Analytics for Value Investing Assignment 1

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Q1

a) Source of data:

Source: Yahoo Finance

Data: S&P500 (^GSPC) monthly adjusted close price index

Period: Dec 1985 to Dec 2020 (421 months)

bii) Compute descriptive statistics of those monthly returns

After droping NA:

count 420.000000
mean 0.007854
std 0.043864
min -0.217630
25% -0.017368
50% 0.011959
75% 0.035292
max 0.131767

biii) Test the hypotheses that the returns for the month of January are higher or more positive, while those for the month of September are lower or more negative, than those for other months,

c) Briefly explain your statistical results and their implications for the hypotheses. Your answer should also include an explanation of what the intercept (β ") of the above regression implies.

t-test of difference in mean returns between January and non-January months

Test conducted

Step 1: To determine if the population is equal, we conduct a Levene's test at 5% significance level with the hypotheses that:

H0: Population variance of Jan return = Population variance of non-Jan return

H1: Population variance of Jan return != Population variance of non-Jan return

Step 2: Using scipy.stats.levene, we obtain p-value=0.361. This means that the evidence is insufficient to reject the null hypothesis that Population variance of Jan return = Population variance of non-Jan return.

	Step 3: Hence, we assume the population variances are unknown but equal: H0: Average January Return – Average non-January Return <= 0 H1: Average January Return – Average non-January Return > 0 Step 4: Perform an independent one-tailed t-test of difference in
	means at 5% significance level.
Results	Result:
obtained	Degree of freedom= 418
	t statistics= 0.286964965739269
	Critical Value= 1.648507149425946
	p-value= 0.3871406696022036
Conclusion	Given that the p-value of 0.387 is greater than the significance level of
	0.05, there is insufficient evidence to support the claim that the
	average return in January is higher than that of the other months.

• *t*-test of difference in mean returns between September and non-September months;

t-test of differe	nce in mean returns between September and non-September months
Test conducted	Step1: To determine if the population is equal, we conduct a Levene's test at 5% significance level with the hypotheses that:
	H0: Population variance of Sep return = Population variance of non-Sep return
	H1: Population variance of Sep return != Population variance of non- Sep return
	Step2: Using scipy.stats.levene, we obtain p-value=0.280. This means that the evidence is insufficient to reject the null hypothesis that Population variance of Sep return = Population variance of non-Sep return.
	Step 3: Hence, we assume the population variances are unknown but equal
	H0: Average September Return – Average non-September Return >= 0 H1: : Average September Return – Average non-September Return < 0
	Step4: Perform an independent one-tailed t-test of difference in means at 5% significance level.

Results obtained	Degree of freedom=418 t statistics=-1.9200309830619031 Critical Value=1.648507149425946 p-value=0.0277671060025958
Conclusion	Given that the p-value of 0.0278 is smaller than the significance level of 0.05, there is sufficient evidence to support that claim that the average return in September is smaller than that of the other months.

• an OLS regression of monthly S&P500 returns as dependent variable and two dummy independent variables

$$R_t = \beta_0 + \beta_1 D_{Jan} + \beta_2 D_{Sep} + \varepsilon_t$$

	Results							
OLS Regression	OLS Regression Results							
	Dep. Variable Model: Method: Date: Time: No. Observati Df Residuals: Df Model: Covariance Ty	Fr .ons:	mktret OLS Least Squares i, 26 Feb 2021 03:15:30 420 417 2 nonrobust	F-stat Prob (red: R-squared: Listic: F-statistic kelihood:):	0.009 0.004 1.845 0.159 719.59 -1433. -1421.	
	Intercept djan dsep	coef 0.0090 0.0009 -0.0147	std err 0.002 0.008 0.008	t 3.850 0.114 -1.899	P> t 0.000 0.909 0.058	[0.025 0.004 -0.014 -0.030	0.975] 0.014 0.016 0.001	
	Omnibus: 60.413 Durbin-Watson: Prob(Omnibus): 0.000 Jarque-Bera (JB): Skew: -0.750 Prob(JB): Kurtosis: 5.366 Cond. No.						1.960 137.381 1.47e-30 3.85	
	Warnings: [1] Standard Errors assume that the covariance matrix of the errors is correctly specified.							
	The model is given by: $R_t = \beta_0 + \beta_1 D_{Jan} + \beta_2 D_{Sep} + \varepsilon_t$							
	Where: Beta0 = 0.0 Beta1 = 0.0 Beta2 = 0.0)090)009	•					

Implications for the hypotheses

The intercept beta0 has a value of 0.009 which is the return when the variables djan and dsep are 0 i.e. when the month is not January or September. Beta0 also has a low p-value of 0.00 which means the evidence is sufficient to reject the null hypothesis that beta0 coefficient is 0.

