Time series

EC 421, Set 7

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Prologue

Schedule

Last Time

Asymptotics, probability limits, and consistency

Today

Time series

Upcoming

Assignment, survey, midterm

Survey

These is your opportunity to provide feedback about your learning experience in this class

Please give it 5 minutes at home to complete it

DuckWeb -> Course Surveys

About our class

1. EC 421 is a **hard class**.

2. EC 421 requires **more math/theory** than most other classes.

3. This **theory is important**—why/when you can trust OLS/regression.

4. With all of this theory, we get **fewer traditional examples**. Proofs and simulations *are* our examples.

5. Midterm will **mix theory, intuition, and application**.

Example questions

Theory

In our proof of the consistency of the OLS estimator for β_1 (for simple linear regression), we got to the point where we had

$$\operatorname{\mathbf{plim}} \hat{\beta}_1 = \beta_1 + \frac{\mathbf{Cov}(x_1,\, u)}{\mathbf{Var}(x_1)} \tag{1}$$

What does the right-hand side of (1) need to simplify to for the OLS estimator $\hat{\beta}_1$ to be consistent?

Example questions

Intuition

We've shown that omitted variables can cause OLS to be biased and inconsistent.

1. What are the two requirements for an omitted variable to cause bias/inconsistency in OLS?

2. Provide an example of a regression that would suffer from omitted variable bias. Explain why it could be biased.

3. Does leaving out a variable from a regression **always** bias OLS? Explain your answer.

Example questions

Application

Your friend is concerned about heteroskedasticity in the regression below.

$$y_i = \hat{\beta}_0 + \hat{\beta}_1 x_{1i} + \hat{\beta}_2 x_{2i} + e_i \tag{2}$$

$$e_i^2 = \hat{\beta}_0 + \hat{\beta}_1 x_{1i} + \hat{\beta}_2 x_{2i} + v_i \tag{3}$$

$$e_i^2 = \hat{\beta}_0 + \hat{\beta}_1 x_{1i} + \hat{\beta}_2 x_{2i} + \hat{\beta}_3 x_{1i}^2 + \hat{\beta}_4 x_{2i}^2 + \hat{\beta}_5 x_{1i} x_{2i} + w_i \tag{4}$$

Because you are such a great friend, you estimated regressions $(\mathbf{3})$ and $(\mathbf{4})$.

The regression in $(\mathbf{3})$ has and R^2 of 0.20, and the regression in $(\mathbf{4})$ has and R^2 of 0.30. You have 100 observations.

1. Calculate the Breusch-Pagan test statistic testing heterosk. in $(\mathbf{1})$.
2. The critical value for the Breusch-Pagan test is 6. Finish the B-P test (state your hypotheses; determine your conclusion).

Asymptotics and consistency

Review

Asymptotics and consistency

Review

1. Compare/contrast the concepts *expected value* and *probability limit*.

2. What does it mean if the estimator $\hat{\theta}$ is consistent for θ?

3. What is required for an omitted variable to bias OLS estimates of β_j?

4. Does omitted-variable bias affect the consistency of OLS for β_j?

5. What can we know about the direction of omitted-variable bias?

6. How does measurement error in an explanatory variable affect the OLS estimate for that variable's effect on the outcome variable?

