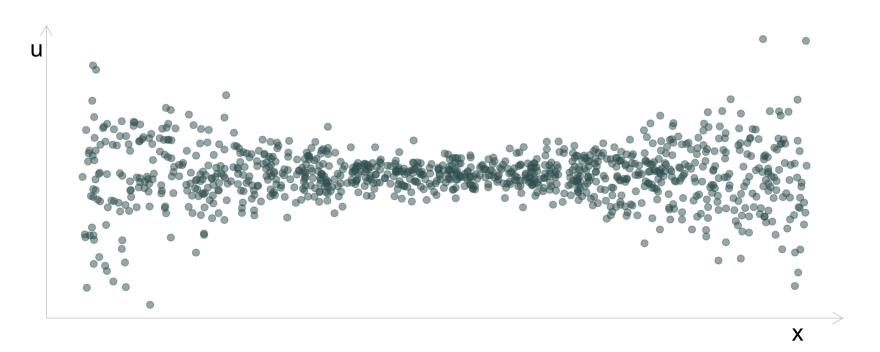
Variance of u increases with x



Review

Another example of heteroskedasticity: (double funnel?)

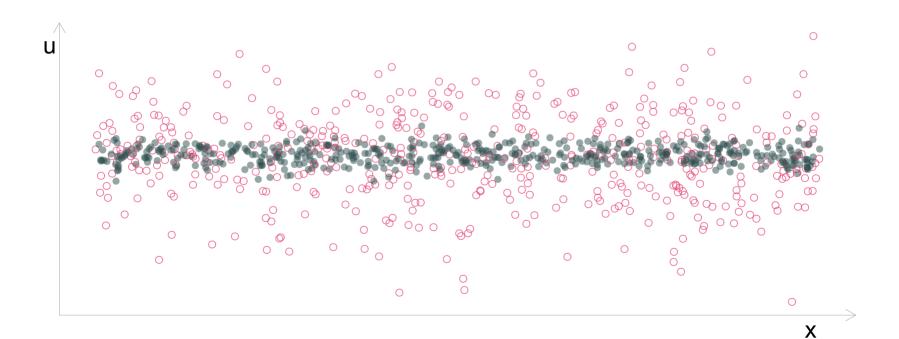
Variance of u increasing at the extremes of x



Review

Another example of heteroskedasticity:

Differing variances of u by group



Review

Heteroskedasticity is present when the variance of u changes with any combination of our explanatory variables x_1 through x_k .

Testing for heteroskedasticity

We have some tests that may help us detect heteroskedasticity.

- Goldfeld-Quandt
- Breusch-Pagan
- White

What do we do if we detect it?

In the presence of heteroskedasticity, OLS is

- still unbiased
- no longer the most efficient unbiased linear estimator

On average, we get the right answer but with more noise (less precision). *Also*: Our standard errors are biased.

Options:

- 1. Check regression **specification**.
- 2. Find a new, more efficient **unbiased estimator** for β_j 's.
- 3. Live with OLS's inefficiency; find a **new variance estimator**.
 - Standard errors
 - Confidence intervals
 - Hypothesis tests

Misspecification

As we've discussed, the specification[†] of your regression model matters a lot for the unbiasedness and efficiency of your estimator.

Response #1: Ensure your specification doesn't cause heteroskedasticity.

Misspecification

Example: Let the population relationship be

$$y_i=eta_0+eta_1x_i+eta_2x_i^2+u_i$$

with $[u_i|x_i]=0$ and $\mathrm{Var}(u_i|x_i)=\sigma^2$.

