### Bayesian Historical Monte Carlo Simulation

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#### Outline

- The current Bootstrapping method
- The Entropy pooling method
  - Time-conditioning
  - View-conditioning
  - The Entropy-minimization algorithm
- Practical application: Monte Carlo simulation with time/views-conditional probabilities



#### The Bootstrapping method

- Create future return predictions by random sampling from a time series
- Bootstrapping is a method to obtain a return confidence interval for a certain time horizon



#### The Bootstrapping method: an example

In this framework, users can express their views on the market through a two-step procedure:

- Standardize historical returns
- Incorporate views in the standardized returns:

$$r_{std} = \frac{r_j - \bar{r}}{\hat{\sigma}_r}, \quad r_{new} = (r_{std} \times \sigma_r^{view}) + \mathbb{E}^{view}(r)$$
 (1)

- r<sub>std</sub> are standardized returns
- $r_{new}$  are returns incorporating the user's views on the expected value and the volatility  $(\mathbb{E}^{view}(r), \sigma_r^{view})$



### Drawbacks of the current methodology

Currently, we are assigning equal probability to every joint realization of returns:

- Probabilities associated with the most recent scenarios should obtain more weight
- Probabilities associated with scenarios closer to the user's views should obtain more weight



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#### The Entropy pooling method:

Consider a N-dimensional random variable X (stock returns) with T the number of observations:

$$X := \{x_{t,j}\}_{t=1,2,...T}^{j=1,2,...N}$$

$$X \sim f_X, f_X := \{x_t, p_t\}_{t=1,2,...,T}$$

- The distribution of X is defined by a series of joint returns  $x_t$  with associated probabilities  $p_t$
- Every  $x_t$  (N-length vector for every t) is associated with a scalar  $p_t$  (a probability for every joint returns observation)



### The initial model (prior)

The initial reference model takes into account the time-decay effect by assigning higher weights to the most recent observations:

• 
$$f_X := \{x_t, p_t\}_{t=1,2,...,T}$$

• 
$$p = [p_1 ... p_t ... p_T]'$$

• 
$$p_{t-1} < p_t$$



### Exponential-decay model for probabilities

The differential equation describing the Exponential-decay<sup>1</sup>:

$$\frac{dp}{dt} = -\lambda p$$

is solved by:

$$p(t) = p_{ew}e^{-\lambda t}$$

- $\bullet$   $\lambda$  is the decay rate
- $au_{HL}=rac{\ln(2)}{\lambda}$ ,  $\lambda=rac{\ln(2)}{ au_{HL}}$



<sup>&</sup>lt;sup>1</sup>Exponential-decay

#### Time-conditioned probabilities

Each entry  $p_t$  of the vector p can be defined as follows:

- $\bullet \ p_t|\tau_{HL}:=p_{\mathrm{ew}}e^{-\frac{\ln(2)}{\tau_{HL}}(|t-T|)}$
- $p_{ew} := 1/\sum_{t} e^{-\frac{\ln(2)}{\tau_{HL}}(|t-T|)}$
- $p_{ew}$  is the equally-weighted probability
- $\tau_{HL}$  is approximately the time required for the probability of a scenario to decrease to half of its maximum value in T
- The lower is the half-life parameter  $\tau_{\it HL}$ , the higher is the decay rate  $\lambda$



## Time-conditioned probabilities: an example

t	$e^{-\lambda t-T }$	р
1	0,5	0.333
2	1	0.666

- $\bullet$   $au_{HI}=1$
- $\lambda = 0,693$
- T = 2

It is worth noting:

$$p(\tau_{HL}) = \frac{1}{2}p(T)$$



#### Prior visualization

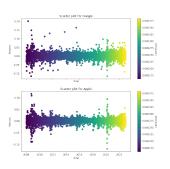


Figure: Google and Iberdrola SA returns scatter plot

From fig.(1) one can notice that most recent returns are getting higher probabilities.

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#### Defining the user's views

To mitigate numerical instability, the time series of each instrument will be normalized with the z-score method:

- $Z_j = \frac{X_j \bar{X}_j}{\sigma(X_j)}$
- Denoting with  $X_j$  the time series of the j-th portfolio's instrument returns
- $Z := \{z_{t,j}\}_{t=1,2,...T}^{j=1,2,...N}$
- $f_Z := \{z_t, p_t\}_{t=1,2,...,T}$



#### Defining the user's views

 Views (V) are represented as expressions of the expectation of arbitrary functions v(X) of returns<sup>2</sup>

$$V := \left\{ \mathbb{E}_p \bigg( v(X) \bigg) \ge v_*^{std} \right\} \tag{2}$$

• Where  $v_*^{std}$  is a threshold value that determines the intensity of the view



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 $<sup>^2</sup>$ in [1] several instances of how such functions v(X) can be defined are presented.