The coefficients of djan and dsep are 0.0009 and -0.0147 respectively. For January return, this supports the hypothesis because the market return would be 0.009 + 0.0009*1 = 0.0099 which is higher than returns for other months.

For September return, this supports the hypothesis because market return would be 0.009 - 0.0147*1 = -0.0057 which is lower than the returns for other months

However, the p-values of djan and dsep are 0.909 and 0.058 which are higher than 0.05. This suggests that the evidence is not sufficient to reject the null hypotheses that the coefficients of these two variables are 0. This causes the coefficient estimates of djan and dsep to be unreliable.

R-squared

The adjusted R-squared is only 0.009 which means only 0.9% of the variance in the S&P500 monthly return is explained by the two dummy variables.

Biases:

The OLS coefficient estimates remain unbiased. However, without adjusting for heteroskedasticity, there is risk of t-statistics of coefficient estimates being biased upwards, causing a risk of rejecting the null hypothesis when it is in fact true.

There also exists a slightly positive autocorrelation among the error terms as shown by the Durbin-Watson value of 1.96. Hence, the t-statistics of coefficient estimates are biased upwards, causing a risk of rejecting the null hypothesis when it is in fact true.

White test for heteroskedasticity

```
White test for heteroskedasticity:

0 1 2 3

0 LM-Statistic LM p-value F-Statistic F-Test p-value
1 0.283866 0.867679 0.141014 0.868518
```

The White test has a p-value of 0.869 which is larger than the significance level of 0.05. It means that evidence is insufficient to reject the null hypothesis that the error terms have constant variances. Hence there is likely homoskedasticity.

OLS using HC3

OLS regression of mktret against dummy variables under heteroskedasticity-consistent **OLS Regression Results** Dep. Variable: mktret R-squared: 0.009 Model: Adj. R-squared: 0.004 Least Squares Fri, 26 Feb 2021 F-statistic: Prob (F-statistic): 1.574 0.208 Method: Date: Log-Likelihood: 04:23:20 719.59 No. Observations: Df Residuals: -1433. 420 Df Model: Covariance Type: HC3 [0.025 0.975] std err P>|z| coef 0.014 0.0090 0.000 0.004 0.002 3.900 Intercept djan dsep -0.01470.008 -1.7580.079 -0.0310.002 Durbin-Watson: Jarque-Bera (JB): Omnibus: Prob(Omnibus): 1.960 137.381 60.413 0.000 Prob(JB): Kurtosis: 5.366 Cond. No. 3.85 [1] Standard Errors are heteroscedasticity robust (HC3)

The OLS estimator with HC3 produces similar result as normal OLS regression as above. However, the z-scores of the coefficients of djan and dsep are adjusted downward compared to normal OLS regression (from 0.114 and -1.899 to 0.106 and -1.758).

This is because HC3 takes into accounts of heteroskedasticity which could inflate the t-statistics of the coefficient estimates.

Despite the fact that the White test ascertains that there is sufficient evidence for homoskedasticity, the variances of the error terms might still be slightly inconsistent. HC3 adjusts for the small heteroskedasticity to deflate the t-statistics of djan and dsep coefficient estimates.

Implication for hypotheses

However, the p-values of djan and dsep are 0.916 and 0.079 which are higher than 0.05. This suggests that the evidence is not sufficient to reject the null hypotheses that the coefficients of these two variables are 0. This causes the coefficient estimates of djan and dsep to be unreliable.