7. How does measurement error in an outcome variable affect OLS?

Time-series data

Time-series data

Introduction

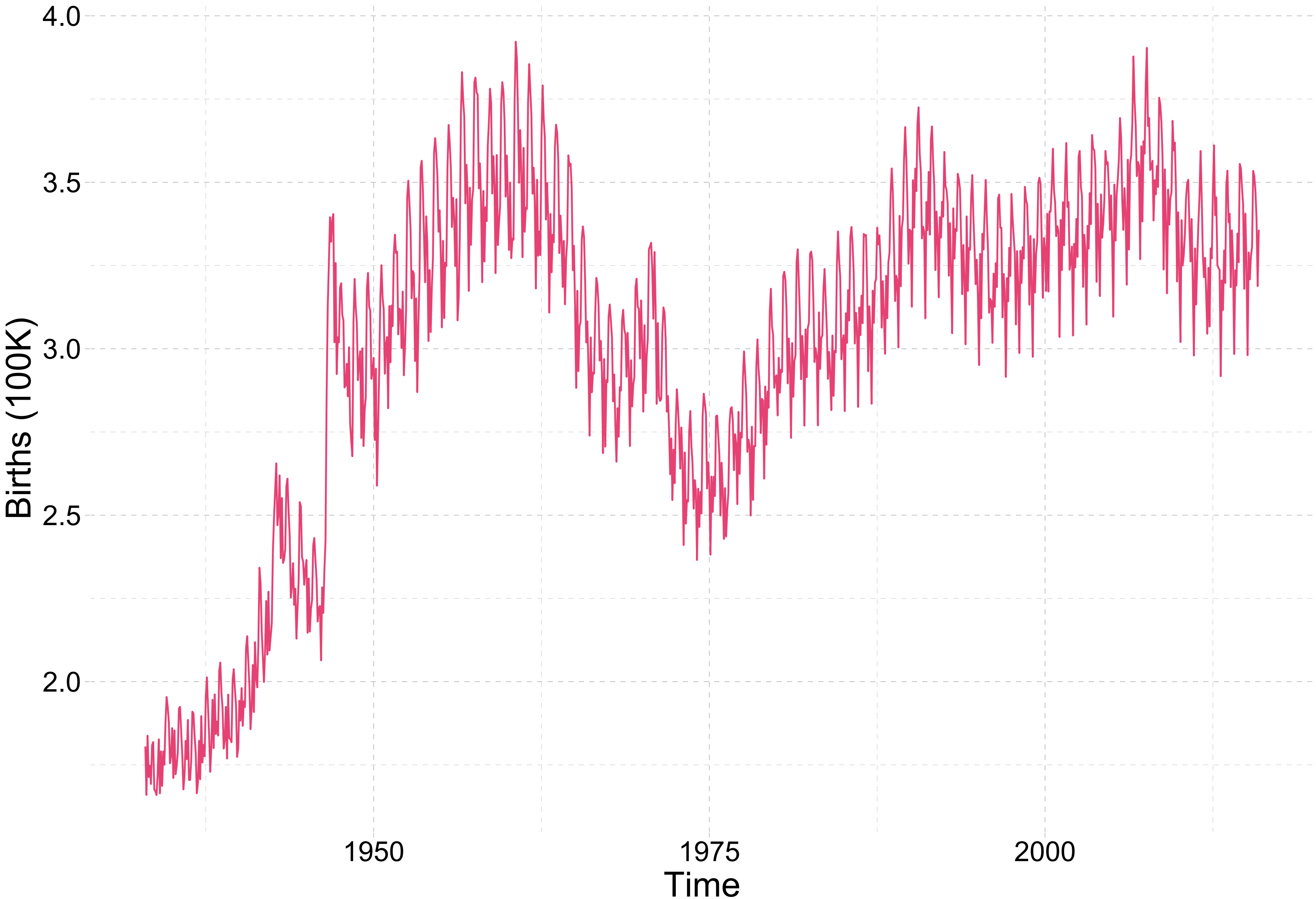
Up to this point, we focused on **cross-sectional data**.

- Sampled *across* a population (*e.g.*, people, counties, countries).
- Sampled at *one moment* in time (*e.g.*, Jan. 1, 2015).
- We had n *individuals*, each indexed i in $\{1, \ldots, n\}$.