However, we omit x^2 and estimate

$$y_i = \gamma_0 + \gamma_1 x_i + w_i$$

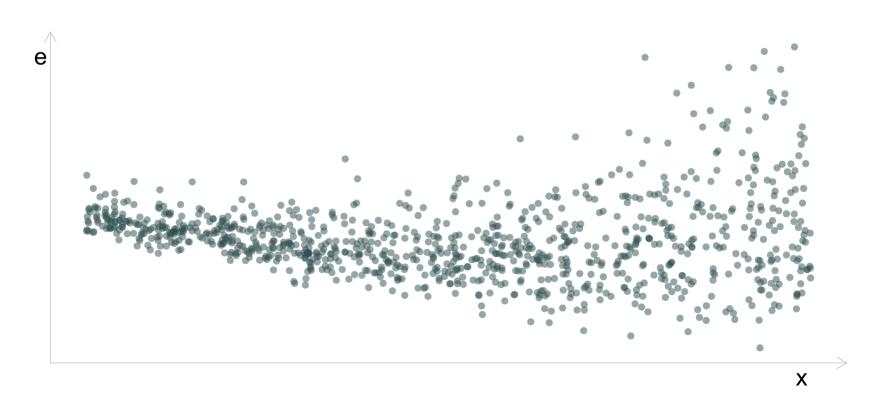
Then

$$w_i = u_i + eta_2 x_i^2 \implies ext{Var}(w_i) = f(x_i)$$

I.e., the variance of w_i changes systematically with x_i (heteroskedasticity).

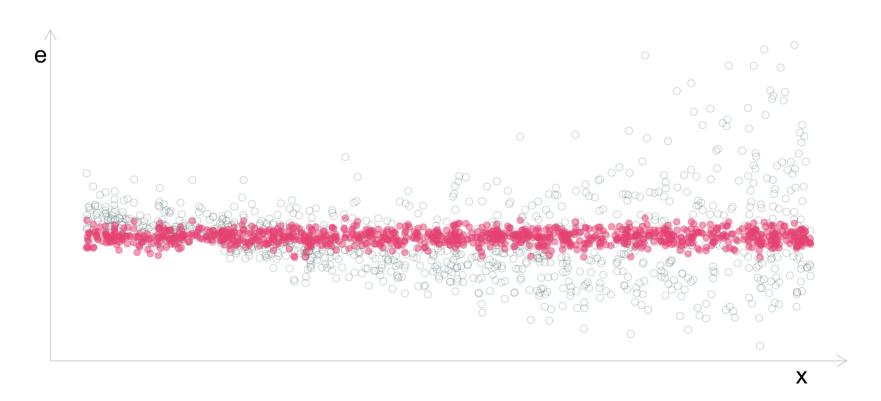
Misspecification

Truth: $\log(y_i) = \beta_0 + \beta_1 x_i + u_i$ Misspecification: $y_i = \beta_0 + \beta_1 x_i + v_i$



Misspecification

Truth: $\log(y_i) = \beta_0 + \beta_1 x_i + u_i$ Misspecification: $y_i = \beta_0 + \beta_1 x_i + v_i$



Misspecification

More generally:

Misspecification problem: Incorrect specification of the regression model can cause heteroskedasticity (among other problems).

Solution: \bigcirc Get it right (e.g., don't omit x^2).

New problems:

- We often don't know the *right* specification.
- We'd like a more formal process for addressing heteroskedasticity.

Conclusion: Specification often will not "solve" heteroskedasticity. However, correctly specifying your model is still really important.

Weighted least squares

Weighted least squares (WLS) presents another approach.

Response #2: Increase efficiency by weighting our observations.

Let the true population relationship be

$$y_i = \beta_0 + \beta_1 x_i + u_i \tag{1}$$

with $u_i \sim N(0,\,\sigma_i^2)$.

Now transform (1) by dividing each observation's data by σ_i , i.e.,

Weighted least squares

$$y_i = \beta_0 + \beta_1 x_i + u_i \tag{1}$$

$$\frac{y_i}{\sigma_i} = \beta_0 \frac{1}{\sigma_i} + \beta_1 \frac{x_i}{\sigma_i} + \frac{u_i}{\sigma_i} \tag{2}$$

Whereas (1) is heteroskedastic, (2) is homoskedastic.

 \therefore OLS is efficient and unbiased for estimating the eta_k in (2)!

Why is (2) homoskedastic?

$$\operatorname{Var}igg(rac{u_i}{\sigma_i}\Big|x_iigg) = rac{1}{\sigma_i^2}\operatorname{Var}(u_i|x_i) = rac{1}{\sigma_i^2}\sigma_i^2 = 1$$

Weighted least squares

WLS is great, but we need to know σ_i^2 , which is generally unlikely.

We can slightly relax this requirement—instead requiring

1.
$$\operatorname{Var}(u_i|x_i) = \sigma_i^2 = \sigma^2 h(x_i)$$

2. We know h(x).

As before, we transform our heteroskedastic model into a homoskedastic model. This time we divide each observation's data[†] by $\sqrt{h(x_i)}$.

[†] Divide all of the data by $\sqrt{h(x_i)}$, including the intercept.