		OLS Regres				=======
Dep. Variable: Model:		mktret OLS	Adi R-	squared:		0.009 0.004
Method: Date:	Fri	Least Squares i, 26 Feb 2021 04:23:20	F-statis	stic: -statistic):		1.638 0.196
Time: No. Observatio		04:23:20	Log-Like	elihood:		719.59 -1433.
Df Residuals:		417				-1421.
Df Model: Covariance Typ	e:	2 HAC				
		std err		P> t		_
Intercept djan dsep	0.0090 0.0009 -0.0147	0.002 0.008 0.008 -	3.842 0.109 1.790	0.000 0.913 0.074	0.004 -0.015 -0.031	0.014 0.017 0.001
Omnibus: Prob(Omnibus): Skew: Kurtosis:		60.413 0.000 -0.750	Durbin-V	Watson: Bera (JB):):		1.960 137.381 1.47e-30 3.85
Warnings: [1] Standard E		heteroscedastic	======== city and a		ion robust	======
OLS regression consistent std		t against dummy			tocorrelati	ion-
		===========	ession Re			
Dep. Variable: Model:		mktret OLS	Adi R	-squared:		0.009 0.004
Method: Date:	Fr	Least Squares i, 26 Feb 2021 04:23:20	F-stat: Prob (istic: F-statistic) :	1.739 0.177
Time: No. Observation	nc.	04:23:20	Log-Li	kelihood:		719.59 -1433.
Df Residuals: Df Model: Covariance Typ		417 2 HAC	BIC:			-1421.
==========		std err		P> t		
		0.002 0.008 0.008			0.005 -0.015 -0.030	
Omnibus:		60.413	Durbin-	-Watson:		1.960
Prob(Omnibus): Skew:		0.000 -0.750	Jarque Prob(JI	-Bera (JB): B):		137.381 1.47e-30
Kurtosis:		5.366 ======		No. =======		3.85
using 3 lags a	nd withou	heteroscedast: t small sample	correction	on les under a		
consistent sto		OLS Regi	ression R	esults =======	=======	
consistent sto	d errors:	mktre	======= t R-squ	======= ared:	=======	
consistent sto	d errors: ======= :	mktre 0LS Least Squares	t R-squ Adj. F-sta	======= ared: R-squared: tistic:		0.004 1.843
consistent sto	d errors: : : F	mktre 0LS Least Squares ri, 26 Feb 2022 04:23:20	R-squ Adj. F-sta Prob Log-L	======= ared: R-squared:	.c):	0.004 1.843 0.160 719.59
consistent sto ===================================	d errors: 	mktre 0L9 Least Squares ri, 26 Feb 2021	R-squ Adj. F-sta Prob Log-L AIC: BIC:	======= ared: R-squared: tistic: (F-statisti	.c):	0.004 1.843 0.160
consistent sto	d errors: 	mktre OLS Least Squares ri, 26 Feb 202: 04:23:20 42(41: HA(R-squ Adj. F-sta Prob Log-L AIC: BIC:	======= ared: R-squared: tistic: (F-statisti	.c):	0.004 1.843 0.160 719.59 -1433. -1421.
consistent sto	d errors:	mktref OLS Least Squares ri, 26 Feb 202: 04:23:20 41: HAC std err 0.002 0.008 0.008	t R-squ S Adj. S F-sta Prob Log-L 0 AIC: 2 2 4.284 0.110 -1.905	ared: ared: R-squared: tistic: (F-statisti ikelihood:	(c): [0.025 0.005 -0.015 -0.030	-1421. 0.975] 0.013 0.017 0.000
consistent sto	coef 0.0090 0.0099 -0.0147	mktret 0.18 Least Squares ri, 26 Feb 202: 04:23:20 41: 41: 2 HAC std err 0.002 0.008 0.008	### R-squ ### R-	ared: R-squared: tistic: (F-statisti ikelihood: P> t 0.000 0.912 0.057 n-Watson:	(0.025 0.005 -0.015 -0.030	0.004 1.843 0.160 719.59 -1433. -1421. 0.975] 0.013 0.017 0.000
consistent sto	coef 0.0090 0.0099 -0.0147	mktret OLS Least Squares ri, 26 Feb 202: 04:23:20 41: HAO std err 0.002 0.008 0.008	t R-squ S Adj. S F-sta L Prob AIC: T BIC: C C C C C C C C C C C C C C C C C C C	ared: R-squared: tistic: (F-statisti ikelihood: P> t 0.000 0.912 0.057	(0.025 0.005 -0.015 -0.030	0.004 1.843 0.160 719.59 -1433. -1421. 0.975]