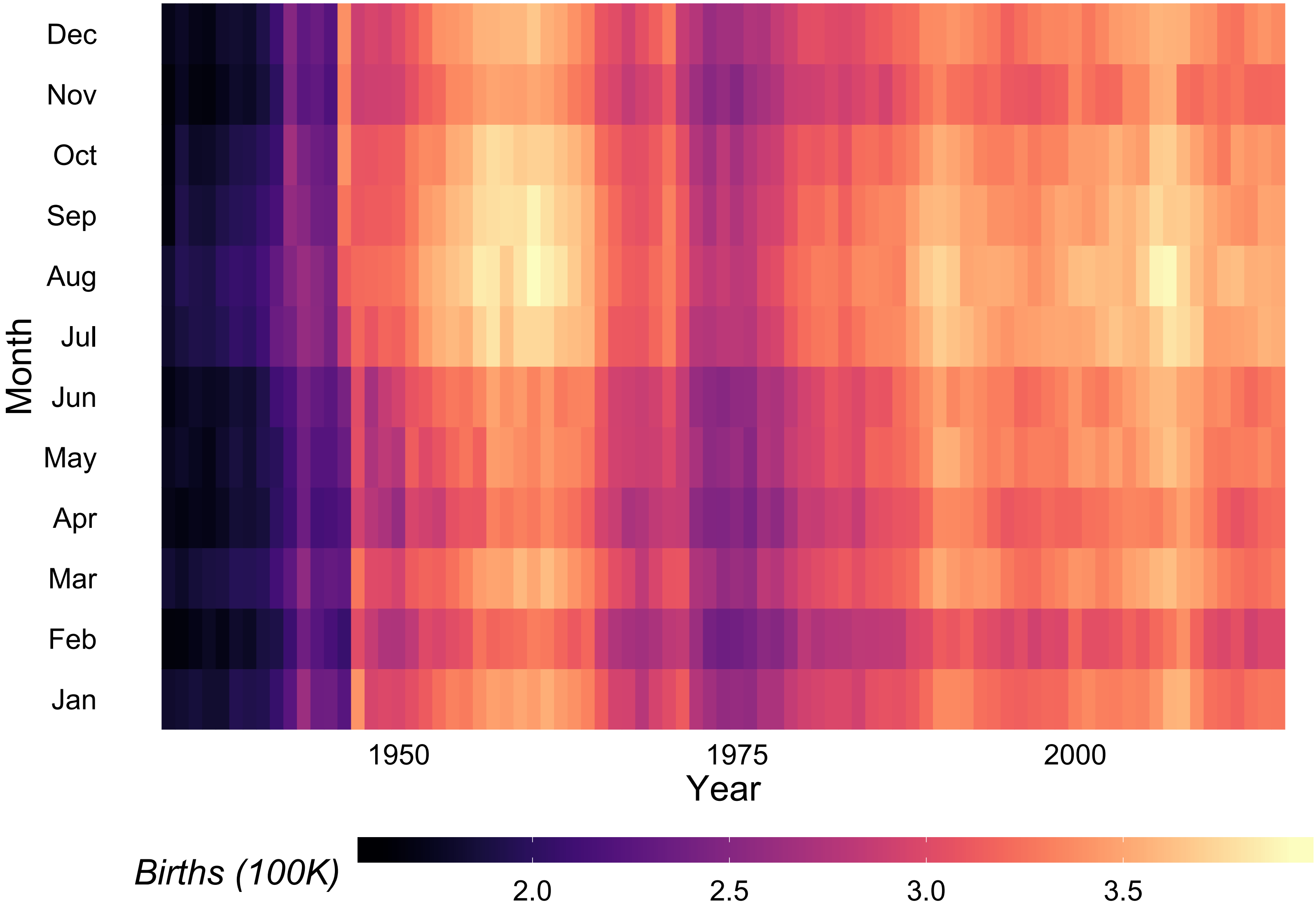
Today, we focus on a different type of data: **time-series data**.

- Sampled within one unit/individual (*e.g.*, Oregon).
- Observe multiple times for the same unit (*e.g.*, Oregon: 1990–2020).
- We have T time periods, each indexed t in $\{1, \ldots, T\}$.