Weighted least squares

$$y_i = \beta_0 + \beta_1 x_i + u_i \tag{1}$$

$$rac{y_i}{\sqrt{h(x_i)}} = eta_0 rac{1}{\sqrt{h(x_i)}} + eta_1 rac{x_i}{\sqrt{h(x_i)}} + rac{u_i}{\sqrt{h(x_i)}} \qquad (2)$$

with $\operatorname{Var}(u_i|x_i) = \sigma^2 h(x_i)$.

Now let's check that (2) is indeed homoskedastic.

$$\operatorname{Var}\!\left(rac{u_i}{\sqrt{h(x_i)}}igg|x_i
ight) = rac{1}{h(x_i)}\operatorname{Var}(u_i|x_i) = rac{1}{h(x_i)}\sigma^2 h(x_i) = rac{\sigma^2}{\sigma^2}$$

Homoskedasticity!

Weighted least squares

Weighted least squares (WLS) estimators are a special class of **generalized least squares** (GLS) estimators focused on heteroskedasticity.

$$y_i = eta_0 + eta_1 x_{1i} + u_i \quad ext{ vs. } \quad rac{y_i}{\sigma_i} = eta_0 rac{1}{\sigma_i} + eta_1 rac{x_{1i}}{\sigma_i} + rac{u_i}{\sigma_i} .$$

Notes:

- 1. WLS **transforms** a heteroskedastic model into a homoskedastic model.
- 2. **Weighting:** WLS downweights observations with higher variance u_i 's.
- 3. **Big requirement:** WLS requires that we *know* σ_i^2 for each observation.
- 4. WLS is generally **infeasible**. Feasible GLS (FGLS) offers a solution.
- 5. Under its assumptions: WLS is the **best linear unbiased estimator**.

Heteroskedasticity-robust standard errors

Response #3:

- Ignore OLS's inefficiency (in the presence of heteroskedasticity).
- Focus on unbiased estimates for our standard errors.
- In the process: Correct inference.

Q: What is a standard error?

A: The standard deviation of an estimator's distribution.

Estimators (like $\hat{\beta}_1$) are random variables, so they have distributions.

Standard errors give us a sense of how much variability is in our estimator.

Heteroskedasticity-robust standard errors

Recall: We can write the OLS estimator for β_1 as

$$\hat{\beta}_{1} = \beta_{1} + \frac{\sum_{i} (x_{i} - \overline{x}) u_{i}}{\sum_{i} (x_{i} - \overline{x})^{2}} = \beta_{1} + \frac{\sum_{i} (x_{i} - \overline{x}) u_{i}}{SST_{x}}$$
(3)

Let $\operatorname{Var}(u_i|x_i)=\sigma_i^2$.

We can use (3) to write the variance of $\hat{\beta}_1$, i.e.,

$$\operatorname{Var}(\hat{\beta}_1 | x_i) = \frac{\sum_i (x_i - \overline{x})^2 \sigma_i^2}{\operatorname{SST}_x^2} \tag{4}$$

Heteroskedasticity-robust standard errors

If we want unbiased estimates for our standard errors, we need an unbiased estimate for

$$\frac{\sum_{i}\left(x_{i}-\overline{x}\right)^{2}\sigma_{i}^{2}}{\mathrm{SST}_{x}^{2}}$$

Our old friend Hal White provided such an estimator:

$$\widehat{ ext{Var}}\left(\hat{eta}_1
ight) = rac{\sum_i \left(x_i - \overline{x}
ight)^2 e_i^2}{ ext{SST}_x^2}$$

where the e_i comes from the OLS regression of interest.

[†] This specific equation is for simple linear regression.

Heteroskedasticity-robust standard errors

Our heteroskedasticity-robust estimators for the standard error of β_i .

Case 1 Simple linear regression, $y_i = \beta_0 + \beta_1 x_i + u_i$

$$\widehat{ ext{Var}}\left(\hat{eta}_{1}
ight) = rac{\sum_{i}\left(x_{i}-\overline{x}
ight)^{2}e_{i}^{2}}{ ext{SST}_{x}^{2}}$$

Case 2 Multiple (linear) regression, $y_i = eta_0 + eta_1 x_{1i} + \dots + eta_k x_{ki} + u_i$

$$\widehat{ ext{Var}}\left(\hat{eta}_{j}
ight) = rac{\sum_{i}\hat{r}_{ij}^{2}e_{i}^{2}}{ ext{SST}_{x_{i}^{2}}}$$

where \hat{r}_{ij} denotes the ith residual from regressing x_j on all other explanatory variables.

