ADVANCED DATA ANALYSIS FOR PSYCHOLOGICAL SCIENCE

Part 1. Introduction to multilevel modeling

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Outline of Part 1

- LM recap: Short recap of linear regression modeling 🚨 🍳
- LMER: Introduction to multilevel modeling (linear mixed-effects regression)
- Data processing: How to approach a multilevel data structure?
 How to manipulate and pre-process multilevel data?
- Descriptives: Which descriptive stats should be reported from a multilevel dataset? How to compute and interpret them?
- Model fit: How to fit a multilevel model in R? How to inspect, report, visualize, and interpret the results of a multilevel model? •
- Model evaluation: Which are the assumptions of multilevel models? How to evaluate them? How to compare multiple models and select the best model?
- Related: Summaries & in-depth topics related to multilevel modeling (e.g., generalized and Bayesian LMER, power analysis)

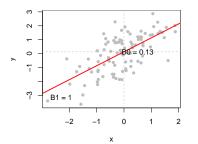
_b = not for the exam

Q = exercises with R (bring your laptop!)

LM recap: Linear regression models

Linear models (LM) allow to determinate the link between two variables as expressed by a linear function: $y_i = \beta_0 + \beta_1 x_i + \epsilon_i$ Such a function can be graphically represented as a **straight line**, where:

- β_0 is the **intercept** (value assumed by y when x = 0)
- β_1 is the **slope** (predicted change in y when x increases by 1 unit)
- ϵ_i are the **errors** (distance between observation i and the regression line)



 x_i and y_i are the values of observation i for the casual variables x and y

 β_0 , β_1 , and ϵ_i are called "parameters", or "coefficients". They are *estimated* from the sampled data and *generalized* to the whole population.

Fitting linear models in R

```
data("children", package = "npregfast") # loading children dataset from npregfast pkg
```

R uses the lm() function to fit linear models with the arguments formula $(y \sim x1 + x2 + ...)$ and data (identifying the dataframe with the model variables).

Null model

Children' height is only predicted by the model intercept $\beta_0 = \text{expected}$ (i.e., mean) value of height in the sample. σ^2 is the variance of the residuals ϵ_i (deviations from the intercept).

```
m0 <- lm(formula = height - 1,
data = children)

coefficients(m0) # model parameters

(Intercept)
153.4013

summary(m0)$sigma^2 # residual variance

[1] 243.9085
```

Simple regression model

height is now predicted by the intercept β_0 (mean value when age is 0), the slope β_1 (expected change for 1-unit increase in age), and the residual variance σ^2 .

```
m1 <- lm(formula = height - age,
data = children)

coefficients(m1) # model parameters

(Intercept) age
94.904099 4.388803

summary(m1)$sigma^2 # residual variance

[1] 56.19656
```

Multiple regression & interactions

LM also allow to include **multiple predictors** and the **interactions**¹ among them. This is done by estimating a separate slope (thus, a separate line) for each predictor by *holding constant* the value of the other predictors, which are fixed to zero.

Multiple regression model

```
eta_0 = 	ext{expected value in girls with age} = 0
eta_1 = 	ext{age effect}^2 	ext{ within the same sex}
eta_2 = 	ext{sex difference when age} = 0
	ext{m2} < - 	ext{lm(formula = height - age + sex,}
	ext{data = children)}
	ext{coefficients(m2)}
	ext{(Intercept)} 	ext{age sexmale}
	ext{95.0075706} 	ext{4.3887983} 	ext{-0.2001025}
```

Interactive model $\beta_1 = \text{age effect in girls}$

104.25

3.70

-19.04

1.41

¹The interaction between x_1 and x_2 is computed as the product of x_1 and x_2 .

 $^{^2}$ In this context, "effect" is used as a synonym of "relationship" (not a causal effect).