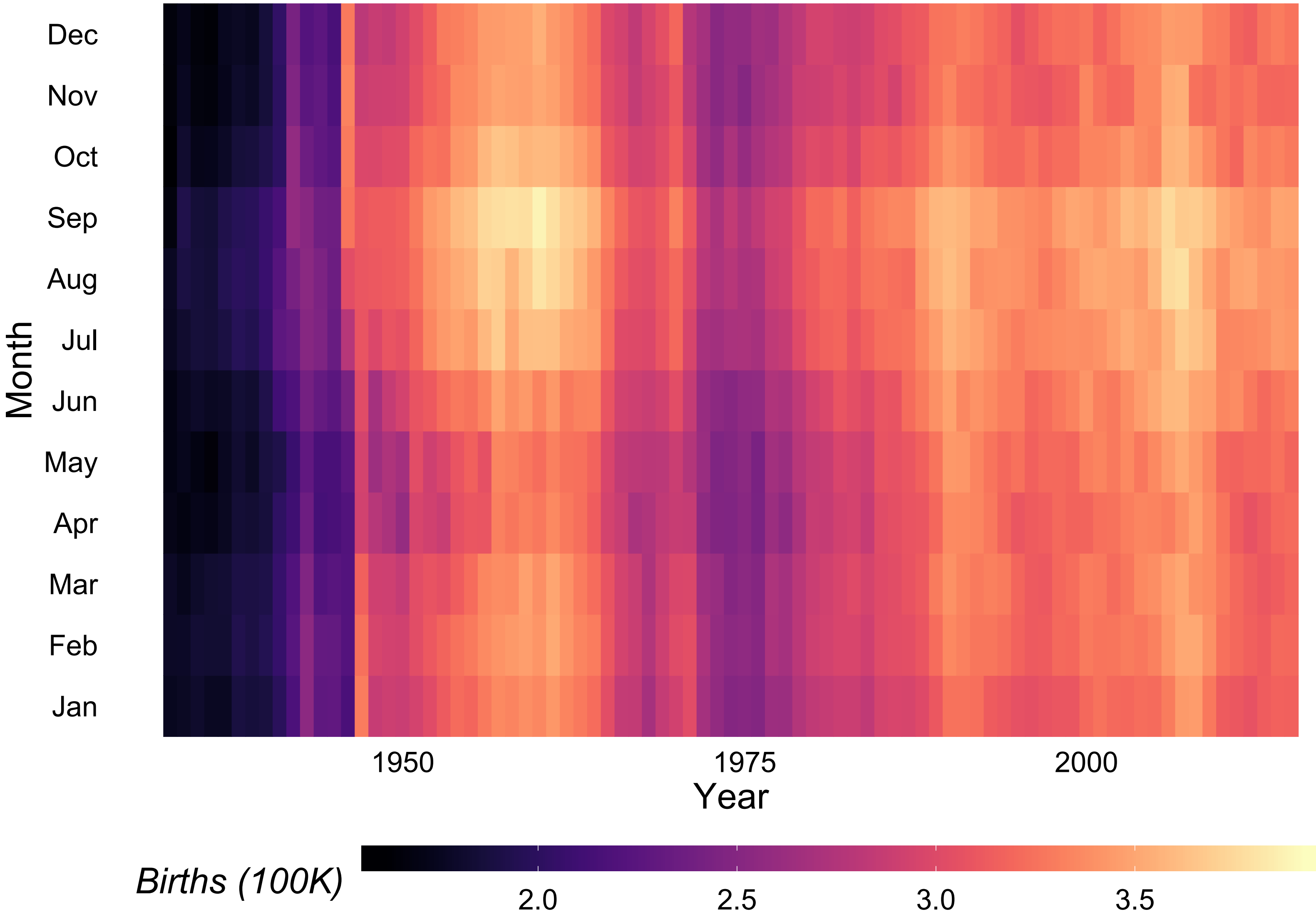
US monthly births, 1933–2015: Classic time-series graph



US monthly births, 1933–2015: Newfangled time-series graph



US monthly births per 30 days, 1933–2015: Newfangled time-series graph



Time-series models

Time-series models

Introduction

Our model now looks something like

$$\text{Births}_t = \beta_0 + \beta_1 \text{Income}_t + u_t$$

or perhaps

$$\text{Births}_t = \beta_0 + \beta_1 \text{Income}_t + \beta_3 \text{Income}_{t-1} + u_t$$

maybe even

$$\text{Births}_t = \beta_0 + \beta_1 \text{Income}_t + \beta_3 \text{Income}_{t-1} + \beta_4 \text{Births}_{t-1} + u_t$$

where $t-1$ denotes the time period prior to t (*lagged* income or births).