Heteroskedasticity-robust standard errors

With these standard errors, we can return to correct statistical inferencel

E.g., we can update our previous t statistic formula with our new heteroskedasticity-robust standard erros.

$$t = \frac{\text{Estimate} - \text{Hypothesized value}}{\text{Standard error}}$$

Heteroskedasticity-robust standard errors

Notes

- We are still using **OLS estimates for** β_i
- Our het.-robust standard errors use a different estimator.
- Homoskedasticity
 - Plain OLS variance estimator is more efficient.
 - Het.-robust is still unbiased.
- Heteroskedasticity
 - Plain OLS variance estimator is biased.
 - Het.-robust variance estimator is unbiased.

Heteroskedasticity-robust standard errors

These standard errors go by many names

- Heteroskedasticity-robust standard errors
- Het.-robust standard errors
- White standard errors
- Eicker-White standard errors
- Huber standard errors
- Eicker-Huber-White standards errors
- (some other combination of Eicker, Huber, and White)

Do not say: "Robust standard errors". The problem: "robust" to what?

Examples

Examples

Back to our test-scores dataset...

```
# Load packages
library(pacman)
p_load(tidyverse, Ecdat)
# Select and rename desired variables; assign to new dataset; format as tibble
test_df 		 Caschool %>% select(
    test_score = testscr, ratio = str, income = avginc, enrollment = enrltot
) %>% as_tibble()
# View first 2 rows of the dataset
head(test_df, 2)
```

```
#> # A tibble: 2 x 4
#> test_score ratio income enrollment
#> <dbl> <dbl> <dbl> <int>
#> 1 691. 17.9 22.7 195
#> 2 661. 21.5 9.82 240
```

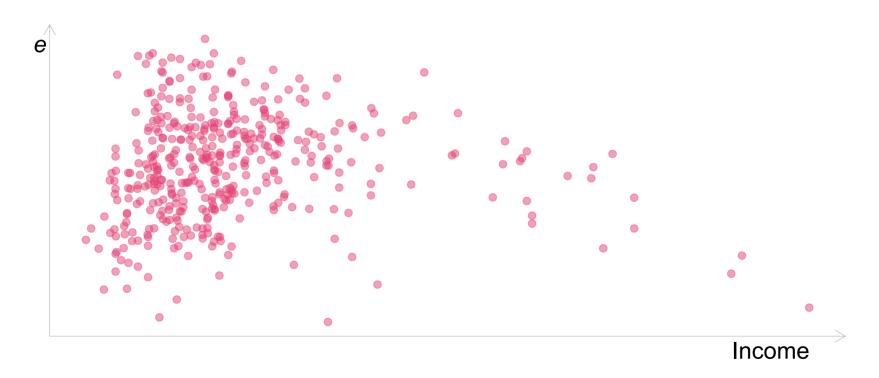
Example: Model specification

We found significant evidence of heteroskedasticity.

Let's check if it was due to misspecifying our model.

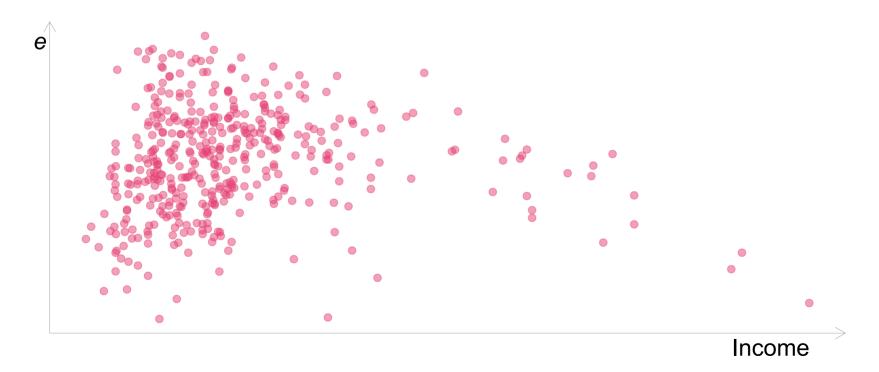
Example: Model specification

```
	ext{Model}_1: \mathbf{Score}_i = eta_0 + eta_1 \mathbf{Ratio}_i + eta_2 \mathbf{Income}_i + u_i 	ext{lm(test\_score} \sim \mathsf{ratio} + \mathsf{income}, \; \mathsf{data} = \mathsf{test\_df})
```