Model comparison & model selection

Likelihood ratio test

#Df

Compares the fit of two nested models (i.e., predicting the same y variable, with the more complex model including all predictors included in the simpler model).

```
library(lmtest)
lrtest(m0.m1.m2.m3) # returns Chisa statistic
```

```
Chisa
     LogLik Df
                         Pr(>Chisa)
2 -10417.84 NA
                    NA
   -8582.42 1 3670.84 0.000000e+00
4 -8582 19 1
                  0.45 5.046155e-01
```

226.67 3.176229e-51

-8468.86 1

Information criteria

The Akaike (AIC) and the Bayesian Information Criterion (BIC) compare multiple models in terms of fit & parsimony (the lower number of parameters the better)

```
AIC(m0,m1,m2,m3) # AIC: the lower the better
[1] 20839.68 17170.83 17172.39 16947.72
# Akaike weights: from 0 (-) to 1 (+)
MuMIn::Weights(AIC(m0,m1,m2,m3))
model weights
[1] 0 0 0 1
```

Here, model fit to the data is expressed by its likelihood = probability of observing the sampled data given the parameters estimated by the model, sometimes referred as the evidence of a model, or its ability to predict/forecast new data that are similar to the sampled data (see interactive visualization by Kristoffer Magnusson).

Parameter estimation in linear regression models

 β_0 , β_1 , and ϵ must be estimated based on data sampled from a population:

$$\hat{\beta}_0 = b_0; \, \hat{\beta}_1 = b_1; \, \hat{\epsilon} = e$$
).

⚠ There are several methods to estimate unknown parameters, such as:

- Ordinary least squares (OLS): finds the parameter values that minimize the sum
 of the squared residuals (default LM estimator)
- Maximum likelihood estimator (MLE): finds the parameter values that maximize
 the model likelihood, making the observed data the most probable under that model
- Bayesian estimator: finds the parameter posterior distributions based on prior knowledge/beliefs (prior) and observed data (likelihood)

Regardless of the used method, parameters values (or distributions) are always accompanied with a measure of the uncertainty/precision associated with their estimate:

Standard errors (SE) = predicted *variability* in the parameter estimate if the data were collected from different random samples from the same population.

SE are used for computing test statistics (Est/SE) & confidence intervals (Est \pm 1.96 \times SE)

[▶] In LM, under the assumption of normally distributed residuals, OLS = MLE

What are residuals?

Residuals are the model-based estimates of the population errors.

Linear model:

$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i$$

Predicted values:

$$\hat{y}_i = \beta_0 + \beta_1 x_i$$

Observed values:

$$y_i = \hat{y}_i + \hat{\epsilon}_i$$

Residuals = observed - predicted

$$\hat{\epsilon}_i = y_i - \hat{y}_i$$

```
observed predicted residuals squared
   150.77
            152.90
                      -2.13
                              4.55
 170.59
            156.61
                      13.98 195.33
3 167.31
            160.31
                       7.00
                            49.01
4 165.72
            165.52
                       0.20
                              0.04
5 171.67
            160.31 11.36 129.06
            151.07
                      -7.33
  143.74
                            53.74
```

sum(residuals(m3)^2) # sum of squared (SS) residuals

```
## [1] 128188.3
```

```
var(residuals(m3)) # residual variance SIGMA2
## [1] 51.29585
```

In LM. model parameters include:

- (1) intercept, (2) slope(s), and (3) residual variance σ^2
- \rightarrow How many parameters in the previous models? (= No. predictors + 2)

Statistical inference on regression coefficients

In the NHST approach, we can **test the statistical significance** of regression coefficients (*two-tail t-test*). This is automatically done by R in the model summary.

Summary(m3) # model results

Estimate Std. Error t value Pr(>|t|)

(Intercept) 104.25 0.88 118.22 0.000000e+00
age 3.70 0.06 57.45 0.000000e+00
sexmale -19.04 1.26 -15.14 1.237494e-49

0.09

- Estimate = estimated parameter
- ${\tt Std}$. ${\tt Error} = {\tt parameter} \ {\tt standard} \ {\tt error}$
- ${\tt t}$ value = test statistic computed as
- t = Estimate/Std.Error
- p-value = p corresponding to the t-value with No. Obs. No. Coeff. 1
 degrees of freedom

Effect size:

age:sexmale

Coefficient of determination

1.41

$$R^2 = 1 - SS_{residuals} / SS_{total}$$

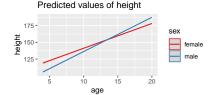
[1] 0.79

The model explains 79% of the variance in height.