Assumptions

1. **New: Weakly persistent outcomes**—essentially, x_{t+k} in the distant period $t+k$ is weakly correlated with period x_t (when k is "big").

2. y_t is a **linear function** of its parameters and disturbance.

3. There is **no perfect collinearity** in our data.

4. The u_t have conditional mean of zero (**exogeneity**), $\boldsymbol{E}[u_t|X] = \mathbf{0}$.

5. The u_t are **homoskedastic** with **zero correlation** between u_t and u_s, *i.e.*, $\mathrm{Var}(u_t|X) = \mathrm{Var}(u_t) = \sigma^2$ and $\mathrm{Cor}(u_t,\, u_s|X) = 0$.

6. **Normality of disturbances**, *i.e.*, $u_t \overset{\text{iid}}{\sim} N\big(0,\, \sigma^2\big)$.

Time-series models

Model options

Time-series modeling boils down to two classes of models.

1. **Static models:** Do not allow for persistent effect.

2. **Dynamic models:** Allow for persistent effects.

    - Models with **lagged explanatory** variables

    - **Autoregressive, distributed-lag** (ADL) models

Model options

Option 1: Static models

Static models assume the outcome depends upon only the current period.

$$\text{Births}_t = \beta_0 + \beta_1 \text{Income}_t + u_t$$

Here, we must believe that income **immediately** affects the number of births and does not affect on the numbers of births in the future.

We also need to believe current births do not depend upon previous births.

Can be a very restrictive way to consider time-series data.

Model options

Option 2: Dynamic models

Dynamic models allow the outcome to depend upon other periods.

Time-series models

Model options

Option 2a: **Dynamic models** with lagged explanatory variables

These models allow the outcome to depend upon the explanatory variable(s) in other periods.

$$\begin{aligned}\text{Births}_t =& \beta_0 + \beta_1 \text{Income}_t + \beta_2 \text{Income}_{t-1} + \\ & \beta_3 \text{Income}_{t-2} + \beta_4 \text{Income}_{t-3} + u_t\end{aligned}$$

Here, income **immediately** affects the number of births *and* affects **future** numbers of births. In other words: Births today depend today's income and *lags* of income—*e.g.*, last month's income, last year's income, ...

Estimate *total* effects by summing lags' coefficients, *e.g.*, $\beta_1 + \beta_2 + \beta_3 + \beta_4$.

Note: We still assume current births don't affect future births.

Model options

Option 2b: Autoregressive distributed-lag (ADL) models

These models allow the outcome to depend upon the explanatory variable(s) and/or the outcome variable in prior periods.

$$\text{Births}_t = \beta_0 + \beta_1 \text{Income}_t + \beta_2 \text{Income}_{t-1} + \beta_3 \text{Births}_{t-1} + u_t$$

Here, current income affects affects **current** births and **future** births.

In addition, current births affect future births—we're allowing lags of the outcome variable.

Autoregressive distributed-lag models

Numbers of lags

ADL models are often specified as $\text{ADL}(p,\ q)$, where

- p is the (maximum) number of **lags** for the outcome variable.
- q is the (maximum) number of **lags** for explanatory variables.

Example: $\text{ADL}(1,\ 0)$

$$\text{Births}_t = \beta_0 + \beta_1 \text{Income}_t + \beta_2 \text{Births}_{t-1} + u_t$$

Example: $\text{ADL}(2,\ 2)$

$$\begin{aligned}\text{Births}_t =& \beta_0 + \beta_1 \text{Income}_t + \beta_2 \text{Income}_{t-1} + \beta_3 \text{Income}_{t-2} \\ &+ \beta_4 \text{Births}_{t-1} + \beta_5 \text{Births}_{t-2} + u_t\end{aligned}$$

Autoregressive distributed-lag models

Complexity

Due to their lags, ADL models actually estimate even more complex relationships than you might first guess.