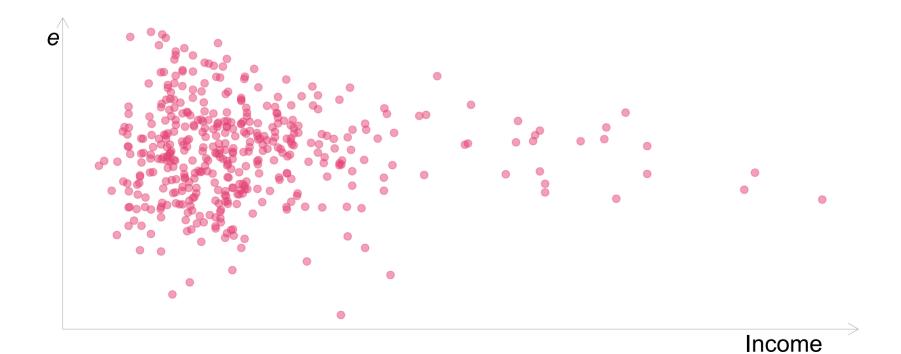
Example: Model specification

```
Model_2: log(Score_i) = \beta_0 + \beta_1 Ratio_i + \beta_2 Income_i + u_i lm(log(test\_score) \sim ratio + income, data = test\_df)
```



Example: Model specification

```
Model<sub>3</sub>: \log(\text{Score}_i) = \beta_0 + \beta_1 \text{Ratio}_i + \beta_2 \log(\text{Income}_i) + u_i
\text{lm(log(test\_score)} \sim \text{ratio} + \log(\text{income}), \text{ data = test\_df)}
```



Example: Model specification

Let's test this new specification with the White test for heteroskedasticity.

$$\mathsf{Model}_3: \mathsf{log}(\mathsf{Score}_i) = \beta_0 + \beta_1 \mathsf{Ratio}_i + \beta_2 \mathsf{log}(\mathsf{Income}_i) + u_i$$

The regression for the White test

$$e_i^2 = lpha_0 + lpha_1 ext{Ratio}_i + lpha_2 \log(ext{Income}_i) + lpha_3 ext{Ratio}_i^2 + lpha_4 (\log(ext{Income}_i))^2 \ + lpha_5 \left(ext{Ratio}_i imes \log(ext{Income}_i)
ight) + v_i$$

yields $R_e^2 pprox 0.029$ and test statistic of $\widehat{
m LM} = n imes R_e^2 pprox 12.2$.

Under H₀, LM is distributed as $\chi_5^2 \implies p$ -value \approx 0.033.

∴ **Reject H₀. Conclusion:** There is statistically significant evidence of heteroskedasticity at the five-percent level.

Example: Model specification

Okay, we tried adjusting our specification, but there is still evidence of heteroskedasticity.

Next: In general, you will turn to heteroskedasticity-robust standard errors.

- OLS is still unbiased for the **coefficients** (the β_i 's)
- Heteroskedasticity-robust standard errors are unbiased for the standard errors of the $\hat{\beta}_j$'s, i.e., $\sqrt{\mathrm{Var} \Big(\hat{\beta}_j \Big)}$.

Example: Het.-robust standard errors

Let's return to our model

$$Score_i = \beta_0 + \beta_1 Ratio_i + \beta_2 Income_i + u_i$$

We can use the lfe package in **R** to calculate standard errors.

Example: Het.-robust standard errors

$$Score_i = \beta_0 + \beta_1 Ratio_i + \beta_2 Income_i + u_i$$

1. Run the regression with felm() (instead of lm())

```
# Load 'lfe' package
p_load(lfe)

#> package 'zoo' successfully unpacked and MD5 sums checked
#> package 'Formula' successfully unpacked and MD5 sums checked
#> package 'sandwich' successfully unpacked and MD5 sums checked
#> package 'lfe' successfully unpacked and MD5 sums checked
#> The downloaded binary packages are in
#> C:\Users\Luciana\AppData\Local\Temp\RtmpYFn5aa\downloaded_packages
```

```
# Regress log score on ratio and log income
test_reg ← felm(test_score ~ ratio + income, data = test_df)
```