The function v(X) maps the risk drivers X in their standardized version Z:

• 
$$v(X) := Z'$$

• 
$$\mathbb{E}_p\bigg(v(X)\bigg):=v(X)p$$



Considering also views for volatilities,  $\Sigma_k(Z)$  denotes the matrix of volatilities over a rolling window k for instruments' standardized time series:

- $v_1(X) := Z'$  (constraints matrix rows for views on expected returns)
- $v_2(X) := \Sigma_k(Z)$  (constraints matrix rows for views on expected volatilities )
- $\mathbb{E}_p\bigg(v_j(X)\bigg):=v_j(X)p,\ j=1,2$



Assuming two stocks with an expected return of -16% and -20%, and an expected volatility of 26% and 29% respectively. One can express the views as follows:

- $v_* = [-0.16, 0.26, -0.20, 0.29]'$  is the user views' vector
- $v_*(\bar{X}) = [-0.16, -0, 20]$  is the returns' view intensity
- $v_*(\sigma) = [0.26, 0.29]$  is the volatilities' view intensity
- Then is necessary to define:  $f(v_*) \rightarrow v_*^{std}$

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To define  $f(v_*) \rightarrow v_*^{std}$ , one can follow the following steps:

- Fit a probability distribution for expected returns and volatilities
  - A chi-square  $\chi_{df,\lambda}$  for expected returns (fitted on X rolling average time series)
  - A log-normal  $\mathcal{LN}_{u,\sigma}$  for rolling volatilities (fitted on X rolling volatility time series)
- Take the value of the cumulative distribution function evaluated at the view intensity to extract the quantile:

• 
$$q_u, q_\sigma = \chi_{df,\lambda}\Big(v_*(\bar{X})\Big), \mathcal{LN}_{u,\sigma}\Big(v_*(\sigma)\Big)$$

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Then  $v_*^{std}$  can be defined as the value corresponding to quantiles  $q_u, q_\sigma$  in  $U_k(Z)$  and  $\Sigma_k(Z)$ :

$$v_*^{std} = [F_{U_k}^{-1}(q_u), F_{\Sigma_k}^{-1}(q_\sigma)]$$

- $F_{U_k}$  is the empirical cumulative distribution function for Z rolling average  $(U_k)$
- $F_{\Sigma_k}$  is the empirical cumulative distribution function for Z rolling volatilities

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The constraints matrix is:

$$v(X) = \begin{bmatrix} v_1(x_{1,G}) & \cdots & v_1(x_{t,G}) & \cdots & v_1(x_{T,G}) \\ v_2(x_{1,G}) & \cdots & v_2(x_{t,G}) & \cdots & v_2(x_{T,G}) \\ v_1(x_{1,I}) & \cdots & v_1(x_{t,I}) & \cdots & v_1(x_{T,I}) \\ v_2(x_{1,I}) & \cdots & v_2(x_{t,I}) & \cdots & v_2(x_{T,I}) \end{bmatrix}$$

- G, I= Google, Iberdrola SA
- $v_1(x_{t,j}) = z_{t,j}$  (standardized return of j at time t)
- $v_2(x_{t,j}) = \sum_k (z_{t,j})$  (rolling volatility of  $z_j$  at time t)

The constraint for Google's (bearish) view on expected return is defined in this way:

• 
$$v_1(X_G) = [v_1(x_{1,G}) \cdots v_1(x_{t,G}) \cdots v_1(x_{T,G})]$$

$$\bullet \; \mathbb{E}_p\bigg(v_1(X_G)\bigg) \leq v_*^{std}(\bar{X}_G)$$

That can also written as follows:

• 
$$Z'_{G}p \leq F_{U_{k,G}}^{-1}\bigg(\chi_{df_{G},\lambda_{G}}(-0.13)\bigg)$$

• Similarly we can define all the remaining constraints

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#### Posterior Visualization

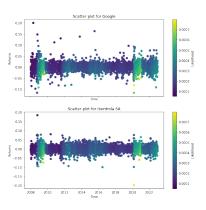


Figure: Google and Iberdrola SA return scatter plot

From fig.(2) one can notice that returns closer to the user views are getting higher probabilities

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### Computing the posterior probability distribution