Consider ADL(1, 0): $\textbf{Births}_t = \beta_0 + \beta_1 \textbf{Income}_t + \beta_2 \textbf{Births}_{t-1} + u_t$

Write out the model for period $t-1$:

$$\textbf{Births}_{t-1} = \beta_0 + \beta_1 \textbf{Income}_{t-1} + \beta_2 \textbf{Births}_{t-2} + u_{t-1}$$

which we can substitute in for $\textbf{Births}_{t-1}$ in the first equation, *i.e.*,

$$\begin{aligned}\textbf{Births}_t =& \beta_0 + \beta_1 \textbf{Income}_t + \\ & \beta_2 \underbrace{\left(\beta_0 + \beta_1 \textbf{Income}_{t-1} + \beta_2 \textbf{Births}_{t-2} + u_{t-1}\right)}_{\text{Births}_{t-1}} + u_t\end{aligned}$$

Complexity

Continuing...

$$
\begin{aligned}
\text{Births}_t = & \beta_0 + \beta_1 \text{Income}_t + \\
& \beta_2 \underbrace{\left(\beta_0 + \beta_1 \text{Income}_{t-1} + \beta_2 \text{Births}_{t-2} + u_{t-1}\right)}_{\text{Births}_{t-1}} + u_t \\
= & \beta_0 \left(1 + \beta_2\right) + \beta_1 \text{Income}_t + \beta_1 \beta_2 \text{Income}_{t-1} + \\
& \beta_2^2 \text{Births}_{t-2} + u_t + \beta_2 u_{t-1}
\end{aligned}
$$

We could then substitute in the equation for Births_{t-2}, Births_{t-3}, ...

Complexity

Eventually we arrive at

$$
\begin{aligned}
\text{Births}_t = & \beta_0 \left(1 + \beta_2 + \beta_2^2 + \beta_2^3 + \cdots\right) + \\
& \beta_1 \left(\text{Income}_t + \beta_2 \text{Income}_{t-1} + \beta_2^2 \text{Income}_{t-2} + \cdots\right) + \\
& u_t + \beta_2 u_{t-1} + \beta_2^2 u_{t-2} + \cdots
\end{aligned}
$$

The point?

By including just **one lag of the dependent variable**—as in a ADL(1, 0)—we implicitly include for *many lags* of the explanatory variables and disturbances.[†]

† These lags enter into the equation in a very specific way—not the most flexible specification.

The partial-adjustment model

There are times that actually want to model an individual's **desired amount**, rather than her ***actual* amount**, but we are unable to observe the desired level.

Partial-adjustment models help us model this situation.

Autoregressive distributed-lag models

The partial-adjustment model

Example

We want to know how the **desired number of cigarettes**, $\widetilde{\text{Cig}}_t$, changes with the current period's cigarette tax, *e.g.*,

$$\widetilde{\text{Cig}}_t = \beta_0 + \beta_1 \text{Tax}_t + u_t \tag{A}$$

Imagine actual cigarette consumption, Cig_t, doesn't change immediately (*e.g.*, habit persistence). Instead, consumption depends upon current desired level and previous consumption level

$$\text{Cig}_t = \lambda \widetilde{\text{Cig}}_t + (1 - \lambda)\, \text{Cig}_{t-1} \tag{B}$$

The partial-adjustment model

Example, continued

$$\widetilde{\text{Cig}_t} = \beta_0 + \beta_1 \text{Tax}_t + u_t \tag{A}$$

$$\text{Cig}_t = \lambda \widetilde{\text{Cig}_t} + (1 - \lambda)\, \text{Cig}_{t-1} \tag{B}$$

Substituting $\widetilde{\text{Cig}_t}$ from $(\mathbf{A})$ into $(\mathbf{B})$ yields

$$\begin{aligned} \text{Cig}_t &= \lambda \left(\beta_0 + \beta_1 \text{Tax}_t + u_t\right) + (1 - \lambda)\, \text{Cig}_{t-1} \\ &= \lambda\beta_0 + \lambda\beta_1 \text{Tax}_t + (1 - \lambda)\, \text{Cig}_{t-1} + \lambda u_t \end{aligned} \tag{C}$$

The equation in $(\mathbf{C})$ is ADL(1, 0).