Example: Het.-robust standard errors

$$\mathrm{Score}_i = eta_0 + eta_1 \mathrm{Ratio}_i + eta_2 \mathrm{Income}_i + u_i$$

2. Estimate het.-robust standard errors with robust = T option in summary()

Example: Het.-robust standard errors

Ceofficients and heteroskedasticity-robust standard errors:

```
#> Estimate Robust s.e t value Pr(>|t|)
#> (Intercept) 638.7292    7.3012   87.482   <2e-16 ***
#> ratio         -0.6487    0.3533   -1.836   0.0671 .
#> income         1.8391   0.1147   16.029   <2e-16 ***</pre>
```

Ceofficients and plain OLS standard errors (assumes homoskedasticity):

Example: WLS

We mentioned that WLS is often not possible—we need to know the functional form of the heteroskedasticity—either

A.
$$\sigma_i^2$$

or

B.
$$h(x_i)$$
, where $\sigma_i^2 = \sigma^2 h(x_i)$

There are occasions in which we can know $h(x_i)$.

Example: WLS

Imagine individuals in a population have homoskedastic disturbances.

However, instead of observing individuals' data, we observe (in data) groups' averages (e.g., cities, counties, school districts).

If these groups have different sizes, then our dataset will be heteroskedastic—in a predictable fashion.

Recall: The variance of the sample mean depends upon the sample size,

$$\operatorname{Var}ig(\overline{x}ig) = rac{\sigma_x^2}{n}$$

Example: Our school testing data is averaged at the school level.

Example: WLS

Example: Our school testing data is averaged at the school level.

Even if individual students have homoskedastic disturbances, the schools would have heteroskedastic disturbances, *i.e.*,

Individual-level model:
$$Score_i = \beta_0 + \beta_1 Ratio_i + \beta_2 Income_i + u_i$$

School-level model:
$$\overline{\text{Score}}_s = \beta_0 + \beta_1 \overline{\text{Ratio}}_s + \beta_2 \overline{\text{Income}}_s + \overline{u}_s$$

where the s subscript denotes an individual school (just as i indexes an individual person).

$$\mathrm{Var}ig(\overline{u}_sig) = rac{\sigma^2}{n_s}$$

Example: WLS

For WLS, we're looking for a function $h(x_s)$ such that $\mathrm{Var}ig(\overline{u}_s|x_sig)=\sigma^2 h(x_s)$.

We just showed
$†$
 that $\mathrm{Var}ig(\overline{u}_s|x_sig)=rac{\sigma^2}{n_s}$.

Thus, $h(x_s) = 1/n_s$, where n_s is the number of students in school s.

To implement WLS, we divide each observation's data by $1/\sqrt{h(x_s)}$, meaning we need to multiply each school's data by $\sqrt{n_s}$.

The variable enrollment in the test_df dataset is our n_s .

[†] Assuming the individuals' disturbances are homoskedastic.

Example: WLS

Using WLS to estimate $Score_i = \beta_0 + \beta_1 Ratio_i + \beta_2 Income_i + u_i$

Step 1: Multiply each variable by $1/\sqrt{h(x_i)} = \sqrt{\mathrm{Enrollment}_i}$

```
# Create WLS transformed variables, multiplying by sqrt of 'pop'
test_df ← mutate(test_df,
   test_score_wls = test_score * sqrt(enrollment),
   ratio_wls = ratio * sqrt(enrollment),
   income_wls = income * sqrt(enrollment),
   intercept_wls = 1 * sqrt(enrollment)
)
```

Notice that we are creating a transformed intercept.

Example: WLS

Using WLS to estimate $Score_i = \beta_0 + \beta_1 Ratio_i + \beta_2 Income_i + u_i$

Step 2: Run our WLS (transformed) regression

```
# WLS regression
wls_reg ← lm(
  test_score_wls ~ -1 + intercept_wls + ratio_wls + income_wls,
  data = test_df
)
```

Note: The -1 in our regression tells **R** not to add an intercept, since we are adding a transformed intercept (intercept_wls).