- Our ultimate goal is to compute a posterior distribution  $f_X^{post}$ , that departs from the prior:  $f_X := \{x_t, p_t | \tau_{HI}\}_{t=1,2,...,T}$
- To take into account the views, the posterior distribution  $f_X^{post}$  is defined by new probabilities  $p_t^{post}$  for the same scenario outcomes  $x_t$ :

$$f_X^{post} := \{x_t, p_t^{post}\}_{t=1,2,...,T}$$



### Computing the posterior probability distribution

To compute  $p^{post}$ , we must rely on the relative entropy  $\xi(p^{post}, p)$  as a measure of the distance between p and  $p^{post}$ :

$$\xi(p^{post},p) := (p^{post})' \bigg( ln(p^{post}) - ln(p) \bigg)$$
 (3)



### Computing the posterior probability distribution

We then define the posterior as the distribution that is closest to the prior, as measured by (3), which satisfies the views (2):

$$\underset{p^{post} \in V}{\operatorname{argmin}_{p^{post}}} \ \xi(p^{post}, p), \tag{4}$$

using the Exponential-decay model as the prior for probabilities ( $p = p|\tau_{HL}$ ).



### Entropy minimization: defining constraints

To formulate the optimization problem, we must define both inequality and equality constraints:

$$V:=\bigg\{Fq\geq f, Hq=h\bigg\},$$

using a vector q as the counterpart of  $p^{post}$  posterior probabilities.



### Entropy minimization: defining constraints

- F = v(X)
- $q = [q_1 \dots q_t \dots q_T]'$  is a vector collecting the probability for each scenario  $x_t$  at time step t
- H is the counterpart of F for equality constraints:  $\sum_t q_t = 1$  is specified, using the vector  $H = \begin{bmatrix} 1 & \dots & 1 \end{bmatrix}$  and the scalar h = 1
- $f = v_*$  is the vector collecting the value for each view intensity



The optimization problem is:

$$\underset{Fq \leq f, Hq=h}{\operatorname{argmin}_{q}} \sum_{t=1}^{T} q_{t} (\ln(q_{t}) - \ln(p_{t}))$$
 (5)

and the Lagrangian function can be expressed in the vectorial notation as:

$$L(q, \lambda_1, \lambda_2) = q'(\ln(q) - \ln(p)) + \lambda'_1(Fq - f) + \lambda'_2(Hq - h)$$
(6)

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The Lagrange multipliers  $\lambda_1'$  and  $\lambda_2'$  are row vectors where the number of rows equals the number of inequality and equality constraints.



The first order condition for g reads:

• 
$$\frac{dL}{dq} = ln(q) - ln(p) + 1 + F'\lambda_1 - H'\lambda_2$$

• 
$$\frac{dL}{dq} = [0 \dots 0 \dots 0]'$$

and solving for q:

$$q(\lambda_1, \lambda_2) = e^{\ln(p) - 1 - F'\lambda_1 - H'\lambda_2}$$
 (7)



### The Duality principle

Given the convexity of  $\xi(q, p)$  in (4) for a posterior q and a fixed prior p, then:

$$\xi(q,p) \leq L(q,\lambda_1,\lambda_2), \forall \lambda_1 \geq 0.$$

Taking the minimum of both sides with respect to q, we get

$$\min_{q} \xi(q, p) \leq \min_{q} L(q, \lambda_1, \lambda_2).$$

Taking the maximum of both sides with respect to  $\lambda$ , we get:

$$\max_{\lambda_1 \geq 0, \lambda_2} \min_{q} \xi(q, p) \leq \max_{\lambda_1 \geq 0, \lambda_2} \min_{q} L(q, \lambda_1, \lambda_2)$$



#### The Duality principle

It is worth noting that, according to (5):

- If  $\lambda_1 < 0$  then the constraint Fq < f is violated
- $\lambda_2$  does not need to be constrained in order to satisfy Hq=h



#### The Duality principle

The dual function is given by definition:

$$G(\lambda_1,\lambda_2) := \min_{q} L(q,\lambda_1,\lambda_2)$$

therefore, we have:

$$\max_{\lambda_1 \geq 0, \lambda_2} G(\lambda_1, \lambda_2) \geq \min_q \xi(q, p)$$

- The solution of the dual problem is an upper bound for the solution of the primal problem
- If the objective function is strictly convex, then the minimization problem has a unique solution:

$$\max_{\lambda_1 \geq 0, \lambda_2} G(\lambda_1, \lambda_2) = \min_{q} \xi(q, p)$$



According to (7) the Lagrange dual function can be expressed as:

$$G(\lambda_1, \lambda_2) := L(q(\lambda_1, \lambda_2), \lambda_1, \lambda_2). \tag{8}$$