We can also estimate/recover the speed-of-adjustment coefficient λ.

OLS in time series

OLS in time series

Unbiased coefficients

As before, the unbiased-ness of OLS is going to depend upon our exogeneity assumption, *i.e.*, $\boldsymbol{E}[u_t|X] = \mathbf{0}$.

We can split this assumption into two parts.

1. The disturbance u_t is independent of the explanatory variables in the **same period** (*i.e.*, X_t).

2. The disturbance u_t is independent of the explanatory variables in the **other periods** (*i.e.*, X_s for $s \neq t$).

We need both of these parts to be true for OLS to be unbiased.

OLS in time series

Unbiased coefficients

We need both parts of our exogeneity assumption for OLS to be unbiased:

$$\boldsymbol{E}\left[\hat{\beta}_1 \middle| X\right] = \beta_1 + \boldsymbol{E}\left[\frac{\sum_t \left(x_t - \overline{x}\right) u_t}{\sum_t \left(x_t - \overline{x}\right)^2} \middle| X\right]$$

I.e., to guarantee the numerator equals zero, we need $\boldsymbol{E}[u_t|X] = 0$—for both $\boldsymbol{E}[u_t|X_t] = 0$ *and* $\boldsymbol{E}[u_t|X_s] = 0$ $(s \neq t)$.

The second part of our exogeneity assumption—requiring that u_t is independent of all regressors in other periods—fails with dynamic models with lagged outcome variables.

Thus, **OLS is biased for dynamic models with lagged outcome variables**.

OLS in time series

Unbiased coefficients

To see why dynamic models with lagged outcome variables violate our exogeneity assumption, consider two periods of our simple ADL(1, 0) model.

$$\text{Births}_t = \beta_0 + \beta_1 \text{Income}_t + \beta_2 \text{Births}_{t-1} + u_t \tag{1}$$

$$\text{Births}_{t+1} = \beta_0 + \beta_1 \text{Income}_{t+1} + \beta_2 \text{Births}_t + u_{t+1} \tag{2}$$

In (1), u_t clearly correlates with Births_t.

However, Births_t is a regressor in (2) (lagged dependent variable).

$\therefore$ The disturbance in t (u_t) correlates with a regressor in $t+1$ (Births_t).

This correlation violates the second part of our exogeneity requirement.

Consistent coefficients

All is not lost.

For OLS to be **consistent**, we only need **contemporaneous exogeneity**.

Contemporaneous exogeneity: each disturbance is uncorrelated with the explanatory variables in the same period, *i.e.*,

$$\boldsymbol{E}[u_t|X_t] = 0$$

With contemporaneous exogeneity, OLS estimates for the coefficients in a time series model are consistent.

Consistent coefficients

To see why OLS is consistent with contemporaneous exogeneity, consider the OLS estimate for β_1 in

$$\text{Births}_t = \beta_0 + \beta_1 \text{Births}_{t-1} + u_t$$

which we've shown (a few times) can be written

$$\hat{\beta}_1 = \beta_1 + \frac{\sum_t \left(\text{Births}_{t-1} - \overline{\text{Births}}\right) u_t}{\sum_t \left(\text{Births}_{t-1} - \overline{\text{Births}}\right)^2}$$

Consistent coefficients

$$\begin{aligned}
\text{plim}\,\hat{\beta}_1 &= \text{plim}\left(\beta_1 + \frac{\sum_t \left(\text{Births}_{t-1} - \overline{\text{Births}}\right) u_t}{\sum_t \left(\text{Births}_{t-1} - \overline{\text{Births}}\right)^2}\right) \\
&= \beta_1 + \frac{\text{plim}\left[\sum_t \left(\text{Births}_{t-1} - \overline{\text{Births}}\right) u_t / T\right]}{\text{plim}\left[\sum_t \left(\text{Births}_{t-1} - \overline{\text{Births}}\right)^2 / T\right]} \\
&= \beta_1 + \frac{\text{Cov}(\text{Births}_{t-1},\, u_t)}{\text{Var}(\text{Births}_t)} \\
&= \beta_1 \quad \textbf{if } \text{Cov}(\text{Births}_{t-1},\, u_t) = 0
\end{aligned}$$

Contemporaneous exogeneity gives us $\text{Cov}(\text{Births}_{t-1},\, u_t) = 0$.