Example: WLS

The WLS estimates and standard errors:

```
#> Estimate Std. Error t value Pr(>|t|)
#> intercept_wls 618.78331 8.26929 74.829 <2e-16 ***
#> ratio_wls -0.21314 0.37676 -0.566 0.572
#> income_wls 2.26493 0.09065 24.985 <2e-16 ***</pre>
```

The **OLS estimates** and **het.-robust standard errors**:

```
#> Estimate Robust s.e t value Pr(>|t|) #> (Intercept) 638.7292 7.3012 87.482 <2e-16 *** #> ratio -0.6487 0.3533 -1.836 0.0671 . #> income 1.8391 0.1147 16.029 <2e-16 ***
```

Example: WLS

Alternative to doing your own weighting: feed lm() some weights.

```
lm(test_score ~ ratio + income, data = test_df, weights = enrollment)
```

In this example

- **Heteroskedasticity-robust standard errors** did not change our standard errors very much (relative to plain OLS standard errors).
- WLS changed our answers a bit—coefficients and standard errors.

These examples highlighted a few things:

- 1. Using the correct estimator for your standard errors really matters. †
- 2. Econometrics doesn't always offer an obviously correct route.

To see #1, let's run a simulation.

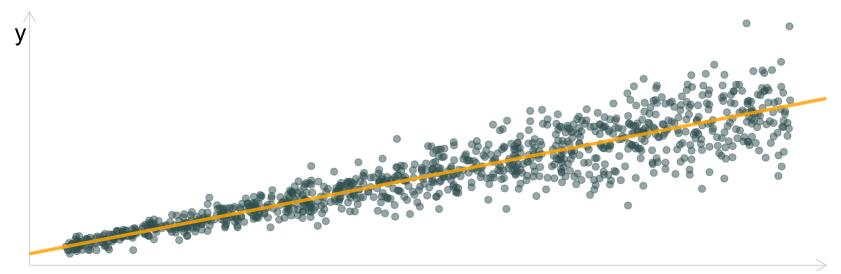
[†] Sit in on an economics seminar, and you will see what I mean.

Simulation

Let's examine a simple linear regression model with heteroskedasticity.

$$y_i = \underbrace{eta_0}_{=1} + \underbrace{eta_1}_{=10} x_i + u_i$$

where $\operatorname{Var}(u_i|x_i) = \sigma_i^2 = \sigma^2 x_i^2$.

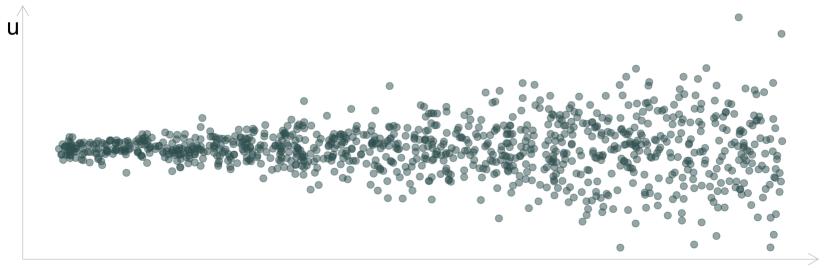


Simulation

Let's examine a simple linear regression model with heteroskedasticity.

$$y_i = \underbrace{eta_0}_{=1} + \underbrace{eta_1}_{=10} x_i + u_i$$

where $\operatorname{Var}(u_i|x_i) = \sigma_i^2 = \sigma^2 x_i^2$.



Simulation

Note regarding WLS:

Since
$$\operatorname{Var}(u_i|x_i) = \sigma^2 x_i^2$$
,

$$\operatorname{Var}(u_i|x_i) = \sigma^2 h(x_i) \implies h(x_i) = x_i^2$$

WLS multiplies each variable by $1/\sqrt{h(x_i)}=1/x_i$.

Simulation

In this simulation, we want to compare

- 1. The **efficiency** of
 - OLS
 - \circ WLS with correct weights: $h(x_i) = x_i$
 - \circ WLS with incorrect weights: $h(x_i) = \sqrt{x_i}$
- 2. How well our **standard errors** perform (via confidence intervals) with
 - Plain OLS standard errors
 - Heteroskedasticity-robust standard errors
 - WLS standard errors

Simulation

The simulation plan:

```
Do 10,000 times:
```

- 1. Generate a sample of size 30 from the population
- 2. Calculate/save OLS and WLS (\times 2) estimates for β_1
- 3. Calculate/save standard errors for β_1 using
 - Plain OLS standard errors
 - Heteroskedasticity-robust standard errors
 - WLS (correct)
 - WLS (incorrect)

Simulation

For one iteration of the simulation:

Code to generate the data...

```
# Parameters
b0 ← 1
b1 ← 10
s2 ← 1
# Sample size
n ← 30
# Generate data
sample_df ← tibble(
    x = runif(n, 0.5, 1.5),
    y = b0 + b1 * x + rnorm(n, 0, sd = s2 * x^2)
)
```

Simulation

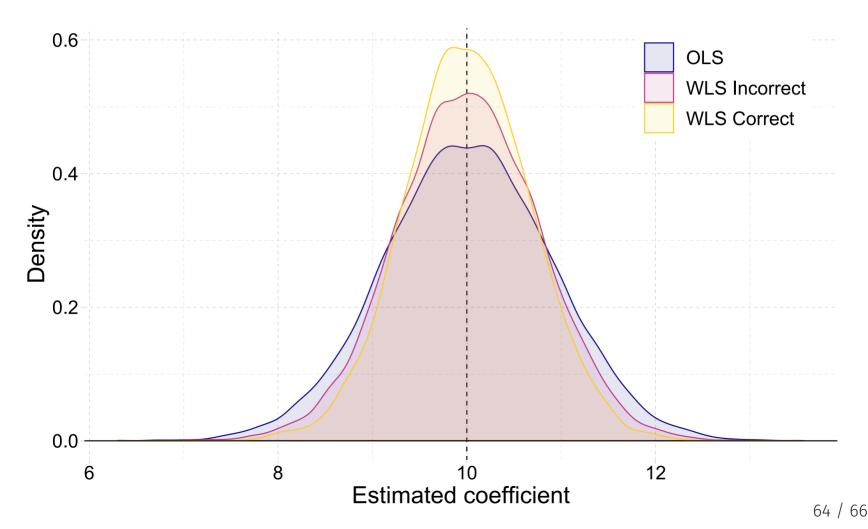
For one iteration of the simulation:

Code to estimate our coefficients and standard errors...

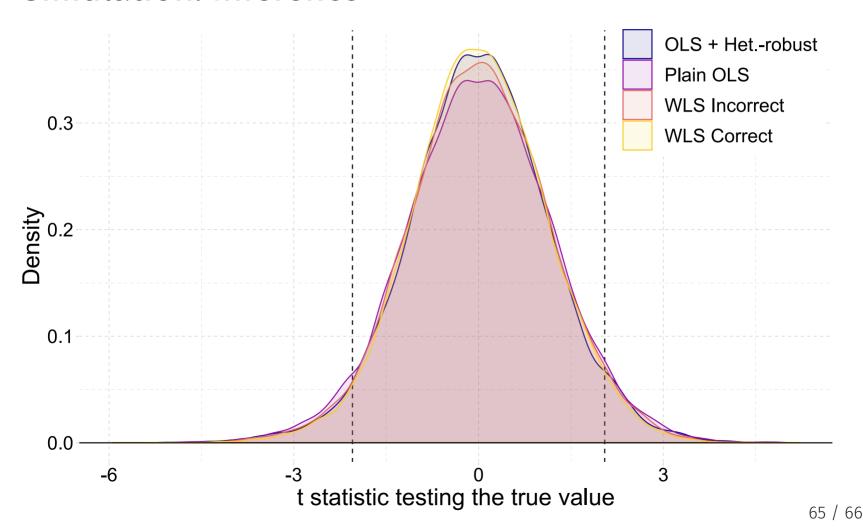
```
# OLS
ols ← felm(y ~ x, data = sample_df)
# WLS: Correct weights
wls_t ← lm(y ~ x, data = sample_df, weights = 1/x^2)
# WLS: Correct weights
wls_f ← lm(y ~ x, data = sample_df, weights = 1/x)
# Coefficients and standard errors
summary(ols, robust = F)
summary(ols, robust = T)
summary(wls_t)
summary(wls_f)
```

Then save the results.

Simulation: Coefficients



Simulation: Inference



Simulation: Results

Summarizing our simulation results (10,000 iterations)

Estimation: Summary of $\hat{\beta}_1$'s

Estimator	Mean	S.D.
OLS	10.014	0.885
WLS Correct	10.012	0.667
WLS Incorrect	10.013	0.757

Inference: % of times we reject β_1

Estimators	% Reject
OLS + Hetrobust	7.3
OLS + Homosk.	8.5
WLS Correct	6.1
WLS Incorrect	7.2