Plotting effects:

15.39 3.897810e-51

sjPlot::plot_model(m3,type="pred",terms=c("age","sex"))



Hands on **R**

1. Download & read the dataset from the "Pregnancy during pandemics" study



depr = postnatal depression, age = mother's age, NICU = intensive care, threat = fear of COVID library(osfr) # package to interact with the Open Science Framework platform proj <- "https://osf.io/ha5dp/" # link to the OSF project osf download(osf ls files(osf retrieve node(proj))[2,],conflicts="overwrite") # download preg <- na.omit(read.csv("OSFData Upload 2023 Mar30.csv", stringsAsFactors=TRUE)) # read data colnames(preg)[c(2,5,12,14)] <- c("age", "depr", "NICU", "threat") # set variable names

- 2. Explore the the variables depr, threat, NICU, and age (descr., corr., & plots)
- 3. Fit a null model m0 of depr
- 4. Fit a simple regression model m1 with depr being predicted by threat
- 5. Fit a multiple regression model m2 also controlling for NICU and age
- 6 Fit an interactive model m3 to check whether age moderates the relationship between threat and depr.

- 7. Compare the models with AIC and likelihood ratio test: which is the best model?
- 8. Print & interpret the coefficients estimated by the selected model
- 9. Print & interpret the statistical significance of the estimated coefficients
- 10 Plot the effects of the selected model.
- 11. Compute the determination coefficient of the selected model

One step back: Linear model assumptions

Core assumptions:

- 1. Linearity: x_i and y_i are linearly associated \rightarrow the expected (mean) value of ϵ_i is zero
- 2. Normality: residuals ϵ_i are normally distributed with $\epsilon_i \sim \mathcal{N}(0, \sigma^2)$
- 3. Homoscedasticity: ϵ_i variance is constant over the levels of x_i (homogeneity of variance)
- 4. Independence of predictors & errors: predictors x_i are unrelated to residuals ϵ_i
- 5. Independence of observations: for any two observations i and j with $i \neq j$, the residual terms ϵ_i and ϵ_j are independent (no common disturbance factors)

Additional assumptions:

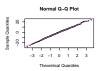
- 6. Absence of influential observations (multivariate outliers)
- 7. Absence of multicollinearity (for multiple regression):

lack of linear relationship between x_1 and x_2

Model diagnostics: Assessing LM assumptions

Normality & linearity ©
hist(residuals(m3))
qqnorm(residuals(m3)); qqline(residuals(m3))





Homoscedasticity & independence x,ϵ Θ

plot(residuals(m3) ~ children\$sex)
plot(residuals(m3) ~ children\$age)

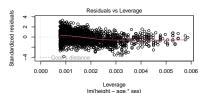




Independence of observations ?

Absence of influential cases ©

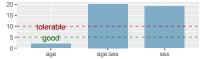
plot(m3,which=5)



Absence of multicollinearity

sjPlot::plot model(m3, "diag")[[1]]

Variance Inflation Factors (multicollinearity)



Are the unmeasured factors influencing y unrelated from one individual to another?

That's all for now!

Questions?

Homework (optional):

- read the slides presented today and write in the Moodle forum if you have any doubts
- refresh your familiarity with \mathbf{Q} : R-intro.pdf
- exe Cises 1-3 from exeRcises.pdf

For each exercise, the solution (or one of the possible solutions) can be found in dedicated chunk of commented code within the exercises.Rmd file

Achronyms & Greek letters

- AIC = Akaike Information Criterion
- BIC = Bayesian Information Criterion
- LM = linear models
- CI = confidence intervals
- MLE = maximum likelihood estimator
- OLS = ordinary least squares
- NHST = null hypothesis significance testing
- SE = standard error
- SS = sum of squares

- β = beta, used to index population-level intercept (β₀) and slope (β₁, β₂, etc.)
 parameters
- ε = epsilon, used to index
 population-level errors to be estimated
 based on model residuals
- σ = sigma, used to index the variance
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