consistent sto		OLS Re	gression R	esults				
======== Dep. Variable:		mktr	et R–sau	======== ared:		0.009		
Model:				R-squared:		0.004		
Method:		Least Squar	,.	tistic:		1.898		
Date:	Fr	i, 26 Feb 20		(F-statistic):	0.151		
Time:		04:23:	20 Log-L	ikelihood:		719.59		
No. Observatio	ns:	4	20 AIČ:			-1433.		
Df Residuals:		4	17 BIC:			-1421.		
Df Model:			2					
Covariance Typ	e:	Н	AC					
	coef	std err	t	P> t	[0.025	0.975]		
 Intercept	0.0090	0.002	4.369	0.000	0.005	0.013		
djan	0.0009	0.008	0.110	0.912	-0.015	0.017		
dsep	-0.0147	0.008	-1.913	0.056	-0.030	0.000		
========= Omnibus:		 60 ₋ 4	13 Durbi	======== n_Watson:	=======	1.960		
Prob(Omnibus):				e-Bera (JB):		137.381		
Skew:			50 Prob(1.47e-30		
Kurtosis:		5.3	66 Cond.	No.		3.85		

The t-statistics of all coefficient estimates increase from 1 to 12 lags.

The OLS estimator using HAC adjusts for both heteroskedasticity and autocorrelation of the error terms. It seems that HAC with lag=1 is more robust/aggressive in eliminating heteroskedasticity and autocorrelation because its t-statistics of the coefficient estimates are the highest, which means there is the lowest risk of rejecting the null hypothesis.

Implication for hypotheses

However, under HAC with lag=1, the p-values of djan and dsep are 0.913 and 0.074 which are higher than 0.05. This suggests that the evidence is not sufficient to reject the null hypotheses that the coefficients of these two variables are 0. This causes the coefficient estimates of djan and dsep to be unreliable.

Conclusion from OLS regression analyses:

The OLS regression analyses do not support the hypotheses even after adjusting for heteroskedasticity and autocorrelation under HC3 and HAC.

The p-values of djan and dsep from all the models above are greater than 0.05, which suggests that the coefficient estimates are statistically insignificant i.e. they are very likely to be 0 which is the null hypothesis. Hence, there is insufficient evidence to support the hypotheses.

a) Source of data

Source: Compustat

Data: Annual Financial Statements of US Companies Conditions: "gsecotr" != 40 AND "mktvalt" >= 1,000

Period: Dec 2003 to Dec 2019 (16 years)

b) Test the hypothesis that operating profitability for the year is positively correlated with the previous- year operating profit margin and operating asset turnover, using the three regression models below:

	Result
Pooled OLS	Pooled OLS regression of future operating profitability against accounting variables with HC std errors OLS Regression Results
	Dep. Variable: op_prof R-squared: 0.003 Model: 0LS Adj. R-squared: 0.003 Method: Least Squares F-statistic: 17.99 Date: Fri, 26 Feb 2021 Prob (F-statistic): 1.56e-08 Time: 14:39:31 Log-Likelihood: -20283. No. Observations: 20045 AIC: 4.057e+04 Df Residuals: 20042 BIC: 4.060e+04 Df Model: 2 Covariance Type: HC3
	coef std err t P> t [0.025 0.975] const 0.0383 0.012 3.165 0.002 0.015 0.062 op_margin 0.2726 0.046 5.966 0.000 0.183 0.362 aturn 0.0090 0.003 2.842 0.004 0.003 0.015
	Omnibus: 1970.232 Durbin-Watson: 1.949 Prob(Omnibus): 0.000 Jarque-Bera (JB): 13853.759 Skew: 0.184 Prob(JB): 0.00 Kurtosis: 7.056 Cond. No. 18.5
	Warnings: [1] Standard Errors are heteroscedasticity robust (HC3) The model produced is: $ \%\Delta EBIT_{i,t} = \beta_0 + \beta_1 \frac{EBIT_{i,t-1}}{Rev_{i,t-1}} + \beta_2 \frac{Rev_{i,t-1}}{NOA_{i,t-1}} + \varepsilon_{it} $ Where:
	Beta0 = 0.0383 Beta1= 0.2726 Beta2= 0.0090
	Implication of hypothesis: Op_margin and aturn have coefficients of 0.2726 and 0.0090 which means it supports the hypothesis that operating profitability for the year is positively correlated with the previous- year operating profit margin and operating asset turnover.

The p-values of op_margin and aturn are 0.000 and 0.004 which are lower than 0.05. This suggests that the evidence is sufficient to reject the null hypotheses that the coefficients of these two variables are 0. Hence, the coefficients are reliable.