The two vectors of Lagrange multipliers  $(\lambda_1, \lambda_2)$  result from maximizing the Lagrange dual function:

$$\left(\lambda_1^*,\lambda_2^*\right) := \operatorname*{argmax}_{\lambda_1 \geq 0,} \left\{ \textit{G}(\lambda_1,\lambda_2) \right\}$$



#### KKT conditions

To ensure that  $\lambda_1, \lambda_2$  are solving the optimization problem (5), one need to check the KKT conditions:

$$\max_{\lambda_1 \geq 0, \lambda_2} \min_q L(q(\lambda_1 \lambda_2), p),$$

such that:

- $\frac{dL(q,p)}{dq} = 0$
- $\lambda_1(Fq f) = 0$
- $Fq f \le 0$
- Hq h = 0



#### KKT conditions

In the minimization case we have:

$$\min_{\lambda_1 \leq 0, \lambda_2} - \min_q L(q(\lambda_1 \lambda_2), p)$$

such that:

• 
$$\lambda_1(Fq - f) = 0$$

• 
$$Fq - f \ge 0$$

• 
$$Hq - h = 0$$



Finally, we can define the set of posterior probabilities as:

$$p^{post} := q(\lambda_1^*, \lambda_2^*) \tag{9}$$



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# Monte Carlo simulation with historical bootstrapping

Assuming we want to evaluate the expected performance of our portfolio over the next year:

- Generate  $\{x_{n,t}, p_t^{post}\}_{t=1,2,\dots,252}^{n=1,2,\dots,N}$  scenarios, for t time steps and n simulations
- For each time step t, calculate the portfolio return for each simulated scenario n

  - 2 with  $w_t$  portfolio weights at time step t
- Ompute the cumulative return over the time period, for each simulated scenario:



# Monte Carlo simulation with historical bootstrapping

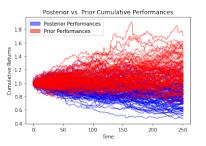


Figure: 100 paths of 1-euro investment in the Google-Iberdrola portfolio over 1 year, views' vector=  $v_* = [-0.16, 0.26, -0.20, 0.29]'$ 

For this simulation, the portfolio weights are constant and equally weighted  $(\frac{1}{2} \text{ Google}, \frac{1}{2} \text{ Iberdrola SA})$ .

# Monte Carlo simulation with historical bootstrapping

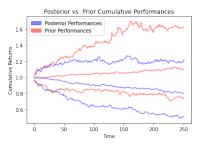


Figure: Quantiles of 1-euro investment in the Google-Iberdrola portfolio over 1 year

The last two figures, outline that negative views have a detrimental influence on portfolio performances → Fincite

There can be an edge case in which there is not an optimal solution for the user views:

- Consider a negative view of Google's expected return (-0.20%) and a positive one for Iberdrola SA (+0.20%)
- The optimization algorithm fails to find an optimal solution that fits both views
- In this case, a probability distribution will be fitted for every instrument's time series of returns



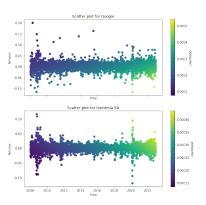


Figure: Google and Iberdrola SA return scatter plot

In Fig.(5), the probability distributions assign different weights to positive and negative returns.

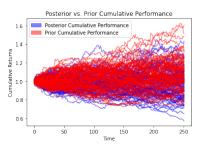


Figure: 100 paths of 1-euro investment in the Google-Iberdrola portfolio over 1 year

From Fig.(6), one can notice that a positive view for Iberdrola SA partially mitigates the negative view for Google.

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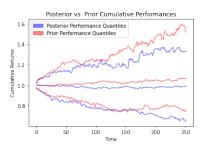


Figure: Quantiles of 1-euro investment in the Google-Iberdrola portfolio over 1 year

Additionally one can also observe the quantile comparison (7) for the new views on the portfolio's instruments.



#### Advantages of Entropy pooling approach

- Incorporate views
  - Entropy pooling can incorporate subjective views or beliefs about different risk factors
- Considers Non-Normal Distributions
  - It does not rely on normality in the risk factors distributions, allowing for non-Gaussian return distribution assumptions
- Addresses Multidimensionality
  - It accounts for multiple risk factors simultaneously, which can be especially useful in complex portfolios



[1] Attilio Meucci. "Mixing probabilities, priors and kernels via entropy pooling". In: *GARP Risk Professional* (2011), pp. 32–36.