Consistent coefficients

Thus, if we assume **contemporaneous exogeneity**, **OLS is consistent** for the coefficients, *even for models with lagged dependent variables.*

The end.

Table of contents

Admin

Time series

Autoregressive distributed-lag models

Equilibrium effects

ADL models also offer interesting insights for long-run/equilibrium effects.

$$\text{Births}_t = \beta_0 + \beta_1 \text{Income}_t + \beta_2 \text{Births}_{t-1} + u_t$$

In this ADL(1, 0) model, β_1 gives the **short-run effect** of income on the number of births. *I.e.*, how income in time t affects births in time t.

Equilibrium effects

Starting with

$$\text{Births}_t = \beta_0 + \beta_1 \text{Income}_t + \beta_2 \text{Births}_{t-1} + u_t$$

we move into equilibrium, *i.e.*, $\text{Births}_t = \text{Births}^\star$, *i.e.*,

$$\text{Births}^\star = \beta_0 + \beta_1 \text{Income}^\star + \beta_2 \text{Births}^\star$$

Now rearrange...

$$\begin{aligned}
\text{Births}^\star - \beta_2 \text{Births}^\star &= \beta_0 + \beta_1 \text{Income}^\star \\
\left(1 - \beta_2\right) \text{Births}^\star &= \beta_0 + \beta_1 \text{Income}^\star \\
\text{Births}^\star &= \frac{\beta_0}{\left(1 - \beta_2\right)} + \frac{\beta_1}{\left(1 - \beta_2\right)} \text{Income}^\star
\end{aligned}$$

Equilibrium effects

Short-run effect of income on births:

$$\text{Births}_t = \beta_0 + \beta_1 \text{Income}_t + \beta_2 \text{Births}_{t-1} + u_t$$

Long-run effect of income on births:

$$\text{Births}^\star = \frac{\beta_0}{(1-\beta_2)} + \frac{\beta_1}{(1-\beta_2)} \text{Income}^\star$$

Equilibrium effects

Another way to see this result:

We already showed

$$\text{Births}_t = \beta_0 + \beta_1 \text{Income}_t + \beta_2 \text{Births}_{t-1}$$

gives us

$$\begin{aligned}\text{Births}_t = &\beta_0 \left(1 + \beta_2 + \beta_2^2 + \beta_2^3 + \cdots\right) + \\ &\beta_1 \left(\text{Income}_t + \beta_2 \text{Income}_{t-1} + \beta_2^2 \text{Income}_{t-2} + \cdots\right) + \\ &u_t + \beta_2 u_{t-1} + \beta_2^2 u_{t-2} + \cdots\end{aligned}$$

In equilibrium: $\text{Income}_t = \text{Income}_{t-k} = \text{Income}^\star$ for all k.

Autoregressive distributed-lag models

Equilibrium effects

Substituting $\mathbf{Income}_t = \mathbf{Income}^\star$ for all k
(and assuming no disturbances in equilibrium):

$$\begin{aligned}\mathbf{Births}_t =& \beta_0 \left(1 + \beta_2 + \beta_2^2 + \beta_2^3 + \cdots\right) + \\ & \beta_1 \left(\mathbf{Income}^\star + \beta_2 \mathbf{Income}^\star + \beta_2^2 \mathbf{Income}^\star + \cdots\right) + \\ =& \beta_0 \left(\frac{1}{\beta_2}\right) + \\ & \beta_1 \left(\frac{1}{\beta_2}\right) \mathbf{Income}^\star\end{aligned}$$

So long as $-1 < \beta_2 < 1$.[†]

† This simplification comes from $\sum_{k=0}^{\infty} p^k = \frac{1}{p}$ for $-1 < k < 1$.