R-squared

The adjusted R-squared is only 0.003 which means only 0.3% of the variance in the future operating profitability is explained by current operating profit margin and asset turn.

Biases:

However, without adjusting for heteroskedasticity, there is risk of t-statistics of coefficient estimates being biased upwards, causing a risk of rejecting the null hypothesis when it is in fact true.

There also exists a positive autocorrelation among the error terms as shown by the Durbin-Watson value of 1.949. Hence, the t-statistics of coefficient estimates are biased upwards, causing a risk of rejecting the null hypothesis when it is in fact true.

Year-by-year regressions

```
Coefficients of year-by-year regressions
          const op_margin
                               aturn
fyear
2003
      0.003484
                0.878796 0.006357
      0.067969 0.896456 0.012460
2004
2005
      0.147332 0.400387 -0.004640
2006
      0.085794
                 0.433101 0.001260
                0.223379 0.007689
2007
      0.064683
2008 -0.001878 0.409437
                           0.020506
2009 -0.111508 -0.023692 0.011437
2010 0.042498 1.093718 0.034851
2011
      0.062079 0.571334 0.000691
                0.190401 0.031559
2012 -0.059425
      0.048396 -0.047039 -0.007697
2013
      0.118278 0.087991 -0.014723
2014
      0.088395 -0.684804 0.013977
0.031582 0.155108 -0.000517
2015
2016
                 0.155108 -0.000517
2017 -0.066317 0.775148 0.024280
      2018
2019
```

T-statistics and p-values of year-by-year coefficients:

```
t-stat p-value
const 2.2898 0.0360
op_margin 2.9806 0.0088
aturn 2.1686 0.0455
```

Implication for hypothesis

The year-by-year OLS regression shows that the operating profitability has different correlations with operating profit margin and asset turn for each year. For example, in 2005, op_prof is negatively correlated with aturn with a coefficient of -0.00464 and

this does not support the hypothesis that op_prof is positively correlated with aturn.

Through conducting a 1-sample two-tailed t-test, the p-values of the samples of op_margin and aturn are 0.0088 and 0.0455 which are lower than 0.05. This suggests that the evidence is sufficient to reject the null hypotheses that the year-by-year coefficients estimates of these two variables are 0.

The year-by-year sample of intercept const gives a p-value of 0.036. Given a significance level of 5%, it means that the evidence is sufficient to reject the null hypothesis that the year-by-year const coefficient estimates are 0.

Logistic regression

Optimization terminated successfully.

Current function value: 0.649450

Iterations 5
Logit regression of operating profitability dummy against accounting variables
Logit Regression Results

Dep. Variabl	e:	d_op_p	rof No. Ob	servations:		20045
Model:		Lo	,	iduals:		20042
Method: Date:			MLE Df Mod	Df Model:		2
		Fri, 26 Feb 2021 Pseudo R-squ.:				0.005758
Time:		14:40	:53 Log-Li	kelihood:		-13018.
converged:		T	rue LL-Nul	l:		-13094.
Covariance T	ype:	nonrob	ust LLR p-	-value:		1.818e-33
	coef	std err	z	P> z	[0.025	0.975]
const	0.3430	0.027	12.868	0.000	0.291	0.395
op_margin	1.3140	0.110	11.995	0.000	1.099	1.529
aturn	0.0418	0.009	4.501	0.000	0.024	0.060

The model produced is:

$$D_{i,t} = \beta_0 + \beta_1 \frac{EBIT_{i,t-1}}{Rev_{i,t-1}} + \beta_2 \frac{Rev_{i,t-1}}{NOA_{i,t-1}} + \varepsilon_{it}$$

Where beta0 = 0.343

Beta1 = 1.314

Beta2 = 0.0418

Implication for hypothesis

Op_margin and aturn have coefficients of 1.31 and 0.0418 which means it supports the hypothesis that operating profitability for the year is positively correlated with the previous- year operating profit margin and operating asset turnover.

The p-values of the coefficient estimates are all 0, which means that the evidence is sufficient to reject the null hypothesis that the coefficients estimates are 0.

Pseudo R squared:

The model has a low pseudo R-squared of 0.00576 which suggests a poor fit and would lead to inaccurate prediction.

Biases:

However, heteroskedasticity and auto-correlation on logistic regression model will cause the coefficient estimates to be biased and the impact on t-statistics is unknown.

Inferences of statistical significance of coefficient estimates of logit regression model are therefore incorrect.